Restoration of turbulence profile from scintillation indices

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ABSTRACT
An algorithm that permits one to measure atmospheric turbulence by statistical analysis of light flux fluctuations in four concentric-ring apertures is described in detail. It consists of computing the scintillation indices for each aperture and pairwise aperture combination and in fitting the set of measured indices to a model with a small number of turbulent layers. The performance of this method is analysed by means of simulations and using the real data from a multi-aperture scintillation sensor. It is shown that a turbulence profile with a vertical resolution of \( \Delta h/h \sim 0.5 \) can be reconstructed and that the errors of the measured intensities of turbulent layers are typically around 10 per cent of the integrated intensity. The integral parameters such as the seeing and the isoplanatic angle are measured with few per cent accuracy.

Key words: turbulence – atmospheric effects – instrumentation: adaptive optics – site testing.

1 INTRODUCTION

Knowledge of the vertical distribution of optical turbulence in the terrestrial atmosphere is essential for understanding the phenomenon of seeing, for selecting good astronomical sites and for predicting the performance of various imaging techniques. For example, the size of a field corrected by adaptive optics (both classical and multiconjugate) depends on the turbulence profile (Roddier 1999). If only the ground turbulent layer is adaptively corrected, as suggested by Rigaut (2002), the resulting improvement of seeing is also profile-dependent.

Stellar scintillation is a natural way to sense turbulence remotely. It results from the propagation of wave-fronts distorted by turbulent layers: phase perturbations are converted to intensity fluctuations. The amplitude and characteristic size of intensity fluctuations increase with the propagation distance. Hence, it is possible to perform remote turbulence sounding by a scintillation analysis, as suggested by Peskoff (1968).

Existing techniques for turbulence profile (TP) measurement such as balloon microthermal sounding or remote optical sounding with double stars, Scintillation Detection and Ranging (SCIDAR) (Fuchs, Tallon & Vernin 1998), are expensive and not suitable for continuous monitoring. Tokovinin & Kornilov (2002) proposed to obtain low-resolution profiles from the statistical analysis of light fluxes in four concentric-ring apertures. Such a multi-aperture scintillation sensor (MASS) was actually built and tested by Kornilov et al. (2003). Compared with the previous single-star scintilometer of Ochs et al. (1976), MASS has a better vertical resolution of \( \Delta h/h \sim 0.5 \) and good absolute calibration. This has been achieved by means of careful signal processing and interpretation. In this paper we describe this algorithm step by step. Although the algorithm was developed for a specific instrument, MASS, it is more general and can be applied to other instruments that would use the same principle.

A simple and robust TP monitor such as MASS is of interest for modern site testing and for support of adaptive optics and interferometry operations at the existing observatories. It is already being used to study the statistics of TP at Cerro Tololo (Tokovinin, Baumont & Vasquez 2003) and La Silla observatories in Chile and at Mauna Kea (Hawaii); more applications are planned in the near future.

The principle of MASS operation is briefly introduced in Section 2. Fluctuations of stellar light caused by scintillation are characterized by the scintillation indices (SIs) that are computed by the algorithm presented in Section 3. Turbulent layers in the atmosphere contribute to a given SI according to the corresponding weighting function (WF, Section 4). Knowing both indices and WFs, we can restore the TP if a simple model of atmosphere with few discrete layers is adopted, as explained in Section 5. In Section 6, we study the vertical resolution and precision of such a restoration technique and show examples of its application. Alternatively, some atmospheric parameters of interest can be derived directly from SIs without explicit profile restoration (Section 7). The conclusions are given in Section 8.

2 THE MASS INSTRUMENT

In Fig. 1, the principle of MASS and its data processing are illustrated. Upon propagation through a turbulent layer and further down to the ground, light from a stellar source acquires amplitude and phase fluctuations. Amplitude fluctuations are the cause of scintillation (twinkling) of stars, they resemble a pattern of ‘flying shadows’
on the ground. The characteristic size of the flying shadows is of the order of the Fresnel radius \( \sqrt{\lambda z} \) for a propagation distance \( z \) and a wavelength \( \lambda \). Thus, the spatial structure of single-star scintillation contains information on the distance to turbulent layers (Peskoff 1968).

In MASS, the flux of a bright star in four concentric apertures A, B, C and D is registered. Those apertures together act as a spatial filter. The smallest aperture A has a diameter of 2 cm, it is sensitive to turbulence at altitudes above 0.5 km. The largest aperture D has an outer diameter of 13 cm and senses mostly the high layers. Statistical analysis of the scintillation signals in all four apertures permits one to localize turbulent layers and to measure their intensities.

