



The Impact of Glycated Hemoglobin (HbA_{1c}) on Cardiovascular Disease Risk: A Mendelian Randomization Study Using UK Biobank

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OBJECTIVE

Glycated hemoglobin (HbA_{1c}) is positively associated with cardiovascular disease (CVD), although evidence is primarily observational. Mendelian randomization studies have only examined its relation with subtypes of CVD. We examined the relation of HbA_{1c} with CVD and its subtypes in the UK Biobank using Mendelian randomization.

RESEARCH DESIGN AND METHODS

We used 38 genetic variants strongly and independently related to HbA_{1c} ($n = 123,665$) applied to the UK Biobank ($n = 392,038$). We used inverse variance weighting (IVW) to obtain the associations of HbA_{1c} with CVD, coronary artery disease (CAD), and stroke (overall and stroke subtypes). Sensitivity analyses included Mendelian randomization (MR)-Egger, a weighted median, and exclusion of potentially invalid single nucleotide polymorphisms (SNPs). We also applied the same genetic instruments to CARDIoGRAMplusC4D (Coronary ARtery Disease Genome wide Replication and Meta-analysis [CARDIoGRAM] plus The Coronary Artery Disease [C4D] Genetics) 1000 Genomes–based genome-wide association study ($n = 184,305$) as a validation for CAD.

RESULTS

In the UK Biobank, HbA_{1c} was not associated with CVD using IVW (odds ratio [OR] 1.11 per %, 95% CI 0.83–1.48). However, HbA_{1c} was associated with increased CAD risk (OR 1.50 per %, 95% CI 1.08–2.11) with directionally consistent results from MR-Egger and weighted median. The positive association with CAD was more pronounced when we excluded potentially invalid SNPs (OR 2.24 per %, 95% CI 1.55–3.25). The positive association was replicated in CARDIoGRAM (OR 1.52 per %, 95% CI 1.03–2.26). The association of HbA_{1c} with stroke and its subtypes was less clear given the low number of cases.

CONCLUSIONS

HbA_{1c} likely causes CAD. The underlying mechanisms remain to be elucidated.

Observational studies strongly suggest a link between type 2 diabetes and coronary artery disease (CAD), but these observations could be confounded by lack of physical activity and obesity (1–4). Randomized controlled trials, such as the ACCORD (Action to Control Cardiovascular Risk in Diabetes) trial, unexpectedly showed

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(coronary angiography results or bypass surgery), use of medications or symptoms that indicate angina, or self-report of a doctor's diagnosis, as described elsewhere (24). CARDIoGRAMplusC4D 1000 Genomes-based GWAS was adjusted for study-specific covariates (e.g., age and sex) and genomic control.

Exposure

The exposure was genetically predicted HbA_{1c} (%).

Outcomes

The primary outcomes were prevalent CVD (defined as ICD-9 401–459.9 and ICD-10 I10–I99) and its subtypes, including CAD (ICD-9 410–414.9 and ICD-10 I20–I25.9), ischemic stroke (ICD-9 434 and 436 and ICD-10 I63–I64), and hemorrhagic stroke (ICD-9 430–431 and ICD-10 I60–I61), based on self-reports and hospital episodes, and death from CVD (ICD-10 I10–I99) or its subtypes CAD (ICD-10 I20–I25.9), ischemic stroke (ICD-10 I63–I64), and hemorrhagic stroke (ICD-10 I60–I61) from death records, following the recommended definitions of the UK Biobank Stroke Outcomes Group (27). For completeness, we also considered other CVD (i.e., all CVD excluding stroke and CAD, which was mainly hypertensive diseases) and other stroke (i.e., all stroke excluding ischemic and hemorrhagic stroke). CVD, CAD, and stroke mortality, based on primary cause of death, were also considered separately as secondary outcomes.

Statistical Analysis

We assessed departure from Hardy-Weinberg equilibrium for each SNP using χ^2 tests with Bonferroni correction to correct for multiple comparison (0.05/39 = 0.00128). We used ANOVA (continuous) and χ^2 test (categorical) to examine whether the genetic variants were associated with factors potentially confounding the association of HbA_{1c} with CVD, including Townsend deprivation index, education, age, BMI, smoking, and alcohol drinking, in the UK Biobank, with Bonferroni correction to correct for multiple comparisons (0.0002, based on 0.05/234 derived from 39 SNPs \times 6 traits). We obtained the association of each SNP with CVD and its subtypes using multivariable logistic regression in the UK Biobank, adjusted for age, sex, genotyping array, and 10 principal components.

