



# A Type 1 Diabetes Genetic Risk Score Can Identify Patients With GAD65 Autoantibody–Positive Type 2 Diabetes Who Rapidly Progress to Insulin Therapy

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## OBJECTIVE

Progression to insulin therapy in clinically diagnosed type 2 diabetes is highly variable. GAD65 autoantibodies (GADA) are associated with faster progression, but their predictive value is limited. We aimed to determine if a type 1 diabetes genetic risk score (T1D GRS) could predict rapid progression to insulin treatment over and above GADA testing.

## RESEARCH DESIGN AND METHODS

We examined the relationship between T1D GRS, GADA (negative or positive), and rapid insulin requirement (within 5 years) using Kaplan-Meier survival analysis and Cox regression in 8,608 participants with clinical type 2 diabetes (onset >35 years and treated without insulin for ≥6 months). T1D GRS was both analyzed continuously (as standardized scores) and categorized based on previously reported centiles of a population with type 1 diabetes (<5th [low], 5th–50th [medium], and >50th [high]).

## RESULTS

In GADA-positive participants (3.3%), those with higher T1D GRS progressed to insulin more quickly: probability of insulin requirement at 5 years (95% CI): 47.9% (35.0%, 62.78%) (high T1D GRS) vs. 27.6% (20.5%, 36.5%) (medium T1D GRS) vs. 17.6% (11.2%, 27.2%) (low T1D GRS);  $P = 0.001$ . In contrast, T1D GRS did not predict rapid insulin requirement in GADA-negative participants ( $P = 0.4$ ). In Cox regression analysis with adjustment for age of diagnosis, BMI, and cohort, T1D GRS was independently associated with time to insulin only in the presence of GADA: hazard ratio per SD increase was 1.48 (1.15, 1.90);  $P = 0.002$ .

## CONCLUSIONS

A T1D GRS alters the clinical implications of a positive GADA test in patients with clinical type 2 diabetes and is independent of and additive to clinical features.

Type 2 diabetes is a progressive disease due to a gradual reduction in the capacity of the pancreatic islet cells ( $\beta$ -cells) to produce insulin (1). The clinical course of this progression is highly variable, with some patients progressing very rapidly to requiring insulin treatment, whereas others can be successfully treated with lifestyle changes or oral agents for many years (1,2). Being able to identify patients likely to rapidly

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progress may have clinical utility in prioritization monitoring and treatment escalation and in choice of therapy.

It has previously been shown that many patients with clinical features of type 2 diabetes have positive GAD65 autoantibodies (GADA) and that the presence of this autoantibody is associated with faster progression to insulin (3,4). This is often termed latent autoimmune diabetes in adults (LADA) (5,6). However, the predictive value of GADA testing is limited in a population with clinical type 2 diabetes, with many GADA-positive patients not requiring insulin treatment for many years (4,7). Previous research has suggested that genetic variants in the HLA region associated with type 1 diabetes are associated with more rapid progression to insulin in patients with clinically defined type 2 diabetes and positive GADA (8).

We have recently developed a type 1 diabetes genetic risk score (T1D GRS), which provides an inexpensive (\$70 in our local clinical laboratory and <\$20 where DNA has been previously extracted), integrated assessment of a person's genetic susceptibility to type 1 diabetes (9). The score is composed of 30 type 1 diabetes risk variants weighted for effect size and aids discrimination of type 1 diabetes from type 2 diabetes. The T1D GRS has advantages over HLA typing alone, as it includes more genetic information, is cheaper than conventional HLA typing, and represents a continuous scale of likelihood of type 1 diabetes susceptibility. In young-onset adults (diagnosed between 20 and 40 years of age), it can predict insulin dependence and is independent of and additive to islet autoantibodies and clinical features (9). It is not known if the T1D GRS will improve the prediction of insulin requirement by GADA in clinically defined type 2 diabetes.

We aimed to determine if the T1D GRS could predict rapid progression to insulin (within 5 years of diagnosis) over and above GADA testing in patients with a clinical diagnosis of type 2 diabetes treated without insulin at diagnosis.

