measures (35-38). Therefore, in a secondary analysis of an existing data set, we investigated the psychometric and test characteristics of three measures of breast cancer risk perception in a random sample of women from a primary care experiment. population. In addition to overall performance, we explicitly considered how the different items perform when identifying groups of particular interest for health behavior research and interventions, such as individuals who perceive themselves at very high or very low risk. Items were compared using 8

Measuring Perceptions of Breast Cancer Risk

Andrea Gurmankin Levy, 1,2 Judy Shea, 3,4 Sankey V. Williams, 3,4 Alex Quistberg,³ and Katrina Armstrong^{3,4,5,6}

¹Center for Community Based Research, Dana-Farber Cancer Institute; ²Department of Society, Human Development and Health, Harvard School of Public Health, Boston, Massachusetts; 'Department of Medicine, 'Leonard Davis Institute of Health Economics, University of Pennsylvania; and 'Center for Clinical Epidemiology and Biostatistics, 'Abramson Cancer Center, University of Pennsylvania School of Medicine, Philadelphia, Pennsylvania

Abstract

Background: Accurate measurement of people's risk perceptions is important for numerous bodies of research and in clinical practice, but there is no consensus about the best measure.

Objective: This study evaluated three measures of women's breast cancer risk perception by assessing their psychometric and test characteristics.

Design: A cross-sectional mailed survey to women from a primary care population asked participants to rate their chance of developing breast cancer in their lifetime on a 0% to 100% numerical scale and a verbal scale with five descriptive categories, and to compare their risk to others (seven categories). Six hundred three of 956 women returned the survey (63.1%), and we analyzed surveys from the 566 women without a self-reported personal history of breast or ovarian cancer.

Results: Scores on the numeric, verbal, and comparative measures were correlated with each other (r > 0.50), worry

(r > 0.51), the Gail estimate (r > 0.26), and family history (r > 0.25). The numerical scale had the strongest correlation with annual mammogram (r = 0.19), and its correlation with the Gail estimate was unassociated with participants' sociodemographics. The numerical and comparative measures had the highest sensitivity (0.89-0.90) and specificity (0.99) for identifying women with very high risk perception. The numerical and comparative scale also did well in identifying women with very low risk perception, although the numerical scale had the highest specificity (0.96), whereas the comparative scale had the highest sensitivity

Conclusion: Different measures of women's perceptions about breast cancer risk have different strengths and weaknesses. Although the numerical measure did best overall, the optimal measure depends on the goals of the measure (i.e., avoidance of false positives or false negatives). (Cancer Epidemiol Biomarkers Prev 2006;15(10):1893–8)

The number of different risk perception measures used by different investigators highlights the lack of consensus regarding the best risk perception measure. This lack of consensus arises, in part, from how little is known about the psychometric properties and relative merits of the various measures (35-38). Therefore, in a secondary analysis of an

Introduction

Research in fields as varied as medicine, psychology, and marketing involve assessment of risk perceptions: beliefs about the likelihood of experiencing various adverse outcomes. For instance, many studies have assessed risk perception to test models that posit an association between risk perception and health behavior (1-14). Other research assesses risk perception to detect errors and biases in risk judgments (15-18), to assess the association between risk judgments and emotion (19-25), and to increase the accuracy of people's risk perceptions (11, 26-28). In sum, several important and large bodies of research depend on obtaining accurate assessments of people's risk perception.

However, there is little consistency in the approach used to measure risk perception. The most common measures include a 0% to 100% numerical measure, a verbal measure such as "not at all likely" to "extremely likely" or "very low" to "very high," and a comparative measure for which respondents compare their risk to that of the average person using a scale such as "much lower than average" to "much higher than average" scale (29-34). Still other measures include "1 in x" scales (e.g., 1 in 200; ref. 35) or verbal response scales with different labels for each step on the scale (35) and different numbers of steps on the scale (36).

Materials and Methods

The study protocol was approved by the Institutional Review Board at the University of Pennsylvania.

traditional measurement theory on score distributions and

construct validity (39). In addition, because there is no gold

standard for perceived risk, we used latent class models to

estimate the test characteristics (specificity and sensitivity) of

each item for identifying women who perceive themselves at

high risk and women who perceive themselves at low risk.

Data Collection. A random sample of 1,200 adult women who had been seen by a University of Pennsylvania Health System primary care provider between 1996 and 1999 was identified through a billing database managed by the University of Pennsylvania Office of the Associate Dean for Health Services Research. From December 1999 to August 2003, surveys were mailed to the 1,016 of these patients who were not excluded by their primary care provider on the basis

Received 6/27/05; revised 7/20/06; accepted 8/8/06.

