

Research Article

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Serum Antioxidant Nutrients, Vitamin A, and Mortality in U.S. Adults

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Abstract

Background: Observational studies have suggested that antioxidant nutrients may reduce cancer and overall mortality risks. However, most randomized trials have failed to show survival benefits. Examining nonlinear associations between antioxidant levels and health outcomes may help to explain these discrepant findings.

Methods: We evaluated all-cause, cancer, and cardiovascular mortality risks associated with quintiles (Q1–Q5) of serum antioxidant (vitamins C and E, β -carotene, and selenium) and vitamin A levels, in 16,008 adult participants of The Third National Health and Nutrition Examination Survey (NHANES III; 1988–1994).

Results: Over a median follow-up period of 14.2 years, there were 4,225 deaths, including 891 from cancer and 1,891 from cardiovascular disease. We observed a dose–response decrease in cancer and overall mortality risks with higher vitamin C levels. In contrast, for vitamin A, risk of cancer death decreased from Q1–Q2, with no further decline in risk at higher levels. For vitamin E, having levels in Q4 was associated with the lowest cancer mortality risk. Both vitamin A and E had U-shaped associations with all-cause mortality. Cancer mortality risks decreased from Q1–Q2 for β -carotene and from Q1–Q4 for selenium. However, for β -carotene and selenium, overall mortality risks decreased from Q1–Q2 but then did not change significantly with higher levels.

Conclusions: Antioxidant supplement use should be studied in the context of overall mortality and other competing mortality risks.

Impact: These data suggest the need for novel intervention studies where doses of these agents are individualized based on their serum levels, and possibly, markers of oxidative stress and systemic inflammatory response. *Cancer Epidemiol Biomarkers Prev*; 22(12); 2202–11. ©2013 AACR.

Introduction

Recent data show that between 28% and 30% of U.S. adults use supplements containing vitamins A, C, and E, and 18% to 19% report using selenium (1, 2). Higher serum antioxidant levels have been associated with lower overall mortality risk (3–15), and fruit and vegetable intake (a rich source of antioxidants) has also been shown to predict better health outcomes in observational studies (16–20). In contrast, most randomized controlled trials have either failed to show significant health benefits from taking these supplements or reported possible harm. A recent systematic review of 78 primary and secondary

prevention trials by the Cochrane Collaboration concluded that use of β -carotene, vitamin E, or high doses of vitamin A supplements was associated with increased all-cause mortality, while the role of vitamin C or selenium supplementation was not clear (21). Although the Physicians' Health Study II (22) recently showed a small reduction in total cancer incidence with multivitamin supplement use, most large clinical trials assessing cancer incidence or mortality have failed to show any beneficial effects of taking these supplements (23–29).

Many authors have suggested that the reason for these discrepant findings is that antioxidants have different effects on various disease processes, and that beyond a certain "threshold" level, they can be potentially toxic (30–40). Surprising little literature exists, however, examining nonlinear effects between serum antioxidant nutrient levels and all-cause and cause-specific mortality outcomes. There are also few studies examining competing risks of mortality in the context of antioxidant nutrient use.

Using data from The Third National Health and Nutrition Examination Survey (NHANES; 1988–1994), a large nationally representative cohort of U.S. adults, we examined whether serum levels of micronutrients with antioxidant properties (vitamins C and E, β -carotene, and

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Note: Supplementary data for this article are available at Cancer Epidemiology, Biomarkers & Prevention Online (<http://cebp.aacrjournals.org/>).

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doi: 10.1158/1055-9965.EPI-13-0381

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selenium) and vitamin A, predicted the risks of all-cause, cancer, and cardiovascular disease mortality outcomes. Furthermore, by assessing nonlinear associations, we explored some of the reasons behind the inconsistency in findings between observational studies and randomized trials about the role of these agents on health outcomes.

Materials and Methods

Study population

NHANES III was conducted from October 1988 through October 1994 by the Centers for Disease Control and Prevention (CDC) to provide national health estimates of the United States' civilian population (41). The overall sample size was 39,695; interview and examination response rates were 86% (33,974 participants) and 78% (30,818 participants), respectively (42). Follow-up was from the date of survey participation (1988–1994) to December 31, 2006. Of the 16,573 NHANES III participants (ages 20 years and above) who underwent medical examination, our study included 16,008 (97%) participants for whom data on serum antioxidant nutrient levels and vital status were available. Missing data on covariates ranged from less than 1% to 8% of the study participants. In the analysis, we included only those individuals with complete information available for these covariates. NHANES III was approved by the Institutional Review Board at the CDC. Written informed consent was obtained from all study participants (43).

Measures

Our primary exposures of interest were serum levels of vitamin A, and micronutrients with antioxidant properties. Levels of vitamin A (retinol), vitamin C, vitamin E (α -tocopherol), and β -carotene were measured by isocratic high-performance liquid chromatography (44). Serum selenium levels were measured with atomic absorption spectrometry (44). NHANES III laboratory procedures including quality control systems have been described elsewhere (45).

Our primary outcomes of interest were cause-specific and all-cause mortality. Mortality data were primarily obtained by probabilistic matching to the National Death Index records using the 2010 public release version of the NHANES III Linked Mortality File (46). NHANES III also used various other sources of information to determine the final mortality status and causes of death for survey participants including death certificates, Social Security Administration data, and records from the Centers for Medicare and Medicaid Services. The 9th and 10th revisions of the International Statistical Classification of Diseases (ICD-9 and ICD-10) were used to classify deaths due to cancer (C00–C97) and cardiovascular disease (I00–I78; refs. 46–48). Final mortality status was determined for more than 99% of the study participants (46).