## RESEARCH DESIGN AND METHODS

We examined the relationship between GADA, T1D GRS, and progression to insulin therapy using survival analysis in 8,608 participants with clinical type 2 diabetes initially treated without insulin therapy.

### Study Population

Included participants had a clinical diagnosis of type 2 diabetes after 35 years of age, were treated without insulin for the first 6 months from diagnosis, and were of white European origin. The study complies with the Declaration of Helsinki.

Participants were identified in the following cohorts: Genetics of Diabetes Audit and Research in Tayside Scotland (GoDARTS) (10), Hoorn Diabetes Care System (DCS) (11), Diabetes Alliance for Research in England (DARE) (12), Predicting Response to Incretin Based Agents in Type 2 Diabetes (PRIBA) (13), and MRC MASTERMIND Progressors (14). These cohorts were studies of participants with a clinical diagnosis of type 2 diabetes recruited from primary and secondary care and are population based, with the exception of PRIBA and MRC MASTERMIND Progressors, which account for <10% of participants. Summaries of the cohort recruitment and data collection methods are shown in Supplementary Table 1, and a flow diagram of sample selection is shown in Supplementary Fig. 1.

Participants known to have had GADA testing either performed in clinical practice or prior to diagnosis (through review of electronic laboratory records) were excluded due to the risk of the result influencing the clinician's treatment decision.

In the GoDARTS cohort, participants diagnosed with diabetes before 1 January 1994 were excluded; due to insufficient prescribing information, we were unable to define time to insulin prior to this date. In the DARE cohort, only the participants recruited in the Exeter center with saved serum were included.

### Assessment of Diabetes Progression (Time to Insulin)

For GoDARTS and DCS cohorts, time to insulin was defined from electronic prescription records. For Exeter cohorts (DARE, PRIBA, and MRC MASTERMIND Progressors), insulin treatment, date of commencing insulin, and date of diagnosis were self-reported at a single visit.

### Laboratory Measurement

The Academic Department of Blood Sciences at the Royal Devon and Exeter Hospital measured GADA for all five cohorts at a median diabetes duration of 6.1 years, using the same assay from biobanked samples stored at  $-80^{\circ}\text{C}$ .

GADA was performed using the RSR Limited ELISA (RSR Limited, Cardiff, U.K.) on the Dynex DS2 ELISA Robot (Dynex Technologies, Worthing, U.K.). The cutoff for positivity was  $\geq 11$  units/mL, based on the 97.5th centile of 1,559 control participants without diabetes (15). The lowest reportable value (lowest calibrant) was 5.0 units/mL. The laboratory participates in the International Autoantibody Standardization Program.

The HbA<sub>1c</sub> value at latest follow-up (closest available result, median 10.6 years diabetes duration) was obtained from electronic health care records or measured on a research sample by the Academic Department of Blood Sciences at the Royal Devon and Exeter Hospital.

### Assessment of T1D GRS

The development of the T1D GRS has been described previously (9). In brief, T1D GRS consists of 30 common type 1 diabetes genetic variants (single nucleotide polymorphisms [SNPs]) from HLA and non-HLA loci; each variant is weighted by their effect size on type 1 diabetes risk from previously published literature, with weights for DR3/DR4-DQ8 assigned based on imputed haplotypes. The combined score represents an individual's genetic susceptibility to type 1 diabetes. Variants used to derive the score are shown in Supplementary Table 2. For ease of clinical interpretation, the score is presented in this article as the centile position of the distribution in the Wellcome Trust Case Control Consortium population with type 1 diabetes (16).

In the Exeter cohorts, genotyping was performed using the KASP genotyping assay by LGC Genomics (Hoddesdon, U.K.) as previously described (9). Genotyping in the GoDARTS cohort was performed using custom genotyping arrays (including Immunochip, Cardio-MetaboChip [MetaboChip], and Human Exome array) from Illumina as previously described (17). Genotyping in the DCS cohort was performed with Illumina's HumanCoreExome Array and imputed using IMPUTE2 (18) into the 1,000 Genomes March 2012 reference panel. All SNPs had an INFO (Information content metric in PLINK)  $> 0.8$ .

T1D GRS calculation was not performed if genotyping results were missing for either of the two alleles with the greatest weighting (DR3/DR4-DQ8 or

HLA\_DRB1\_15) or if more than two of any other SNPs were missing.

