Temporal Characterization of the Neural Correlates of Perceptual Decision Making in the Human Brain

Marios G. Philiastides and Paul Sajda

Laboratory for Intelligent Imaging and Neural Computing, Department of Biomedical Engineering, Columbia University, New York, NY 10027, USA

Single and multi-unit recordings in primates have identified spatially localized neuronal activity correlating with an animal’s behavioral performance. Due to the invasive nature of these experiments, it has been difficult to identify such correlates in humans. We report the first non-invasive neural measurements of perceptual decision making, via single-trial EEG analysis, that lead to neurometric functions predictive of psychophysical performance for a face versus car categorization task. We identified two major discriminating components. The earliest correlating with psychophysical performance was consistent with the well-known face-selective N170. The second component, which was a better match to the psychometric function, did not occur until at least 130 ms later. As evidence for faces versus cars decreased, onset of the later, but not the earlier, component systematically shifted forward in time. In addition, a choice probability analysis indicated strong correlation between the neural responses of the later component and our subjects’ behavioral judgements. These findings demonstrate a temporal evolution of component activity indicative of an evidence accumulation process which begins after early visual perception and has a processing time that depends on the strength of the evidence.

Keywords: choice probability, decision making, electroencephalogram, neurometric function, perceptual discrimination, psychometric function, single-trial

Introduction

Identifying neural activity directly responsible for perceptual decision making is a major challenge for systems and cognitive neuroscience. A number of investigators have studied the neural correlates of decision making in awake behaving animals, in particular primates, where single and multi-unit recordings have been analyzed using signal detection theory (Green and Swets, 1966) and subsequently correlated with the animal’s observed behavior/decisions. For instance, Newsome and colleagues (Newsome et al., 1989; Britten et al., 1992, 1996) used visual stimuli consisting of varying amounts of coherent motion and showed that neurometric functions constructed from the activity of specific individual and small populations of neurons were indistinguishable from the animal’s psychometric functions. They also computed choice probabilities and showed that the neural responses had a small but significant association with the animal’s decisions. This approach of comparing neurometric and psychometric functions and considering choice probabilities enables one to directly address the decision making process since it explicitly relates the variability of the neural activity to the variability observed in the behavioral response. The technique has been applied in a variety of perceptual decision making paradigms, including discrimination of visual objects such as faces (Keysers et al., 2001) and tactile discrimination tasks (Hernandez et al., 2000; Romo et al., 2002). The approach, though powerful, has been limited to animal studies which use invasive recordings of single-trial neural activities. Yet to be demonstrated is whether decision making could be studied in a similar fashion, though under the constraint that single-trial neural activity be measured non-invasively in humans.

Decision making during face processing has been extensively studied in both primates and humans, in particular within the context of face categorization and identification. Human subject studies using trial-averaged electroencephalography (EEG) have identified waveforms at specific times (e.g. N170, N200) that are well-correlated with the presentation of faces compared with nonface objects (Jeffreys, 1996; Halgren et al., 2000; Liu et al., 2000; Rossion et al., 2003). Recent studies using magnetoencephalography (MEG) (Liu et al., 2002) have found earlier trial-averaged activity (M100) which is correlated with face categorization but not identification of individual faces, suggesting a two-stage processing strategy in face perception. However, these experiments do not directly address the detailed temporal aspects of the decision making process during human face perception since they do not consider single-trial variability of the activity relative to the variability of the response.

Keysers et al. (2001), using rapid serial visual presentation (RSVP), identify activity of individual neurons in macaque temporal cortex which predict whether the monkey responds that it saw a face in the stimulus. These investigators attempt to relate this neural activity to human subject performance by comparing neurometric functions for single-neurons in macaque with psychometric functions of human subjects’ doing a similar task. Their comparison, though indirect and qualitative, indicates that the monkey’s neurometric function has a similar shape to the human subject’s psychometric function, though the curves themselves are quantitatively very different. This approach is problematic, for activities and decisions are not only being compared across species but also across subjects and different experimental sessions. Thus the inter-subject and intra-subject variability in the decision making process cannot be captured/measured. Rather this would require simultaneous measurement of the neuronal and psychophysical performance while the human subject performs the task.