The MASS instrument is described by Kornilov et al. (2003). Briefly, it consists of a feeding telescope with a clear (unobstructed) aperture of 14-cm diameter and a detector box. Within the box, the image of the exit pupil is segmented into rings by a system of concentric mirrors with different tilts – a segmentator. The four beams A, B, C and D are then directed to the four photomultipliers (PMTs), which work in photon-counting mode with a pulse resolution of approximately 15 ns. The spectral response of MASS is determined by a glass filter, which passes the wavelength band from 400 to 550 nm, with a full-width at half maximum (FWHM) bandwidth of 100 nm. A star of \( B = 0 \) mag gives a count rate in the smallest aperture of approximately 1 count ms\(^{-1}\). The series of photon counts are used to compute the 10 SIs (four normal and six differential) and to restore the TP as explained below.

MASS in its normal mode is not sensitive to the near-ground turbulence because the latter does not produce any scintillation. When a positive lens is placed in the focal plane, the image of a pupil on the segmentator is defocused and corresponds to the additional turbulence because the latter does not produce any scintillation.

The background \( B \) is the number of background counts in channel \( i \). For any two channels \( x \) and \( y \) \( (x = y \) for auto-covariance) and for a time lag \( k \) \( (k = 0, 1, 2) \) the covariance \( \rho'_{xy,k} \) is found as

\[
\rho'_{xy,k} = \frac{1}{2(N-k)} \sum_{i=1}^{N-k} ((x_i - \bar{x})(y_{i+k} - \bar{y}) + (x_{i+k} - \bar{x})(y_{i} - \bar{y})).
\]

3 SCINTILLATION INDICES AND THEIR ERRORS

In this section we explain how the series of raw photon counts are processed to obtain SIs. Suppose that the instantaneous light flux received by a given MASS aperture is \( X \). The SI \( s_x \) is defined as a variance of flux normalized by the square of the average flux:

\[
s_x = \frac{(X - \bar{X})^2}{\bar{X}^2} = \frac{(X - \bar{X})^2}{\bar{X}^2} - 1.
\]

The differential SI \( s_{xy} \) (Tokovinin 1998, 2002) is defined similarly as a variance of the difference of normalized fluxes \( X \) and \( Y \) in a pair of apertures:

\[
s_{xy} = \frac{(X - \bar{X} - Y + \bar{Y})^2}{\bar{X} \bar{Y}} = s_x + s_y - 2X\bar{Y} - \frac{(X - \bar{X})^2}{\bar{X}^2} + 2.
\]

Differential SIs are related to covariances.

The instantaneous fluxes \( X \) and \( Y \) are not known. Instead, we measure the series of photon counts in the four channels with a finite microexposure time \( t \). Let \( x_i \) and \( y_j \) be these counts in two channels, where \( N \) is the number of samples during a 1-s exposure \( (N = 1000 \) for \( t = 1 \) ms). To obtain correct SIs from this data, we must take into account the non-linearity of photon counters, subtract the contribution from photon noise, and extrapolate the indices to zero exposure time, thus removing the bias caused by finite \( t \). A correction for the background (dark count of PMTs and light from the sky) must be made as well. These operations are performed in several steps.

Step 1. Compute the mean count values \( \bar{x} \) for all channels.

Step 2. Compute the raw auto- and cross-covariances of the counts in all channels. Auto-covariances are needed for time lags of 0, 1, 2 sampling intervals, cross-covariances – only for lags 0 and 1. For any two channels \( x \) and \( y \) \( (x = y \) for auto-covariance) and for a time lag \( k \) \( (k = 0, 1, 2) \) the covariance \( \rho'_{xy,k} \) is found as

\[
\rho'_{xy,k} = \frac{1}{N} \sum_{i=1}^{N-k} (x_i - \bar{x})(x_{i+k} - \bar{x}) + (y_i - \bar{y})(y_{i+k} - \bar{y}).
\]

Step 3. Correct the raw covariances and mean counts (designated here by a prime superscript) for the non-linearity of photon counters (Kornilov & Pogrosheva 1989). The non-linearity parameter \( L = \tau / t \) is the ratio of counter dead time \( \tau \) to the microexposure time \( t \).

\[
\bar{x} = \bar{x}(1 + L\bar{x}), \quad \rho'_{xy} = \frac{\rho'_{xy}}{(1 - 2L\bar{x})(1 - 2L\bar{y})}.
\]

Step 4. Compute the SIs that are free of photon noise, corrected for background and extrapolated to zero microexposure time (1 ms is not short enough to freeze scintillations in the smallest apertures). The background \( B \) is the number of background counts in channel \( x \) per microexposure; it is almost always less than 1 per cent of the signal and not critical for data reduction. The non-Poisson parameter \( p \) is the ratio of the variance of photon counts to their mean value at constant light, \( p = (\bar{x}^2 - \bar{x})/\bar{x} \). For an ideal photon counter \( p = 1 \), but for real PMTs at low fluxes the measured \( p \) is typically in the range from 1 to 1.1. The covariances are already corrected for non-linearity at this point. The normal and differential SIs are:

\[
s_x = \frac{1.5(\rho_{xx,0} - \rho'_{xx}) - 0.5\rho_{xx,1}}{(\bar{x} - B)^2},
\]

---

**Figure 1.** Principle of turbulence profile measurement with MASS. This figure is available in colour in the online version of the journal on Synergy.
\[ s_{xy} = s_x + s_y - 2 \frac{1.5 \rho_{xy,0} - 0.5 \rho_{xy,1}}{(\bar{T} - B_i)(\bar{T} - B_i)} . \]  \hspace{1cm} (6)

The derivation of these formulae is given in Appendix A. The set of 10 SIs is computed every second. Thus, slow flux variations (below 1 Hz) caused by unstable transparency, changing zenith distance, etc. do not contribute to the SIs.