We conducted our main analysis using inverse variance weighting (IVW) with multiplicative random effects, which is a weighted regression of gene-outcome associations on gene-exposure associations for UK Biobank and CARDIoGRAMplusC4D 1000 Genomes-based GWAS. Given IVW assumes no horizontal pleiotropy, which cannot be empirically assessed, we used the I^2 of the Wald estimates (SNP-outcome association divided by SNP-exposure association) to indicate the presence of invalid instruments. In the presence of invalid SNPs (i.e., SNPs that have effect on the outcome not via HbA_{1c}), IVW will be invalid. As such, we also conducted several sensitivity analyses to assess the robustness of our results to potential violations of the Mendelian randomization assumptions because these analyses have different assumptions for validity, as described below. Although these approaches may have different statistical power (e.g., wider CIs for Mendelian randomization [MR]-Egger), the rationale is that if these approaches give a similar conclusion regarding the association of HbA_{1c} with the outcomes, then we are more confident in inferring that the positive findings are unlikely driven by violation of the Mendelian randomization assumptions (28).

Instrument Strength

To assess instrument strength, we computed the F statistic for the association of genetic instruments with HbA_{1c}, assuming the included genetic variants explained at least 4.2% of the HbA_{1c} variance (the lower bound of variance explained based on the previous GWAS) (21). Higher F statistics indicate a stronger instrument.

Sensitivity Analysis

1) MR-Egger. We conducted MR-Egger regression, which produces valid estimates even if all the genetic instruments are invalid, as long as the InSIDE (Instrument Strength Independent of Direct Effect) assumption holds (29). We also presented the intercept P value from the MR-Egger regression because a significant intercept indicates the IVW estimate may be invalid due to horizontal pleiotropy.

2) Weighted Median. We also used a weighted median, which produces valid estimates as long as more than 50% of the information is derived from valid SNPs (30).

3) Exclusion of Potentially Invalid SNPs. We identified potentially invalid genetic instruments (SNPs) in two ways. First, we excluded SNPs related to potential causes of CAD based on the GWAS Catalog (31)/PhenoScanner (32) and SNPs associated with potential confounders in the UK Biobank (set 1). Second, we additionally excluded SNPs defined as "erythrocytic" in the original GWAS because they did not predict diabetes and hence may be irrelevant to glycemic exposure (i.e., invalid) (21) (set 2). Supplementary Table 1 summarizes the choice of SNPs in different sensitivity analyses.

To rule out the possibility of false positives due to inclusion of related individuals, we also repeated the analyses including only unrelated participants.

All analyses were performed using R version 3.3.2 (R Development Core Team, Vienna, Austria) with the R package (TwosampleMR).

Ethics Approval

UK Biobank received ethics approval from the National Health Service National Research Ethics Service, and participants provided written informed consent. No ethics approval was required for the analysis using publicly available data (CARDIoGRAMplusC4D 1000 Genomes-based GWAS).

RESULTS

Among 502,642 participants in the UK Biobank, 442,698 (88%) were British white. After excluding those who had permanently withdrawn, with poor quality or missing genotype, with a mismatch between self-reported and genetic sex or ancestry, had sex chromosome aneuploidy, or extensive relatedness, 392,038 participants remained for subsequent analyses. Among these 392,038 people, 158,601 had prevalent CVD, 29,293 prevalent CAD, and 9,042 prevalent stroke (3,707 ischemic and 1,655 hemorrhagic), with some participants having more than one condition according to the data available in April 2018. Since the baseline recruitment in March 2006, there were 2,313 CVD deaths, including 1,294 CAD and 356 stroke deaths, where the latest date of death was 16 February 2016.

Of the 39 SNPs for HbA_{1c}, 1 SNP (rs1800562, $P = 0.0006$) violated Hardy-Weinberg equilibrium (Supplementary Table 2) and was discarded. Supplementary

Table 1—Association of HbA_{1c} (%) with CVD, CAD, and stroke (overall and stroke subtypes) using Mendelian randomization in the UK Biobank