In this paper we used single-trial analysis of the EEG to identify the cortical correlates of decision making during face perception in human subjects. We used a machine learning approach to identify linear spatial weightings of the EEG sensors for specific temporal windows which optimally discriminate between target (faces) and non-target (cars) trials during a simple categorization task. From these discriminating components we
constructed neurometric curves, as a function of stimulus evidence (phase coherence) and compared them to subjects' psychometric functions. We analyzed the temporal characteristics of those components which were strongly correlated with psychophysical performance and considered how changes in their temporal onset was affected by the stimulus evidence for the decision.

Materials and Methods

Subjects
Six people (three females and three males, age range 21–37 years) participated in the study. All had normal or corrected to normal vision and reported no history of neurological problems. Informed consent was obtained from all participants in accordance with the guidelines and approval of the Columbia University Institutional Review Board.

Stimuli
We used a set of 12 face (Max Planck Institute face database) and 12 car grayscale images (image size 512 × 512 pixels, 8-bits/pixel). Both image types contained equal numbers of frontal and side views (up to ±45°). All images were equated for spatial frequency, luminance and contrast. They all had identical magnitude spectra (average magnitude spectrum $\cos(w)$ = $100\%$ phase coherence, and $\sin(w)$). We used a set of 12 face (Max Planck Institute face database) and 12 car grayscale images (image size 512 × 512 pixels, 8-bits/pixel). Both image types contained equal numbers of frontal and side views (up to ±45°). All images were equated for spatial frequency, luminance and contrast. They all had identical magnitude spectra (average magnitude spectrum $\cos(w)$ = $100\%$ phase coherence, and $\sin(w)$).

Data Acquisition
EEG data was acquired simultaneously in an electrostatically shielded room (ETS-Lindgren, Glendale Heights, IL) using a Sensorium EPA-6 Electric Physiological Amplifier (Charlotte, VT) from 60 Ag/AgCl scalp electrodes and from three pericortical electrodes placed below the left eye and at the left and right outer canthi. All channels were referenced to the left mastoid with input impedance <15kΩ and chin ground. Data were sampled at 1000 Hz with an analog pass band of 0.01–300 Hz using 12 dB/octave high pass and eighth-order Elliptic low pass filters. Subsequently, a software based 0.5 Hz high pass filter was used to remove DC drifts and 60 and 120 Hz (harmonic) notch filters were applied to minimize line noise artifacts. These filters were designed to be linear-phase to minimize delay distortions. Motor response and stimulus events recorded on separate channels were delayed to match latencies introduced by digitally filtering the EEG.

Movement Artifact Removal
Prior to the main experiment, subjects completed an eye muscle calibration experiment during which they were instructed to blink repeatedly upon the appearance of a white-on-black fixation cross and then to make several horizontal and vertical saccades according to the position of the fixation cross subtending 1" × 1" of visual field. Horizontal saccades subtended 35° and vertical saccades subtended 22°. The timing of these visual cues was recorded with EEG. This enabled us to determine linear components associated with eye blinks and saccades (using principal component analysis) that were subsequently projected out of the EEG recorded during the main experiment (Parra et al., 2003). Trials with strong eye movement or other movement artifacts were manually removed by inspection. There were at least 40 artifact-free trials for any given condition (i.e. at least 80 trials for both sets of face and car trials at each phase coherence level).

