

# Use and abuse of detrital zircon U-Pb geochronology—A case from the Río Orinoco delta, eastern Venezuela

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## ABSTRACT

Advancements in mass spectrometry methods over the past two decades have allowed for rapid measurement of (U-Th)/Pb isotopes for geochronologic applications, a tool that has deeply influenced the way sediment provenance and paleo-tectonic reconstructions are approached. Geochronology-based studies of sediment provenance typically rely on dating  $n \approx 100$ –150 single detrital zircon crystals from individual samples, where the sample age distributions are used to make inferences about the parent age distributions, make qualitative geologic interpretations, and/or perform quantitative intersample comparisons. Most efforts to quantitatively compare detrital zircon age spectra make use of non-parametric dissimilarity statistics. Here, we use laser ablation–inductively coupled plasma–mass spectrometry (LA-ICP-MS) U-Pb detrital zircon moderate- $n$  ( $n \approx 100$ ) and large- $n$  ( $n \approx 1000$ ) results from unconsolidated fluvial sediments of the Río Orinoco delta, eastern Venezuela, to highlight the concealed pitfalls of making geological interpretations based on quantitative comparisons of U-Pb age distributions alone. Three samples analyzed at large  $n$ , selected from contrastingly different mean sediment grain sizes along the active channel of the Río Orinoco, yield large dissimilarities amongst their age spectra, resulting in the misleading conclusion that these were likely not sourced from the same parent distribution. We demonstrate that statistically significant differences amongst detrital zircon samples derived from the same (integrated) source region can be introduced by the dynamics of sediment transport, which may in turn lead to erroneous geologic interpretations arising from the inaccurate assumptions that, at present, condition the quantitative treatment of detrital zircon data.

## INTRODUCTION

“The fascinating impressiveness of rigorous mathematical analysis, with its atmosphere of precision and elegance, should not blind us to the defects of the premises that condition the whole process.”—Thomas C. Chamberlin (Chamberlin, 1899).

Most sediment provenance studies that apply quantitative detrital zircon U-Pb geochronology rely on the assumption that direct intersample comparisons using dissimilarity metrics allows for determination of similar and/or distinct sediment sources (e.g., Saylor and Sundell, 2016; Vermeesch, 2018, and references therein). There are, however, two inconvenient and commonly ignored issues in geochronology-based sediment

provenance: (1) detrital zircon U-Pb ages can cluster by grain size (Sircombe et al., 2001; Garzanti et al., 2009; Lawrence et al., 2011); and (2) grain-size sorting is strongly influenced by the dynamics of sediment transport (e.g., Udden, 1914; Hajek et al., 2010). The size, shape, and density of detrital minerals are the primary physical properties that determine the settling and deposition of detritus entrained within tractive currents (e.g., Schuiling et al., 1985; Malusà et al., 2016). Thus, even if sediments are derived from the same catchment area, sedimentary rocks containing detrital zircon crystals with vastly different physical properties may appear to have distinguishable provenance characteristics based on comparison of detrital zircon age distributions alone.

To determine the provenance of sediments, geologists are often tasked with making geological interpretations based on comparison of samples collected tens to thousands of kilometers apart, sometimes across terrane boundaries and modern or ancient ocean basins (e.g., Rainbird et al., 1992). One presumption that has led workers astray in detrital mineral geochronology is the belief that if a sufficient number of age observations are made in a sample, then the measured distribution of ages will approach the “true” distribution present in a given sedimentary system with sufficient accuracy to make geologically sound interpretations (e.g., Pullen et al., 2014). This premise remains ignorant of how detrital minerals fractionate within sedimentary environments, and also presumes that biases introduced through sample selection and processing are negligible (Sircombe and Stern, 2002; Sláma and Košler, 2012).

Here, we document the fractionation of detrital zircon age distributions within the natural fluvial system of the Río (River) Orinoco delta, eastern Venezuela (Fig. 1), and illustrate the issues with current quantitative approaches to detrital zircon age comparisons. We show that single samples, even if investigated at large numbers of single detrital zircon crystals ( $n \approx 1000$ ), are insufficient to adequately characterize the complete spectrum of detrital zircon ages available to a fluvial system from any given sample. In turn, our results demonstrate that physico-mechanical biasing (e.g., size and morphological sorting) can influence the distribution of ages found amongst samples deposited in contrasting energy environments within the same sedimentary system, thus affecting the accuracy of quantitative provenance interpretations built exclusively upon detrital

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**Figure 1. A: Composite satellite image of Río Orinoco delta, eastern Venezuela (base images from Google Earth). Yellow box indicates location of panel B. B: Sampling localities across Río Orinoco delta; white circles and stars denote samples investigated using  $n \approx 100$  and  $n \approx 1000$  single detrital zircon crystals, respectively.**

zircon U-Pb dates. These findings are used to (1) point out pitfalls of current approaches and identify where the greatest gains can be made in the future, and (2) remind practitioners of this technique that “statistically significant” is not necessarily the same as “geologically significant” (e.g., Vermeesch, 2009).

