Estimating Anesthesia Time Using the Medicare Claim

A Validation Study

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ABSTRACT

Introduction: Procedure length is a fundamental variable associated with quality of care, though seldom studied on a large scale. The authors sought to estimate procedure length through information obtained in the anesthesia claim submitted to Medicare to validate this method for future studies.

Methods: The Obesity and Surgical Outcomes Study enlisted 47 hospitals located across New York, Texas, and Illinois to study patients undergoing hip, knee, colon, and thoracotomy procedures. A total of 15,914 charts were abstracted to determine body mass index and initial patient physiology. Included in this abstraction were induction, cut, close, and recovery room times. This chart information was merged to Medicare claims that included anesthesia Part B billing information. Correlations between chart times and claim times were analyzed, models developed, and median absolute differences in minutes calculated.

Results: Of the 15,914 eligible patients, there were 14,369 for whom both chart and claim times were available for analysis. For these 14,369, the Spearman correlation between chart and claim time was 0.94 (95% CI 0.94, 0.95), and the median absolute difference between chart and claim time was only 5 min (95% CI: 5.0, 5.5). The anesthesia claim can also be used to estimate surgical procedure length, with only a modest increase in error.

Conclusion: The anesthesia bill found in Medicare claims provides an excellent source of information for studying surgery time on a vast scale throughout the United States. However, errors in both chart abstraction and anesthesia claims can occur. Care must be taken in the handling of outliers in these data.

What We Already Know about This Topic

- Procedure length is a fundamental variable associated with quality of care, but the accuracy of using anesthesia claims from Medicare for this variable is not well known

What This Article Tells Us That Is New

- In a review of more than 14,000 charts, the anesthesia claim regarding procedure length provided a good estimate of the procedure length obtained from chart review

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PROCEDURE length is a fundamental variable used to describe surgical performance and even quality of care because it has been shown to be associated with postoperative complications[1–17] and is an integral part of any measurement of efficiency.[18–28] In previous research we reported on a method to estimate both anesthesia and surgical procedure length using the anesthesia Medicare claim based on 1,931 high-risk general surgery and orthopedics cases performed during 1995 and 1996 in Pennsylvania. We found that we could achieve an excellent prediction of anesthesia chart time using anesthesia claims data ($R^2 = 0.89$).[29] We subsequently used the method to estimate procedure times in the 20 most frequent orthopedic and 20 most frequent general surgical procedures in Pennsylvania during that period.[30] Other investigators have used this technique to answer questions re-

garding procedure length, yet no new validations with large-scale chart abstraction have been attempted.[10,31–38]

Obtaining the anesthesia chart time from the Medicare claim is not straightforward because the Medicare variable was not developed with this purpose in mind. As will be described, there is considerable opportunity for the anesthesia claim time to diverge from the chart time for a number of reasons: (1) the anesthesia claims do not always specify the exact surgical procedure associated with that claim, so matching anesthesia procedure to surgical procedure is not simple; (2) there may be mistakes in the claim; (3) there may be mistakes in the chart abstraction; (4) there may be confusion concerning times when more than one anesthesia provider was involved with the same surgical case and billed for overlapping time periods (such as a physician and nurse anesthetist billing for the same case as part of the anesthesia team). For all these reasons, the claim-derived time may not necessarily be correct. The intent of this article is to demonstrate that using our proposed algorithm, the Medicare anesthesia claim can be used to accurately obtain procedure time information.

In this report we present chart abstraction data on a far larger data set (more than sevenfold larger than our previous study), over three different states, across four types of surgery. Our original report, which was based on a case control study of mortality in Pennsylvania, analyzed patients who were uniformly very ill (all died within 60 days of admission, and control subjects were matched to these patients based on similar comorbidities and age). The current report, using data that are a decade more recent than those of the original study, provides an update using a population far more representative of patients undergoing the procedures studied, and uses a new and better methodology to estimate procedure length from Medicare claims while providing a more detailed account of the errors in measurement.

Establishing that anesthesia procedure time can be accurately estimated from Medicare claims may facilitate study of both surgical and anesthesia quality throughout the entire Medicare system. It also may aid in studying important clinical questions concerning cumulative anesthesia exposure time and its relationship to outcomes.

Materials and Methods

Study Perspective

The aim of this study is to inform researchers and policy analysts about the validity of using anesthesia claims data to determine procedure time when chart data are not available. We take the perspective that we wish to evaluate the proposed claims algorithm using chart data as a "gold standard." However, we have seen that both claims and charts have errors in measurement and recording. In an ideal world, one could have both claims and charts to inform the presence of errors from either source. For example, chart data may have transcription errors. If a chart time was 5 min for a colectomy, it is likely an error, especially if the claim time was 205

min. Because we wish to evaluate these two measures, in reporting these data we generally do not use chart data to inform or correct claims data, and we do not use claims data to inform or correct chart data, unless specifically stated.

Study Overview

The Obesity and Surgical Outcomes Study (OBSOS) is a study of surgery at 47 hospitals located throughout Illinois, New York, and Texas (appendix 1). Using Medicare claims, patients who underwent one of five types of surgery between 2002 and 2006 were identified in each study hospital: (1) hip replacement or revision excluding fracture (International Classification of Diseases, Ninth Revision [ICD-9] CM Principal Procedure codes 81.51–81.53); (2) knee replacement or revision (ICD-9 CM Principal Procedure 81.54, 81.55); (3) colectomy for cancer (ICD-9 CM Principal Procedure codes 45.7–45.79, 45.8 and ICD-9 CM Principal Diagnosis codes 153–153.9, 154–154.8, 230.3–6); (4) colectomy not for cancer (ICD-9 CM Principal Procedure 45.7–45.79, 45.8 and ICD-9 CM Principal Diagnosis codes 562.1–562.13); and (5) thoracotomy (ICD-9 CM Principal Procedure codes 32–32.9).

Hospitals were approached by the Oklahoma Foundation for Medical Quality and requested to abstract from 300 to 400 charts to collect baseline information including body mass index, admission vital signs and laboratory tests, and information on the surgical procedure including time of induction, initial surgery, closure, and time to recovery room. All data collected were deidentified and merged with encrypted Medicare files and sent to the study investigators for analysis. Approval was obtained from The Children's Hospital of Philadelphia Institutional Review Board (IRB) (Philadelphia, Pennsylvania), the IRB associated with the principal investigator of the study, as well as hospital-specific IRB approval when requested.