### Statistical Analysis

We assessed the relationship between time to insulin treatment and each of GADA and T1D GRS using survival analysis. For this analysis, T1D GRS was categorized based on centiles of a population with type 1 diabetes (Wellcome Trust Case Control Consortium [16]): <5th centile (<0.234 [low]), 5th–50th centile ( $\geq 0.234$  and  $\leq 0.280$  [medium]), and >50th centile ( $> 0.280$  [high]) as previously reported (9,19). GADA was dichotomized into negative or positive based on the cutoff for positivity. Participants were then classified into six risk groups from these categories: 1) GADA negative, low T1D GRS, 2) GADA negative, medium T1D GRS, 3) GADA negative, high T1D GRS, 4) GADA positive, low T1D GRS, 5) GADA positive, medium T1D GRS, and 6) GADA positive, high T1D GRS.

Time to insulin data were censored at 5 years (or the latest available time point not on insulin, if earlier). Survival distributions for time to insulin, stratified by risk groups, were estimated using the Kaplan-Meier product limit estimator (20). The proportional hazard assumption was checked visually and failed. Differences in time to insulin among risk groups were therefore compared using the Wilcoxon (Breslow) test. Positive predicted values were obtained from the product limit estimator, which makes allowances for censored observations.

To assess whether clinical characteristics were different across risk groups, we performed the Wilcoxon test for trend (21) on the continuous variables and Pearson  $\chi^2$  test for categorical variables.

To assess whether GADA, T1D GRS (as a continuous covariate), age of diagnosis, and BMI (closest available to diagnosis, median 3 years diabetes duration) are independent predictors of rapid progression to insulin, we performed multivariate Cox proportional hazards regression analysis (22). When T1D GRS was used as a continuous covariate, the proportional hazard assumption was satisfied. T1D GRS and GADA were added in as separate variables and as an interaction term. The log-linearity assumption was checked by examining Martingale-based residual plots and

considered valid. Study of origin was included as a strata variable to control for effects of cohort differences.

As a 10 SNP T1D GRS combining the 10 alleles with the greatest weightings ordered by published odds ratios (Supplementary Table 3) has also been proposed for clinical practice, we repeated survival analysis using T1D GRS defined by this 10 SNP score using the same centile cutoffs for categorization (9). We also estimated survival distributions for risk groups based on imputed HLA-DR3/DR4 genotypes, individually and grouped by number of copies of at-risk alleles.

Median follow-up time was calculated using the reverse Kaplan-Meier method (23). All analysis was performed in Stata/SE 15.1 (StataCorp, College Station, TX).

### RESULTS

We identified 8,608 participants with a clinical diagnosis of type 2 diabetes meeting all of our inclusion criteria. Table 1 shows the characteristics for these participants. A total of 79.9% ( $n = 6,879$ ) had been followed for at least 5 years; median follow-up time, calculated as the median time to censoring (insulin treatment or latest follow-up), was 10.5 (95% CI 10.3, 10.6) years. A total of 7.8% ( $n = 533$ ) of those participants with over 5 years' follow-up had progressed to insulin  $\leq 5$  years, and 3.3% ( $n = 280$ ) of participants were GADA positive (measured at a median 6.1 years diabetes duration). The distribution of participants by low, medium, and high T1D GRS category was 53.2% ( $n = 4,580$ ), 40.7% ( $n = 3,504$ ), and 6.1% ( $n = 524$ ),

respectively. Characteristics of the participants stratified by cohort are shown in Supplementary Table 4.

### High T1D GRS Is Associated With Markedly Higher Rates of Rapid Insulin Requirement in Participants With Positive GADA, but Is Not Associated in Those Who Are GADA Negative

T1D GRS was strongly predictive of rapid insulin requirement in participants with positive GADA (Fig. 1). In GADA-positive participants, those with higher T1D GRS progressed to insulin more quickly ( $P = 0.001$ ); the probability of requiring insulin at 5 years postdiagnosis (positive predictive value) (95% CI) was: 47.9% (35.0%, 62.78%) (high T1D GRS) vs. 27.6% (20.5%, 36.5%) (medium T1D GRS) vs. 17.6% (11.2%, 27.2%) (low T1D GRS).

