

GSTM1 and *CYP1A1* Polymorphisms, Tobacco, Air Pollution, and Lung Cancer: A Study in Rural Thailand

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Abstract

Objectives: The Lampang Province is situated in the northern region of Thailand. Incidence rates of lung cancer are high for Asian standards, particularly in women. This study was conducted to quantify the risk of lung cancer associated with exposures prevalent in the area and to investigate possible interactions with genetic susceptibility. The presence of several large open-cast coal mines from 1955 close to electricity-generating plants was a particular focus of concern.

Methods: Two-hundred and eleven cases of primary lung cancers diagnosed in 1993 to 1995 and residents in the province were recruited at the Lampang Provincial Hospital (main referral center for treatment of the disease). Two sets of controls, frequency-matched to the cases by sex and age, were recruited (a) from the resident population (202 interviewed) and (b) from patients admitted to the hospital for diseases predominantly unrelated to tobacco smoking (211 interviewed). Sociodemographic information, complete residential history, and characteristics of the household (place of cooking, cooking fuel, and heating fuels), occupational history, and history of tobacco smoking were obtained by interview. Cases and controls (~50% of the population-based series) provided a blood sample. A point source air pollution exposure index was calculated for each village/township reported in residential histories based on the linear distance from the Mae Moh Center (the area of the electricity-generating plants), the year-specific gaseous (SO₂ and NO₂) or total suspended particulate emissions from the Mae Moh Power Plant, and the percentage of wind from the center. Odds ratios (OR) for the disease associated with categorical variables were estimated within unconditional logistic regression. Extraction of genomic DNA and genotyping of variants in *CYP1A1* and *GSTM1* were conducted to

assess the extent of modification of risk by these genes that are involved in the metabolism of polycyclic aromatic hydrocarbons, a common component of the exposures.

Results: Overall, there was no evidence of relevant differences in the socioeconomic level of the three groups. The two control sets were similar with respect to lifelong tobacco habit and were subsequently pooled in analyses. Never-smokers were 7% of men and 33% of women. Smoking of local traditional products unfiltered and high in tar content is a common habit in the rural female population. ORs associated with smoking increased with duration of the habit and average daily amount, being 4.9 [95% confidence interval (95% CI), 2.5-9.7] for smokers of ≥7 cigarettes/d and 3.3 (95% CI, 1.7-6.2) for duration of 41 years or longer compared with nonsmokers. Smoking of local products was associated with an independent OR of 3.1 (95% CI, 1.7-5.6) adjusted for lifelong cumulative amount of tobacco smoked. Although most smokers had the habit for at least 16 years, the daily consumption was low compared with Western standards. Other potential sources of exposure to lung carcinogens (emission from the power-generating plants and domestic burning of coal and wood for cooking and heating) were not associated with increased risk of lung cancer. None of the three polymorphisms examined increased the risk of lung cancer or modified the risk associated with tobacco smoking. **Conclusion:** In this rural population, 96% of male and 64% of female lung cancer incidence were explained by tobacco smoking. None of the potential sources of air pollution deriving from the combustion of coal and wood, or polymorphisms in the *CYP1A1* gene or deletion of the *GSTM1* had an effect on the risk of lung cancer, either together or separately. (Cancer Epidemiol Biomarkers Prev 2006;15(4):667-74)

Introduction

Overwhelming evidence of the role of tobacco smoking in the occurrence of lung cancer has been accumulated since the 1960s (1). Quantitatively, it accounts for 37% of all cancer death in men and 9% in women in high-risk countries (2). Although declines in the prevalence of smoking are observed among men living in Western countries, rapid increases are reported in many developing countries where it remains an uncommon habit only in women (1). Further, many carcinogens of the

respiratory tract have been identified as well as environmental sources of exposure to these agents.⁵ These account for a relatively small proportion of the burden relative to tobacco in regions with widespread smoking but may have a significant effect in populations where tobacco smoking is a recent or rare habit. A high proportion of lung cancer is attributable to known environmental causes, raising the question of putative modulation of risk by genetic variants in metabolic genes (3, 4). Several polymorphisms in genes implicated in the metabolism of lung carcinogens have been identified and variation in their prevalence has been shown across ethnic groups (5).