Sources of Error in Time Collection

The fundamental question we seek to answer is whether our algorithm for anesthesia claims data can provide valid anesthesia times. In this study we have the luxury of collecting chart-derived anesthesia time to aid in validation. However, even chart-derived anesthesia times are not perfect. In a study of this size, there are occasional mistakes in the abstraction of chart information when collecting anesthesia time and these mistakes may contribute to the appearance of mistakes in the claims algorithm. To make sense of these potential errors, we developed definitions for chart- and claim-derived variables based on an algorithm used by each definition. We define three times:

(a) "Isolated" Chart Time. This is a chart time cleaned in isolation from the claim time information. Changes in obviously incorrect dates (for example, off by 1 yr or 1 month) were corrected using only chart information, not claims information. If times were obviously incorrect ($\leq$ 30 min or $\geq$ 24 h), these were either fixed with internal information

from the chart or, if no time could be determined using only the chart, the chart time was coded as missing.

(b) "Isolated" Claim Time. This is a time derived from using the claim only and not using any chart information. To clean the isolated claim information, we used our claim algorithm (described below), or where the claim was obviously too long (≥ 24 h) or where the claim was obviously too short (≤ 30 min) for the procedures we coded the time as missing.

(c) "Best" Chart Time. This is a time derived using the best information available. Isolated chart time information was augmented with claim information. Correlating claim time with the "best" chart time is obviously tautologic, but there will be instances where we provide this information to place an informal upper bound on the quality of the claims information. The terms "chart time" and "claim time" always refer to "isolated chart" and "isolated claim" times unless otherwise noted.

Defining Anesthesia and Surgical Chart Time

Chart data on induction, incision, closure, and recovery room times were defined for the principal procedure in a standard manner as reported previously.[29] We collected chart time and date for start of induction, start of incision, end of closure, and entrance to the recovery room. Claim time can only directly provide anesthesia time, because that is how anesthesiologists bill Medicare. Anesthesia time refers to time from induction to recovery room. Surgical time is only available from the chart, and is defined as cut to close time because surgeons do not bill Medicare by the minute.[39]
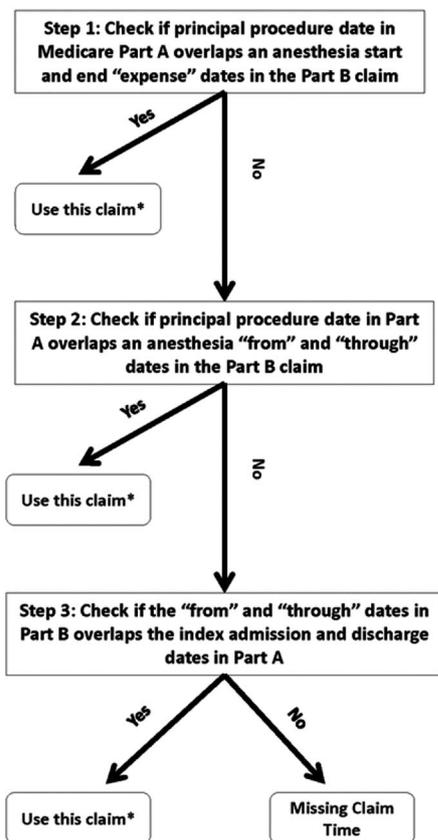
Statistical Methods

The correlations between chart and claim times were assessed with Pearson, Spearman, and Kendall correlation coefficients.[40] When performing multiple regression models, we used the Huber robust m-estimation as implemented in SAS Version 9 (SAS Institute, Inc., Cary, NC) using the bisquare weight function.[41–43] In our robust regressions, we report as R^2 (or rank R^2) the square of the Spearman rank correlation

Fig. 1. Algorithm for defining anesthesia claim time.

Step 1: Check if principal procedure date in Medicare Part A overlaps an anesthesia start and end "expense" dates in the Part B claim

Yes → Use this claim*

No ↓

Step 2: Check if principal procedure date in Part A overlaps an anesthesia "from" and "through" dates in the Part B claim

Yes → Use this claim*

No ↓

Step 3: Check if the "from" and "through" dates in Part B overlaps the index admission and discharge dates in Part A

Yes → Use this claim*

No → Missing Claim Time

*Each step may result in multiple qualifying anesthesia claims. Choose the claim with the longest claim time as defined by anesthesia time units.

Table 1. Patient and Hospital Characteristics in the Study Population

	Study Population	
Number of patients	15,914	—
Demographics	—	—
Age, ys (SD)	72.77	(4.10)
Male sex (%)	42.06	—
Race (White, %)	91.89	—
Race (Black, %)	4.90	—
Race (other, %)	3.21	—
Procedure	—	—
Hip (%)	24.63	—
Knee (%)	36.98	—
Colectomy for cancer (%)	15.02	—
Colectomy not for cancer (%)	7.34	—
Thoracotomy (%)	16.03	—
Hospital characteristics	—	—
Location	—	—
Illinois (%)	38.59	—
Texas (%)	43.19	—
New York (%)	18.22	—
Number of hospitals	47	—
Size (beds, SD)	534	(623)
Size distribution	—	—
0–200 (%)	31.92	—
201–400 (%)	27.66	—
401–600 (%)	10.64	—
601–800 (%)	8.51	—
>800 (%)	21.28	—
Nurse to bed ratio (SD)*	1.63	(1.01)
Nurse mix (SD)†	0.91	(0.53)
Technology index (%)‡	86.03	—
Teaching intensity based on RB ratio (%)	—	—
Nonteaching (RB ratio = 0)	46.81	—
Very minor (0 < RB < 0.05)	17.02	—
Minor (0.05 < RB < 0.25)	12.77	—
Major (0.25 < RB < 0.6)	10.64	—
Very major (0.6 < RB < 1.1)	12.77	—

* Full-time equivalent registered nurses/number of beds. † Registered nurses/(registered nurses + licensed practical nurses). ‡ Technology index = 1 if hospital performs open heart surgery, organ transplantation, or has a burn unit; otherwise index = 0. RB = resident-to-bed ratio.

between the observed and expected y variables, which is analogous to the square of the Pearson correlation between the observed and predicted ranks of y = chart time. This prevents one or two peculiar claims from greatly increasing or decreasing the R^2.