Instrument	Outcome*	IVW with multiplicative random effects				MR-Egger				Weighted median	
		OR	95% CI	I ² of Wald estimates	OR	95% CI	Intercept	P value for intercept	OR	95% CI	
All SNPs (38)	CVD	1.11	0.83–1.48	0.89	1.12	0.64–1.96	0.000	0.95	1.25	1.07–1.47	
	CAD	1.50	1.08–2.11	0.71	1.25	0.66–2.36	0.004	0.50	1.41	1.02–1.94	
	Stroke	1.39	0.93–2.07	0.41	0.95	0.45–2.01	0.008	0.25	1.23	0.77–1.95	
	Ischemic	1.04	0.62–1.75	0.17	0.49	0.19–1.26	0.017	0.07	0.65	0.33–1.27	
	Hemorrhagic	1.27	0.58–2.81	0.20	1.25	0.28–5.68	0.000	0.98	0.82	0.26–2.59	
	Others	1.84	0.99–3.39	0.41	1.54	0.47–5.03	0.004	0.73	1.42	0.65–3.11	
	Other CVD	0.97	0.75–1.26	0.85	1.07	0.65–1.75	–0.002	0.65	1.08	0.91–1.30	
	Excluded SNPs if associated with potential causes of CAD or confounders (23)†										
	CVD	1.20	1.02–1.42	0.57	1.10	0.81–1.50	0.002	0.50	1.26	1.07–1.47	
	CAD	1.47	1.10–1.97	0.53	0.97	0.59–1.60	0.011	0.07	1.27	0.89–1.81	
Stroke	1.47	1.04–2.07	0.00	0.81	0.43–1.52	0.016	0.04	1.22	0.76–1.95		
Ischemic	0.90	0.53–1.53	0.00	0.53	0.20–1.40	0.014	0.22	0.64	0.30–1.39		
Hemorrhagic	1.69	0.66–4.35	0.30	0.80	0.14–4.47	0.020	0.32	0.86	0.27–2.75		
Others	2.14	1.20–3.81	0.17	1.25	0.43–3.60	0.014	0.25	1.42	0.66–3.07		
Other CVD	1.06	0.90–1.25	0.53	1.15	0.84–1.56	–0.002	0.56	1.11	0.93–1.32		
Excluded SNPs if classified as erythrocytic, associated with potential causes of CAD or confounders (14)‡											
CVD	1.27	1.03–1.56	0.43	0.82	0.51–1.31	0.009	0.07	1.20	0.96–1.51		
CAD	2.24	1.55–3.25	0.39	2.10	0.80–5.50	0.001	0.89	1.90	1.25–2.88		
Stroke	1.61	0.98–2.65	0.00	0.88	0.25–3.03	0.013	0.31	1.40	0.71–2.74		
Ischemic	1.14	0.53–2.47	0.00	0.45	0.07–3.02	0.020	0.31	0.97	0.35–2.73		
Hemorrhagic	1.26	0.40–3.93	0.00	1.89	0.10–35.05	–0.009	0.77	0.72	0.16–3.36		
Others	2.32	0.96–5.59	0.25	1.14	0.12–10.76	0.015	0.51	1.20	0.40–3.61		
Other CVD	0.97	0.82–1.14	0.00	0.65	0.43–0.98	0.008	0.06	0.95	0.76–1.18		

*Definitions of disease as below: prevalent CVD (defined as ICD-9 401–459.9, ICD-10 I10–I99) and its subtypes, including CAD (defined as ICD-9 410–414.9, ICD-10 I20–I25.9), stroke (defined as ICD-9 430, 431, 434, 436; ICD-10 I60, I61, I63, I64), ischemic stroke (defined as ICD-9 434, 436, ICD-10 I63–I64), and hemorrhagic stroke (defined as ICD-9 430, 431, ICD-10 I60–I61). Cases and mortality were both included. †Set 1: SNP excluded if associated with potential causes of CAD or confounders based on public data sources (PhenoScanner and GWAS Catalog) and UK Biobank. ‡Set 2: SNP excluded if classified as erythrocytic based on the HbA_{1c} GWAS or associated with potential causes of CAD or confounders based on public data sources (PhenoScanner and GWAS Catalog) and UK Biobank.

Table 3 lists the 38 SNPs used as genetic instruments. The *F* statistic for the association of the 38 SNPs on HbA_{1c} was 142, suggesting little weak instrument bias. A few SNPs were associated with potential confounders, two SNPs with education (rs9818758 and rs11964178), five SNPs with BMI (rs8192675, rs7756992, rs17747324, rs10774625, and rs1558902), and two SNPs with smoking (rs10774625 and rs17509001) after Bonferroni correction (Supplementary Table 2). According to GWAS Catalog or PhenoScanner, 13 SNPs were related to potential causes of CAD. Based on this information, we repeated the analyses using the two exclusion criteria for choice of SNPs. First, we excluded 15 SNPs related to potential confounders or causes of CAD, leaving 23 SNPs (set 1). Second, we additionally excluded SNPs defined as “erythrocytic” in the original GWAS (9 SNPs) among the 23 SNPs because they did not predict diabetes and hence may be irrelevant to glycemic exposure (i.e., invalid) (21), leaving 14 SNPs (set 2).