Grant support: American Cancer Society Research Training Grant and Robert Wood Johnson Generalist Faculty Scholar Award (K. Armstrong).

The costs of publication of this article were defrayed in part by the payment of page charges. This article must therefore be hereby marked advertisement in accordance with 18 U.S.C. Section 1734 solely to indicate this fact.

Requests for reprints: Andrea Gurmankin Levy, Dana-Farber Cancer Institute, 44 Binney Street, Smith 253, Boston, MA 02115. Phone: 617 582 7942; Fax: 617 632 5690. E-mail: andrea_gurmankin@dfci.harvard.edu

Copyright © 2006 American Association for Cancer Research. doi:10.1158/1055-9965.EPI-05-0482

of being deceased, non-English speaking, too ill, or male. Of these 1,016 patients, 47 surveys were undeliverable, and an additional 5 were deceased and 8 were too sick to participate. Of the remaining 956, 603 returned their survey (63.1%). Of the 603, the 566 without a self-reported personal history of breast or ovarian cancer were eligible for the study.

Measures

Risk Perception. Risk perception was assessed using the following measures: (a) a numerical measure ["What do you think your chance is of developing breast cancer in your lifetime? Please choose a number between 0% (no chance of breast cancer) and 100% (definitely will get breast cancer)"], (b) a verbal measure ("How would you rate your chance of developing breast cancer? Please check very low, moderately low, neither high nor low, moderately high or very high"), and (c) a comparative measure ("Overall, how do you think your chance of developing breast cancer compares to the average woman your age?" 1, much lower; 4, about the same; 7, much higher).

Breast Cancer Worry. Breast cancer worry was assessed with a two-item measure that includes an item assessing frequency and effect on daily life ("How often do you worry about developing breast cancer?" and "How much does worrying about developing breast cancer interfere with your everyday life?" both on a scale from 1, not at all, to 7, all the time). This measure has been previously validated (40).

Mammography Adherence. Mammography adherence was assessed by asking whether the participant had ever had a screening mammogram and if so, the month and year of the last screening mammogram following the procedures used in the Behavioral Risk Factor Surveillance System (41). All women 50 years of age and older who reported having a mammogram within the 12 months before receiving the completed questionnaire were coded as adherent to current recommendations for mammography. Women who provided only the year of the past mammogram were assumed to have undergone screening in the middle of that year (i.e., June). Analyzing mammogram data with and without this assumption yielded the same pattern of results.

Breast Cancer Risk Factors. Lifetime breast cancer risk was calculated using the Gail model (42). The risk factors included in the Gail model are age, degree of family history, age at menarche, age at first live birth and history of breast biopsy. Degree of family history was categorized by $(a) \ge 1$ first- and second-degree relatives with breast or ovarian cancer; $(b) \ge 1$ first-degree relatives, no second-degree relatives; $(c) \ge 1$ second-degree relatives, no first-degree relatives; and (d) no first- or second-degree relatives. Age at first live birth was categorized as in the Gail model: none, <20 years, 20 to 24 years, 25 to 30 years, and >30 years.

Sociodemographic Characteristics. Age, educational attainment, and household income were measured using items from the Behavioral Risk Factor Surveillance System 1998 questionnaire (41). Race and ethnicity were collapsed into Black, White, and other.

Statistical Analyses. The data were analyzed using STATA 7.0 and additional software that estimates the error rates of diagnostic tests or measurements when there is no gold standard by applying maximum likelihood estimation methods to latent class models representing the observed data.⁷ Descriptive analyses were used to examine the characteristics of the study sample. The normality of the distribution of risk-

perception scores for each measure was assessed using the Shapiro-Wilks test.

We assessed the validity of each measure in several ways. Although the nomenclature for validity assessment is inconsistent, we used the traditional categories of construct validity, focusing on convergent validity (the degree to which each measure correlates with other measures of the same construct), discriminant validity (the degree to which each measure does not correlate with measures of different constructs), and predictive validity (the relationship between the measure and the criterion it is supposed to predict; ref. 39). Convergent validity was assessed by examining the correlation of each risk perception measure with the other risk perception measures. Discriminant validity was assessed by examining the correlation of each risk perception measure with a measure of breast cancer worry. Predictive validity was assessed by examining the correlation between each risk perception measure and measures of actual risk and adherence to screening mammography. The extent of family history was included in addition to the Gail estimate because, although the former is included in the latter, patients are aware of their family history but generally unaware of their Gail estimate. Analyses of annual mammography included only those respondents ≥50 years (n = 318). We also assessed the degree to which the correlation with the primary measure of absolute risk varied by race (Black versus White), household income (>\$50,000 versus \leq \$50,000), age (<50 versus \geq 50 years), and education (completed college or more versus some college or less). This was tested by examining the interaction term of the risk perception measure and each sociodemographic variable when predicting the Gail estimate.