The Claims Algorithm

In order to ascertain the anesthesia time from the Medicare claims, we linked to the index admission in the Inpatient file all the claims in Part B that pertain to that patient. Then we selected only bills that identified an anesthesia service, these are bills with Healthcare Common Procedure Coding System codes in the range of 00100–01999. We applied the algorithm presented in (fig. 1) that ensured that we match the principal surgical procedure in the hospital's inpatient claim with the appropriate anesthesia bill from the provider in Part B. The first step was to align the dates in the Inpatient and in Part B files. We tried to match the anesthesia date in Part B to the surgical procedure date in the Inpatient file by choosing the anesthesia bill in Part B with the "first expense date" and the "last expense date" that included the "procedure date" of the principal procedure. If there was no overlap between any of the expense dates and the procedure date of the principal procedure (step 2), we used the interval between the "from date" and the "through date" in Part B that included the "procedure date" of the principal procedure. If there was no overlap between "from" and "through" dates of anesthesia bills and the procedure date of the principal procedure (step 3), we broadened the time frame in the hospital file so that the index admission and discharge dates in the hospital bill would overlap the "from"-"through" date interval in the provider bill. If multiple anesthesia bills were found that matched in terms of the time frame, we calculated each length and chose the bill with the longest time. If more than one provider reported the same longest time, we did not want to double count time. However, it was possible that both providers worked sequentially and did not perform services concurrently as assumed by our algorithm. As will be seen, the algorithm performs well despite the potential undercounting of time when anesthesia providers worked sequentially, leading us to conclude that for most cases the longest

Table 3. Relationship between Available Anesthesia Data Found in Claim and Chart

	Claim Time		
	Present	Missing	Totals
Chart time			
Present	14,369 (90.3%)	1,160 (7.3%)	15,529 (97.6%)
Missing	358 (2.2%)	27 (0.2%)	385 (2.4%)
Totals	14,727 (92.5%)	1,187 (7.5%)	15,914 (100%)

anesthesia time billed reflects the total time needed for the entire anesthesia procedure.

The length of the anesthesia is calculated by multiplying the anesthesia time unit variable in the Physician Part B file by 15 min per unit. The time units are identified by the variable "mile/time/units/services indicator" code; when this variable equals 2, it identifies anesthesia. For example, a time unit value of "25" implies 2.5 time units (Centers for Medicare and Medicaid Services reports units starting at the tenth place and do not provide the decimal point). We therefore multiply 2.5 units by 15 min/unit to get 37.5 min billed by the anesthesia provider.

Results

Description of Population and Setting

Table 1 displays the distribution of patient and hospital characteristics in the OBSOS study population. The OBSOS study was not a random sample of hospitals in the three states, but did provide a representative cross-section of hospitals and patients.

Table 2 describes comorbidities in each of the five procedure categories by study and nonstudy hospital groups. Again, the sample of 47 study hospitals is comprised of patients who look fairly similar to nonstudy hospital patients.

Description of Time Differences

Table 3 provides a comparison of missing data as defined by the claim time and chart time variables from the 15,914 patients evaluable in the OBSOS study. Here we see that there were missing data in both the claims- and chart-derived

Table 2. Comorbidities by Procedure Category in Study Population Compared with Nonstudy Hospital Population

	Procedure Category									
	Hip		Knee		Colectomy for Cancer		Colectomy Not for Cancer		Thoracotomy	
	In Study	Not In Study	In Study	Not In Study	In Study	Not In Study	In Study	Not In Study	In Study	Not In Study
Comorbidities	N = 3,920	N = 34,163	N = 5,885	N = 84,141	N = 2,391	N = 16,440	N = 1,168	N = 8,171	N = 2,550	N = 11,480
Mean number of comorbidities	2.45	2.46	2.36‡	2.43	3.91§	4.06	3.22	3.34	4.72	4.73
Hx CHF (%)	9.26*	8.32	7.39‡	8.52	17.61†	19.93	16.87‡	20.51	14.39‖	17.46
Hx AMI (%)	5.97	5.63	4.38	4.50	7.40	6.85	5.74	6.02	8.51	8.88
Hx Diabetes (%)	19.36	19.15	24.91	25.48	27.65†	30.38	25.51	26.44	23.10	23.48
Hx cancer (%)	18.72	18.29	14.43	14.47	N/A	N/A	27.91	25.82	90.35#	87.54

In Study N = 15,914; Not In Study N = 154,395 in Texas, Illinois, or New York with study eligible procedures and not in the 47 study hospitals.
P values compare In Study to Not in Study * < 0.05; † < 0.01; ‡ < 0.005; § < 0.001; ‖ < 0.0005; # <0.0001.
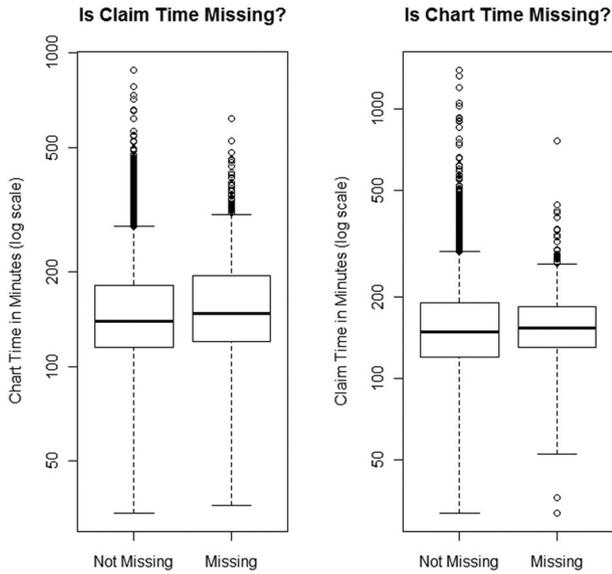
Fig. 2. Distribution of anesthesia claim and chart times by missing status.

times. There were 1,187 patients or 7.5% of the claims with missing times, and 385 patients or 2.4% of abstracted charts with missing times, but there was almost no overlap between the patients with missing claim time and those with missing chart time, with only 27 patients missing time data from both claim and chart. The distribution of anesthesia chart times in those patients who were missing anesthesia claim times was almost identical to the distribution of anesthesia chart times in those not missing anesthesia claim times. Similarly, the distribution of anesthesia claim times in those patients who were missing anesthesia chart times was almost identical to the distribution of anesthesia claim times in those patients who were not missing anesthesia chart times (fig. 2). Figure 2 suggests no interesting relationship between missing times on one variable and recorded times on another.