The Lampang Province is situated in the northern region of Thailand. It contains several large open-cast coal mines opened from 1955 onward in proximity to functioning electricity-generating plants. The first started production in 1978 and

Received 8/30/05; revised 1/14/06; accepted 1/31/06.

Grant support: Electricity Generating Authority of Thailand and Cancer Research UK.

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doi:10.1158/1055-9965.EPI-05-0667

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by the enzyme resulting in fragments of 163 and 32 bp. DNA carrying the variant is not cut, resulting in the presence of the 195 bp amplified fragment. One microliter of extracted DNA (~25-50 ng) is amplified using 0.625 unit AmpliTaq DNA polymerase in the supplied buffer containing 1.5 mmol/L MgCl₂, 200 μmol/L deoxynucleotide triphosphates, and 50 ng of forward and reverse primers (CYPIA1E7-F, 5'-GAACTGC-CACTTCAGCTGTCT; CYPIA1E7-R, 5'-CCAGGAAGAGAAA-GACCTCCCAGCGGGCCA) in a 50 μL reaction. PCRs were done on an ABI9600 Thermal cycler at an annealing temperature of 59°C for 34 cycles and the PCR products were digested with 5 units NcoI in the supplied restriction buffer (New England Biolabs). DNA fragments were then visualized on an 8% Tris-borate EDTA acrylamide gel.

Index of Exposure to Emissions from the Electricity-Generating Plants. The Ecology and Environment Division, Survey and Ecology Department, Electricity Generating Authority of Thailand provided historical information on wind speed and direction and measurements of air pollution. A point source air pollution exposure index was calculated for each village/township reported in residential histories based on three variables: linear (*D*) distance from the Mae Moh Lignite Center (the area of coal exploitation and electricity production); the year-specific gaseous (SO₂ and NO₂) or total suspended particulate emissions from the Mae Moh Power Plant (*E*_{year}) for years 1978 (first year of activity of the plant) to 1994; and the percentage of wind coming from the township direction (*T*_w) based on its angular distance from the Mae Moh Center. The point source air pollution index (*P*_e) was computed according to the following formula, which was derived from Brown et al. (8) and Buffler et al. (9):

$$P_e = \sum_{\text{year}} E_{\text{year}} \times 1/D^{1/2} \times T_w$$

Places of residence in surrounding provinces were estimated to be at the highest concentration of towns within the boundaries of the map that covered the whole of Lampang Province but only part of the neighboring provinces. Emission data were only available for years 1985 to 1994. Data for SO₂, NO₂, and total suspended particulate emission were extrapolated to years 1978 to 1984. Average percentage of wind in 16 directions was calculated using 1991, 1993, and 1994 wind rose diagrams. The percentages were then multiplied by each of the four wind speed measurements in each of the four main directions for each of the 3 years and then averaged. A cumulative pollution exposure index was calculated for each subject, summing point exposures across all places of residence reported in the residential history from 1978 using the above formula. For periods during which subjects lived outside Lampang province and outside neighboring districts within the relevant time frame of exposure, emissions were set to zero.

Index of Exposure to Domestic Fumes. Exposure to domestic fumes was estimated based on the type of fuel used to heat the residence and to cook (defined as coal or wood versus no heating or gas/electrical supplies), taking into account whether cooking was inside or outside, and the number of years living in this condition. These variables were combined in an exposure index calculated as the total time (years) spent using polluting fuel (coal or wood). Years of exposure due to cooking by coal or wood were halved if cooking was done outdoors.

Smoking History. The smoking histories covered different periods if changes in the number of cigarettes smoked per day or type of tobacco products occurred. Cigarette types were classified as filtered or unfiltered commercial products and local traditional hand-made khii yo and yamuan, both unfiltered. Traditional cigarettes are composed of dry uncured tobacco rolled in a dry banana leaf and seasoned with tamarind bark.

The lifelong cumulative amount of tobacco smoked was calculated as follows:

$$\sum_j \sum_i (\text{daily number})_{ij} \times (\text{days})_{ij}$$

where *i* is the type of product and *j* is the period.

To allow for their bigger size, daily amounts of khii yo were multiplied by 1.5.