Statistical analysis

To accommodate the complex survey design of NHANES III, we applied appropriate statistical weights

in our analyses (42). Cox Proportional Hazards Models were used to estimate the HRs, and the proportional-hazards assumption was tested using the Kolmogorov-type supremum test (49). To examine whether the associations had a dose–response relationship, we modeled serum levels of antioxidants as quintiles, using the first quintile as the reference group.

Covariates used in the analysis were selected *a priori* based on their suspected roles as confounders. We first fit a basic model adjusted for age and gender (Model 1). For multivariable analysis, Model 2 additionally included race–ethnicity, education, income, body mass index (BMI), smoking status, serum cotinine levels, alcohol consumption, fruit and vegetable intake, physical activity, serum total cholesterol levels, hypertension status (systolic or diastolic blood pressure ≥ 140 or ≥ 90 mm Hg, respectively, use of antihypertensive drugs, or hypertension medical history), diabetes mellitus status (glycosylated hemoglobin $\geq 6.5\%$, use of antidiabetic drugs, or diabetes medical history), history of heart attack, congestive heart failure, stroke or cancer, hormone use among women (use of any estrogen or progesterone including oral contraceptive pills in the past one month), and use of vitamin or mineral supplements (in the past one month). Categories of variables used in the analysis were consistent with the NHANES III survey design (42).

Additional analyses examined these associations in the first 5 or 10 years of follow-up, after excluding deaths within the first 3 years of the survey, and after excluding current smokers. We also assessed the independent effects of these agents by adjusting for other micronutrients in the multivariable models. Because systemic inflammatory response has been shown to affect plasma micronutrient measurements, in an additional analysis, we further adjusted for C-reactive protein levels (50). As participants with self-reported history of comorbidities may have changed their dietary habits and supplement usage, we conducted separate analyses after excluding those with known history of heart attack, congestive heart failure, stroke, or cancer. Finally, to investigate dose–response associations using vitamin A and antioxidant nutrient levels as continuous variables, we created restricted cubic spline functions for all-cause, cancer, and cardiovascular disease mortality outcomes (51). All analyses were done using SAS (Version 9.3; SAS Institute Inc.).

Results

Of the 16,008 study participants, 4,225 died over a median follow-up period of 14.2 years. Eight hundred and ninety-one deaths were due to cancer and 1,891 were due to cardiovascular disease. Table 1 shows the mean baseline serum levels subdivided by sociodemographic, lifestyle, and health-related variables. There were significant differences in serum micronutrient levels for different participant characteristics, especially for BMI categories and smoking status, with current smokers and those with BMI ≥ 30 kg/m² having