Using the 14,369 patients who had both chart and claim times, we next studied the correlations between chart and claim times. Table 4 provides these correlations using the Spearman, Pearson, and Kendall τ statistics with their CIs and P values. The correlations between chart and claim times were very high, ranging from 0.85 for the Kendall τ to 0.94 for the Spearman correlation. We also provide the probabil-

Table 4. Associations between Anesthesia Claim and Chart Times (N = 14,369)

	Value	95% CI	P Value
Pearson correlation	0.87	0.84, 0.91	< 0.0001
Spearman correlation	0.94	0.94, 0.95	< 0.0001
Kendall correlation (τ)	0.85	0.85, 0.86	< 0.0001
Probability of concordance*	0.93	0.92, 0.93	< 0.0001
Median absolute difference (min)	5.0	5.0, 5.5	—

* Probability of concordance = (1 + Kendall τ)/2.

Table 5. Associations between Anesthesia Claim and "Best" Chart Times (N = 14,370)

	Value	95% CI	P Value
Pearson correlation	0.89	0.85, 0.92	< 0.0001
Spearman correlation	0.94	0.94, 0.95	< 0.0001
Kendall correlation (τ)	0.85	0.85, 0.86	< 0.0001
Probability of concordance*	0.93	0.92, 0.93	< 0.0001
Median absolute difference (min)	5.0	5.0, 5.5	—

* Probability of concordance = (1 + Kendall τ)/2.

ity of concordance associated with the Kendall coefficient τ, which is equal to $(\tau + 1)/2$; for two patients, it is the probability that the chart and claim will agree about which patient had the longer anesthesia time. The probability of concordance was 0.93. The median absolute difference between chart and claim was only 5 min (95% CI 5.0, 5.5), and the median difference was 4.5 min.

Table 5 provides information similar to that in table 3 but compares claim time to "best" chart time. As expected, because of the definition of "best" chart time, there is better correlation between the claim time and the "best" chart time—although this is presented just to help bound the correlation because this calculation information from the claim was used to correct the chart time as was described in the Methods section.

Table 6 presents the distribution of claim times, chart times, and best chart times for each of the five procedure groups in the study (hip replacement or revision, knee replacement or revision, colectomy for cancer, colectomy not for cancer, and thoracotomy).

A Bland-Altman plot is presented in figure 3, which displays the difference between anesthesia chart and anesthesia claim times *versus* the average value of each pair. Most points show little difference between chart and claim, but there are a few outliers with respect to both measures. Because there are 14,369 points on this graph, it should be remembered that outliers represent only a very small fraction of patients. The "wings" of the Bland-Altman plot do show rare large outliers. Eighty percent of the pairs showed differences between −16 and 0.5 min, and 95% of the pairs show differences between −49 and 16 min.

Regression Modeling: The Influence of Patient, Procedure, and Hospital

We next asked whether we could detect any appreciable difference in the discrepancy between claim and chart depending on type of surgical procedure or on the specific study hospital. For each regression using m-estimation our dependent variable is chart time, and the independent variables are claim time as well as hospital identifiers and/or procedure types, depending on the model. We use m-estimation because we observe that there are some extreme outliers in both the claim and the chart, and as our work has suggested in the

Table 6. Distribution of Anesthesia Claim Time, Chart Time, and Best Chart Time in the Five Procedure Categories

Procedure Category	Time Variables (min)	N	5th %ile	25th %ile	50th %ile	75th %ile	95th %ile
Hip without fracture	Claim time	3,620	90	119	143	177	255
N = 3,920	Chart time	3,832	85	114	138	171	238
—	Best estimate chart	3,832	85	114	138	171	237
Knee	Claim time	5,491	78	114	137	165	233
N = 5,885	Chart time	5,753	73	108	130	159	220
—	Best estimate chart	5,753	74	108	130	159	220
Colectomy for cancer	Claim time	1,084	95	132	165	220	324
N = 1,168	Chart time	1,131	91	126	160	210	314
—	Best estimate chart	1,131	91	126	160	210	315
Colectomy not for cancer	Claim time	2,223	90	122	161	210	323
N = 2,391	Chart time	2,324	84	117	152	204	312
—	Best estimate chart	2,325	84	117	152	204	312
Thoracotomy	Claim time	2,309	104	146	185	230	326
N = 2,550	Chart time	2,489	95	135	175	221	310
—	Best estimate chart	2,489	95	135	175	220	309

past,[29] m-estimation is less sensitive to such errors.[41,42] Table 7 displays four models. Model 1 simply predicts chart time using claim time. Model 2 adds into model 1 individual hospitals. Model 3 adds procedure type to model 1. Finally, model 4 adds both hospital and procedure variables to model 1.

Model 1 suggests that we can estimate chart time very well with claim time. The coefficient on the claim time was nearly 1, and the intercept was −1.21 min. The model R^2 was 0.89. Model 2 asks whether the relationship between claim and chart changes with the hospital. We do observe that the hospital does have a significant influence on the model but the effects were extremely small. The hospital with the largest effect increased the difference between the claim and the chart only by 15 min (results not shown). We next asked if the individual procedure influenced the relationship between claim and chart. Again, we observed statistically significant but clinically insignificant ef-

fects from procedure, with effects on the order of only 1–2 min. Finally, model 4 includes both hospital and procedure variables. Again, we see no appreciable difference in the estimates. Hence, a single formula that is not adjusted for procedure or hospital appears reasonable for this data set.

In a related set of analyses we ran a series of models to explore whether patient procedure, patient characteristics, and hospital characteristics could predict the difference between anesthesia chart time and anesthesia claim time. Although some results were statistically significant, all effects were very small and not of clinical interest. The median absolute error for the model using procedure as an independent variable was 4.39 min, with only 2.39 min separating the most extreme procedures. Adding patient characteristics did not improve the median absolute error, and adding hospital characteristics only reduced the median absolute error to 3.88 min. Hence, patient, procedure, and hospital characteristics did not influence the errors between anesthesia chart time and claim time in this data set.