Statistical Analyses. Odds ratios (OR) for the disease associated with categorical variables were estimated with unconditional logistic regression. Continuous variables were categorized by percentiles of the distribution of population controls. The overall population-attributable risk was calculated as the weighted sum of age- and sex-specific attributable risks (10) weighted for the proportion of cases in the sex-age stratum according to the incidence recorded by the local cancer registry (6). Weights were the sex-specific proportions of cases in each age group. The observed genotype distributions for the CYPIA1 variants were tested for consistency with Hardy-Weinberg equilibrium among controls (only two phenotypes were available for GSTM1, making such a test uninformative).

Stata version 8 (Stata Corporation, College Station, TX) was used for all statistical analyses using standard methodology. Haplotype frequencies were estimated using the hapipf instruction.

Results

Table 1 compares demographic and exposure characteristics of the three study groups. Three hundred and ninety (63%) of the study population were male. Cases and controls were similar with respect to level of education. Over 95% of all three sets reported speaking the native language, being of Buddhist religion, and married or widowed (data not shown). Over 70% of cases and controls were farm laborers and 17% worked in service activities (drivers, street vendors, maids, and cooks). Less than 2% of either cases or controls were used in clerical or professional jobs. Overall, there was no evidence of relevant differences in the socioeconomic level of the three groups.

Four percent and 13% of the male and female cases, respectively, had never smoked. The prevalence of never-smokers in women was greater than in men in both control sets, 33% and 37% in hospital and population controls, respectively, versus 10% and 6%. Eleven percent of the smokers consumed only manufactured commercial cigarettes, whereas all others reported mixed consumption of manufactured and traditional products (88%). The consumption of commercial cigarettes was strongly associated with sex (being 1% in women versus 13% in men) and age in the two control sets (55% of controls up to age 55 years and 20% in age group ≥72 years). Fifty-one percent of hospital controls reported being ex-smoker compared with 38% in the population-based set. The mean duration of regular smoking among smokers was 32 years in population controls and 28 years in hospital controls. The average number of cigarettes consumed per day was low compared with Western levels with less than one third of the controls consuming ≥7 cigarettes/d. The median consumption in this upper category was 10 cigarettes/d, cases and controls pooled. The median daily consumption among long-term smokers (41+ years) was 5 cigarettes/d. The two control sets did not differ by cumulative amount of cigarettes smoked (*P* = 0.58). All controls were therefore pooled in the subsequent analyses except in those involving air pollution in which case only population controls were used.

Chewing of Betel nut was rare in men (overall 4%) and thrice more popular in women (overall 15%).

Table 2 shows ORs associated with tobacco smoking adjusted for sex and age. The risk of lung cancer increased with increasing duration of the habit and reached an OR of 3.3

[95% confidence interval (95% CI), 1.7-6.2] for >40 years of regular smoking compared with nonsmokers. Relatively low amounts of daily consumption were reported with only 14% of all smokers estimating consumption greater than the equivalent of 10 cigarettes/d. The ORs associated with ≥ 8 cigarettes/d was 4.9 (95% CI, 2.5-9.7). Seventy-eight percent of the smokers consumed the unfiltered traditional products, which were associated with an OR of 2.9 (95% CI, 1.7-5.2). Lifelong cumulative number of cigarettes smoked satisfactorily described the risk associated with tobacco smoking for various patterns of duration and amount. The OR increased regularly up to 4.9 (95% CI, 2.5-9.4) with increasing lifelong cumulative consumption. This variable was used to control confounding by tobacco in all the following analyses. The exclusion of 33 hospital controls admitted for cardiovascular or respiratory diseases did not materially modify the risks described (Table 2). When risk was analyzed separately by sex, ORs associated with cumulative exposure appeared greater in women compared with men although this difference was not statistically significant ($P = 0.18$).

Squamous cell carcinoma were more strongly associated with tobacco smoking than adenocarcinomas (Table 3). For squamous cell carcinoma, the trend of increasing OR was significant and reached an OR of 6.8 (95% CI, 2.8-16.8) in the higher category of exposure. There was a nonsignificant trend ($P = 0.13$) in the ORs among adenocarcinoma cases with increasing levels of exposure but none of the smoking categories was statistically significant. The risks in the group containing the remaining small cell, large cell, and mixed types, were of the same magnitude as those of squamous cell carcinoma and had similar significance.