## GEOLOGIC SETTING AND DETRITAL ZIRCON GEOCHRONOLOGY

Uranium-lead (U-Pb) dates were measured for eight unconsolidated sediment samples from an active delta channel of the Río Orinoco (Fig. 1), the fourth-largest river on Earth by discharge volume of water. Zircon U-Pb dates were determined by laser ablation–inductively

coupled plasma–mass spectrometry (LA-ICP-MS) following methods outlined in Pullen et al. (2014) (see the GSA Data Repository<sup>1</sup> for results and analytical methods). The laser beam diameter for these experiments was 20  $\mu\text{m}$  and thus defined the analytical limit on the zircon crystal size that could be dated. Samples were collected over a relatively small area ( $\sim 26 \text{ km}^2$ ) across the same channel, compared to the overall magnitude of the river length ( $\sim 2.1 \times 10^3 \text{ km}$ ) and drainage basin area ( $\sim 8.8 \times 10^5 \text{ km}^2$ ). This narrowly defined sampling was employed to focus on the channel-perpendicular fractionation of detrital zircon ages (i.e., differences in zircon age distributions occurring along the “cross-sectional energy profile” of the channel), as opposed to the fractionation of detrital zircon age populations that would occur along the thousands-of-kilometers-long path of the river (i.e., channel-parallel fractionation). A moderate number of dates ( $n \approx 100$ ) were determined for each of the eight samples (Fig. 2A). Three of these samples were selected for large- $n$  dating based on their contrasting mean grain sizes (Figs. 2B and 2C); these samples include an overbank deposit (sample RORDVN05, mean zircon size =  $75 \pm 18 \mu\text{m}$ ), a mid-channel sandy bar deposit (sample RORDVN02, mean zircon size =  $130 \pm 30 \mu\text{m}$ ), and a coarse-grained mid-channel bed load sample (sample RORDVN03, mean zircon size =  $220 \pm 45 \mu\text{m}$ ; Figs. 1 and 2B).

## RESULTS AND DISCUSSION

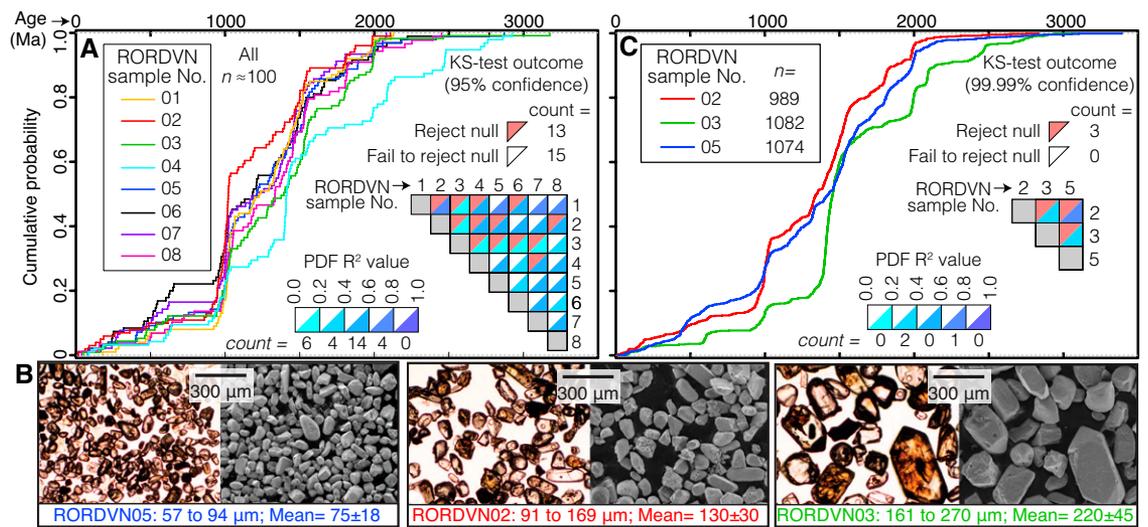
The sediment load of a continental-scale river system like the Río Orinoco is expected to include detrital zircon crystals of multiple definable age components (or modes). Major age components from a sample (e.g., fractional

abundance  $f \geq 0.15$ ) have a reasonably high probability of being recognized if  $n = 100$ –120 individual crystals are dated (Vermeesch, 2004; Andersen, 2005). Nevertheless, the relative proportions of those major age components at this level of observation are unlikely to be determined with accuracy, nor could that accuracy be assessed (Pullen et al., 2014). Additionally, there is a high probability that lower-abundance (e.g.,  $f < 0.1$ ), albeit potentially diagnostic, age modes would remain unobserved or poorly represented with moderate- $n$  studies ( $n \approx 100$ ).