Estimating Surgical Chart Time from Anesthesia Claim Time

One very likely application of the algorithm we use to derive information from the anesthesia claim is the estimate of surgical time. Because only anesthesiologists, and not surgeons, bill by the minute, we do not have a direct bill for surgical time. We can, however, observe how well the Medicare anesthesia claim information can describe surgical time. We might imagine that the difference between the anesthesia claim time and the surgical chart time may be more susceptible to the influence of procedure type and hospital than when using anesthesia claim time to predict anesthesia chart time. This is because the style of practice in a hospital may dictate different styles of coordination between the surgeon and the anesthetist. Table 8 displays the exact models as seen in table 7, but here we have substituted anesthesia chart time with surgical chart time. There are some immediate differences between table 7 and table 8. We see that there is gen-
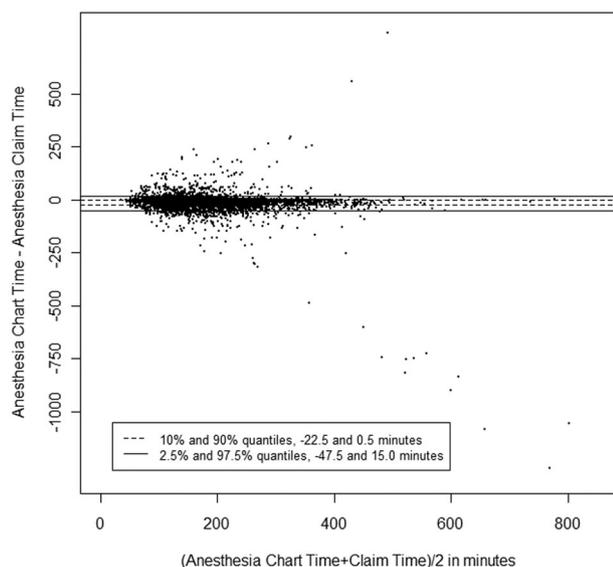

Fig. 3. Bland-Altman plot: Anesthesia chart time – anesthesia claim time *versus* (anesthesia chart time + anesthesia claim time)/2 (n = 14,369 observations).

Legend:
--- 10% and 90% quantiles, -22.5 and 0.5 minutes
— 2.5% and 97.5% quantiles, -47.5 and 15.0 minutes

Table 7. M-Estimation Regression Models to Predict Anesthesia Chart Time Using Anesthesia Claim Time

Regression Models	Model 1: Claim Time (min)	Model 2: Claim Time + Hospital Identifier	Model 3: Claim Time + Surgical Procedure	Model 4: Claim Time + Hospital Identifier + Surgical Procedure
Intercept	−1.21	(N/A)	−1.25	(N/A)
Claim time	0.977	0.989	0.979	0.990
(SE)	(0.0008)	(0.0006)	(0.0009)	(0.0007)
(CI)	(0.976–0.979)	(0.987–0.990)	(0.978–0.981)	(0.988–0.991)
Claim time P value	< 0.0001	< 0.0001	< 0.0001	< 0.0001
N	14,369	14,369	14,369	14,369
Model R^2	0.89	0.90	0.89	0.90
Median absolute error	4.33	2.80	4.32	2.79
P value (wald test) *vs.* model 1	—	< 0.0001	< 0.0001	< 0.0001

erally a 22-min gap between the total surgical time and the total anesthesia time. Luckily for the patient, the intercept term is negative, suggesting anesthesia time is longer than surgical time! We also observe that the influence of hospital style does play a slightly larger role in the regression, as does the influence of the procedure. However, as before, both effects were quite small, typically amounting to only a few minutes difference by institution or procedure. When we reran all models, substituting anesthesia chart time for anesthesia claim time, we obtained almost identical coefficients with slightly smaller median errors.

Figure 4 describes the relationship between surgical chart time and predicted surgical time derived from model 1 in table 8. Using anesthesia claims to predict surgical time was not as accurate as using anesthesia claims to predict anesthesia claim time, yet 80% of paired differences were between −24 and 19 min.

Best Anesthesia Chart Time versus Claim Time

Finally, we wish to describe the relationship between the anesthesia claim time and the anesthesia chart time that is corrected by claims when the chart time is missing. Although the relationship is tautologic in that we used some claim time information to "correct" obvious chart errors, we only cor-

rected these errors when we had no consistent chart information to make a judgment. In other words, for 146 patients we corrected the charts by using the claims, and the odds are great that these were fairly close (because there is only a 5-min median absolute time difference). We present this information to better describe how well an individual using the anesthesia claims could mimic the actual anesthesia time as determined as best as possible. As can be seen in table 9, the results were quite similar to the previous findings. Hence, using the claim time does an excellent job at predicting chart time. Figure 5 displays a Bland-Altman plot for the Best Anesthesia Chart Time and the Anesthesia Claim Time. These plots look almost identical to the Anesthesia Chart Time *versus* Anesthesia Claim Time displayed in figure 3.

Discussion

The OBSOS study provided us with a unique opportunity to examine how Medicare claims can be used to estimate procedure length, because the study was designed to measure surgical time and entailed the merging of chart information with Medicare claims. Procedure length is a fundamental variable associated with quality and outcomes. Many have published articles on procedure length,[1–17] often using chart

Table 8. M-Estimation Regression Models to Predict Surgical Chart Time Using Anesthesia Claim Time

| Regression Models	Model 1: Claim Time (min)	Model 2: Claim Time + Hospital Identifier	Model 3: Claim Time + Surgical Procedure	Model 4: Claim Time + Hospital Identifier + Surgical Procedure
Intercept	−21.77	(N/A)	−21.61	(N/A)
Claim time	0.805	0.850	0.803	0.848
(SE)	(0.0020)	(0.0019)	(0.0020)	(0.0019)
(CI)	(0.801–0.809)	(0.847–0.854)	(0.799–0.807)	(0.844–0.851)
Claim time P value	< 0.0001	< 0.0001	< 0.0001	< 0.0001
N	14,371	14,371	14,371	14,371
Model R^2	0.78	0.80	0.79	0.81
Median absolute error	9.63	8.33	9.25	8.12
P value (wald test) *vs.* model 1	—	< 0.0001	< 0.0001	< 0.0001
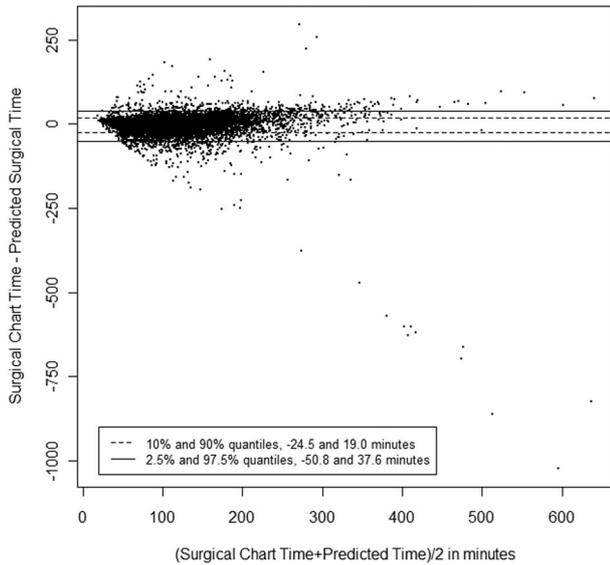
Fig. 4. Bland-Altman plot: Surgical chart time – predicted surgical chart time *versus* (surgical chart time + predicted surgical chart time)/2 (n = 14,371 observations).
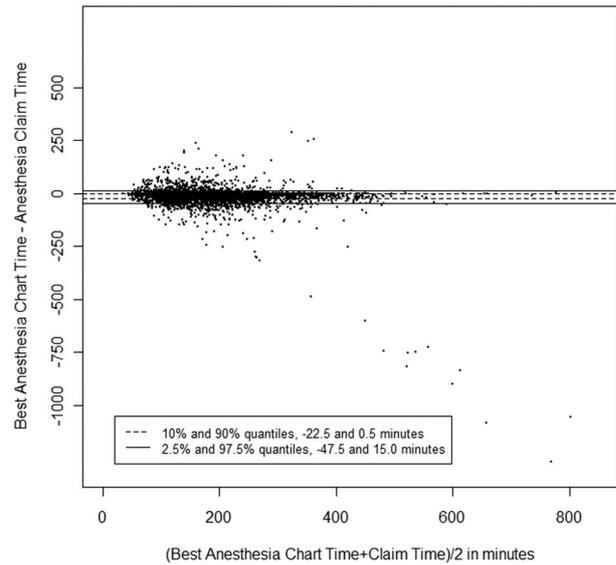