Data Analysis
We used a single-trial analysis of the EEG to discriminate between face versus car trials at each phase coherence level. Logistic regression was used to find an optimal basis for discriminating between two conditions over a specific temporal window (Parra et al., 2002). Specifically, we defined a training window starting at a poststimulus onset time τ, with a duration of δ, and used logistic regression to estimate a spatial weighting vector $w_\omega$ which maximally discriminates between sensor array signals $X$ for two conditions (e.g. target versus non-target trials):

$$y = w_\omega^T X \quad (2)$$

where X is an $N \times T$ matrix ($N$ sensors and $T$ time samples). The result is a 'discriminating component' $y$ which is specific to activity correlated with condition 1 (faces) while minimizing activity correlated with both task conditions such as early visual processing. We use the term 'component' instead of 'source' to make it clear that this is a projection of the 'sensor projections' $y$.

Equation 3 describes the electrical coupling $a$ of the discriminating component $y$ that explains most of the activity $X$. Strong coupling indicates low attenuation of the component and can be visualized as the intensity of the 'sensor projections' $a$. $a$ can also be seen as a forward model of the discriminating component activity (Parra et al., 2002).

At a given coherence level we also constructed discriminant component maps. We aligned our trials to the onset of visual stimulation as shown in Figure 2a (black vertical line) and sorted them by reaction time (sigmoidal curves). Each row of the discriminant component map represents a single trial across time. Discriminant components are represented by the $y$ vectors. To construct this map we chose a training window, indicated by the white vertical bars (for this example starting...
at 180 ms post-stimulus), during which we trained our linear discriminator to estimate the weighting vector \( w \) for all of our sensors in \( X \), such that \( y \) is maximally discriminating between face and car trials. Once the \( w \) vector has been estimated using data derived only from the training window, we applied \( w \) to the data across all trials and all time. The resultant discriminant component map is shown in Figure 2a. Red represents positive and blue negative activity. For this example the resultant discriminating component appears ~180 ms after the onset of visual stimulation and it is stimulus-locked. Response-locked components are expected to have a sigmoidal profile similar to the subject’s actual response times.

Training the discriminator over several temporal windows allowed us to visualize the temporal evolution of the discriminating component activity. We used a coarse-to-fine approach to identify the two most discriminating time windows at every phase coherence level. Initially we trained the discriminator while sliding our training window in non-overlapping segments of 30 ms from the onset of visual stimulation to the earliest response time. We subsequently re-trained the discriminator by sliding the window in finer steps (10 ms) only around the two most discriminating regions as identified by the coarser search.

Taking advantage of the linearity of our model, we subsequently used the forward model to project this discriminating activity back to the sensors. The resultant scalp projections \( a \) are shown in Figure 2b and are used for interpreting the neuroanatomical significance of the resultant discriminating component.

We quantified the performance of the linear discriminator by the area under the receiver operator characteristic (ROC) curve, referred to as \( A_z \), with a leave-one-out approach (Duda et al., 2001). We used the ROC \( A_z \) metric to characterize the discrimination performance while sliding our training window from stimulus onset to response time (varying \( t \)). Finally in order to assess the significance of the resultant discriminating component we used a bootstrapping technique to compute an \( A_z \) value.

Figure 2. Summary of our single-trial EEG analysis. (a) Example of a discriminant component map resulted from our single-trial linear discrimination analysis. (b) Dorsal view of sensor projections \( a \) and component significance as quantified by an ROC analysis (see text for details).
leading to a significance level of $P = 0.01$. Specifically we computed a significance level for $A_z$ by performing the leave-one-out test after randomizing the truth labels of our face and car trials. We repeated this randomization process 100 times to produce an $A_z$ randomization distribution and compute the $A_z$ leading to a significance level of $P = 0.01$ shown in Figure 2b by the dotted red line.

**Psychometric and Neurometric Functions**

We constructed psychometric curves to characterize our subjects behavioral performance. The proportion of correct choices was plotted against the percentage of phase coherence of our stimuli. To characterize neuronal performance, in a manner comparable to the description of psychophysical sensitivity as captured in the psychometric function, we constructed an EEG-based neurometric function by displaying the $A_z$ values (Green and Swets, 1966) obtained during single-trial discrimination of face versus car trials at each phase coherence level. We used a maximum-likelihood method (Watson, 1979) to compute the best-fitting cumulative Weibull distribution (Quick, 1974) for each data set:

$$p = 1 - 0.5 e^{\frac{-L}{\beta}}$$

(4)

where $p$ is the proportion of correct choices computed as a function of the phase coherence, $c$, of our stimuli. The fitted parameters $\alpha$ and $\beta$ are the phase coherence supporting threshold performance (82% correct) and the slope of the curve respectively.