Table 1. Serum antioxidant nutrient levels in NHANES III participants

| Characteristic | n | Mean serum levels (95% CI) | | | | |
|--------------------------------|--------|------------------------------------|----------------------------------|------------------------------------|--|---------------------------------|
| | | Vitamin A ($\mu\text{mol/L}$) | Vitamin C (mmol/L) | Vitamin E ($\mu\text{mol/L}$) | β -Carotene ($\mu\text{mol/L}$) | Selenium (nmol/L) |
| Age | | | | | | |
| 20–39 years | 6,425 | 1.93 (1.90–1.96) | 39.77 (37.95–41.60) | 22.84 (22.45–23.22) | 0.29 (0.28–0.31) | 1.58 (1.55–1.60) |
| 40–59 years | 4,252 | 2.08 (2.05–2.11) | 41.94 (39.90–43.97) | 28.80 (27.89–29.72) | 0.39 (0.37–0.41) | 1.60 (1.57–1.62) |
| ≥ 60 years | 5,331 | 2.24 (2.20–2.28) | 49.52 (47.35–51.69) | 32.85 (32.21–33.49) | 0.49 (0.47–0.52) | 1.59 (1.56–1.61) |
| Gender | | | | | | |
| Male | 7,510 | 2.18 (2.15–2.21) | 38.24 (36.58–39.90) | 26.38 (25.82–26.93) | 0.31 (0.30–0.33) | 1.61 (1.58–1.63) |
| Female | 8,498 | 1.93 (1.90–1.95) | 46.59 (44.60–48.58) | 27.52 (26.95–28.09) | 0.42 (0.41–0.43) | 1.57 (1.54–1.59) |
| Race-ethnicity | | | | | | |
| Non-Hispanic White | 6,783 | 2.09 (2.07–2.12) | 43.80 (41.70–45.89) | 27.77 (27.19–28.35) | 0.37 (0.36–0.39) | 1.60 (1.57–1.63) |
| Non-Hispanic Black | 4,270 | 1.90 (1.88–1.93) | 34.27 (33.27–35.26) | 22.91 (22.51–23.30) | 0.36 (0.33–0.38) | 1.50 (1.49–1.52) |
| Mexican-American | 4,325 | 1.86 (1.83–1.89) | 39.34 (37.61–41.08) | 25.06 (24.64–25.49) | 0.32 (0.29–0.34) | 1.57 (1.55–1.59) |
| Other | 630 | 1.92 (1.84–2.00) | 43.68 (39.63–47.73) | 25.77 (24.48–27.06) | 0.39 (0.36–0.43) | 1.57 (1.54–1.60) |
| Level of education | | | | | | |
| Less than high school | 6,556 | 2.06 (2.02–2.09) | 37.85 (34.73–40.97) | 26.77 (26.36–27.18) | 0.35 (0.33–0.37) | 1.56 (1.54–1.58) |
| High school or more | 9,344 | 2.05 (2.02–2.07) | 44.08 (42.57–45.59) | 27.03 (26.47–27.60) | 0.37 (0.36–0.39) | 1.60 (1.57–1.62) |
| Annual family income | | | | | | |
| <\$20,000 | 7,730 | 2.02 (1.98–2.05) | 38.54 (36.67–40.40) | 25.71 (25.31–26.11) | 0.34 (0.33–0.36) | 1.56 (1.54–1.58) |
| \geq \$20,000 | 8,018 | 2.06 (2.04–2.09) | 44.53 (42.75–46.30) | 27.59 (27.02–28.17) | 0.38 (0.37–0.40) | 1.60 (1.57–1.62) |
| BMI | | | | | | |
| <25 kg/m^2 | 6,314 | 1.98 (1.95–2.01) | 45.07 (43.27–46.87) | 25.37 (24.78–25.96) | 0.41 (0.40–0.43) | 1.60 (1.57–1.62) |
| 25–29.9 kg/m^2 | 5,609 | 2.13 (2.10–2.15) | 42.44 (40.51–44.37) | 28.27 (27.70–28.85) | 0.36 (0.35–0.37) | 1.58 (1.56–1.61) |
| ≥ 30 kg/m^2 | 4,042 | 2.06 (2.03–2.10) | 37.79 (35.60–39.97) | 28.27 (27.50–29.03) | 0.29 (0.27–0.31) | 1.57 (1.54–1.60) |
| Smoking status | | | | | | |
| Current smoker | 4,086 | 1.99 (1.96–2.03) | 31.67 (29.28–34.05) | 24.27 (23.82–24.72) | 0.25 (0.23–0.26) | 1.56 (1.53–1.58) |
| Former smoker | 4,018 | 2.18 (2.15–2.21) | 45.99 (44.31–47.66) | 29.68 (28.91–30.44) | 0.41 (0.38–0.43) | 1.61 (1.59–1.64) |
| Never smoker | 7,903 | 2.01 (1.98–2.03) | 47.45 (45.79–49.10) | 27.12 (26.54–27.70) | 0.42 (0.41–0.44) | 1.59 (1.57–1.61) |
| Alcohol consumption | | | | | | |
| Yes | 7,598 | 2.09 (2.06–2.12) | 41.50 (39.94–43.05) | 26.03 (25.54–26.53) | 0.34 (0.32–0.35) | 1.60 (1.57–1.62) |
| No | 8,410 | 1.99 (1.97–2.02) | 43.95 (41.62–46.29) | 28.15 (27.45–28.86) | 0.41 (0.39–0.43) | 1.57 (1.55–1.60) |
| Physical activity ^a | | | | | | |
| More active | 4,958 | 2.11 (2.08–2.14) | 46.11 (44.26–47.96) | 28.52 (27.71–29.34) | 0.44 (0.42–0.46) | 1.60 (1.57–1.62) |
| Less active | 3,518 | 2.00 (1.96–2.03) | 39.04 (37.59–40.49) | 25.66 (25.12–26.20) | 0.30 (0.29–0.32) | 1.57 (1.54–1.60) |
| About the same | 7,217 | 2.03 (2.00–2.05) | 41.44 (39.27–43.61) | 26.48 (25.96–26.99) | 0.35 (0.33–0.36) | 1.59 (1.56–1.61) |
| Hypertension status | | | | | | |
| Yes | 5,484 | 2.23 (2.20–2.26) | 42.90 (40.71–45.09) | 30.88 (30.29–31.46) | 0.39 (0.37–0.42) | 1.59 (1.57–1.62) |
| No | 10,522 | 1.98 (1.96–2.00) | 42.45 (40.78–44.12) | 25.47 (25.01–25.94) | 0.36 (0.35–0.37) | 1.58 (1.56–1.61) |
| Diabetes mellitus status | | | | | | |
| Yes | 1,770 | 2.13 (2.07–2.19) | 38.37 (36.00–40.74) | 32.24 (30.79–33.68) | 0.37 (0.35–0.40) | 1.60 (1.57–1.62) |
| No | 14,238 | 2.04 (2.02–2.07) | 42.89 (41.12–44.66) | 26.57 (26.10–27.04) | 0.37 (0.36–0.38) | 1.59 (1.56–1.61) |
| Hypercholesterolemia | | | | | | |
| Yes | 4,875 | 2.25 (2.23–2.28) | 44.53 (42.28–46.78) | 33.91 (33.12–34.69) | 0.44 (0.41–0.46) | 1.61 (1.59–1.64) |
| No | 11,124 | 1.95 (1.93–1.98) | 41.69 (39.99–43.38) | 23.81 (23.40–24.22) | 0.34 (0.32–0.35) | 1.57 (1.55–1.60) |
| Hormone use in women | | | | | | |
| Yes | 818 | 2.30 (2.25–2.36) | 51.16 (47.51–54.80) | 30.93 (29.14–32.72) | 0.38 (0.34–0.41) | 1.65 (1.62–1.68) |
| No | 7,494 | 2.03 (2.01–2.05) | 41.89 (40.17–43.60) | 26.66 (26.23–27.10) | 0.37 (0.36–0.38) | 1.58 (1.56–1.60) |
| Supplement use | | | | | | |
| Yes | 6,050 | 2.13 (2.10–2.16) | 52.43 (50.31–54.55) | 31.22 (30.46–31.98) | 0.46 (0.44–0.48) | 1.59 (1.56–1.61) |
| No | 9,948 | 1.99 (1.96–2.01) | 35.33 (33.56–37.11) | 23.85 (23.50–24.21) | 0.30 (0.29–0.31) | 1.59 (1.56–1.61) |

^aCompared with others of same age.

significantly lower vitamin C and β -carotene levels. NHANES III Analytic and Reporting Guidelines provide further details about the characteristics of the study population (42).