Because we had no gold standard of very low or high risk perception, we used latent class analysis to assess the sensitivity and specificity of each measure as described by Walter and Irwig (43). In this method, all of the measures are assumed to be subject to error (i.e., misclassification) that is independent across measures. Initial estimates of the "true" classification of each participant (e.g., very high risk perception or not) are revised iteratively using maximum likelihood estimation until convergence occurs. The estimated variables are the false-positive and false-negative rates of each item and the prevalence of the outcome in the population. This methodology has been previously applied to measurement issues in epidemiologic studies (43, 44). To conduct these analyses, each risk perception measure was categorized into five categories with four cutoff points. For the comparative measure, the seven response options were collapsed into five categories: 1 and 2, 3, 4, 5, and 6 and 7. The five categories of the verbal measure remained the same. We categorized the numerical measure into five groups corresponding to the average lifetime risk of breast cancer: 0% to 5%, 6% to 10%, 11% to 15%, 16% to 50%, and >50%. Because there is debate about whether women are able to use a numerical scale as corresponding to the predicted risk of breast cancer, we examined three other categorizations of the 0% to 100% scale; the first using categories of 0% to 19%, 20% to 39%, 40 to 59%, 60 to 79%, 80 to 100%; the second using categories of 0 to 5%, 6% to 10%, 11% to 15%, 16% to 20%, and 21% to 100%; and the third using the original categories but omitting the women who chose 50% as this choice may reflect uncertainty rather than a specific level of risk perception (45).

We calculated sensitivity and specificity for two outcomes: very high risk perception and very low risk perception. In our main approach (approach A), very high risk perception was defined as responses above the highest cutoff point on each measure (>50% for numeric, very high for verbal, >5 on the seven-point comparative measure) and very low risk perception was defined as responses below the lowest cutoff point on each scale (<6% for numeric, very low for verbal, and <3 on the 7-point comparative measure). However, multiple alternative

 $^{^7 \, \}mathrm{S}.$ Williams, personal communication, with data transfer via 3.5'' disk containing programs and examples.

approaches were used to define very high risk perception for the numerical measure to determine the effect of these alternative definitions on the results. In approach B, we used 50% as the cutoff and excluded participants with a 50% perceived risk on the numerical measure for predicting high risk because of the known tendency for people to use 50% to mean "I don't know" rather than a 50/50 chance (45, 46). In approach C, we used a cutoff of 20% for the numerical measure without this additional exclusion, and in approach D, we used a cutoff of 20% for the numerical measure and again excluded those with a 50% perceived risk on the numerical measure.

Results

Respondents. The study population had a mean age of 52 years (with a range from 20 to 82 and 50% being over 50 years of age). Sixty-six percent of study respondents were White. Respondents encompassed a wide range of educational attainment and household income. Respondent characteristics and Gail model breast cancer risk factors are shown in Table 1. On average, respondents had a Gail lifetime risk of 8% with a range of 0.8% to 29%.

Distribution of Scores. As shown in Table 2, the Shapiro-Wilks tests indicate that none of the risk perception measures are normally distributed.

Construct Validity. Table 2 reports our assessments of convergent, discriminant, and predictive validity. Each risk

Table 1. Respondents' demographic characteristics and breast cancer risk factors

	(n = 566)
Age, y (%)	
18-30	7
31-40	17
41-50	25
51-60	20
>60	30
Race/ethnicity (%)	
Black	34
White	66
Education (%)	
High school or less	29
Some college	29
College or higher	42
Annual household income (%)	12
≤\$30,000	22
\$30,001-\$50,000	22
\$50,001-\$70,000	20
>\$70,000	36
Family history of breast or ovarian cancer (%)	50
None	66
≥1 second-degree relatives	00
No first-degree relatives	21
≥1 first-degree relatives	21
No second-degree relatives	11
≥1 first- and second-degree relatives	3
	3
Age first live birth, y (%) No live births	22
<20	10
20-24	19
25-30	22
≥30	26
	20
History of breast biopsy (%) Yes	22
No	78
- 10	76
Age first period, y (%)	10
9-11 12-13	18 55
	27
≥ 14	27
Gail estimate (%)	0.4
Average or below ($\leq 12\%$)	84
Above average (>12%)	16