None of the modestly elevated risks associated with the two indices of exposure to air pollutants from the Mae Moh area reached statistical significance (Table 4). ORs were slightly reduced when adjusted for tobacco smoking, indicating some confounding. Among never-smokers, the OR associated with nondomestic air pollution was more elevated but not statistically significantly (OR, 4.0; 95% CI, 0.7-23.0; Table 4). Cumulative exposure to domestic fumes from burning coal or wood was not associated with excess risk overall nor among

Table 1. Distribution of cases and controls by selected characteristics

| | Cases | Hospital controls | Population controls | P of log-likelihood ratio test* |
|---|-------|-------------------|---------------------|---------------------------------|
| No. subjects | 211 | 211 | 197 | |
| Men (%) | 66.4 | 59.7 | 62.9 | NS |
| Education (%) | 100 | 100 | 100 | |
| Illiterate | 33.2 | 38.9 | 35.1 | |
| Primary school | 64 | 59.7 | 61.9 | NS |
| Vocational and high school | 2.8 | 1.4 | 3.0 | |
| Occupation (%) | 100 | 100 | 100 | |
| Agricultural workers | 71.6 | 70.1 | 73.3 | |
| Service workers† | 16.6 | 20.9 | 13.3 | |
| Professional and clerical | 1.9 | 0.9 | 0.6 | NS |
| Industry crafts and trade | 4.3 | 3.8 | 9.4 | |
| Other | 5.7 | 4.3 | 3.3 | |
| Smoking status (%) | 100 | 100 | 100 | |
| Never | 7.1 | 19.0 | 16.8 | |
| Occasional | 0.5 | 0.9 | 2.0 | |
| Ex-smoker | 48.3 | 50.7 | 37.6 | <0.01 |
| Current | 44.1 | 29.4 | 43.6 | |
| %Nonsmokers | | | | |
| Men | 4.3 | 10.3 | 7.1 | NS [‡] |
| Women | 14.1 | 34.1 | 39.7 | <0.01 [‡] |
| Years of regular smoking [§] | | | | |
| 1-15 | 2.1 | 5.3 | 6.1 | |
| 16-40 | 48.2 | 58.0 | 43.9 | 0.02 |
| 41+ | 49.7 | 36.7 | 50.0 | |
| Average cigarettes/d [§] | | | | |
| 1-3 | 30.1 | 32.1 | 45.0 | |
| 4-7 | 24.4 | 32.1 | 26.9 | <0.01 |
| 7+ | 45.6 | 35.7 | 28.1 | |
| %Smokers of manufactured cigarettes only | | | | |
| Men | 7.9 | 15.9 | 21.0 | <0.01 [‡] |
| Women | 5.6 | 3.5 | 4.1 | NS [‡] |
| %Smokers consuming commercial cigarettes by age | | | | |
| Age (y) | | | | |
| <55 | 16.1 | 44.6 | 41.8 | |
| 56-64 | 32.3 | 29.2 | 29.9 | $P < 0.01$ |
| 65-71 | 25.8 | 15.4 | 20.9 | |
| 72+ | 25.8 | 10.8 | 7.5 | |
| %Regular chewers of miang [¶] | | | | |
| Men | 71.4 | 72.2 | 72.6 | NS |
| Women | 81.8 | 62.3 | 72.5 | 0.03 |
| %Regular chewers of betel nut** | | | | |
| Men | 2.1 | 5.6 | 4.8 | NS |
| Women | 12.7 | 14.1 | 17.8 | NS |

Abbreviation: NS, not significant.

*Test for independence of case/control status and variable on the row unless otherwise specified.

†Includes street vendors.

‡P for the difference of the three proportions within group.

§Smokers excluded.

||P for trend by age.

¶Mixture of fermented tea leaves, coconut, peanut, and sugar.