Moderate- $n$  U-Pb dating of detrital zircon crystals from all eight samples collected from the Río Orinoco shows the presence of multiple ( $>10$ ) age modes within all samples, with three of these (i.e., ca. 1000, ca. 1400, and ca. 1500 Ma) representing  $f > 0.15$  each in most cases (Fig. 2A). Because of undersampling, moderate- $n$  analyses cannot determine with accuracy the percent that each of these age modes compose of the “true” distribution of zircon ages within a sample (Andersen, 2005). Furthermore, because of the large intersample variability, it is similarly unrealistic to accurately determine the “true” distribution of zircon ages available to the bulk sedimentary system, or address whether the complete age spectrum could be determined from the analysis of a single sample even if a large number of single-grain observations were made (see discussion below).

To quantitatively assess the correlations that may exist among samples and/or evaluate the likelihood that a given set of samples could be drawn from the same parent distribution (i.e., the “null” hypothesis), most studies make use of non-parametric similarity-dissimilarity metrics for intersample comparisons (see Vermeesch, 2018,

**Figure 2. Cumulative density functions of samples analyzed in this study and graphical representation of upper-half portion of dissimilarity matrices showing results of Kolmogorov-Smirnov (KS) test and probability density function (PDF) cross-correlation ( $R^2$ ) values. Numbers bordering the matrices are sample numbers. A: Results of eight samples analyzed using  $n \approx 100$  single detrital zircon crystals. B: Representative transmitted light microscope images (left) and scanning electron microscope–backscattered electron images (right) of zircon crystals from three samples investigated at  $n \approx 1000$ ; mean short-axis and long-axis lengths of zircons found in each sample are shown below each panel. C: Results of three samples analyzed at  $n \approx 1000$ . See Data Repository (see footnote 1) for PDFs and kernel density estimate plots of all samples, as well as complete dissimilarity matrices.**



<sup>1</sup>GSA Data Repository item 2018388, U-Pb geochronology methods, analytical results, PDF and KDE figures for all samples, dissimilarity matrices, and KS test results from bootstrap subsamples, is available online at <http://www.geosociety.org/datarepository/2018/> or on request from [editing@geosociety.org](mailto:editing@geosociety.org).

and references therein). Traditionally, most researchers apply a two-sided Kolmogorov-Smirnov (KS) or Kuiper test to the cumulative density functions (CDF), and use the test's  $D$ - and/or  $p$ -values to quantify dissimilarity and/or to reject (or fail to reject) the null hypothesis, respectively (e.g., Berry et al., 2001; Lawrence et al., 2011; Laskowski et al., 2013). Applying a two-sided KS test to our moderate- $n$  Orinoco data (Fig. 2A) indicates, at 95% confidence, that only about half of the possible sample pairs fail to reject the null hypothesis. This is counter to the geological observations conditioning this experiment, as failure to reject the null hypothesis is the *a priori* expectation based on the knowledge that all analyzed sediments were collected from the same channel of this active drainage basin (Fig. 1; see also Bonich et al., 2017).

Recent studies have suggested that the KS and Kuiper tests are poorly suited as quantitative descriptors of dissimilarity at low observation numbers (i.e.,  $<1000$  for KS and  $<475$  for Kuiper), and instead have suggested that cross-correlation coefficients ( $R^2$  values) of probability density functions (PDFs) might be better suited for this purpose at moderate- $n$  values, even if they are not suitable for hypothesis testing (Saylor and Sundell, 2016). Application of the PDF cross-correlation test to our moderate- $n$  data set also reveals, as the KS test did, the differences amongst the measured spectra, as 86% of all possible intersample pairs ( $n = 24$ ) yield  $R^2$  values  $<0.60$ , and 36% of pairs ( $n = 10$ ) yield  $R^2$  values  $<0.40$  (Fig. 2A); nevertheless, due to the low sampling numbers, these coefficients from PDF cross-correlations cannot reliably discriminate identical from different populations (Saylor and Sundell, 2016).