Step 5. Estimate average SIs and their errors from \( K = 60 \) 1-s measurements accumulated during 1 min. If 1-s measurements were uncorrelated, the error of the mean would be simply \( \sqrt{K} \) times less than the rms scatter of the data. Partial correlation of data is taken into account by the following formula for the variance \( \sigma^2 \) of the average SI:

\[ \sigma^2 = \frac{1}{K} \left[ \rho_{s,0} + 2 \sum_{k=1}^{K-1} (1 - k/K) \rho_{s,k} \right] , \hspace{1cm} (7) \]

where \( \rho_{s,k} \) is the auto-covariance of SI \( s \) for the lag \( k \) computed by the same formula (3) as used for photon counts. We found that the relative errors \( \sigma^2/\bar{s} \) of all normal and differential SIs estimated by the above formula are approximately 2 per cent, except for the smallest differential SI \( s_{AB} \) where the photon noise is important and the relative error is \( 3 \)–\( 7 \) per cent, depending on the photon flux. Below we explain why all other SIs have the same relative errors independent of the scintillation strength and the stellar flux.

The variance of the normalized light intensity from which SI is derived is a sum of the scintillation variance \( s_s \) and the photon noise variance \( 1/\bar{T} \); the latter is subtracted (equation 5), but its statistical uncertainty still contributes to the SI error. The relative proportion of those two terms depends on the photon flux, TP and type of the index considered. For example, a 1-ms flux in the smallest aperture A for a 2.5-mag star is \( \bar{T} = 20 \) photons, while typically \( s_s < 0.1 \), making the photon noise contribution unimportant. This applies to other indices as well. Only the differential SI in the smallest aperture pair under good (0.5-arcsec) seeing is smaller than the photon noise variance \( s_{AB} \approx 0.006 \), explaining why this SI is measured with a larger error.

The contribution of these terms to the error of SI also depends on the correlation time of these signals, which is different. It is well known that a signal with a correlation time \( \tau \) measured during acquisition time \( T \) would have a relative error of its variance estimate of \( \sqrt{T/\tau} \approx 0.013 \) for \( \tau \approx 0.01 \) s. The photon noise has a white spectrum, so for a sampling time \( t = 1 \) ms the relative error of the photon noise variance estimate is only \( \sqrt{T/\tau} \approx 0.004 \), or \( \sim 3 \) times less than for the same scintillation signal. In fact, turbulence is not quite stationary even during 1 min, increasing the scatter of 1-s SIs and the errors of mean SIs compared with the estimates for a stationary case given above; this is why we rely on the experimental estimates of SI errors.

If the relative noise in the measured SIs is more or less constant, the relative noise of TPs restored from this data will also be constant. It means that the sensitivity of MASS improves when turbulence is weak and degrades under strong turbulence, as confirmed below by the analysis of real data.

4 WEIGHTING FUNCTIONS

Atmospheric turbulence can be considered as a collection of independent turbulent layers at altitudes \( h_i \). The intensities of those layers \( J_i \) are the integrals of refractive index structure constant \( C^2_n(h) \) over the layers

\[ J_i = \int_{h_{i+1}}^{h_i} C^2_n(h) \, dh . \hspace{1cm} (8) \]

they are measured in m\(^{-1/3}\). The TP is completely defined by the set of \( \{ h_i, J_i \} \) values. This might appear different from the classical view of TP as a continuous function \( C^2_n(h) \), although it is nothing else but a discrete representation of such a function. The optical effect of turbulence is directly related to the intensities \( J_i \) rather than to \( C^2_n(h) \).

The theory of weak perturbations (Tatarsky 1961; Roddier 1981) establishes a linear relation between some scintillation index \( s_s \) and the intensities of all layers:

\[ s_s = \sum_{i=1}^{I} J_i W_x(h_i \sec \gamma) \sec \gamma , \hspace{1cm} (9) \]

where \( \gamma \) is the zenith distance of a star, \( z = h \sec \gamma \) is the propagation distance, and the altitudes \( h_i \) are counted from the observatory level rather than from sea level. The coefficients \( W_x(h) \) are called weighting functions (WFs), they depend on wavelength and aperture size. Equation (9) is valid only for weak (non-saturated) scintillations, which means practically \( s_s \lesssim 0.3 \). Application of the linear theory to saturated scintillations will result in an underestimation of \( J_i \).