**Betel nut, tobacco, red lime, and fresh leaf of piper betel.

Table 2. ORs associated with selected variables describing tobacco habit

| | No. cases | No. controls | All controls, OR* (95% CI) | CVD and respiratory diseases excluded (33 hospital controls), OR* (95% CI) |
|--|-----------|--------------|----------------------------|--|
| Smoking status | | | | |
| Never | 16 | 80 | 1 | |
| Ex-regular smokers | 102 | 183 | 2.6 (1.4-4.8) | 2.7 (1.5-5.0) |
| Current regular smokers | 93 | 145 | 3.2 (1.7-6.0) | 3.0 (1.6-5.6) |
| Duration, y | | | | |
| 1-15 | 4 | 19 | 0.8 (0.2-3.0) | 1.1 (0.3-3.8) |
| 16-40 | 94 | 167 | 2.9 (1.6-5.3) | 2.9 (1.6-5.4) |
| 41+ | 97 | 142 | 3.3 (1.7-6.2) | 3.0 (1.6-5.8) |
| Average no. cigarettes/d | | | | |
| 1-3 | 58 | 126 | 2.2 (1.2-4.2) | 2.2 (1.1-4.1) |
| 4-7 | 47 | 97 | 2.5 (1.2-5.0) | 2.6 (1.3-5.4) |
| 7+ | 88 | 105 | 4.9 (2.5-9.7) | 4.7 (2.4-9.3) |
| Types of cigarettes | | | | |
| Manufactured filtered only | 9 | 26 | 1.6 (0.6-4.2) | 1.6 (0.6-4.1) |
| Only traditional and manufactured unfiltered | 158 | 245 | 2.9 (1.7-5.2) | 2.9 (1.7-5.1) |
| Smokers of all types | 26 | 51 | 2.4 (1.2-4.9) | 2.4 (1.1-5.0) |
| Lifelong cumulative tobacco smoked (000s) | | | | |
| <60 | 37 | 100 | 1.9 (1.0-3.8) | 1.9 (1.0-3.6) |
| 61-126 | 56 | 110 | 2.8 (1.4-5.3) | 2.8 (1.4-5.4) |
| 127+ | 100 | 118 | 4.9 (2.5-9.4) | 4.9 (2.5-9.5) |
| Ever traditional products | | | | |
| Smokers of manufactured products only | 15 | 52 | 1.4 (0.6-3.1) | 1.3 (0.6-3.0) |
| Smokers of traditional products | 180 | 276 | 3.1 (1.7-5.6) | 3.1 (1.7-5.5) |

Abbreviation: CVD, cardiovascular disease.

*Reference category for all variables listed are never-smokers. ORs are adjusted for sex and age.

nonsmokers (Table 4). None of the individual variables contributing to the cumulative index was significantly associated with the disease (details not shown).

The risk of lung cancer associated with smoking was analyzed in combination with the three genetic variants examined in this study. These variants had similar frequency distributions in the two control sets and their pooled distributions were in Hardy-Weinberg equilibrium (Table 5). The *GSTM1*-/- genotype occurred in about two thirds of the population, one third were homozygous carriers of the *CYP1A1**2A variant, whereas homozygous carriers of the *CYP1A1**2C variant were 6% and 8% in cases and controls, respectively. None of the variants was associated with the disease or with tobacco smoking. The possible effect modification of the genotypes on the risk due to tobacco smoking is reported in Table 6. This shows that the magnitude of the tobacco-associated risks is constant across the three genotypes examined, which, therefore, do not show any evidence of acting as modifiers. The same results were found when cases were restricted to squamous cell carcinoma and to adenocarcinoma (details not shown).

Based on the figures in Table 2, the proportion of the lung cancer incidence attributable to tobacco smoking was 98% and 64% of the cases were attributable to tobacco in men and women, respectively.

The frequency of *CYP1A1* haplotypes was estimated from both control groups; there was clear evidence of strong linkage

disequilibrium between the *CYP1A1**2A variant and the *CYP1A1**2C variant ($P < 0.0001$) but no evidence for any difference in haplotype frequencies between the two control groups ($P = 0.62$). Finally, there was no evidence of any haplotype frequency differences between cases and controls ($P = 0.32$). For the combined control group, the estimated haplotype frequencies were 0.42 for the wild-type haplotype (Ile for *CYP1A1**2C, T for *CYP1A1**2A), 0.26 for the haplotype carrying the *CYP1A1**2A variant only (Ile, C), 0.02 for the *CYP1A1**2C variant only (Val, T), and 0.30 for the haplotype carrying both variants (Val, C), which is called the *CYP1A1**2B allele.