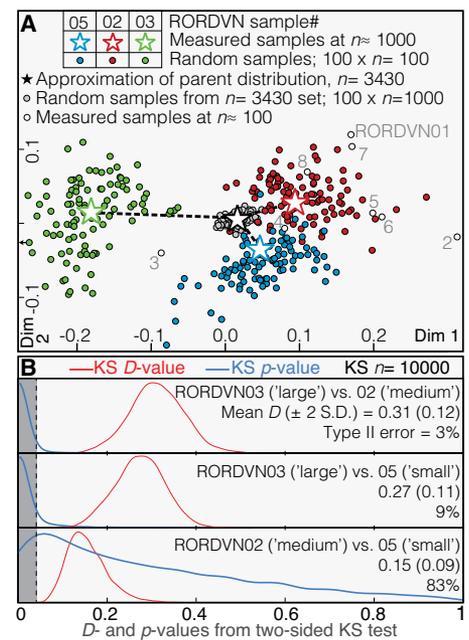
The statistical tests described above using moderate- $n$  results have inherently low precision (due to random sampling; Andersen, 2005), but application to samples with a larger number ( $n \approx 1000$ ) of single-crystal U-Pb observations can be considered statistically accurate (Saylor and Sundell, 2016). Application of the KS test to the large- $n$  data set yields large dissimilarities and indicates that the null hypothesis can in all cases be rejected with  $>99.99\%$  confidence (i.e.,  $p$ -values  $<1 \times 10^{-4}$ ). On the other hand, PDF cross-correlation coefficients are 0.30 and 0.38 for comparison of the large-grained sample with the small- and medium-grained samples, respectively, whereas comparison of the medium- and fine-grained samples yields a coefficient of  $\sim 0.76$  (Fig. 2C). Low  $R^2$  values of 0.30 and 0.38 indicate that, statistically speaking, intersample comparisons between the large- and the small- and medium-grained samples both confidently allow rejecting the (null) hypothesis that these were derived from the same parent distribution. The  $R^2$  of  $\sim 0.76$ , on the other hand, would be indicative of mixing from different sources, but cannot confidently be used to reject or fail to

reject the hypothesis that these sediments were derived from the same source or parent distribution (see Saylor and Sundell, 2016, their figure 12, for interpretation).

Although the tests described above—particularly for our large- $n$  data set—can be considered statistically accurate, they are unlikely to result in geologically accurate conclusions; many researchers would interpret these statistical observations to indicate derivation from different sources or “parent” distributions.  $D$ -values in the KS and Kuiper tests and  $R^2$  values in PDF cross-correlations are strongly controlled by the presence or absence and the abundance of particular age modes that compose the spectra, both of which can be strongly controlled by natural sorting processes and are thus not necessarily indicative of different provenance, source areas, or tectonic phenomena (e.g., Fig. 2C). The calculated dissimilarities amongst samples as described above are arguably *not* the result of derivation from different source areas (i.e., we know they derive from the same modern drainage basin), but instead arise from the systematic sorting and biasing experienced by detrital zircon crystals during transport. We conclude, then, that the critical disadvantage that any generalized dissimilarity approach has for comparing detrital mineral geochronologic data lies in that such tests remain ignorant to, and uninformed by, the underlying geology.

To further illustrate the limitations that traditional  $n \approx 100$  geochronologic data sets have at accurately representing (graphically and quantitatively) detrital zircon provenance, we performed a sensitivity test using random sampling methods. Starting with the measured large- $n$  data sets for each of our three measured samples (Fig. 2C), 100 subsamples of  $n = 100$  dates were drawn from each of these “parent” distributions using bootstrap without substitution. Shown in multidimensional scaling space using KS  $D$ -values as dissimilarity estimator (KS-MDS; see Vermeesch, 2018), these random subsamples plot as “clouds” around their respective  $n \approx 1000$  “parents” (Fig. 3A). The “parent” age distribution for our large-grained sample is so dissimilar to that of the medium- and small-grained samples that their mutual random subsamples do not overlap in KS-MDS space.