perception measure was significantly associated with each of the other measures ($r \ge 0.60$ for all) and slightly less correlated with a measure of breast cancer worry ($r \ge 0.51$ for all). Each risk perception measure was also significantly correlated with the Gail estimate ($r \ge 0.26$ for all) and degree of family history ($r \ge 0.25$ for all). However, only the numerical measure was significantly correlated with adherence to annual mammography (r = 0.19). The verbal and comparative measures were only marginally correlated with mammography adherence.

The correlation between the risk perception measures and Gail estimate was significantly stronger for Whites than Blacks for the verbal measure and the comparative measure. The correlation between the Gail estimate and the risk perception measure was not affected by income or education for the numerical measure, but was significantly stronger for those with higher income (>\$50,000) and who completed college or more for the comparative measure. The correlation between Gail risk and the verbal measure was marginally stronger among those with higher income (>\$50,000). The associations with Gail risk did not differ by age (being ≥50 versus <50 years) for any of the three risk perception measures.

Sensitivity and Specificity. The test characteristics of each $\frac{8}{8}$ measure for identifying women with very high risk perception and women with very low risk perception are reported in § Table 3.

The numerical and comparative measures had the highest sensitivity (0.89-0.90) and the verbal measure had the lowest sensitivity (0.37) for identifying women with very high risk perception. Similarly, the numerical and comparative measures exhibited the highest specificity (0.99) for excluding women who did not have very high risk perception. The verbal measure had the lowest specificity for excluding women who did not have very high risk perception, although it was still high (0.93). Overall, for identifying women with very high risk perception, the numerical and comparative measures had higher sensitivity and specificity than the verbal measure. Using these cutoffs, the $\frac{1}{6}$ latent class models estimated that 5% of respondents had a very high risk perception and 15% had a very low risk perception.

Alternative approaches for categorizing the numerical scale and some effect on the relative sensitivity and specificities of the different measures for identifying women with high risk perception (Table 3). In general, the numerical measure retained high sensitivity and specificity unless the threshold for high risk perception on the numerical measure was set at 20% (approach C). The verbal measure had lower sensitivity but relatively high 💆 specificity across the approaches tested. The comparative § measure had relatively high sensitivity and specificity, with levels slightly above that of the numerical measure when a 🤄 cutoff of 20% for high risk was used on the numerical scale 8 (approach C) or when women who responded 50% as their risk on the numerical scale were omitted from the analysis.

For identifying women with very low risk perception, the numerical measure had the lowest sensitivity (0.74) and highest specificity (0.96), whereas the comparative measure had the highest sensitivity (0.89) and the lowest specificity (0.91). The verbal measure had an intermediate sensitivity at 0.81 and a relatively high specificity at 0.95.

Discussion

The results of this study offer some information to guide the decision of how to best measure cancer risk perception. Each of the three measures of risk perception that we included in this study had significant strengths and weaknesses. The comparative measure of perceived risk, which asked women to compare their risk to the average woman on a seven-point scale, had a nonnormal distribution, was strongly correlated with the other measures, and moderately correlated with measures of actual risk. However, the comparative measure of

Table 2. Normality and convergent, discriminant and predictive validity of the risk-perception measures (n = 566)

Criteria	Risk-perception measure						
	Numerical (0-100)	Verbal (1-5)	Comparative (1-7)				
Normality							
Shapiro Wilks test (W)*	0.976	0.992	0.981				
1 ' '	P < 0.0001	P = 0.006	P < 0.0001				
Convergent validity (correlation with other risk-perception measure	res)						
Verbal	0.62		_				
	P < 0.0001	_					
Comparative	0.60	0.72					
ī	P < 0.0001	P < 0.0001	_				
Discriminant and predictive validity							
Correlation with breast cancer worry	0.51	0.51	0.52				
, the state of the	P < 0.0001	P < 0.0001	P < 0.0001				
Correlation with measures of actual risk Gail model estimate	0.26	0.35	0.33				
	P < 0.0001	P < 0.0001	P < 0.0001				
Degree of family history	0.28	0.27	0.25				
, ,	P < 0.0001	P < 0.0001	P < 0.0001				
Correlation with breast cancer screening annual mammogram	0.19	0.12	0.10				
	P = 0.0019	P = 0.053	P = 0.110				
Association with Gail risk affected by: Race (Black vs White)	Not significant	White > Black	White > Black				
	P = 0.103	P = 0.034	P = 0.014				
Income $(>$50,000 \text{ vs} \le $50,000)^{\top}$	Not significant	Not significant	Higher income > lower income				
	P = 0.444	P = 0.061	P = 0.032				
Age $(<50 \text{ vs } \ge 50 \text{ y})^{\top}$	Not significant	Not significant	Not significant				
	P = 0.887	P = 0.555	P = 0.823				
Education (completed college or more vs less than college)	Not significant	Not significant	More education > less education				
	P = 0.525	P = 0.175	P = 0.021				