The formulae to compute WFs for normal SIs and monochromatic light can be found in many publications. Tokovinin (1998, 2002) gives these formulae for both normal and differential SIs and for apertures of arbitrary shape. However, it turns out that for differential SIs the effect of spectral bandpass is significant and the polychromatic WFs cannot be computed as a weighted average of monochromatic WFs. Formulae for polychromatic WFs are provided in Tokovinin (2003) and reproduced in Appendix B for completeness.

Polychromatic WFs computed for the two smallest MASS apertures are plotted in Fig. 2. The MASS bandpass is wide enough, so we have to take into account the variations of the energy distribution in the stellar spectrum within the bandpass. Stars are subdivided into 12 groups according to their spectra and the WFs are pre-computed for each spectral group separately. During observations, an adequate table of WFs is selected depending on the star.

Equation (9) can be written in a matrix form by considering the vector of indices \( X \) of length \( M = 10 \) (\( M = 13 \) in generalized mode),
the vector of layer intensities $J$ of length $I$ and the matrix $W$ with elements $W_{mi} = W_m(h_i \text{sec } \gamma) \text{sec } \gamma$: $S = WJ$.

This formula is the basis of TP restoration.

5 RESTORATION OF THE TURBULENCE PROFILE

5.1 The model

We measure only a small number of indices, much fewer than any realistic number of atmospheric layers. However, in order to obtain some information on TP we fit the data (indices) to a model with only a few layers. It will be shown that such a procedure leads to a low-resolution TP, which adequately represents any real TP. The ‘layers’ restored in this way are in fact thick slabs of turbulence. Despite the crudeness of such a model, layer intensities are measured correctly.

Two alternative restoration techniques are used. The first fixed-layer method assumes layers at some pre-defined altitudes $h_i$. In this case the relation between data $S$ and unknowns $J$ is linear (equation 10). The second floating-layer method uses a model with only three layers that can be located at any altitude and solves equation (10) for the unknowns $h_i$, $J_i$. In both models the number of unknowns must be less than the number of measurements $M$. In fitting models to the data, we search for model parameters that minimize the $\chi^2$ function. Let $\bar{S}_m$ be the measured average indices, $\sigma_m$, their variances (as estimated from equation 7), $\hat{S}_m$, the expected (or model) indices computed from equation (10), then

$$\chi^2 = \sum_{m=1}^{M} \frac{(\bar{S}_m - \hat{S}_m)^2}{\sigma_m^2}.$$  

The minimum value of $\chi^2$ is a measure of the correspondence between the model and the data. In theory, its expected value is $M - I$, where $I$ is the number of model parameters. In practice it is almost always larger for many reasons (e.g. inadequate model, biased data). The data (indices) are not mutually independent because 10 SIs are computed from only four light fluxes.

Below we consider each restoration method in more detail.

5.2 Fixed-layers restoration

For fixed-altitude layers the system (10) is linear and a solution $\hat{J}$ that minimizes $\chi^2$ is found by the least-squares method (Press et al. 1992):

$$\hat{J} = (W^TW)^{-1}WS,$$  

where the system matrix $W$ and the data are renormalized by the errors, $W_{mi} = W_m / \sigma_m$ and $S_m = S_m / \sigma_m^2$.

The noise variance of the $i$th restored layer intensity $\sigma_{\hat{J}i}$ is then equal to the $i$th diagonal element of the inverse normal matrix:

$$\sigma_{\hat{J}i}^2 = (W^TW)^{-1}_{ii}.$$  

Given that the least-squares method is not quite applicable in our case (correlated data, non-negativity constraint), equation (13) can be used only for a rough estimate of the resulting errors.

It is well known that the normal matrix of the least-squares problem can be ill-conditioned. This happens when two or more unknowns have a similar influence on the data. For example, if we select two pre-defined layers at 10 and 11 km, the WFs for these layers will be very similar, hence the matrix $W$ will be almost degenerate. A reasonable choice of layer altitudes is essential for successful restoration. We have chosen a logarithmic altitude grid (1, 2, 4, 8, 16 km) that matches the character of WFs. A sixth layer at 0.5 km was added after it became clear that it was required to model data in some cases. A layer at 0 km is added in the generalized mode. On the other hand, if the model contains too few layers the fit to the data will be poor (large $\chi^2$).

Linear restoration does not take into account the fact that layer intensities $J_i$ must be non-negative. So, instead of solving (10) by the least-squares method, we minimize $\chi^2$ (equation 11) directly with this additional constraint. Technically, we change variables to $J_i = y_i^2$ and minimize $\chi^2$ over $y_i$ using the standard Powell method.

5.3 Floating-layers restoration

The second restoration method uses a model with only three layers and six parameters $\{h_i, J_i\}$. Real TPs often contain strong layers that can be localized with this technique. Of course, the continuously distributed turbulence that is also always present in the atmosphere is included into the three layers as well.