Table 2 shows the HRs and 95% confidence intervals (CI) for all-cause mortality by quintiles (Q1–Q5) of micronutrient levels. Table 2 also provides the cutoff points used for dividing serum levels into quintiles. HRs for cancer and cardiovascular disease mortality are shown in Table 3.

We observed U-shaped associations between serum levels of vitamins A and E, and all-cause mortality (Fig. 1) with those with levels in Q1 or Q5 having higher mortality risks compared with those with levels in Q2–Q4. For vitamin A, risk of cancer death decreased from Q1 to Q2, with no further decline in risk at higher levels, whereas for vitamin E, having levels ≥ 26.08 $\mu\text{mol/L}$ (Q4–Q5) were associated with the lowest cancer mortality risk. The increased all-cause mortality risk for those with levels in Q5 for vitamin A was mainly driven by higher cardio-

vascular disease mortality (Fig. 2), and for vitamin E, by higher stroke mortality.

For vitamin C (Model 2), all-cause mortality risk decreased with increases in serum levels from Q1 to Q4 ($P_{\text{trend}} < 0.001$) with no further decline in risk with higher levels (≥ 60.19 mmol/L). For cancer mortality, we observed a dose–response relationship between higher levels and reduced risk ($P_{\text{trend}} < 0.001$). Cardiovascular disease mortality decreased with higher vitamin C levels, except for those with levels in Q5.

For β -carotene and selenium, there was a significant decrease in the overall mortality risk from Q1 to Q2 (HR for Q2 vs. Q1 for β -carotene: 0.75; 95% CI, 0.61–0.93; for selenium: 0.74; 95% CI, 0.63–0.87); however, higher levels (≥ 0.15 $\mu\text{mol/L}$ for β -carotene; ≥ 1.40 nmol/L for selenium) did not significantly change the risk. Cancer mortality risks decreased from Q1 to Q2 for β -carotene and from Q1 to Q4 for selenium. Beyond Q1, higher levels seemed to increase the cardiovascular disease mortality risk for both β -carotene and selenium.

Table 2. HRs for all-cause mortality by quintiles of serum levels

| | Quintiles of serum levels | | | | | P_{trend} |
|---|---------------------------|------------------|------------------|------------------|------------------|--------------------|
| | Q1 | Q2 | Q3 | Q4 | Q5 | |
| Vitamin A levels, $\mu\text{mol/L}$ | ≤ 1.50 | 1.54–1.78 | 1.82–2.06 | 2.09–2.41 | ≥ 2.44 | |
| All-cause mortality, no. of events | 641 | 629 | 804 | 916 | 1218 | |
| Model 1 ^a | 1 (Reference) | 0.80 (0.67–0.96) | 0.73 (0.61–0.86) | 0.70 (0.59–0.82) | 0.83 (0.70–0.98) | 0.16 |
| Model 2 ^b | 1 (Reference) | 0.89 (0.73–1.09) | 0.82 (0.68–0.98) | 0.82 (0.69–0.97) | 0.95 (0.78–1.15) | 0.93 |
| Vitamin C levels, mmol/L | ≤ 15.33 | 15.90–31.80 | 32.36–45.42 | 45.99–59.62 | ≥ 60.19 | |
| All-cause mortality, no. of events | 949 | 691 | 651 | 662 | 838 | |
| Model 1 | 1 (Reference) | 0.78 (0.65–0.93) | 0.63 (0.51–0.78) | 0.53 (0.45–0.62) | 0.55 (0.47–0.65) | <0.01 |
| Model 2 | 1 (Reference) | 0.90 (0.77–1.06) | 0.80 (0.63–1.01) | 0.75 (0.61–0.91) | 0.75 (0.62–0.91) | <0.01 |
| Vitamin E levels, $\mu\text{mol/L}$ | ≤ 18.65 | 18.67–22.08 | 22.11–26.05 | 26.08–32.16 | ≥ 32.18 | |
| All-cause mortality, no. of events | 555 | 620 | 803 | 979 | 1251 | |
| Model 1 | 1 (Reference) | 0.80 (0.66–0.97) | 0.69 (0.57–0.84) | 0.69 (0.56–0.84) | 0.71 (0.60–0.84) | <0.01 |
| Model 2 | 1 (Reference) | 0.84 (0.67–1.04) | 0.81 (0.68–0.97) | 0.84 (0.66–1.07) | 0.89 (0.70–1.13) | 0.77 |
| β -Carotene levels, $\mu\text{mol/L}$ | ≤ 0.13 | 0.15–0.20 | 0.22–0.32 | 0.34–0.50 | ≥ 0.52 | |
| All-cause mortality, no. of events | 693 | 662 | 814 | 912 | 1127 | |
| Model 1 | 1 (Reference) | 0.72 (0.61–0.85) | 0.65 (0.56–0.76) | 0.60 (0.51–0.72) | 0.58 (0.49–0.69) | <0.01 |
| Model 2 | 1 (Reference) | 0.75 (0.61–0.93) | 0.76 (0.65–0.89) | 0.75 (0.63–0.89) | 0.75 (0.63–0.90) | 0.01 |
| Selenium levels, nmol/L | ≤ 1.38 | 1.40–1.50 | 1.51–1.60 | 1.61–1.71 | ≥ 1.73 | |
| All-cause mortality, no. of events | 930 | 858 | 763 | 679 | 856 | |
| Model 1 | 1 (Reference) | 0.70 (0.60–0.82) | 0.67 (0.57–0.79) | 0.58 (0.48–0.70) | 0.66 (0.57–0.76) | <0.01 |
| Model 2 | 1 (Reference) | 0.74 (0.63–0.87) | 0.77 (0.66–0.91) | 0.69 (0.56–0.85) | 0.79 (0.68–0.92) | 0.03 |

NOTE: Data are given as: HR (95% CI).