**Choice Probabilities**

We computed choice probabilities at selected phase coherence levels for all our subjects. We labeled all trials based on the response choice rather than the type of stimulus (i.e. all face responses - correct face and incorrect car trials - versus all car responses - correct car and incorrect face trials). We then repeated the ROC analysis described in Data Analysis and classified between ‘face’ and ‘car’ choice trials such that the new set of $A_z$ values now represented choice probabilities.

**Statistical Tests**

We used a conventional $\chi^2$ test to evaluate the goodness of fit to our data. Using the criterion ($\chi^2$, $P < 0.05$), none of the fits of all the subjects in our main data set were rejected. In order to quantify the degree of similarity between the psychometric and neurometric functions we fit the best single Weibull function jointly to the two data sets in addition to the individual fits. The likelihoods ($L$) obtained from these two conditions were transformed by:

$$\lambda = -2ln\left(\frac{L_{\text{jointed curve}}}{L_{\text{individual curves}}}\right)$$

(5)

so that $\lambda$ is distributed as $\chi^2$ with two degrees of freedom (Hoel et al., 1971). If $\lambda$ does not exceed the criterion value (for $P = 0.05$), we concluded that a single function fits the two data sets as well as two separate functions.

**Results**

**EEG-derived Neurometric Functions**

We measured the psychophysical performance of six subjects to a face versus car categorization task (Fig. 1a) while simultaneously recording neuronal activity using a high-density EEG electrode array. We changed the stimulus evidence by manipulating the phase coherence of our images (Fig. 1b). Within a block of trials, face and car images over a range of phase coherences were presented in random order. All subjects performed nearly perfect at the highest phase coherence but performed near chance at the lowest coherence.

We compared the EEG activity obtained for the two types of images at each phase coherence level on a single-trial basis in order to capture the response variability which is normally concealed by trial averaging. Using a linear discriminator which integrates EEG over space rather than across trials, we identified components that maximally discriminated between the two experimental conditions. We used ROC analysis to quantify the discriminator’s performance. Furthermore, taking advantage of the linearity of our model, we computed sensor projections of the discriminating component activity. These scalp projections can provide a forward model for interpreting the neuroanatomical significance of the resultant discriminating components.

At each phase coherence level we identified the two most discriminating components in the interval between the onset of the visual stimulation and the earliest reaction time. In order to characterize the neuronal performance at these two times, in a manner compatible with the description of the psychophysical sensitivity as captured by the psychometric functions (Green and Swets, 1966), we constructed neurometric functions by plotting the area under the ROC curves ($A_z$ values) against the corresponding phase coherence levels and fitting the data with Weibull functions (Quick, 1974). We trained our linear discriminator at these two temporal windows separately and combined. When the discriminator was trained with data integrated across both time windows we generally observed, for the discriminator, improved performance and hence higher $A_z$ values. Figure 3 shows comparisons of the psychometric and neurometric functions for all subjects in our dataset.

In order to quantify the degree of similarity between the psychometric and neurometric functions and demonstrate that our EEG-derived neurometric functions can account for psychophysical performance, we fit the best single Weibull function jointly to the two data sets in addition to the individual fits. We subsequently used a likelihood-ratio test (Hoel et al., 1971) which showed that for all our subjects a single function can fit the behavioral and neuronal data sets as well as the two separate functions. We concluded that these neurometric functions can be used to quantify the relationship between the neuronal signals and behavior by isolating the components that are associated with the perceptual discrimination.