Fig. 5. Bland-Altman plot: Best anesthesia chart time–anesthesia claim time *versus* (best anesthesia chart time + anesthesia claim time)/2 (n = 14,369 observations).

reviews at single institutions.[1–7] If anesthesia claims could be used to reliably provide valid information on procedure length, then many questions now relying on single institution studies with relatively small data sets could be answered with much larger and more representative samples. For example, large-scale, nationwide studies of anesthesia claim time can be used to study a vast assortment of questions involving both clinical and health services research in anesthesiology and surgery. On the clinical side, better measures of anesthesia cumulative exposure may provide methods to study potential toxicities associated with anesthetic agents and may provide us with a better way to study and develop models that assess postoperative risk caused, in part, by deviations from the expected anesthesia time for the actual procedure performed. On the health services side, questions of quality can be studied with benchmarking across all hospitals that care for Medicare patients. Examples include the study of racial disparities in procedure length

inside and between hospitals throughout the United States, again, based on the actual procedures performed.

The results provided in the current study give the potential investigator a higher degree of confidence that anesthesia claims can be used to derive anesthesia time. The data presented in this study represent far more observations than those we reported on 3 yr ago. Previously, using data from 1995 to1996, we had analyzed 1,931 Medicare patients in 187 hospitals in the state of Pennsylvania. When we compared the chart to the claim, we observed a median absolute error of 5.49 min.[29] In the current study, we report on the abstraction of 14,369 Medicare charts in 3 states over 47 hospitals. We find a median absolute difference that was very small, only 5.0 min. In other words, we can be quite certain that for most cases, anesthesia claims work well at estimating anesthesia time.

In the current study, as in our original study, we did observe occasional errors that were substantial. Therefore, as

Table 9. M-Estimation Regression Models to Predict "Best" Anesthesia Chart Time Using Anesthesia Claim Time

Regression Models	Model 1: Claim Time (min)	Model 2: Claim Time + Hospital Identifier	Model 3: Claim Time + Surgical Procedure	Model 4: Claim Time + Hospital Identifier + Surgical Procedure
Intercept	−1.19	(N/A)	−1.23	(N/A)
Claim time	0.977	0.989	0.980	0.990
(SE)	(0.0008)	(0.0006)	(0.0009)	(0.0007)
(CI)	(0.976–0.979)	(0.987–0.990)	(0.978–0.981)	(0.988–0.991)
Claim time P value	<0.0001	< 0.0001	< 0.0001	< 0.0001
N	14,370	14,370	14,370	14,370
Model R^2	0.89	0.90	0.89	0.90
Median absolute error	4.30	2.79	4.30	2.78
P value (wald test) vs. model 1	—	< 0.0001	< 0.0001	< 0.0001

in the past report, we suggest the use of regression techniques that down-weight outliers when fitting models. Such techniques are ideally suited for problems such as ours, where claims information is usually correct but may occasionally fail to reflect the true procedure length because of mistakes in the algorithm that links claim to procedure, mistakes in the algorithm identifying whether anesthesiologists worked sequentially or concurrently, or mistakes in coding. In situations where there is no single member of the anesthesia team that bills for the entire procedure, the claim may underestimate the chart. Furthermore, we may observe situations where the claim overestimates the chart information. These instances may reflect mistaken linkages between the specific procedure for which the claim was made. Because anesthesia bills often use a "from-through" date that encompasses multiple procedures, one may mistakenly assign excess time to a single procedure that mistakenly reflects other procedures' time.

Although this article has focused on the potential use of anesthesia claim time as a dependent variable (an outcome variable) for many analyses, anesthesia claim time can also be used as an independent variable in models designed to predict outcomes.

Just as when a claim time is used as the dependent (y) variable in regression, it is important to fit these models using a robust method such as m-estimation[44] (because claim times closely reproduce chart times with rare but large errors); when a claim time is used as an independent (x) variable in a model, it is similarly important to fit these models using bounded-influence methods.[45,46]

Although we want investigators to be aware of the potential pitfalls in using claims to determine anesthesia and surgical time, we do not want to overstate these problems. The correlations we report, now in two separate studies spanning over 8 yr of data and close to 16,000 observations, are high and will be useful for applying the claims estimates to many important questions being studied concerning procedure time.

It is also interesting to note that billing styles were fairly similar across hospitals. We generally found only small differences between hospitals, with the exception of a few that were associated with 10- to 15-min claim-chart time differences. Furthermore, the median difference between the claim time and the chart time was 5 min. This number would not appear to be a coincidence. As one anesthesia time unit equals 15 min, a policy of always rounding up to the higher unit would lead to an approximate 5-min difference on average (assuming a uniform distribution for the fraction of units remaining before rounding).

In summary, we have demonstrated that the Medicare anesthesia claim can be used to construct an excellent measure of procedure time. Future investigators can be confident that they may use our algorithm to better study procedure length through using the Medicare claim, without the need to collect procedure length information directly from the chart.