^aModel 1: adjusted for age and gender.

^bModel 2: adjusted for variables in Model 1 and race–ethnicity, level of education, annual family income, BMI, smoking status, serum cotinine level, alcohol consumption, fruit and vegetable intake, physical activity, serum total cholesterol levels, hypertension status, diabetes status, history of heart attack, congestive heart failure, stroke or cancer, hormone use in women, and supplement use.

Table 3. HRs for cancer and cardiovascular disease mortality by quintiles of serum levels

| | Quintiles of serum levels | | | | | <i>P</i> _{trend} |
|---|---------------------------|------------------|------------------|------------------|------------------|---------------------------|
| | Q1 | Q2 | Q3 | Q4 | Q5 | |
| Vitamin A levels, $\mu\text{mol/L}$ | ≤ 1.50 | 1.54–1.78 | 1.82–2.06 | 2.09–2.41 | ≥ 2.44 | |
| Cancer, no. of deaths | 150 | 137 | 171 | 215 | 216 | |
| Model 1 ^a | 1 (Reference) | 0.83 (0.58–1.18) | 0.79 (0.61–1.02) | 0.80 (0.61–1.06) | 0.72 (0.53–0.98) | 0.08 |
| Model 2 ^b | 1 (Reference) | 0.89 (0.59–1.34) | 0.87 (0.65–1.15) | 0.96 (0.72–1.28) | 0.91 (0.64–1.28) | 0.89 |
| Cardiovascular disease, no. of deaths | 250 | 265 | 365 | 404 | 596 | |
| Model 1 | 1 (Reference) | 0.87 (0.71–1.07) | 0.90 (0.71–1.13) | 0.78 (0.64–0.96) | 1.08 (0.86–1.36) | 0.24 |
| Model 2 | 1 (Reference) | 0.97 (0.78–1.21) | 0.93 (0.72–1.20) | 0.84 (0.66–1.06) | 1.09 (0.83–1.43) | 0.45 |
| Vitamin C levels, mmol/L | ≤ 15.33 | 15.90–31.80 | 32.36–45.42 | 45.99–59.62 | ≥ 60.19 | |
| Cancer, no. of deaths | 234 | 161 | 147 | 141 | 140 | |
| Model 1 | 1 (Reference) | 0.79 (0.61–1.04) | 0.56 (0.37–0.85) | 0.49 (0.37–0.64) | 0.40 (0.29–0.57) | <0.01 |
| Model 2 | 1 (Reference) | 0.86 (0.64–1.15) | 0.70 (0.45–1.09) | 0.69 (0.51–0.91) | 0.55 (0.39–0.79) | <0.01 |
| Cardiovascular disease, no. of deaths | 409 | 288 | 267 | 281 | 416 | |
| Model 1 | 1 (Reference) | 0.73 (0.56–0.94) | 0.62 (0.48–0.81) | 0.49 (0.39–0.61) | 0.60 (0.49–0.73) | <0.01 |
| Model 2 | 1 (Reference) | 0.83 (0.64–1.08) | 0.76 (0.58–1.01) | 0.66 (0.49–0.89) | 0.77 (0.60–1.00) | 0.04 |
| Vitamin E levels, $\mu\text{mol/L}$ | ≤ 18.65 | 18.67–22.08 | 22.11–26.05 | 26.08–32.16 | ≥ 32.18 | |
| Cancer, no. of deaths | 117 | 155 | 192 | 195 | 230 | |
| Model 1 | 1 (Reference) | 0.98 (0.64–1.51) | 0.85 (0.59–1.22) | 0.68 (0.42–1.09) | 0.68 (0.49–0.94) | <0.01 |
| Model 2 | 1 (Reference) | 1.04 (0.64–1.68) | 1.01 (0.67–1.55) | 0.88 (0.50–1.53) | 0.93 (0.57–1.50) | 0.52 |
| Cardiovascular disease, no. of deaths | 210 | 249 | 356 | 469 | 596 | |
| Model 1 | 1 (Reference) | 1.06 (0.78–1.44) | 0.91 (0.64–1.30) | 1.00 (0.75–1.32) | 1.03 (0.78–1.35) | 0.73 |
| Model 2 | 1 (Reference) | 1.05 (0.76–1.45) | 1.01 (0.69–1.48) | 1.10 (0.83–1.47) | 1.16 (0.84–1.59) | 0.29 |
| β -Carotene levels, $\mu\text{mol/L}$ | ≤ 0.13 | 0.15–0.20 | 0.22–0.32 | 0.34–0.50 | ≥ 0.52 | |
| Cancer, no. of deaths | 173 | 150 | 180 | 181 | 205 | |
| Model 1 | 1 (Reference) | 0.67 (0.49–0.93) | 0.63 (0.47–0.86) | 0.60 (0.46–0.77) | 0.51 (0.39–0.67) | <0.01 |
| Model 2 | 1 (Reference) | 0.74 (0.50–1.09) | 0.76 (0.53–1.08) | 0.83 (0.59–1.16) | 0.74 (0.54–1.01) | 0.25 |
| Cardiovascular disease, no. of deaths | 263 | 265 | 354 | 437 | 561 | |
| Model 1 | 1 (Reference) | 0.73 (0.57–0.93) | 0.71 (0.55–0.92) | 0.70 (0.54–0.89) | 0.69 (0.53–0.90) | 0.03 |
| Model 2 | 1 (Reference) | 0.75 (0.56–1.01) | 0.84 (0.64–1.10) | 0.86 (0.67–1.12) | 0.91 (0.68–1.21) | 0.82 |
| Selenium levels, nmol/L | ≤ 1.38 | 1.40–1.50 | 1.51–1.60 | 1.61–1.71 | ≥ 1.73 | |
| Cancer, no. of deaths | 194 | 191 | 169 | 128 | 182 | |
| Model 1 | 1 (Reference) | 0.79 (0.56–1.12) | 0.77 (0.58–1.02) | 0.49 (0.34–0.69) | 0.73 (0.54–0.97) | 0.01 |
| Model 2 | 1 (Reference) | 0.76 (0.52–1.12) | 0.81 (0.62–1.08) | 0.53 (0.36–0.79) | 0.86 (0.62–1.20) | 0.20 |
| Cardiovascular disease, no. of deaths | 400 | 387 | 335 | 320 | 383 | |
| Model 1 | 1 (Reference) | 0.71 (0.57–0.90) | 0.67 (0.53–0.85) | 0.73 (0.58–0.93) | 0.71 (0.58–0.88) | 0.04 |
| Model 2 | 1 (Reference) | 0.76 (0.61–0.95) | 0.77 (0.63–0.96) | 0.89 (0.71–1.12) | 0.83 (0.67–1.04) | 0.58 |