**Early and Late Face-selective Responses**

In the time interval between the stimulus onset and the earliest response time we identified two face-selective components that were maximally discriminating between face and car trials. To visualize the temporal evolution of these components we constructed discriminant component maps. Figure 4 summarizes the complete results for one subject. The early component was consistent with the well-known N170 and its temporal onset appeared to be consistent across all subjects. The late component, which was of opposite sign, appeared on average 130 ms after the first at the highest coherence level. Its temporal onset varied across subjects in the range 300–450 ms from stimulus onset across all coherence levels.

As can be seen by the single-trial component projections (Fig. 4) from discrimination of stimulus-locked face versus car trials, both face-selective responses appeared to be more correlated with the onset of visual stimulation (black vertical line) rather than the response (sigmoidal curves). To verify this point with respect to possible bias in our stimulus locked analysis, we reanalyzed the results response locked and found the components remained strongly stimulus locked (not shown). This observation indicated that the discriminating activity is not directly predictive of reaction time; rather, it
Figure 3. Comparison of behavioral and neuronal performance. (a–f) Psychometric (red) and neurometric (black) functions for all six subjects. The abscissas represent the percentage of phase coherence of our stimuli and the ordinate indicates the subject’s performance as proportion correct. We fit both data with separate Weibull functions and compute the corresponding threshold (a) and slope (b) parameters. For all six subjects psychophysical and neuronal data were statistically indistinguishable as assessed by a likelihood ratio test after we fit the best single Weibull function jointly to the two data sets. The \( P \)-value in each panel represents the output of this test. A \( P \)-value greater than 0.05 indicates that a single function fits the two data sets as well as the two separate functions. The dotted gray lines connect the \( A_z \) values computed for each of the two training windows separately (earlier window, open circles; later window, open squares).

Figure 4. Discriminant component activity that shows the difference between face versus car trials at each coherence level for one subject for (a) the early (N170) and (b) the late (~300–400 ms) window. Red represents positive and blue negative activity. All trials were aligned to the onset of visual stimulation, as indicated by the vertical black line at time 0 ms, and sorted by response time. The black and magenta sigmoidal curves represent the subject’s response times for face and car trials respectively. We subsequently applied our linear discrimination algorithm to construct discriminant component maps. Each row of these maps represents the output of the linear discriminator for a single trial, using a 60 ms training window (vertical white lines) with onset times specified at the top of each panel. The representation of the topology of the discriminating activity is shown by the scalp plots to the right (dorsal view). Red represents positive correlation of the sensor readings to the extracted activity and blue negative correlation. The \( A_z \) values for each time window at each coherence level are represented by the bar graphs. The significance of the difference activity is represented by the red line (\( P = 0.01 \)). For this subject the discriminant component activity was statistically significant down to 30% phase coherence for both time windows.
appears to be related to the stimuli reaching a perceptual level of processing (Super et al., 2001).

In addition we observed that the late face-selective component resulted in a better match to the psychophysical data as shown in Figure 3. In fact, for four of our six subjects (Fig. 3a–d) using the \( A_z \) values obtained only from the late training window to construct our neurometric function was sufficient to show that the psychometric and neurometric functions were statistically indistinguishable. Interestingly this was never true when neurometric functions were derived solely based on analysis from the early face-selective N170 component.

Evidence Changes Onset of Late Component
We investigated the relationship between the temporal onset of the early and late face-selective responses and task difficulty. The most discriminating time window at each phase coherence level for each one of the two face-selective temporal components appears on the top of each set of projections in Figure 4. We used these times to study the extent to which these face-selective components systematically shifted in time. We used a bootstrapping method, to identify all significantly discriminating components \( (P < 0.01; \text{ red dotted line}) \). Only times from these statistically significant components were used for this analysis. We regressed a line through these time points for each one of the face-selective components and computed their corresponding slopes (Fig. 5a). We found that for the early component (N170) there was no significant shift in time, as the slopes did not differ significantly from zero (two-tailed \( t \)-test, \( P > 0.45 \)). On the other hand, the optimal onset of the late face-selective component had a systematic forward shift in time as the task became more difficult and the subjects took longer to respond. The slopes for the late component were statistically less than zero (left-tailed \( t \)-test, \( P < 0.007 \)) and also statistically more negative from the slopes of the earlier component (right-tailed, paired \( t \)-test, \( P < 0.005 \)).