On the other hand, random subsamples of the medium- and fine-grained samples show a greater degree of overlap; this implies that there is greater probability that random combinations of  $n \approx 100$  subsamples from samples RORDVN05 and RORDVN02 could yield KS distances that fail to reject the null hypothesis and thus result in statistical type II errors (or false positives). In other words, even though we know from the KS test using  $n \approx 1000$  data that the sample distributions of samples RORDVN05 and RORDVN02 are statistically distinct (Fig. 2C), if different researchers take samples from the exact same



**Figure 3. A:** Multidimensional scaling plot using Kolmogorov-Smirnov (KS)  $D$ -values as dissimilarity estimator (KS-MDS) including: (1) eight samples analyzed using  $n \approx 100$  single detrital zircon crystals (open black circles); (2) three samples analyzed at  $n \approx 1000$  (colored stars); (3) 100 random (bootstrap) subsamples of  $n = 100$  drawn from each large- $n$  data set (color-filled circles); (4) reconstructed approximation of Rio Orinoco delta (Venezuela) parent distribution, computed by adding all 3430 zircons analyzed in this study (black star); and (5) 100 random subsamples at  $n = 1000$  (gray circles), shown only to highlight decrease in KS-MDS scatter resulting from use of large- $n$  relative to the moderate- $n$  approach. Axes Dim 1 and Dim 2 are two dimensions onto which sample dissimilarities are projected, normalized to the sum of the squared Euclidean distances within the MDS configuration (see Vermeesch, 2018). **B:** Distribution of KS  $D$ - and  $p$ -values obtained from all possible combinations of random  $n = 100$  subsamples simulated from large- $n$  data; “large,” “medium,” and “small” refer to the samples’ relative mean zircon grain size as shown in Figure 2B. Gray-shaded region shows range of  $p$ -values  $<0.05$ , which indicate rejection of the null hypothesis to 95% or greater confidence. Area under blue curve outside gray region indicates probability of random  $n = 100$  pairs resulting in type II statistical errors (false positives). S.D.—standard deviation.

locations and analyze equivalent samples at  $n \approx 100$ , there is a large probability that they would arrive at geologically contradictory conclusions. To illustrate this, Figure 3B shows the distribution of  $D$ - (red curves) and  $p$ -values (blue curves) that would result from comparison of all random subsamples amongst our three large- $n$  data sets, plotted as kernel density estimates (KDEs) of all possible combinations of two-tailed KS tests amongst subsamples of two given large- $n$  samples. Whereas comparison amongst the  $n = 100$  random subsamples of sample RORDVN03 with respect to RORDVN02 and RORDVN05 results

in low probabilities of type II errors (i.e.,  $p$ -values only fail to reject the null hypothesis in 3%–9% of the cases, at 95% confidence), subsample comparisons amongst samples RORDVN02 and RORDVN05 would result in type II errors in 83% of the instances (at 95% confidence).

The experiments shown here pose the uncomfortable problem: If the questions we formulate in such a simple experiment from an active fluvial system fail to be accurately answered using current data evaluation tools, what confidence do we have in our quantitative detrital zircon provenance interpretations from the ancient geological record? It is becoming increasingly clear that in order to resolve this, simply dating more crystals per sample is not the only answer. Although larger values of  $n$  are evidently fundamental for improving representativity and statistical accuracy (Pullen et al., 2014; Zhang et al., 2016; Nie et al., 2018), continuing to ignore statistically significant sources of natural systematic bias will only result in formulation of systematically misleading geological hypotheses. Thus, in order to improve the quantitative use and application of geochronologic-based detrital mineral provenance approaches, we must better our understanding of how zircon crystals pass through, and fractionate from, sedimentary systems (Hietpas et al., 2011), and use that understanding to inform our statistical treatments accordingly. This is particularly critical as the number of publications including a component of detrital zircon provenance continues to grow, because this is inevitably driving the community toward a “big data” approach that will (by necessity) continue to increase its reliance on robust statistical treatments.

Based on the findings of this work, we make the following observations:

(1) Direct application of dissimilarity metrics (e.g., KS test, PDF cross-correlation, etc.) is a poorly suited approach for quantitatively interpreting detrital zircon data.

(2) Future detrital zircon U-Pb investigations will continue to benefit from the utilization of large- $n$  approaches.

(3) Regardless of  $n$ , further documenting (and quantifying) morphological information for crystals dated in detrital zircon studies will continue to illuminate the systematic relationships between physico-mechanical and age biasing that are prevalent in different sedimentary environments (e.g., Finzel, 2017; Markwitz et al., 2017).

(4) If the desired goal is to comprehensively describe the detrital mineral age spectrum available to a sedimentary system, as opposed to using a single aliquot for a given sedimentary bed, sampling and analysis of multiple samples with vastly different grain sizes within that system may constitute a better strategy (e.g., Fig. 3). We infer that, in most scenarios, this approach will better approximate the complete zircon age

distribution available to the system compared to a single-sample large- $n$  approach, thus allowing for more robust interpretations of sediment provenance and ensuing geological processes.

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