The implementation of the floating-layer method is more complicated and computer-intensive than for fixed layers. The problem is highly non-linear, so minimization of $\chi^2$ by gradient techniques leads to local minima. Instead, we check all plausible combinations of altitudes and for each combination find the best-fitting layer intensities and the corresponding $\chi^2$. The altitude of the lowest layer $h_1$ is selected among the grid points (with a logarithmically uniform step). For each $h_1$, the altitude $h_2$ is selected from the same grid to be larger than $h_1$, then $h_3$ is selected to be larger than $h_2$. This process eliminates duplicate computation. For each altitude combination $\{h_i\}$ layer intensities are found by direct inversion of the linear equation (10) (the matrix $W$ is in fact inverted by singular-value decomposition). Negative $J_i$ that may result from this procedure are set to zero, and the $\chi^2$ that corresponds to the given $\{h_i\}$ is finally computed. When all altitude combinations are checked, the one with minimum $\chi^2$ is taken as a solution. Floating-layer restoration takes less than 1 s on a modern PC computer.

5.4 Example of profile restoration

The example of application of both methods to real data is shown in Fig. 3 (see also Kornilov et al. 2003). The consistency between successive profiles is immediately seen – convincing evidence that restoration is indeed successful. The changing character of turbulence is apparent: after 3 h UT, the floating-layer method does not give consistent layer altitudes, indicating that turbulence is likely to be distributed continuously. The two restoration methods are complementary to each other.

The residuals of three representative SIs to the fixed-layer model are plotted in Fig. 4 for the same data. It is typical that relative residuals are under 10 per cent most of the time. Some systematic deviations from zero (few per cent) are observed in $\sigma_m$, signifying that the actual structure of high-altitude turbulence is not perfectly represented by our simple model. The residuals to the floating-layer model are quite similar. Low residuals translate to low $\chi^2 / N$, typically under 10.

6 PERFORMANCE OF FIXED-LAYER RESTORATION

6.1 Numerical modelling: one and two layers

The fixed-layer restoration was studied by numerical simulations. For some input TP, we computed true SIs, added a realistic noise
Turbulence profile from scintillation indices

Figure 3. Turbulence profile evolution during several hours on 2002 September 11/12 as measured with MASS at Cerro Tololo (Chile): (a) fixed-layer method and (b) floating-layer method. The length of bars is proportional to the intensities of turbulent layers with the scale indicated on the right. This figure is available in colour in the online version of the journal on Synergy.

Figure 4. Relative residuals (observed minus modelled divided by modelled) of scintillation indices in apertures A, D and AB for the data in Fig. 3, fixed-layers restoration. This figure is available in colour in the online version of the journal on Synergy.

Figure 5. (a) Simulation of the response of the fixed-layers restoration to a single turbulent layer located at different altitudes (horizontal axis). The length of bars shows the reaction of slabs in relative units. The upper curve shows the sum of all reactions. (b) Same for two equal layers, one moving and another fixed at 4 km.

6.2 Realistic profiles

The fixed-layer restoration procedure was simulated with real turbulence profiles and was shown to give reasonable results (Tokovinin & Kornilov 2002). Here we repeat this exercise with realistic noise levels appropriate for MASS (see above). We use the same set of 12 real turbulence profiles measured by balloons at Cerro Paranal in Chile during the PARSCA campaigns (Fuchs & Vernin 1993); those profiles are plotted in Le Louarn et al. (2000).

Figs 6 and 7 illustrate the results for two very distinct cases. The scintillation indices were computed from real profiles (with 100-m vertical resolution) and artificially corrupted by noise. Then a restoration procedure was applied several times, with different noise realizations. True intensities of the ‘layers’ were computed by integrating the products of profiles and triangular response functions, to be compared with the results of restoration. Here we address two problems: (i) the reaction of the crude six-layer model to

that the total turbulence intensity (hence free-atmosphere seeing) is measured correctly for all layer altitudes. Noise in the restoration is barely noticeable.

A reaction of MASS to two layers is simulated in Fig. 5(b). Here the input TP consists of two equal layers, one fixed at 4 km and another with variable altitude. The overlapping layers are successfully separated, but the noise is more pronounced. However, the noise mostly tends to redistribute the intensity between adjacent slabs, with less effect on the total intensity.
Figure 6. (a) The turbulence profile number 50 as measured at Paranal (full line) and the idealized response functions of MASS in relative units (dotted lines). (b) The results of 10 restorations with independent noise realizations normalized by $J_{\text{tot}}$: the true profile as a full line and restored profiles as dotted lines.

Figure 7. Same as in Fig. 6 for Paranal profile number 51.

6.3 Noise estimates from real data

The performance of fixed-layer restoration was checked on real data by analysing the reproducibility of the measured profiles. We attempted to separate real variations of the turbulence strength occurring in each slab and the instrumental noise by means of covariance analysis. Let $J(t_i)$ be the strength of some layer (slab) as measured by MASS on moments $t_i$ (with a time resolution of 1 min). Then the auto-covariance $C_{J,J}$ is

$$ C_{J,J} = \frac{1}{N} \sum J(t_i)J(t_i + \Delta t), \quad (14) $$

where $N$ is the total number of samples entering into the sum. An example of the auto-covariance is given in Fig. 8.