These findings seem to suggest an early component related to bottom-up processing of the stimulus and a late component responsible for the evaluation of the evidence (Shadlen and Newsome, 1999; Keysers and Perrett, 2002), occurring later for more ambiguous stimuli. In addition, the timing of the second component (300–450 ms post-stimulus) is consistent with previous findings suggesting recurrent processing of the stimulus/evidence (Super et al., 2001; VanRullen and Koch, 2003), with an average reverberation time of 130 ms. Such reverberatory activity is likely to reflect the integration of information that underlies perception considering that recurrent/feedback connections are shown to mediate processes such as perceptual organization, attention and visual awareness (Lamme et al., 1998; Hupe et al., 1998; Super et al., 2001).

Association between Neuronal Responses and Behavioral Decisions
To address the possibility that the neural responses associated with the discriminating components are correlated with our subjects’ choices, we employed a method based on signal detection theory, analogous to the ROC analysis used earlier, to compute choice probabilities as in Britten et al. (1996). Unlike traditional uses of signal detection theory, however, which establish the relationship between the stimulus and neural responses, this alternative formulation quantifies a relationship between neuronal activity and a subject’s choices/decisions. A choice probability value of 0.5 represents chance performance and a value of 1.0 represents perfect association between neuronal and behavioral responses. In order for the choice probability metric to be meaningful, neural responses from stimuli near threshold are to be used (so that the subjects make a useful number of errors on the psychophysical task).

We pooled data across two coherence levels (30 and 35%) that were near threshold where subjects made both ‘face’ and

Figure 5. (a) Results showing delayed onset of the second window as a function of coherence level. No systematic shifts are seen for the early window. Plots (raw data and linear regression) of onset times for early and late windows for all six subjects are shown. Only significant components are used in the regression analysis. (b) Average scalp topologies across subjects at each phase coherence level for both the early (left) and late (right) components. Each subject’s scalp projections were normalized prior to computing the grand averages. Red represents positive correlation between the sensors and the discriminating components whereas blue represents negative correlation.
‘car’ choices in response to any particular stimulus, and computed a choice probability value for every subject. We computed three sets of choice probabilities using neuronal data from (i) the early component, (ii) the late component and (iii) the early and late components combined. The results are summarized in Table 1. To assess the significance of these choice probabilities we employed a bootstrap technique where we randomly permuted the trial labels 500 times and computed choice probability distributions for every subject. This permutation test ensured that the association between the neuronal and behavioral responses was abolished, while the distributions of neuronal and behavioral judgements remained untouched. We then checked if the observed choice probability values were outside the 95% confidence intervals of these distributions, in which case we concluded that they were statistically significant.

Only three out of six subjects had a choice probability significantly greater than chance when data from only the earlier component were considered. Interestingly, however, the observed choice probabilities for the late component were shown to be statistically significant for all six subjects. These results clearly demonstrate that the neuronal responses, especially of the later component, for all our subjects had a significant positive association with their behavioral choices. Taken together, this finding and the systematic forward shift of the second component as evidence decreased suggest that the second component may be associated with the actual decision making process, or at the very least reflect an intermediate stage of perceptual processing that is situated between purely sensory and decision stages (Super et al., 2001).

**Spatial Distribution of Activity and the Importance of Spatial Integration**

For both the early (N170) and late face-selective responses, at each phase coherence level, we constructed scalp maps of the discriminating components, and the results for one subject are shown in Figure 4a,b. The $A_x$ values which describe the discriminator’s performance at each phase coherence level are also shown. For the subject shown in Figure 4, the discriminant activity was statistically significant down to a 30% phase coherence for both temporally distributed components as assessed by our bootstrapping technique ($P < 0.01$; red dotted line).