Table 1 shows HbA_{1c} was not clearly associated with CVD using all 38 SNPs (odds ratio [OR] 1.11 per %, 95% CI 0.83–1.48). However, higher HbA_{1c} was associated with higher CAD risk using IVW using all 38 SNPs (OR 1.50 per %, 95% CI 1.08–2.11), with directionally consistent results from MR-Egger and weighted median. After we excluded potentially pleiotropic SNPs or those related to confounders (23 SNPs: set 1), the positive associations remained for CAD in IVW (OR 1.47 per %, 95% CI 1.10–1.97), with directionally consistent results from the weighted median method. The results for CAD were most consistent across IVW, MR-Egger, and weighted median when we further restricted SNPs that were nonerythrocytic (set 2). The association of HbA_{1c} with stroke and its subtypes appeared heterogeneous, although these estimates had wide CIs. MR-Egger intercepts suggested little evidence of directional pleiotropy in all analyses. Heterogeneity in the Wald estimates decreased after potentially invalid SNPs were removed.

Table 2 shows HbA_{1c} was positively associated with CAD in CARDIoGRAMplusC4D 1000 Genomes-based GWAS using IVW (OR 1.52 per %, 95% CI 1.03–2.26), with directionally consistent estimates from

Table 2—Association of HbA_{1c} (%) with CAD using Mendelian randomization in CARDIoGRAMplusC4D 1000 Genomes–based GWAS

Instrument	Outcome	IVW weighting with multiplicative random effects			MR-Egger				Weighted median	
		OR	95% CI	<i>I</i> ² of Wald estimates	OR	95% CI	Intercept	<i>P</i> value for intercept	OR	95% CI
All 38 SNPs	CAD	1.52	1.03–2.26	0.74	1.64	0.73–3.71	–0.002	0.83	1.50	1.09–2.05
23 SNPs (set 1)*	CAD	1.30	0.98–1.73	0.36	2.05	1.21–3.48	–0.011	0.06	1.42	1.02–1.95
14 SNPs (set 2)†	CAD	1.27	0.90–1.78	0.14	1.79	0.75–4.27	–0.007	0.41	1.36	0.90–2.06

*Set 1: SNP excluded if associated with potential causes of CAD or confounders based on public data sources (GWAS Catalog and PhenoScanner) and UK Biobank. †Set 2: SNP excluded if classified as erythrocytic based on the HbA_{1c} GWAS or associated with potential causes of CAD or confounders based on public data sources (GWAS Catalog and PhenoScanner) and UK Biobank.

sensitivity analyses including MR-Egger (OR 1.64 per %, 95% CI 0.73–3.71) and the weighted median (OR 1.50 per %, 95% CI 1.09–2.05). There was little evidence of directional pleiotropy based on the MR-Egger intercept (-0.002 , $P = 0.83$). Similar to the results from UK Biobank, the estimates all similarly suggested detrimental effects of HbA_{1c} on CAD regardless of the SNP selection. Heterogeneity in the Wald estimates decreased after potentially invalid SNPs were removed.

The associations of HbA_{1c} with CVD, CAD, and stroke mortality were less clear (Supplementary Table 4), with wide CIs, most likely due to the low mortality rate in the UK Biobank. Similar conclusions were drawn when we restricted our analyses to unrelated participants (Supplementary Table 5).

CONCLUSIONS

To our knowledge, this is the first Mendelian randomization study using the most recently published GWAS of HbA_{1c} applied to both the UK Biobank and CARDIoGRAMplusC4D 1000 Genomes–based GWAS, encompassing more than 700,000 participants. HbA_{1c} was positively associated with CAD in UK Biobank, which was replicated in CARDIoGRAMplusC4D 1000 Genomes–based GWAS, consistent with previous observational studies and an earlier Mendelian randomization study (1–3,15). We cannot exclude HbA_{1c} being associated with CVD. Our study is suggestive of different effects of HbA_{1c} on stroke subtypes, although the number of events in UK Biobank was not enough to allow precise estimates and should be examined further in large stroke consortiums.

Previous observational studies have consistently reported a positive association of HbA_{1c} with CAD, although they are susceptible to confounding (1–3).