References

1. Kessler S, Kinkel S, Käfer W, Puhl W, Schochat T: Influence of operation duration on perioperative morbidity in revision total hip arthroplasty. Acta Orthop Belg 2003; 69:328–33
2. McGillicuddy EA, Schuster KM, Davis KA, Longo WE: Factors predicting morbidity and mortality in emergency colorectal procedures in elderly patients. Arch Surg 2009; 144:1157–62
3. Peersman G, Laskin R, Davis J, Peterson MG, Richart T: Prolonged operative time correlates with increased infection rate after total knee arthroplasty. HSS J 2006; 2:70–2
4. Imai E, Ueda M, Kanao K, Kubota T, Hasegawa H, Omae K, Kitajima M: Surgical site infection risk factors identified by multivariate analysis for patient undergoing laparoscopic, open colon, and gastric surgery. Am J Infect Control 2008; 36:727–31
5. Collins R, Scrimgeour A, Yusuf S, Peto R: Reduction in fatal pulmonary embolism and venous thrombosis by perioperative administration of subcutaneous heparin. Overview of results of randomized trials in general, orthopedic, and urologic surgery. N Engl J Med 1988; 318:1162–73
6. Garibaldi RA, Cushing D, Lerer T: Risk factors for postoperative infection. Am J Med 1991; 91:158S–63S
7. Boruk M, Chernobilsky B, Rosenfeld RM, Har-El G: Age as a prognostic factor for complications of major head and neck surgery. Arch Otolaryngol Head Neck Surg 2005; 131:605–9
8. Scott CF Jr: Length of operation and morbidity: Is there a relationships? Plast Reconstr Surg 1982; 69:1017–21
9. Collins TC, Daley J, Henderson WH, Khuri SF: Risk factors for prolonged length of stay after major elective surgery. Ann Surg 1999; 230:251–9
10. Ong KL, Lau E, Manley M, Kurtz SM: Effect of procedure duration on total hip arthroplasty and total knee arthroplasty survivorship in the United States Medicare population. J Arthroplasty 2008; 23:127–32
11. Campbell DA Jr, Henderson WG, Englesbe MJ, Hall BL, O'Reilly M, Bratzler D, Dellinger EP, Neumayer L, Bass BL, Hutter MM, Schwartz J, Ko C, Itani K, Steinberg SM, Siperstein A, Sawyer RG, Turner DJ, Khuri SF: Surgical site infection prevention: The importance of operative duration and blood transfusion–results of the first American College of Surgeons-National Surgical Quality Improvement Program Best Practices Initiative. J Am Coll Surg 2008; 207:810–20
12. Hollenbeck BK, Miller DC, Taub D, Dunn RL, Khuri SF, Henderson WG, Montie JE, Underwood W 3rd, Wei JT: Identifying risk factors for potentially avoidable complications following radical cystectomy. J Urol 2005; 174:1231–7; discussion 1237
13. Rogers SO Jr, Kilaru RK, Hosokawa P, Henderson WG, Zinner MJ, Khuri SF: Multivariable predictors of postoperative venous thromboembolic events after general and vascular surgery: Results from the patient safety in surgery study. J Am Coll Surg 2007; 204:1211–21
14. Culver DH, Horan TC, Gaynes RP, Martone WJ, Jarvis WR, Emori TG, Banerjee SN, Edwards JR, Tolson JS, Henderson TS: Surgical wound infection rates by wound class, operative

procedure, and patient risk index. National Nosocomial Infections Surveillance System. Am J Med 1991; 91:152S–7S

15. Wallner LP, Dunn RL, Sarma AV, Campbell DA Jr, Wei JT: Risk factors for prolonged length of stay after urologic surgery: The National Surgical Quality Improvement Program. J Am Coll Surg 2008; 207:904–13

16. Khuri SF, Najjar SF, Daley J, Krasnicka B, Hossain M, Henderson WG, Aust JB, Bass B, Bishop MJ, Demakis J, DePalma R, Fabri PJ, Fink A, Gibbs J, Grover F, Hammermeister K, McDonald G, Neumayer L, Roswell RH, Spencer J, Turnage RH, VA National Surgical Quality Improvement Program: Comparison of surgical outcomes between teaching and nonteaching hospitals in the Department of Veterans Affairs. Ann Surg 2001; 234:370–82; discussion 382–3

17. Schwartz SR, Yueh B, Maynard C, Daley J, Henderson W, Khuri SF: Predictors of wound complications after laryngectomy: A study of over 2000 patients. Otolaryngol Head Neck Surg 2004; 131:61–8

18. Abouleish AE, Prough DS, Zornow MH, Hughes J, Whitten CW, Conlay LA, Abate JJ, Horn TE: The impact of longer-than-average anesthesia times on the billing of academic anesthesiology departments. Anesth Analg 2001; 93:1537–43, table of contents

19. Schuster M, Standl T, Reissmann H, Kuntz L, Am Esch JS: Reduction of anesthesia process times after the introduction of an internal transfer pricing system for anesthesia services. Anesth Analg 2005; 101:187–94, table of contents

20. Sandberg WS, Daily B, Egan M, Stahl JE, Goldman JM, Wiklund RA, Rattner D: Deliberate perioperative systems design improves operating room throughput. ANESTHESIOLOGY 2005; 103:406–18

21. Torkki PM, Marjamaa RA, Torkki MI, Kallio PE, Kirvelä OA: Use of anesthesia induction rooms can increase the number of urgent orthopedic cases completed within 7 hours. ANESTHESIOLOGY 2005; 103:401–5

22. Dexter F, Coffin S, Tinker JH: Decreases in anesthesia-controlled time cannot permit one additional surgical operation to be reliably scheduled during the workday. Anesth Analg 1995; 81:1263–8

23. Dexter F: A strategy to decide whether to move the last case of the day in an operating room to another empty operating room to decrease overtime labor costs. Anesth Analg 2000; 91:925–8

24. Abouleish AE, Prough DS, Vadhera RB: Influence of the type of anesthesia provider on costs of labor analgesia to the Texas Medicaid program. ANESTHESIOLOGY 2004; 101:991–8

25. Abouleish AE, Dexter F, Whitten CW, Zavaleta JR, Prough DS: Quantifying net staffing costs due to longer-than-average surgical case durations. ANESTHESIOLOGY 2004; 100:403–12

26. Watcha MF, White PF: Economics of anesthetic practice. ANESTHESIOLOGY 1997; 86:1170–96

27. Abouleish AE: Increasing operating room throughput: Just buzzwords for this decade? ANESTHESIOLOGY 2008; 109:3–4

28. Smith MP, Sandberg WS, Foss J, Massoli K, Kanda M, Barsoum W, Schubert A: High-throughput operating room system for joint arthroplasties durably outperforms routine processes. ANESTHESIOLOGY 2008; 109:25–35