The average scalp projections from significantly discriminating components for the early face-selective component (Fig. 5b) indicated significant differences between face versus car trials at occipito-temporal electrode sites in the left and right hemispheres (negative correlation) and a number of centro-frontal sites (positive correlation). These results are consistent with functional neuroimaging studies (Kanwisher et al., 1996; 1997; Puce et al., 1996; Hasson et al., 2002) and several ERP/MEG studies (Botzel et al., 1995; Bentin et al., 1996; Halgren et al., 2000; Liu et al., 2000; Rossion et al., 2003), where face-sensitive activations are always found relative to objects in occipito-temporal cortex (mainly the inferior occipital and fusiform gyri) bilaterally. Some studies have also identified face-selective responses (Jeffreys, 1989, 1996) and target/nontarget responses (VanRullen and Thorpe, 2001), in addition to the occipito-temporal sites, in centro-frontal locations. These are also consistent with recent findings which identified active regions in the dorsal lateral prefrontal cortex (DLPFC) which are thought to be associated with decision making during a face versus house categorization task (Heekeren et al., 2004). The late face-selective component also demonstrated a very similar activation pattern to the early/N170 component (Fig. 5b), though with opposite sign. This was an interesting observation, though only a simultaneous fMRI study could determine definitively which cortical systems contributed to this component.

To emphasize the importance of spatial integration for identifying discriminating components predictive of the psychophysical sensitivity of our subjects, we used an alternative approach to computing a neurometric function. Instead of using our spatial integration algorithm, which weights the activity across all EEG sensors, we repeated the discrimination using only the activity from a single electrode. We chose electrode PO8, over the right face-selective activation area, where the greatest difference in the EEG signal between face and car trials was identified. The neurometric function that we constructed with this approach was not nearly as predictive of the psychophysical performance as the neurometric function which was computed using our spatial integration technique. Figure 6 illustrates this point.

**Motor Activity not Predictive of Psychophysical Performance**

In some cases we observed small, though significant, differences in reaction time for face versus car responses (e.g. see the sigmoidal reaction time curves in Fig. 4). To test whether the discriminating activity we identified was due to a difference in this reaction time, for example a component associated with motor activity, we generated neurometric functions using temporal windows near the reaction time. Specifically, for each subject we used a training window around the median reaction time at each phase coherence level. A typical curve derived during this period is also shown in Figure 6. It is clear from this plot that discriminating components extracted near reaction times (e.g. components representative of preparatory motor, motor or somatosensory activity) were not predictive of the psychophysical sensitivity of our subjects. Additional evidence that the two components, specifically the second component, are not reflective of response selection/motor programming can be seen by considering the change in the strength of the component as a function of coherence level. If the component were reflective of response selection, then one would expect no difference in the strength of the component at different coherence levels. Finally, the second component is strongly stimulus-locked, providing further evidence that it is not reflective of response selection/motor programming. We can conclude that the earlier and late component activities are not artifacts of reaction time differences but are in fact closely linked to the perception and decision making processes respectively.
Comparison to ERP Component Analysis

To reinforce the main findings of our study, we also present a brief summary of a traditional ERP analysis on our data. ERPs, though the result of trial averaging which compromises single-trial variability, are easier to interpret and can provide useful insights about the time course of the underlying visual processing. In our task we found strong differential activity (i.e. faces -- cars) at virtually all electrode locations which emphasizes the magnitude of the effect. Figure 7 summarizes these results. Specifically, we computed difference ERP waveforms at six centro-frontal locations as well as at six occipito-parietal sites for one subject at the highest coherence level (i.e. 45%). Both waveforms indicate significant differential responses that peak at ~170 ms (early component) and 330 ms (late component) after the onset of visual stimulation. Overlaid are the two 60 ms training windows used to achieve maximum single-trial discrimination for this particular subject (early window, light gray; late window, dark gray).