^{*}Value of 1 indicates complete normality; significant P value indicates nonnormal.

perceived risk was less correlated with adherence to annual mammography and seemed to perform differently among women with different levels of income, educational attainment, and racial groups. The verbal measure of perceived risk, where a woman rated her risk on five-point scale, also had a nonnormal distribution and was strongly correlated with the other measures and moderately correlated with measures of actual risk. However, the verbal measure of perceived risk had a correlation with mammography adherence of borderline statistical significance and its performance was not consistent across all sociodemographic characteristics. The numerical measure of perceived risk also had a nonnormal distribution, strong correlations with the other measures, and moderate correlations with measures of actual risk. In addition, it was significantly correlated with adherence to annual mammography and there was no evidence that its performance differed between women of different socioeconomic groups. Thus, the numerical measure had the best overall performance across these criteria and could be considered the default option for most studies.

The high correlations between the risk perception measures and the convergent validity in their positive associations with

other related variables (e.g., breast cancer worry) also suggest that the three risk perception measures represent similar constructs. Although these high correlations suggest a single underlying construct, it is possible that the differential association of each risk perception measure with other variables indicates that each of the measures may represent a different aspect of risk perception, which may have implications for health behavior. There is a growing recognition that risk perception is a more complex construct than pure expected probability and is likely to include dimensions of affective response (47), as is suggested by the very strong association between the risk perception measures and breast cancer worry. Further research is needed to tease apart how these measures may relate to this more complex model of risk perception.

Although the criteria in Table 2 provide some guidance as to the overall performance of the measures, the calculations of sensitivity and specificity can inform the decision about which measure (and which cutoff) to use when there is a specific need to accurately identify individuals with very high or low risk perception. In these analyses, both the numerical measure and the comparative measure did the best—whether in identifying

Table 3. Sensitivity and specificity of risk-perception measures (n = 566)

	Approach A			1	Approach B		Approach C			Approach D		
	N	V	С	N	V	С	N	V	С	N	V	С
Very high sensitivity Very high specificity Very low sensitivity Very low specificity	0.89 0.99 0.74 0.96	0.37 0.93 0.81 0.95	0.89 0.99 0.74 0.96	0.91 0.99	0.75 0.98	0.86 0.96	0.89 0.36	0.31 1.00	1.00 1.00	0.89 0.90	0.37 1.00	0.90 0.99

NOTE: Approach A: Using cutoff of 50% for numeric, very high for verbal, 6 on seven-point comparative measure (7—much higher than average) for predicting high risk, or 5% for numeric, very low for verbal, 2 on seven-point comparative measure (1—much lower than average) for predicting low risk. Approach B: Using cutoff of 50% numeric, very high for verbal, 6 on seven-point comparative measure (7—much higher than average) for predicting high risk, and dropping participants with a 50% perceived risk on the numerical scale for predicting high risk. Approach C: Using cutoff of 20% for numeric, very high for verbal, 6 on seven-point comparative measure (7—much higher than average) for predicting high risk. Approach D: Using cutoff of 20% for numeric, very high for verbal, 6 on seven-point comparative measure (7—much higher than average) and dropping participants with a 50% perceived risk on the numerical scale for predicting high risk. Abbreviations: N, numeric; V, verbal; C, comparative.

[†]P value for interaction term

women who perceive themselves at very high risk or in identifying women who perceive themselves at very low risk. The test characteristics of the numerical and comparative measures were very similar for identifying women with very high risk perception; both had very low rates of false positives (1%) and reasonably low rates of false negatives (\sim 10%). It is important to note that we calculated the sensitivity and specificity of the measures using different cutoffs for the numerical measure; however, other cutoffs for very high risk or very low risk may be of interest and affect the results. As expected, as the threshold for calling a risk perception "very high" decreased, the specificity of the numerical measure decreased. Furthermore, the inclusion or removal of the 50% response on the numerical scale did not substantively affect our results, which suggests that despite known problems with this response (45, 46), individuals reporting a risk perception of 50% do not need to be excluded for the numerical scale to have reasonable test characteristics.