The idea of noise estimation is to extrapolate $C$ linearly to zero using the first and second points. Then the difference with the actual value will be a rough estimate of the MASS noise variance:

$$ \sigma^2 = C_{J,0} - (2C_{J,1} - C_{J,2}). \quad (15) $$

It should be stressed that $\sigma^2$ is actually an upper limit for noise because it includes natural short-scale variations of the turbulence intensity.

The results of the calculations are given in Table 1. The nights selected for this analysis are different: the first two had an exceptionally calm upper atmosphere, 2002 July 10 was typical and 2002 July 24 was worse than average. The total number of profiles $N_{\text{prof}}$ and the average turbulence integral for each night $J_{\text{tot}}$ are given for reference. In one case the extrapolation method failed to give a non-negative noise estimate, which is not surprising for a statistical technique.

Table 1. Noise of measured layer intensities $\sigma_J$ in $10^{-13} \text{ m}^{1/3}$.

<table>
<thead>
<tr>
<th>Date (2002)</th>
<th>$N_{\text{prof}}$</th>
<th>$J_{\text{tot}}$</th>
<th>Layers (km)</th>
</tr>
</thead>
<tbody>
<tr>
<td>19/06</td>
<td>597</td>
<td>1.0</td>
<td>0.5 1 2 4 8 16</td>
</tr>
<tr>
<td>22/06</td>
<td>473</td>
<td>0.8</td>
<td>0.04 0.07 0.03 0.04 0.05 0.03</td>
</tr>
<tr>
<td>10/07</td>
<td>546</td>
<td>3.2</td>
<td>0.14 0.15 0.45 0.17 0.10 0.06</td>
</tr>
<tr>
<td>24/07</td>
<td>498</td>
<td>8.3</td>
<td>0.60 0.81 0.36 0.27 – 0.03</td>
</tr>
</tbody>
</table>

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The signal-dependent noise characteristics are confirmed by the actual analysis: noise is roughly proportional to \( J_{\text{tot}} \). It never exceeds 0.14\( J_{\text{tot}} \), and this is an upper limit because some fast turbulence variations are included. The absolute sensitivity of MASS is impressively high, especially at high altitudes. A noise-equivalent \( J = 0.03 \times 10^{-13} \) m\(^3\)/s represents a seeing of only 0.04 arcsec. The worst-case error encountered in Table 1 is \( 0.8 \times 10^{-13} \) m\(^3\)/s, or a seeing of 0.3 arcsec. It should remain clear that noise in MASS is not a bias, it must not to be subtracted from the measured layer strengths or seeing.

6.4 Optimization of the aperture geometry

The size of apertures in MASS was selected to match the Fresnel radii for turbulent layers at different altitudes. Now we can refine this choice and look for such an aperture geometry that would lead to most accurate profile restoration. This can be performed with the help of equation (13), under the approximation that input data are independent and the best least-squares restoration corresponds to this choice and look for such an aperture geometry that would lead to the best separation of turbulent layers by MASS apertures. However, the problem of aperture optimization does not have a unique solution. First, the optimum geometry depends on the desired vertical resolution, i.e. the number and altitudes of pre-defined slabs. Secondly, the noise in MASS depends on the turbulence profile.

We selected five fixed-layer altitudes at 1, 2, 4, 8 and 16 km. The noise was modelled as 2 per cent of scintillation indices (except \( s_{\text{AB}} \) where it was 4 per cent) for a fiducial profile with 0.6 of its energy at 1 km and the rest equally divided between 4 and 16 km. The wavelength of 450 nm and spectral bandwidth of 100 nm were assumed, close to real MASS parameters. For each trial aperture geometry the WFs and indices were found and then the noise of \( J_i \) was computed from equation (13). The maximum value of relative noise \( r_i = \sigma_i / J_{\text{tot}} \) was taken as a measure of restoration quality. We found that the best aperture diameters are 2, 3, 6 and 8.5 cm and that the maximum noise of \( r = 0.1 \) occurs at the 4-km layer. Optimum aperture geometry only weakly depends on the input turbulence profile. Restoration quality is quite forgiving of the deviations from the optimum aperture geometry; for the actual MASS apertures it is only marginally worse than for the optimal apertures.

7 DIRECT MEASUREMENT OF SEEING AND ISPLANATIC ANGLE

Seeing is related to the integral of \( C_n^2(h) \) over the whole atmosphere, \( J_{\text{tot}} \). MASS is not sensitive to ground-layer turbulence (unless a generalized mode is used), but it can measure the seeing produced by all higher layers, which we call free-atmosphere seeing. The FWHM of long-exposure stellar image in a large telescope, \( \epsilon \), is the generally accepted seeing characteristic. The standard turbulence theory provides one-to-one correspondence between \( \epsilon \) and the Fried parameter \( r_0 \). \( \epsilon = 0.98g \times r_0 \). The relation of \( r_0 \) to the turbulence integral \( J_{\text{tot}} \) is well known (Roddier 1981).