NOTE: Data are given as: HR (95% CI).

^aModel 1: adjusted for age and gender.^bModel 2: adjusted for variables in Model 1 and race–ethnicity, level of education, annual family income, BMI, smoking status, serum cotinine level, alcohol consumption, fruit and vegetable intake, physical activity, serum total cholesterol levels, hypertension status, diabetes status, history of heart attack, congestive heart failure, stroke or cancer, hormone use in women, and supplement use.

In our analysis, in multivariable models, HRs did not change significantly after further adjusting for serum C-reactive protein levels (Supplementary Table S1) or after adjusting for other micronutrients in the study (Supplementary Table S2). Excluding NHANES III participants

who are current smokers (Supplementary Table S3) or those with history of heart attack, congestive heart failure, stroke, or cancer (Supplementary Table S4) also did not materially change the findings. We observed similar patterns for mortality risks after excluding deaths within the

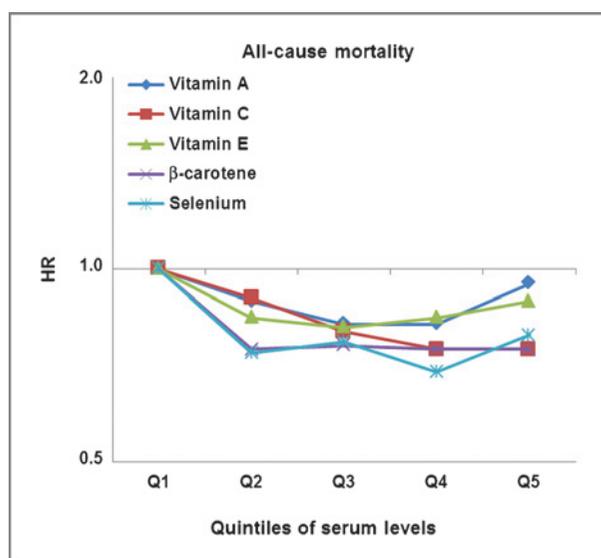


Figure 1. HRs for all-cause mortality. HRs are adjusted for variables in Model 2: age, gender, race-ethnicity, level of education, annual family income, BMI, smoking status, serum cotinine level, alcohol consumption, fruit and vegetable intake, physical activity, serum total cholesterol levels, hypertension status, diabetes status, history of heart attack, congestive heart failure, stroke or cancer, hormone use in women, and supplement use. The y-axis is shown on a log scale.

first 3 years of the survey (Supplementary Table S5) or after limiting follow-up to 5 or 10 years (Supplementary Table S6). We also examined HRs for all-cause mortality stratified by age, gender, race-ethnicity, BMI, and supplement use (Supplementary Table S7). More detailed results for cancer and cardiovascular disease mortality outcomes are provided in Supplementary Tables S8 and

S9, respectively. Restricted cubic splines for all-cause mortality, cancer mortality, and cardiovascular mortality, are presented in Supplementary Figures S1, S2, and S3, respectively.

Discussion

Oxidative stress has been implicated in the pathogenesis of several chronic diseases, including various cancers (52–55). Vitamin A and antioxidant nutrients such as vitamins C and E, β -carotene, and selenium have been hypothesized to prevent or delay this damage (56–61). Consequently, use of these agents is rising, and almost 40% of U.S. adults take them for their perceived health benefits (1, 2).

In this study, using data from a large, nationally representative cohort of 16,008 U.S. adults, we observed threshold effects and nonlinear associations between serum levels of these micronutrients, and all-cause and cause-specific mortality outcomes. These results help to explain some of the discrepant findings between the observational studies and randomized trials about the role of these agents on health outcomes.