The results of these analyses suggest that for the identification of women with very low risk perception, the comparative measure may be preferred if the goal is to minimize the rate of false negatives (classifying a woman who has very low risk perception as having average or high risk perception). This may be the goal in many settings, such as interventions to encourage cancer risk reduction behaviors among women at risk of not participating because of low perceived risk of cancer. However, if the goal is to maximize the specificity of the measure and minimize the number of false positives (classifying a woman with average or high risk perception as having low risk perception), the numerical measure would be best. Although this scenario seems less likely, it may apply to settings such as clinical trials of interventions targeted at low risk women where eligibility is dependent upon having a low

The results of this study contribute to an ongoing discussion about the best approach to measuring risk perception. One previous study compared a comparative measure to a numerical measure and found that the comparative measure did better on various psychometric measures and subject ratings (37). Another study that compared verbal measures to numerical measures concluded that the verbal measures were subjectively better. Specifically, Diefenbach et al. (36) found that subjects rated a verbal measure as easier to use and as a better reflection of how they felt compared with a dichotomous measure and multiple numerical measures. In addition, people tend to have difficulty with numerical probabilities ("innumeracy"; ref. 48) and report a preference for expressing risk using words, not numbers (49, 50). On the other hand, there is substantial variability in how people interpret the verbal expressions of probability used to label the verbal scales, raising questions about the reliability of these measures and making it very difficult to translate verbal expressions into quantitative metrics (49, 51, 52).

Thus, to date, much debate and uncertainty has surrounded the choice of risk perception measures for clinical and research use. This study applies a clinical diagnostic tool to examine the strengths and weaknesses of three different risk perception measures. The results not only contribute to what is known about the strengths and weaknesses of each measure, but also extend this line of work by specifying which measures are best under which circumstances. As noted, previous work has generated mixed recommendations regarding the best approach to risk perception measurement. The approach used in this study suggest that despite the problems revealed with it in other work, the numerical measure performs well (if not best) by most criteria, but that the optimal measure also depends on the relative costs of false negatives and false positives in risk perception measurement.

The results of this study must be considered within its limitations. The order of the risk perception measures was not

randomized, and a previous study found that risk perception measure order affects responses to these measures (53). Because that study found that responses to measures that follow the comparative measure are lower than when they do not follow the comparative measure, responses to our verbal measure (which preceded the comparative measure) may have been higher than they would have been had the verbal measure followed the comparative measure. However, the size of this effect is small and it is unlikely to have substantively affected the overall ranking of the measures. A second limitation is that, as with any survey research, there is the potential for response bias. Although the response rate of 63% is respectable, it is possible that nonresponders differed from responders. However, we do not have information on the nonresponders to enable us to examine this possibility. Another concern is that we did not specifically test whether numeracy or literacy affected the performance of the measures, although we did examine the effect of education. Thus, it is possible that one of the measures performs less well among low numeracy or literacy groups. Finally, we did not assess the performance of multiple items compared with single items. Future work is needed to determine whether multiple item \(\frac{1}{2} \) measures of risk perception are preferable to the single items $\frac{1}{2}$ considered here.

Nevertheless, the results of this study provide new information about the strengths and weaknesses of multiple breast cancer risk perception measures, including information about their relative strengths and weaknesses depending on the goals or circumstances of the inquiry. The numerical measure did well overall, but other measures may be preferred. measure did well overall, but other measures may be preferred depending on the study question and the relative importance of avoiding false positives versus false negatives. The recommendations from this study for cancer risk perception measures can increase the validity of risk perception measures used in clinical practice and research. In addition, this used in clinical practice and research. In addition, this information can contribute to greater consensus about risk-perception measurement and, therefore, greater consistency in measurement across studies.

Acknowledgments

We thank Barbara Weber, Jill Stopfer, Susan Domchek, Ellyn Micco, Amy Carney, and the women who participated in the study.