In addition to these maximally discriminating components, we observed an even earlier component (~100 ms), the magnitude of which was not as strong as the other two. The presence of this component and its spatial distribution (as indicated by the corresponding scalp maps) are consistent with what was reported by Liu et al. (2002). In all cases, however, the A values we computed for this component were not sufficiently high to account for the psychophysical performance of our subjects.

Scalp maps for each component, constructed based on the difference ERPs and our sensor projections, are also shown in Figure 7. Red presents positive correlation between the sensors and the underlying discriminating component and blue represents negative correlation. Note that the sign of the differential ERP activity at the different electrode locations is consistent with the scalp topology. For instance, there is a negative correlation between the N170 component and the occipito-parietal sensors and positive correlations with centro-frontal locations. The signs flip for the later component.

Even though ERP analysis can identify both the early and later components, it could not unequivocally associate, especially the later component, to our subjects’ decision making process. Single-trial analysis, on the other hand, provides a more rigorous, and direct, method to compare neuronal responses to psychophysical performance (which is not obtainable using simple correlation of sorted ERP derived amplitudes) and therefore directly addresses the decision making process.

Discussion

Our results demonstrate that neural correlates of perceptual decision making can be identified using high-spatial density EEG and that the corresponding component activities are temporally distributed. Clearly important to identification of these neural correlates is the spatial, and to a lesser extent the temporal integration of the EEG component activities. This approach is complementary to approaches using single and multi-unit recordings since it sacrifices spatial and some temporal resolution (local field potentials versus spike-trains) for a more spatially distributed view of the neural activity during decision making. The fact that we were able to identify neural correlates of perceptual decision making using relatively poor spatial resolution of EEG suggests that these neural correlates represent strong activities of neuronal populations and not the activity of a small number of neurons.
It is interesting to consider the temporal characteristics of the discriminating components we identified relative to models of evidence accumulation in decision making (Kim and Shadlen, 1999; Shadlen and Newsome, 2001; Mazurek et al., 2003). Unlike studies which use dynamic stimuli, such as moving random dot sequences, where evidence (via analysis of spatio-temporal correlation structure) can accumulate across time, our stimuli are static images with no (or little) spatiotemporal correlation from one to the next. Thus, the temporal nature of our neural correlates is directly related to the underlying nature of the internal processing for static object recognition.

Previous work using trial-averaged ERPs has attempted to identify the timing and activity of early and late components in object recognition. Several studies have claimed a component at 150 ms poststimulus representative of the speed of visual processing (Thorpe et al., 1996) and which is correlated with subject behavior (VanRullen and Thorpe, 2001). Other studies have claimed that a second, later component (~300 ms) is in fact more directly correlated with the recognition process, with the earlier component corresponding to low-level feature discrimination (Johnson and Olshausen, 2003). Our findings provide additional evidence that the later component is more closely linked to a recognition event/decision, while also providing evidence for a cortical processing strategy that enables a trade-off between processing time and accuracy. Assuming a fast-feedforward recognition process within 150 ms of stimulus onset, the early component appears to represent a quick evaluation of the evidence which, while less accurate, could enable a faster response.

Our results complement this study by characterizing the temporal evolution of component activities that are correlates of perceptual decision making. In addition, our approach precisely quantifies the relationship between the neural signal and behavior through the comparison of psychometric and neurometric functions. The construction of neurometric functions was enabled by our single-trial EEG analysis methods. We saw component activity predictive of decision making and consistent with signaling between occipito-temporal and frontal networks. Unlike fMRI, however, EEG does not have sufficient spatial resolution to precisely identify the cortical regions responsible for these components. Simultaneous methods for acquiring EEG and fMRI may provide a better picture of these spatio-temporal network dynamics indicative of cortical processing underlying perceptual decision making.

Notes

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Address correspondence to Paul Sajda, 351 Engineering Terrace Building, Mail Code 8904, 1210 Amsterdam Avenue, New York, NY 10027, USA. Email: ps629@columbia.edu.