The free-atmosphere seeing at zenith is defined to be

\[
\epsilon_f = 5.307 \lambda^{-1/3} \left[ \int_0^\infty C_n^2(h)W_0(h) \, dh \right]^{3/5},
\]

where the function \( W_0(h) \) determines which layers contribute to the seeing; it is equal to 1 at high altitudes and drops to zero at ground level. The integration is performed over all altitudes, from the observatory level up. Free-atmosphere seeing is thus dependent on the somewhat arbitrary definition of \( W_0(h) \). With \( W_0(h) = 1 \) everywhere we obtain the standard relation of seeing to the zeroth moment of turbulence profile \( J_{\text{tot}} \) (Roddier 1981). The constant 5.307 = 0.98 \times 0.432^{3/5} \times (2\pi)^{3/5}.

The isoplanatic angle \( \theta_0 \) characterizes the size of the corrected field of view in adaptive optics (Fried 1982). It is related to the 2\( r \)d moment of turbulence profile. Putting \( W_5(h) = h^{5/3} \), the isoplanatic angle at zenith is

\[
\theta_0 = 0.1334 \left[ \int C_n^2(h)W_5(h) \, dh \right]^{-3/5},
\]

where 0.1334 = (2.91\pi^{3/5})^{-3/5}.

The turbulence moments can be obtained from SIIs directly, without profile restoration. Each SI is an integral over altitude of \( C_n^2 \) multiplied by some WF. Thus, any linear combination of SIs corresponds to the linear combination of WFs. If it approximates the desired moment-generating weights \( W_0 \) and \( W_5 \), the moments can be estimated from linear combinations of indices. For example, in order to estimate \( \epsilon_f \) we search for coefficients \( c_{m,n} \) such that

\[
W_0(z) \approx \sum_{m=1}^M c_{m,n} W_m(z),
\]

where we substitute altitude \( h \) with the propagation distance \( z \). When \( \epsilon_f \) and \( \theta_0 \) are computed, they are transformed to viewing at zenith by multiplication by \( (\cos \gamma)^{3/5} \) and \( (\sec \gamma)^{3/5} \), respectively. However, the lower cut-off of the free-atmosphere seeing weight \( W_0(h) \) does depend on \( \gamma \); somewhat lower layers are included into the \( \epsilon_f \) estimate for non-zenith viewing.

To obtain \( c_{m,n} \), we write equation (18) in matrix form for a discrete set of altitudes \( \{ z_i \} \) and solve for the vector \( c_{m,n} \) by singular-value decomposition of the matrix \( W_m(z_i) \). The same procedure applies for \( W_5(z) \).

The coefficients \( c_{m,n} \) for moment estimation obtained by this procedure depend on the choice of altitude grid and the set of indices that enter into equation (18). Of course, we can use all measured indices and require the approximation to be as precise as possible. However, the coefficients thus obtained are large and of both positive and negative signs. In computation of moments the errors of index measurements would increase to the point that even negative moments can sometimes be obtained.

To avoid noise amplification, we restrict both the set of indices used for moment computation and the finesse of the altitude grid. For free-atmosphere seeing we take only the indices \( s_{AB} \), \( s_B \), and \( s_{AB} \) and select the coarse altitude grid: 1, 2, 4, 8, 16 km. This results in a very reasonable \( W_0(h) \) (Fig. 9) with the coefficients of \(-0.06, 0.09\)

![Figure 9. Ratio of the MASS weighting functions for free-atmosphere seeing and isoplanatic angle W(AB) (full line) and isoplanatic angle W(5) (dashed line) to the ideal \( h^5 \) and \( h^{5/3} \) laws, respectively.](https://academic.oup.com/mnras/article-abstract/343/3/891/1123465/12465)
of turbulent layers. This technique is based on photometry; the only calibration parameters needed for correct data interpretation are the sizes of apertures, the spectral response of the instrument and the characteristics of photon counters. Examples of real data reductions are given.

MASS is an inexpensive instrument that uses only a small telescope. It is well adapted for TP monitoring at existing observatories and at new sites. We have started to use this technique for building extensive data base of TPs at several sites (e.g. Tokovinin et al. 2003). When these data are complemented with meteorological information and three-dimensional computer modelling of turbulence, a better understanding of the ‘seeing’ phenomenon will result, with important practical consequences for site selection. Another use of MASS will be to support the operation of adaptive-optics systems that can be enhanced or made more efficient if TP is known in real time.