Discussion

The study presented was motivated and designed to evaluate the causes of the relatively high incidence of lung cancer in a Thai province where several installations of electricity-generating plants coal powered were a source of public concern. The area is largely rural as shown by the high prevalence of agriculture workers and the very low number of people used in industrial activities (Table 1). Pollution from industrial sources other than the one investigated as well as occupational exposures were judged therefore negligible. Road traffic is extensive only along the main provincial roads but certainly sparse in the countryside where most of the population lives.

Table 3. ORs and 95% CIs associated with tobacco smoking by histologic type (population and hospital controls pooled)

| | Squamous cell carcinoma | Adenocarcinoma | Other and unknown |
|--|-------------------------|----------------|-------------------|
| Never or occasional* No. cases/controls | 1 | 1 | 1 |
| Lifelong cigarettes smoked (000s) | | | |
| <60 | 2.7 (1.1-6.6) | 0.80 (0.2-2.6) | 3.1 (0.6-15.3) |
| 61-127 | 3.1 (1.3-7.8) | 1.3 (0.3-4.0) | 5.7 (1.2-26.8) |
| 128+ | 6.8 (2.8-16.8) | 2.5 (0.8-7.3) | 7.4 (1.5-36.4) |
| Test for trend, <i>P</i> | 0.00 | 0.13 | 0.01 |

NOTE: ORs were adjusted for sex, age, and both.

*Reference category.

Table 4. ORs and 95% CIs associated with estimated exposure to air pollution from the Mae Mo Center and with domestic sources of fumes

| | No. cases | No. controls | OR* (95% CI) | OR [†] (95% CI) |
|--|-----------|--------------|----------------|--------------------------|
| Cumulative index of exposure to SO ₂ or NO ₂ , tons/√km/y | | | | |
| <1,808 | 54 | 62 | 1 [‡] | 1 [‡] |
| 1,808-3,507 | 86 | 72 | 1.4 (0.8-2.2) | 1.3 (0.8-2.1) |
| >3,507 | 71 | 66 | 1.2 (0.8-2.1) | 1.2 (0.7-2.0) |
| Cumulative index of exposure to suspended particulate, tons/√km/y | | | | |
| <214 | 56 | 61 | 1 [‡] | 1 [‡] |
| 214-392 | 84 | 70 | 1.3 (0.8-2.1) | 1.2 (0.7-2.0) |
| >392 | 71 | 67 | 1.1 (0.7-1.9) | 1.1 (0.7-1.8) |
| Cumulative index of exposure to domestic fumes, y | | | | |
| <9 | 51 | 50 | 1 [‡] | 1 [‡] |
| 9-14 | 67 | 50 | 1.3 (0.8-2.2) | 1.3 (0.7-2.2) |
| 15-20 | 43 | 50 | 0.8 (0.5-1.5) | 0.8 (0.4-1.4) |
| 21+ | 50 | 49 | 1 (0.6-1.7) | 0.8 (0.5-1.5) |
| Nonsmokers only | | | | |
| Cumulative index of exposure to SO ₂ or NO ₂ /suspended particulate [§] | | | | |
| <1,808/<214 | 3 | 18 | 1 [‡] | |
| ≥1,809/≥215 | 13 | 26 | 3.4 (0.7-23.3) | |
| Cumulative index of exposure to domestic fumes, y | | | | |
| <15 | 11 | 30 | 1 [‡] | |
| ≥15 | 5 | 14 | 0.4 (0.1-2.0) | |

NOTE: ORs were adjusted for sex, age, and cumulative amount of cigarettes smoked. Only population controls were included.

*Adjusted for sex and age.

[†]Adjusted for sex, age, and cumulative amount of cigarettes smoked.

[‡]Reference category.

[§]The two indexes coincide in this subgroup.

The main sources of potential air pollutants identified were the domestic combustion of coal or wood and the power-generating plants. Information on place of cooking and type of fuel used for cooking and heating was obtained by questionnaire. These variables were considered as proxy estimates of exposure to domestic fumes. Nonexposed were subjects who did not use these types of fuel.