References

References

References

**Text No. 11 to 11

- Kash KM, Holland JC, Halper MS, Miller DG. Psychological distress and surveillance behaviors of women with a family history of breast cancer. J Natl Cancer Inst 1992;84:24-30.
- J Natl Cancer Inst 1992;84:24–50.
 King E, Rimer BK, Balshem A, Ross E, Seay J. Mammography-related beliefs of older women. J Aging Health 1993;5:82-100.
- Harris RP, Fletcher SW, Gonzalez JJ, et al. Mammography and age: are we targeting the wrong women? A community survey of women and physicians. Cancer 1991;67:2010–4.
- Lerman C, Rimer BK, Trick B, Balshem A, Engstrom PF. Factors associated with repeat adherence to breast cancer screening. Prev Med 1990;19:279 - 90.
- McCaul KD, Branstetter AD, Schroeder DM, Glasgow RE. What is the relationship between breast cancer risk and mammography screening? A meta-analytic review. Health Psychol 1996;15:423-9.
- Aiken LS, West SG, Woodward CK, Reno RR. Health beliefs and compliance with mammography screening recommendations in asymptomatic women. Health Psychol 1994;13:122-9.
- Cummings KM, Hellmann R, Emont SL. Correlates of participation in a worksite stop-smoking contest. J Behav Med 1988;11:267-77
- Klesges RC, Brown K, Pascale RW, Murphy M, Williams E, Cigrang JA. Factors associated with participation, attrition and outcome in a smoking cessation program at the workplace. Health Psychol 1988;7:575-89.
- Kviz FJ, Crittenden KS, Belzer LJ, Warnecke RB. Psychosocial factors and enrollment in a televised smoking cessation program. Health Educ Q 1991; 18:445-61.
- Boney-McCoy S, Gibbons FX, Reis TJ, Gerrard M, Luus CA, Sufka AV. Perceptions of smoking risk as a function of smoking status. J Behav Med 1992;15:469-88.
- 11. Weinstein ND, Sandman P, Roberts N. Perceived susceptibility and self protective behavior: a field experiment to encourage home radon testing. Health Psychol 1991;10:25-33.

- Price J. Perceptions of colorectal cancer in a socioeconomically disadvantaged population. J Community Health 1993;18:347–62.
- Weinstein N, Sandman PM, Roberts NE. Determinants of self-protective behavior: home radon testing. J of Applied Social Psychology 1990;20: 783–801.
- Orbell S, Crombie I, Johnston G. Social cognition and social structure in the prediction of cervical screening uptake. Br J Health Psychol 1996;1: 35–50.
- Edwards A, Elwyn G, Covey J, Matthews E, Pill R. Presenting risk information—a review of the effects of "framing" and other manipulations on patient outcomes. J Health Commun 2001;6:61–82.
- Johnson EJ, Hershey J, Meszaros J, Kunreuther H, Framing. probability distortions and insurance decisions. J Risk Uncertain 1993;7:35–51.
- Lichtenstein S, Slovic P, Fischhoff B, Layman M, Combs B. Judged frequency of lethal events. J Exp Psychol Hum Percept Perform 1978;4:551–78.
- Rottenstreich Y, Hsee C. Money, kisses, and electric shocks: on the affective psychology of risk. Psychol Sci 2001;12:185–90.
- Wright W, Bower G. Mood effects on subjective probability assessment. Organ Behav Hum Decis Process 1992;52:276–91.
- Isen A, Patrick R. The effect of positive feelings on risk-taking: when the chips are down. Organ Behav Hum Decis Process 1983;31:194–202.
- Johnson E, Tversky A. Affect, generalization, and the perception of risk. J Pers Soc Psychol 1983;45:20–31.
- Schwarz N, Clore G. Mood, misattribution and judgments of well-being: informative and directive functions of affective states. J Pers Soc Psychol 1983;45:513–23.
- Kavanagh D, Bower G. Mood and self-efficacy: impact of joy and sadness on perceived capabilities. Cognit Ther Res 1985;9:507–25.
- 24. Lerner J, Dacher K. Fear, anger, and risk. J Pers Soc Psychol 2001;8:146–59.
- Lerner J, Gonzalez R, Small D, Fischoff B. Effects of fear and anger on perceived risks of terrorism. A national field experiment. Psychol Sci 2003; 14:144 – 50
- Lerman C, Lustbader E, Rimer B, Daly M, Miller S, Sands C, Balshem A. Effects of individualized breast cancer risk counseling: a randomized trial. J Natl Cancer Inst 1995;87:286–92.
- Lipkus I, Biradavolu M, Fenn K, Keller P, Rimer B. Informing women about their breast cancer risks: truth and consequences. Health Commun 2001;13: 205–26.
- 28. Alexander N, Ross J, Sumner W, Nease R, Littenberg B. The effect of an educational intervention on the perceived risk of breast cancer. J Gen Intern Med 1995:11:92–7.
- Lipkus I, Rimer B, Strigo T. Relationships among objective and subjective risk for breast cancer and mammography stages of change. Cancer Epidemiol Biomarkers Prev 1996;5:1005–11.
- Gurmankin AD, Domchek S, Stopfer J, Fels C, Armstrong K. Patients' resistance to risk information in genetic counseling for BRCA1/2. Arch Intern Med 2005;165:523–9.
- Gurmankin AD, Baron J, Armstrong K. The message sent versus message received in hypothetical physician risk communications: exploring the gap. Risk Anal 2004;24:1337 –47.
- Cohn L, Macfarlane S, Yanez C, et al. Risk perception: differences between adolescents and adults. Health Psychol 1995;14:217 – 22.
- 33. Lerman C, Schwartz MD, Miller SM, Daly M, Sands C, Rimer BK. A