Randomized controlled trials targeting HbA_{1c} reduction are difficult to interpret given the interventions, primarily on lifestyle, may have multiple effects that do not necessarily only reflect the effect of HbA_{1c} on CAD (33,34). Our study adds by showing that higher HbA_{1c} is positively associated with CAD using a Mendelian randomization study in two different large studies as well as using different analytics and SNP selections. Combining the results obtained from UK Biobank and CARDIoGRAMplusC4D did not change the conclusion (Supplementary Table 6). Considering triangulation of the evidence from different designs with different underlying assumptions, HbA_{1c} may be causal in the development of CAD in the general population (28). The exact mechanistic pathways remain to be elucidated.

Our study does not provide strong evidence for the same magnitude of association of HbA_{1c} with CVD, contrary to previous observational studies (17). This discrepancy could indicate potential confounding or selection bias in observational studies. Potentially different associations of HbA_{1c} with CAD and other CVD subtypes is consistent with the argument that CVD subtypes have different etiologies with different contributions of each factor (35). UK Biobank had a low response rate (~5%) at recruitment, although a low response rate at recruitment does not necessarily invalidate causal inference (36). This is evident from the similar estimates obtained from the UK Biobank and CARDIoGRAMplusC4D, which used different study designs and sampling approaches. Recruitment of generally healthier people into the UK Biobank study would also not explain the different findings for CAD and CVD. Alternatively, given the UK Biobank recruited from age 40 years to 69 years, with an

average age of 57 years, a different pattern of death by age from specific types of CVD related to HbA_{1c} genetics would artifactually generate different associations of HbA_{1c} with CVD by subtype (37) because of varying levels of left truncation from the underlying birth cohort.

Although we included more than 700,000 participants in this study and used Mendelian randomization to reduce confounding, some limitations exist. The validity of Mendelian randomization depends on whether the three underlying assumptions, as described in the RESEARCH DESIGN AND METHODS, are satisfied; that is, the instruments predict the exposure, the instruments are not confounded, and the instruments affect the outcome only via the exposure (25). In our study, we used genetic variants predicting HbA_{1c} identified in GWAS of people of European descent to reduce weak instrument bias (indicated by the *F* statistics). Restricting the samples to adults mostly of European descent reduces the likelihood of confounding by population stratification.

We also assessed the associations of the genetic variants with potential confounders and found little association with most confounders (Supplementary Table 2), which would not have been possible using summary statistics from GWAS. Although we could not assess whether the genetic instruments were associated with the outcomes only via their association with HbA_{1c} (exclusion-restriction assumption), we conducted several sensitivity analyses, such as MR-Egger and a weighted median, which have different assumptions for validity, although MR-Egger has reduced statistical power. We also repeated the analyses excluding potentially pleiotropic SNPs, which may violate the exclusion-restriction assumption (38). Given the

consistent results for HbA_{1c} on CAD for these different approaches with different assumptions, the association of HbA_{1c} on CAD is likely to be causal. Repeating the analyses for CAD by sex, as a check, showed similar patterns (Supplementary Table 7). When we repeated the CAD analyses without self-reports, the results were most consistent excluding erythrocytic SNPs (data not shown). We also repeated the analyses restricted to erythrocytic SNPs and found a less clear relation of HbA_{1c} with CAD. This is expected because these SNPs did not predict diabetes, and hence, these SNPs are likely irrelevant to the glycemic exposure (Supplementary Table 8) (21).

Although we used one of the largest possible studies, the relatively low number of stroke cases led to imprecise estimates. The suggestive differences in the relation of HbA_{1c} and stroke subtypes seen in our study should be examined elsewhere using large GWAS consortium or settings where stroke is more prevalent, such as China (39).

Lastly, we were unable to use an allele score approach, which may increase statistical power because HbA_{1c} was not available from the UK Biobank at the time this study was conducted.

Our study provides more evidence of a causal role of HbA_{1c} in CAD. Interventions that target HbA_{1c} reduction may be potential targets for reducing the global burden of CAD. Future studies should also clarify the effect of HbA_{1c} on CVD subtypes, which may provide additional insight into the global distribution of CVD subtypes such as stroke, which is more prevalent in Asians.

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interpreted the results. S.L.A.Y. performed analyses with feedback from S.L. and C.M.S. S.L.A.Y. wrote the first draft of the manuscript with critical feedback and revisions from S.L. and C.M.S. S.L.A.Y., S.L., and C.M.S. gave final approval of the version to be published. S.L.A.Y. is the guarantor of this study and, as such, had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

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