29. Silber JH, Rosenbaum PR, Zhang X, Even-Shoshan O: Estimating anesthesia and surgical procedure times from medicare anesthesia claims. ANESTHESIOLOGY 2007; 106:346–55

30. Silber JH, Rosenbaum PR, Zhang X, Even-Shoshan O: Influence of patient and hospital characteristics on anesthesia time in medicare patients undergoing general and orthopedics surgery. ANESTHESIOLOGY 2007; 106:356–64

31. Kurtz SM, Ong KL, Lau E, Bozic KJ, Berry D, Parvizi J: Prosthetic joint infection risk after TKA in the Medicare population. Clin Orthop Relat Res 2010; 468:52–6

32. Pandit JJ: Similarity of operation times for common general surgical procedures in the United Kingdom and the United States. ANESTHESIOLOGY 2007; 107:512–3

33. Westbury S, Pandit M, Pandit JJ: Matching surgical operating capacity to demand using estimates of operating times. J Health Organ Manag 2009; 23:554–67

34. Stubbs D, Ward ME, Pandit JJ: Estimating hourly anaesthetic and surgical reimbursement from private medical insurers' benefit maxima: Implications for pricing services and for incentives. Anaesthesia 2010; 65:396–408

35. Eijkemans MJ, van Houdenhoven M, Nguyen T, Boersma E, Steyerberg EW, Kazemier G: Predicting the unpredictable: A new prediction model for operating room times using individual characteristics and the surgeon's estimate. ANESTHESIOLOGY 2010; 112:41–9

36. Pandit JJ, Stubbs D, Pandit M: Measuring the quantitative performance of surgical operating lists: Theoretical modelling of 'productive potential' and 'efficiency'. Anaesthesia 2009; 64:473–86

37. Burgess DJ: Are providers more likely to contribute to healthcare disparities under high levels of cognitive load? How features of the healthcare setting may lead to biases in medical decision making. Med Decis Making 2010; 30:246–57

38. Hanss R, Roemer T, Hedderich J, Roesler L, Steinfath M, Bein B, Scholz J, Bauer M: Influence of anaesthesia resident training on the duration of three common surgical operations. Anaesthesia 2009; 64:632–7

39. Centers for Medicare and Medicaid Services: Chapter 12: Physicians/Nonphysician Practitioners; Section 50: Payment for ANESTHESIOLOGY Services; Part A: General Payment Rule, Medicare Claims Processing Manual, 2007, pp 88–9

40. Hollander M, Wolfe DA: Nonparametric Statistical Methods, 2nd Edition. New York, NY, John Wiley & Sons, 1999

41. Huber PJ: Chapter 3. The basic types of estimates, in Robust Statistics. Hoboken, NJ, John Wiley & Sons, 1981, pp 43–55

42. Hampel FR, Ronchett EM, Rousseeuw PJ, Stahel WA: Linear Models: Robust Estimation, in Robust Statistics: The Approach Based on Influence Functions. New York, NY, John Wiley & Sons, 1986, pp 315–28

43. Huber PJ: Robust regression: Asymptotics, conjectures and the Monte Carlo. Ann Stat 1973; 1:799–821

44. Huber PJ: Robust Statistics. Hoboken, NJ, John Wiley & Sons, 1981

45. Markatou M, Hettmansperger TP: Robust bounded-influence tests in linear models. J Am Stat Assoc 1990; 85:187–90

46. Heritier S, Ronchetti E: Robust bounded-influence tests in general parametric models. J Am Stat Assoc 1994;89:897–904

Appendix 1. List of Study Hospitals

	Hospital Name	City and State
1	Advocate Good Samaritan Hospital	Downers Grove, Illinois
2	Advocate Lutheran General Hospital	Park Ridge, Illinois
3	Baptist Health System	San Antonio, Texas
4	Baylor All Saints Medical Center	Dallas, Texas
5	Baylor Medical Center at Garland	Dallas, Texas
6	Baylor Medical Center at Grapevine	Dallas, Texas
7	Baylor Medical Center at Irving	Dallas, Texas
8	Baylor Medical Center at Plano	Dallas, Texas
9	Baylor University Medical Center	Dallas, Texas
10	Blessing Hospital	Quincy, Illinois
11	Carle Foundation Hospital	Urbana, Illinois
12	Central DuPage Hospital	Winfield, Illinois
13	ChristusSpohn Hospital	Corpus Christi, Texas
14	Christus St. Michael Health System	Texarkana, Texas
15	Covenant Health System	Lubbock, Texas
16	Good Shepherd Medical Center	Longview, Texas
17	Harris Methodist Fort Worth	Fort Worth, Texas
18	Harris Methodist HEB	Bedford, Texas
19	Huntington Hospital	Great Neck, Texas
20	Kaleida Health	Buffalo, Texas
21	Lake Forest Hospital	Lake Forest, Illinois
22	Las Palmas Medical Center	El Paso, Texas
23	Loyola University Medical Center	Maywood, Illinois
24	Memorial Sloan-Kettering Cancer Center	New York, New York
25	Methodist Medical Center of Illinois	Peoria, Illinois
26	New York Hospital Queens	Flushing, New York
27	North Shore Long Island Jewish Medical Center	Great Neck, New York
28	North Shore University Hospital	Great Neck, New York
29	Northwestern Memorial Hospital	Chicago, Illinois
30	NYU Langone Medical Center	New York, New York
31	Passavant Area Hospital	Jacksonville, Illinois
32	Proctor Hospital	Peoria, Illinois
33	Providence Health Center	Waco, Texas
34	Riverside Medical Center	Kankakee, Illinois
35	Rockford Memorial	Rockford, Illinois
36	Seton Medical Center Austin	Austin, Texas
37	Sherman Hospital	Elgin, Illinois
38	SID Peterson Hospital	Kerrville, Texas
39	Southwest Texas Methodist	San Antonio, Texas
40	St. James Olympia Fields	Olympia Fields, Illinois
41	St. Joseph Regional Health System	Bryan, Texas
42	St. Anthony's Memorial Hospital	Effingham, Illinois
43	St. Elizabeth	Beaumont, Texas
44	St. John's Hospital	Springfield, Illinois
45	Texas Orthopedic Hospital	Houston, Texas
46	Trinity Medical Center	Rock Island, Illinois
47	Unity Hospital	Rochester, New York

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