We also tested three polymorphisms in two genes involved in the metabolism of lung carcinogens, the *GSTM1* null/null genotype and the two *CYP1A1**2A and *CYP1A1**2C variants. We also examined the *CYP2D6**3, *CYP2D6**4, and *CYP2D6**5 alleles, the common inactive alleles of *CYP2D6* among Caucasians, but found them to be rare in this population and so these are not reported here. *GSTM1* and *CYP1A1* are involved in the metabolism of polycyclic aromatic hydrocarbons. These compounds are environmental contaminants formed during combustion of organic material and are, therefore, particularly relevant in our study where the major

sources of pollution derived from tobacco smoking and the burning of coal and wood. Besides tobacco smoking, which alone explained 96% of male and 64% of female incidence, none of the other environmental factors analyzed was associated with the risk of lung cancer. No excess risk was associated with the cumulative index of exposure to exhausts from the Mae Moh plants, or duration of residence by distance from the production area. The relative risks of tobacco smoking that we found are comparable with those measured in many other populations and reflect the cumulative smoking experience of this community. Similarly, low risks for comparable levels of exposure were reported in studies conducted in China and Japan (11-15). The consumption of Western-style manufactured cigarettes is low in this rural region of Thailand where consumers prefer local traditional products consisting of chopped dry tobacco uncured, wrapped in a dry banana leaf. They are unfiltered and larger in size than standard cigarettes. The temperature of mainstream smoke is

Table 5. ORs and 95% CIs associated with polymorphisms in selected metabolic genes

| | No. cases | No. controls | OR* (95% CI) | OR [†] (95% CI) |
|--|-----------|-----------------|---------------|--------------------------|
| <i>GSTM1</i> | | | | |
| -/+ or +/+ [‡] | 67 | 105 | 1 | 1 |
| -/- | 99 | 184 | 0.9 (0.6-1.3) | 0.8 (0.5-1.2) |
| All | 166 | 283 | | |
| <i>CYP1A1</i> *2A SNP (rs4646903) | | | | |
| T/T [‡] | 26 | 53 | 1 | 1 |
| T/C | 87 | 155 | 1.2 (0.7-2.1) | 1.2 (0.7-2.0) |
| C/C | 55 | 78 | 1.5 (0.8-2.8) | 1.5 (0.8-2.7) |
| All | 168 | 286 | | |
| Hardy Weinberg equilibrium test [§] | | <i>P</i> = 0.15 | | |
| <i>CYP1A1</i> *2C SNP (rs1048943) | | | | |
| Ile/Ile [‡] | 78 | 135 | 1 | 1 |
| Ile/Val | 79 | 129 | 1.1 (0.7-1.6) | 1.0 (0.7-1.5) |
| Val/Val | 10 | 23 | 0.7 (0.3-1.6) | 0.8 (0.3-1.7) |
| All | 167 | 287 | | |
| Hardy-Weinberg equilibrium test [§] | | <i>P</i> = 0.33 | | |

*Adjusted for sex and age. Two control sets pooled.

[†]Adjusted for sex, age, and lifelong total tobacco smoked. Two control sets pooled.

[‡]Reference category.

[§]Exact probability.

Table 6. Effect of polymorphism in selected genes on the effect of lifelong amount of tobacco smoked

| | <i>GSTM1</i> | |
|---------------|------------------|------------------|
| | + / + or - / + | - / - |
| Never-smokers | 1* | 0.8 |
| Medium | 3.1 [†] | 1.8 |
| High | 2.9 [†] | 3.2 [†] |
| Controls (%) | 105 (37%) | 184 (64%) |

| | <i>CYP1A1*2A</i> | | |
|---------------|------------------|-----------|------------------|
| | T/T | T/C | C/C |
| Never-smokers | 1* | 1.7 | 0.5 |
| Medium | 2.0 | 3.8 | 3.7 |
| High | 4.8 | 3.6 | 6.0 [§] |
| Controls (%) | 53 (19%) | 155 (54%) | 78 (27%) |

| | <i>CYP1A1*2C</i> | | |
|---------------|------------------|------------------|-----------------|
| | Ile/Ile | Ile/Val | Val/Val |
| Never-smokers | 1* | 0.9 | NA [‡] |
| Medium | 1.6 | 3.0 [†] | 1.8 |
| High | 3.8 [†] | 2.5 | 2.5 |
| Controls (%) | 213 (74%) | 208 (73%) | 33 (12%) |

Abbreviation: NA, not applicable.