We found that higher levels were generally associated with a modest decrease in all-cancer mortality risk. However, the most significant decline in risk was generally from the first to the second quintile. A number of high profile trials assessing antioxidants as cancer prevention agents have failed to show significant mortality benefits with taking these supplements (23–27) and some have reported possible harm (28, 29). A recent meta-analysis of 22 primary and secondary prevention trials concluded that there was no evidence to support preventive effects of any of the antioxidant supplements on any cancer type

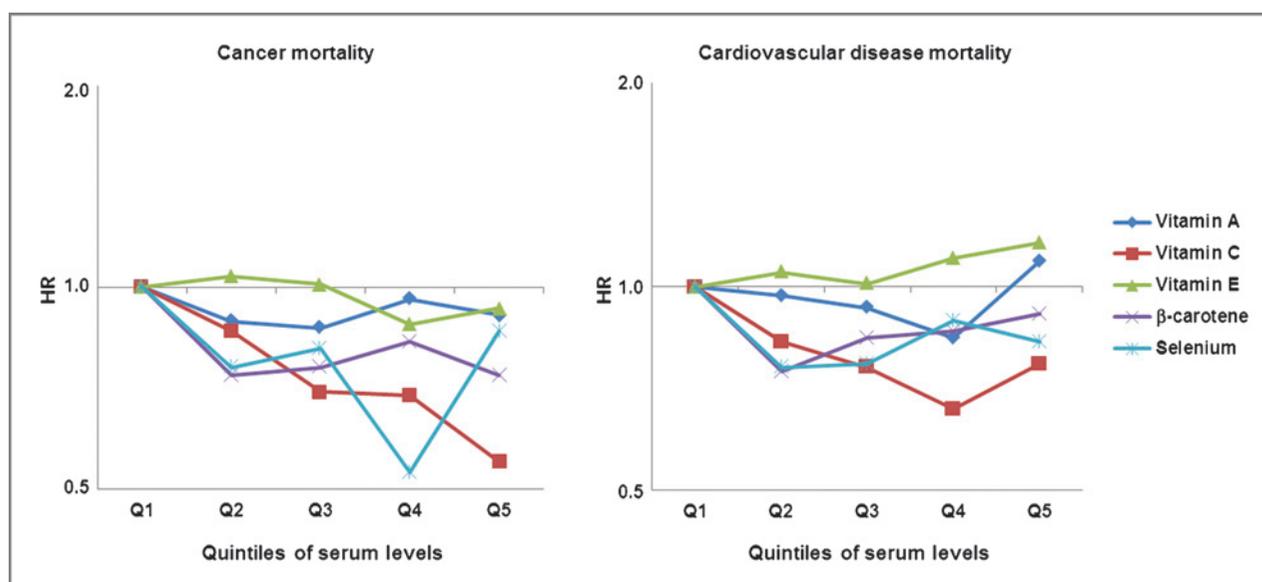


Figure 2. HRs for cancer and cardiovascular disease mortality. HRs are adjusted for variables in Model 2: age, gender, race-ethnicity, level of education, annual family income, BMI, smoking status, serum cotinine level, alcohol consumption, fruit and vegetable intake, physical activity, serum total cholesterol levels, hypertension status, diabetes status, history of heart attack, congestive heart failure, stroke or cancer, hormone use in women, and supplement use. The y-axes are shown on a log scale.

(62). The Physicians' Health Study II also recently showed only a small reduction in total cancer incidence (HR = 0.92; 95% CI, 0.86–0.998) and a nonsignificant reduction in mortality (HR = 0.88; 95% CI, 0.77–1.01) with more than a decade of daily multivitamin use and follow-up (22). Furthermore, most of these large prevention trials did not consider baseline nutrition level in their inclusion criteria (63, 64).

The importance of assessing baseline nutritional status is underscored by the Linxian Study in China, which specifically targeted a geographic area where subjects had chronically low blood levels of multiple micronutrients. It is the only prevention trial to our knowledge so far that has reported statistically significant reductions in all-cancer (RR = 0.87; 95% CI, 0.75–1.00) as well as overall mortality (RR = 0.91; 95% CI, 0.84–0.99) risks with antioxidant supplementation (65).

Interestingly, in the Physicians' Health Study I, there was no decrease in total cancer incidence with β -carotene supplementation. However, in subgroup analyses, lower baseline β -carotene levels were associated with alcohol consumption and higher BMI, and overall cancer incidence was modestly reduced with supplementation in these subgroups (among daily alcohol drinkers, RR = 0.9; 95% CI, 0.8–1.0; among those in the highest BMI quartile, RR = 0.9; 95% CI, 0.7–1.0; ref. 23). Similarly, in the SU.VI. MAX trial, lower total cancer incidence (RR = 0.69; 95% CI, 0.53–0.91) and all-cause mortality (RR = 0.63; 95% CI, 0.42–0.93) with multivitamin use in men was attributable to their lower baseline antioxidant status (27). Although some of these results could be due to testing multiple hypotheses, the findings presented in our study as well as the results from others examining a nutritionally deficient population further support the hypothesis that supplementation might only be useful for those who have low serum antioxidant levels, and that beyond a threshold, higher levels do not lead to additional benefit, and may potentially be toxic.

Our results also have broader implications for assessing overall mortality benefits from the use of dietary supplements. For example, using antioxidant levels as continuous variables, a study in the British National Diet and Nutrition Survey found that higher levels of vitamin C, selenium, and β -carotene were associated with lower overall mortality risk (HR per SD for vitamin C = 0.81, $P < 0.001$; for selenium = 0.76, $P < 0.001$; and for β -carotene = 0.92, $P = 0.08$), whereas the results for vitamins A and E were not significant (HR for vitamin A = 0.96, $P = 0.4$; α -tocopherol = 0.96, $P = 0.4$; ref. 6). If we had used antioxidant levels as continuous variables, we would have obtained similar results. However, by assessing nonlinear associations, we were able to show that for selenium and β -carotene, beyond the first quintile, there was no apparent decrease in the mortality risk with higher levels (Fig. 1). Moreover, we showed a U-shaped association between levels of vitamins A and E, and all-cause mortality.