- randomized trial of breast cancer risk counseling: Interacting effects of counseling, educational level, and coping style. Health Psychol 1996;15: 75–83.
- 34. Lipkus IM, Iden D, Terrenoire J, Feaganes JR. Relationships among breast cancer concern, risk perceptions, and interest in genetic testing for breast cancer susceptibility among African-American women with and without a family history of breast cancer. Cancer Epidemiol Biomarkers Prev 1999;8: 533–9
- Woloshin S, Schwartz L, Byram S, Fischhoff B, Welch HG. A new scale for assessing perceptions of chance: a validation study. Med Decis Making 2000;20:298–307.
- Diefenbach M, Weinstein N, O'Reilly J. Scales for assessing perceptions of health hazard susceptibility. Health Educ Res 1993;8:181–92.
- 37. Woloshin S, Schwartz L, Black W, Welch G. Women's perceptions of breast cancer risk: how you ask matters. Med Decis Making 1999;19:221–9.
- Schapira M, Davids SL, McAuliffe TL, Nattinger AB. Agreement between scales in the measurement of breast cancer risk perceptions. Risk Anal 2004; 24:665–73.
- 39. Nunnally JC. Psychometric theory. New York: McGraw-Hill; 1994.
- Lerman C, Kash K, Stefanek M. Younger women at increased risk for breast cancer: perceived risk, psychological well-being, and surveillance behavior. J Natl Cancer Inst 1994;16:171–6.
- HHS. Centers for Disease Control. Washington (District of Columbia): Centers for Disease Control; 2002.
- Gail MH, Brinton LA, Byar DP, et al. Projecting individualized probabilities
 of developing breast cancer for white females who are being examined
 annually. J Natl Cancer Inst 1989;81:1879–86.
- Walter SD, Irwig LM. Estimation of test error rates, disease prevalence and relative risk from misclassified data: a review. J Clin Epidemiol 1988;41: 923–37.
- Walter SD, Frommer DJ, Cook RJ. The estimation of sensitivity and specificity in colorectal cancer screening methods. Cancer Detect Prev 1991;15:465–9.
- **45.** Fischhoff B, DeBruin W. Fifty-fifty = 50%. J Behav Decis Making 1999;12: 149–63.
- **46.** Fischhoff B, DeBruin W, Millstein S, Halpern-Felsher B. Verbal and numerical expressions of probability: "it's a fifty-fifty chance." Organ Behav Hum Decis Process 2000;81:115–31.
- 47. Loewenstein G, Weber E, Hsee C, Welch N. Risk as feelings. Psychol Bull 2001;127:267–86.
- **48.** Lipkus I, Samsa G, Rimer BK. General performance on a numeracy scale among highly educated samples. Med Decis Making 2001;21:37 44.
- Budescu D, Wallsten T. Consistency in interpretation of probabilistic phrases. Organ Behav Hum Decis Process 1985;36:391–405.
- Olson M, Budescu D. Patterns of preference for numerical and verbal probabilities. J Behav Decis Making 1997;10:117–31.
- Beyth-Marom R, How probable is probable? A numerical translation of verbal probability expressions. Int J Forecast 1982;1:257–69.
- verbal probability expressions. Int J Forecast 1982;1:257–69.

 52. Wallsten T, Budescu D, Papoport A, et al. Measuring the vague meanings of
- probability terms. J Exp Psychol Gen 1986;115:348–65.

 Taylor KL, Shelby RA, Schwartz MD, et al. The impact of item order on
- 1aylor KL, Sheiby KA, Schwartz MD, et al. The impact of item order on ratings of cancer risk perception. Cancer Epidemiol Biomarkers Prev 2002; 11:654 – 9.