*Reference category.

[†]*P* < 0.05.[‡]Six controls but no cases occurred in this stratum.

high so that inhalation is likely to affect predominantly the bronchus and upper sections of the lungs. The high proportion of smokers in women is remarkable particularly by comparison with other Asian countries. Smoking in women is, however, a traditional habit rather than a recently acquired sign of Westernization; >92% of smoking women in fact had smoked only traditional products and had done so for over 15 years.

The *CYP1A1* gene encodes phase I enzymes that convert many lung carcinogens into DNA-binding metabolites with carcinogenic potential (16). The *GST* family produces phase II enzymes that detoxify reactive electrophils that act as mutagens (17). Variants in these genes may, therefore, modulate their action and result in varying susceptibility to develop cancer. Studies of lung cancer and polymorphisms in the *CYP1A1* gene were reviewed by D'Errico et al. (5), Houlston (18), and Bartsch et al. (19). Studies in Japan have consistently shown increased risks associated with the two variants we analyzed, particularly for squamous cell carcinoma. Excess risk confined to squamous cell carcinoma was found in three ethnic groups of Hawaii (20) and one of three case-control studies in China (21) the other two being negative for an association (22, 23). In a Swedish study (24), neither variant was associated with any subtype of lung cancer. Hong et al. (25) reported a significant deficit of the *CYP1A1*2C* polymorphism in Korean lung cancer cases compared with controls, although notably genotypes were not in Hardy-Weinberg equilibrium, even in controls, making this result difficult to interpret. Studies of non-Asians conducted in America and Europe have been generally negative with few exceptions (26-29). However, the best summary of the accumulated evidence comes from a recent pooled analysis of 22 case-control studies that included over 2,400 cases and 3,300 controls and examined genotype in combination with the effect of tobacco smoking (30). This analysis concluded that the homozygous variant *CYP1A1*2A* polymorphism increases the risk of lung cancer in Caucasians, whereas no significant increase was observed in Asians among whom the risk associated with tobacco smoking appeared constant by

genotype. Our results are therefore in full agreement with this overview of the literature.

Published results on lung cancer risk and the *GSTM1*-null genotype are even more heterogeneous. Two meta-analyses estimated a weak but significant increase in risk associated with *GSTM1* deficiency (5, 31). Several case-control studies conducted in countries of diverse ethnicities were published later. Four in 10 found an overall positive association with lung cancer (29, 32-34), whereas the others were negative for an association (23, 35-40). When exposure to environmental carcinogens was taken into account, effect modification was found in some but not all studies. *GSTM1*-deficient subjects have been found at increased risk of lung cancer if moderately exposed to tobacco smoking (32), or never-smokers carrying the *CYP1A1*2C* variant (38), or had a stronger protective effect from vegetable-derived isothiocyanates compared with non-deficient subjects (41). But interaction with tobacco smoking or exposure to asbestos or other environmental pollutant was not detected in other studies (20, 37, 42). In our data, *GSTM1* deficiency and the two *CYP1A1* variants examined were not associated with an increased risk of lung cancer even when polymorphisms in the two genes were combined. Moreover, they did not act as modifiers of the effect of tobacco smoking either overall or in the main histologic subtypes. The distributions of the genotypes that we observed were in Hardy-Weinberg equilibrium in the controls and in agreement with the frequencies estimated for Asians in the Genetic Susceptibility to Environmental Carcinogens (43), an international project that aims at describing the frequency of various alleles in metabolic genes, in different geographic and ethnic groups. We conclude that tobacco smoking remains the only identified risk factor for the disease in this Thai population.

Acknowledgments

We thank the cases and the controls for their participation in the study, the staff of the Lampang Provincial Health Office and Lampang Hospital for carrying out the interviews and collecting the specimens, Miranda Y. Ku for computing pollution exposure estimates, and Dr. Max Parkin for his advice in designing and conducting the study.

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