These results for overall mortality are also consistent with the findings of the recent systematic review by the

Cochrane Collaboration, which concluded that supplementation with β -carotene, vitamin E, or high doses of vitamin A was associated with increased mortality risk, whereas the role of vitamin C or selenium supplement use was not clear (21). Figure 1 in our study shows that except for vitamin C, having antioxidant levels beyond the first quintile did not lead to any further decrease in all-cause mortality risk. Moreover, for vitamins A and E, higher levels increased the overall risk of death.

Our study has a number of strengths compared with previous studies. First, this is the largest study to date assessing serum vitamin A and antioxidant nutrient levels, and cause-specific and all-cause mortality outcomes. Therefore, instead of examining serum levels as continuous variables, we were able to divide them into quintiles, and assess threshold and nonlinear relationships with mortality. Because of small sample sizes and/or short follow-up durations, most observational studies have been unable to show such associations (3–15). Furthermore, NHANES III used a nationally representative sample of the U.S. population, and final mortality status was available for more than 99% of the participants, minimizing the possibility of selection bias. Assessing serum levels of these agents instead of using dietary history, along with standardized and validated laboratory methods, reduced the potential for information bias. Finally, use of appropriate sampling weights in the analysis helped to obtain statistical estimates similar to those if the entire sampling frame (the United States) had been surveyed.

Our study has several limitations. We tested multiple hypotheses. Therefore, results must be interpreted with caution. We limited our analysis to vitamin A and antioxidant nutrients that are commonly used as dietary supplements. Other potential agents with antioxidant properties such as zinc, lycopene, and other carotenoids need further evaluation. As with any observational study, residual confounding by socioeconomic status, lifestyle variables, and other factors cannot be excluded. However, NHANES III assessed a large number of health-related variables, which enabled us to control for many potential confounders. Furthermore, we obtained different results for different micronutrients as well as for different health outcomes. These findings would be difficult to explain entirely on the basis of residual confounding. In addition to assessing cardiovascular mortality, there are many other possible competing risks for mortality that we could have assessed. We assessed cardiovascular mortality as an example of potential different effects of antioxidants on different disease processes.

Finally, in this study, we used a single measurement of serum levels to assess long-term nutritional status. In our analysis (Supplementary Table S6), we observed similar HRs for 5-year, 10-year, and for the entire duration of follow-up. This suggests that at the population level, a single measurement of serum levels may provide a reliable estimate of long-term antioxidant status. A major limitation of randomized trials assessing use of

supplements for primary prevention is that the participants typically have to be kept on intervention for long periods to significantly affect health outcomes. Therefore, observational studies such as this may facilitate the assessment of the relationship between long-term nutritional status and health outcomes.

In summary, using data from a nationally representative cohort of more than 16,000 U.S. adults, we were able to identify specific plasma levels of vitamin A and various antioxidant nutrients, which were associated with maximum cancer-related and overall survival. We found that beyond a certain threshold, there was generally no additional benefit with higher serum levels with respect to overall mortality. Specifically, for vitamins A and E, higher levels increased the overall mortality risk. These data support the findings of recent randomized trials that have generally failed to show health benefits with taking these supplements (21, 62, 66). We also showed that having low serum antioxidant nutrient levels was associated with higher mortality risk, suggesting that supplementation might still be useful for those who are nutritionally deficient.

Our findings underscore the need to assess safety of these agents like other drugs, rather than classifying them as "dietary supplements," which affects their regulatory oversight (67). Although the current Institute of Medicine guidelines provide recommended dietary allowances and tolerable upper intake levels for these agents, we also highlight the potential significance of measuring serum levels to guide their use as supplements (68). Novel intervention studies might then be planned where doses of these agents are individualized on the basis of serum levels, lifestyle behaviors such as smoking which affect

levels, and possibly, markers of oxidative stress and systemic inflammatory response (50, 69, 70). Such a strategy would help ensure that those who are deficient get the required micronutrients, while also preventing toxic levels of antioxidants that could potentially lead to worse health outcomes.

Disclosure of Potential Conflicts of Interest

No potential conflicts of interest were disclosed.

Authors' Contributions

Conception and design: A. Goyal, M.B. Terry, A.B. Siegel

Development of methodology: A. Goyal

Acquisition of data (provided animals, acquired and managed patients, provided facilities, etc.): A. Goyal, A.B. Siegel

Analysis and interpretation of data (e.g., statistical analysis, biostatistics, computational analysis): A. Goyal, M.B. Terry, A.B. Siegel

Writing, review, and/or revision of the manuscript: A. Goyal, M.B. Terry, A.B. Siegel

Administrative, technical, or material support (i.e., reporting or organizing data, constructing databases): A. Goyal, A.B. Siegel

Study supervision: A. Goyal, M.B. Terry, A.B. Siegel

Acknowledgments

The authors thank Dr. Zhezhen Jin from the Department of Biostatistics, Mailman School of Public Health, Columbia University, for his help with restricted cubic spline functions.

Grant Support

This work was supported by the Steven J. Levinson Medical Research Foundation and NIH K23 CA149084-01A1 (to A.B. Siegel).

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Received April 12, 2013; revised June 26, 2013; accepted July 23, 2013; published OnlineFirst August 8, 2013.

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