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APPENDIX A: DERIVATION OF EQUATIONS (5) AND (6)

Here we derive the equations (5) and (6). Let $x$ be the individual photon counts in some aperture, $B$ – average background flux in the same aperture. The scintillation index $s$ cannot be computed from

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its definition (1) by simply replacing the true flux \( X \) with photon counts \( x \), because counts fluctuate even at constant flux owing to the random nature of photodetection. The photon-noise variance is equal to the mean value \( \bar{X} \) for an ideal detector. For real PMT it is \( p \bar{X} \), where the non-Poisson factor \( p \) can be slightly more than 1 for many reasons, e.g. occasional after-pulses. Photon noise is independent from scintillation and must be subtracted from the measured variance. So, in order to obtain SI we compute the variance of photon counts \( \rho_{x,x,0} \) (equation 3), subtract photon noise and normalize to the square of the mean stellar flux \( \bar{X} - B \):

\[
s_{x,0} = \frac{\rho_{x,x,0} - p\bar{X}}{(\bar{X} - B)^2}.
\]  

(A1)

The notation \( s_{x,1} \) reminds one that the SI is computed for a 1-ms sampling time and is not yet corrected to zero microexposure.

Photon noise in any two channels \( x \) and \( y \) is not correlated and does not distort the covariance \( \rho_{x,y,0} \). It follows from equation (2) that

\[
s_{x,y,1} = s_{x,1} + s_{y,1} - 2\frac{\rho_{x,y,0}}{(\bar{X} - B_x)(\bar{X} - B_y)}.
\]  

(A2)

What remains is the extrapolation of indices to zero exposure time. This can be performed by binning the data and computing the index for double microexposure \( s_{x,2} \). The SI for zero exposure \( s_{x,0} \) can then be estimated by linear interpolation,

\[
s_{x,0} = 2s_{x,1} - s_{x,2},
\]  

(A3)

which was shown to be a reasonable approximation (Tokovinin 2002). Such a correction is really important for the differential SIs with the smallest MASS apertures, where 1 ms is not short enough to ‘freeze’ fast layers (for a 30 m s\(^{-1}\) wind speed the layer displacement is 3 cm, larger than the 2-cm aperture). For large apertures, the finite duration of microexposure is not important, although we still apply the correction.

Instead of actually binning the counts and repeating all calculations, we take advantage of pre-computed covariances with time lag. Let \( \sigma^2_z \) be the variance of doubly binned data, \( \sigma^2_1 \) – the variance of unbinned data, and \( \rho_{x,1} \) – the covariance with time lag 1. It can be easily shown that

\[
\sigma^2_z = 0.5(\sigma^2_1 - \rho_{x,1}).
\]  

(A4)

The variance for zero exposure \( \sigma^2_0 \) is extrapolated as \( \sigma^2_0 = 1.5\sigma^2_1 - 0.5\rho_1 \). Now the transition from 1-ms indices as given by equations (A1) and (A2) to 0-ms indices (equations 5 and 6) is straightforward.

**APPENDIX B: FORMULA FOR WEIGHTING FUNCTIONS**

Weighting functions for polychromatic scintillation were derived in Tokovinin (2003). Here we give the computing formulae without derivation. The WF \( W(z) \) is an integral over spatial frequency \( f \):

\[
W(z) = 9.62\int_0^{\infty} f^{-8/3}A(f)S(z, f)\,df,
\]  

(B1)

where the factor \( S(z, f) \) depends on the propagation range \( z \), wavelength \( \lambda \) and the normalized spectral response \( F(\lambda) \) (in photon nm\(^{-1}\) for a photon-counting detector):

\[
S(z, f) = \left[ \int \lambda^{-1}F(\lambda)\sin(\pi\lambda zf^2)\,d\lambda \right]^2.
\]  

(B2)

It can be seen that \( S(z, f) \) is equal to the squared imaginary part of the Fourier transform (FT) of \( F(\lambda)/\lambda \). For monochromatic spectral response it reduces to the known \( \lambda^{-2}\sin^2(\pi\lambda zf^2) \) term, whereas for a real response we compute it using a FT.

The aperture factor \( A \) is the square modulus of the FT of the normalized aperture transmission function \( P(x) \) or of their difference:

\[
A(f) = |\hat{P}(f)|^2 \quad \text{for normal SI}
\]  

(B3)

\[
A(f) = |\hat{P}_1(f) - \hat{P}_2(f)|^2 \quad \text{for differential SI}.
\]  

(B4)

For a circular aperture of diameter \( D \) with a relative central obscuration \( \epsilon \)

\[
P(f) = \frac{1}{1 - \epsilon^2} \left[ \frac{2J_1(\pi D f)}{\pi D f} - \epsilon^2 \frac{2J_1(\epsilon \pi D f)}{\epsilon \pi D f} \right].
\]  

(B5)

For a special ‘quasi-Gaussian’ spectral response with FWHM \( \Lambda \) and central wavelength \( \lambda_0 \) the WF is

\[
W(z) = 9.62\lambda_0^{-2}\int_0^{\infty} df f^{-8/3}A(f)\exp(-1.780\varepsilon^2 f^4\Lambda^2)
\times \sin^2(\pi\lambda_0 zf^2).
\]  

(B6)

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