T1D GRS was not associated with rapid insulin requirement in GADA-negative participants. For the GADA-negative participants, the probability of requiring insulin at 5 years postdiagnosis was similar across all risk groups ( $P = 0.4$ ): 7.4% (5.3%, 10.3%) (high T1D GRS) vs. 7.3% (6.5%, 8.3%) (medium T1D GRS) vs. 6.7% (5.9%, 7.5%) (low T1D GRS).

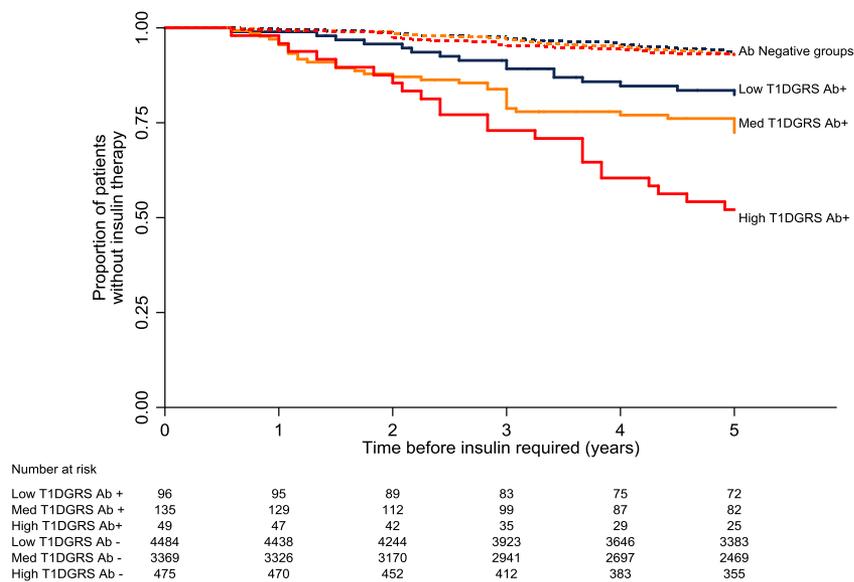
### Differences in T1D GRS Were Associated With Higher HbA<sub>1c</sub> and Lower BMI in GADA-Positive Participants but No Differences in Other Clinical Features

The characteristics of the GADA-positive and -negative participants split by T1D GRS category are shown in Table 2. In GADA-positive participants, HbA<sub>1c</sub> increased ( $P = 0.04$ ) and BMI decreased ( $P = 0.01$ ) with higher T1D GRS category. In GADA-negative participants, clinical

**Table 1—Participant characteristics**

Characteristic	Value
Sex (% male)	56.4
Age at diagnosis (years)	60 (52, 68)
BMI (kg/m <sup>2</sup> )*	30.4 (27.2, 34.7)
Duration of diabetes at latest follow-up (years)	10.6 (6.0, 14.3)
Duration of diabetes at GADA (years)	6.1 (3.3, 10.0)
Insulin treated within 5 years (%)†	7.8
HbA <sub>1c</sub> (%)‡	7.0 (6.4, 8.0)
HbA <sub>1c</sub> (mmol/mol)‡	53 (46, 64)
GADA positive (%)	3.3
T1D GRS centile§	4.2 (0.6, 16.1)

Data are median (interquartile range) or percentage ( $n = 8,608$ ). \*Closest to diagnosis (median 3 years diabetes duration). †Percentage of participants observed for at least 5 years. ‡At latest follow-up. §Centile of participants with type 1 diabetes from the Wellcome Trust Case Control Consortium.



**Figure 1**—Kaplan-Meier plot of probability of requiring insulin therapy during 5-year follow-up by risk group of T1D GRS. Solid lines represent GADA-positive groups, and dashed lines represent GADA-negative groups. Blue indicates low T1D GRS (<5th centile of a population with type 1 diabetes [ $<0.234$ ]), orange is medium T1D GRS (5th–50th centile of a population with type 1 diabetes [ $\geq 0.234$  and  $\leq 0.280$ ]), and red is high T1D GRS (>50th centile of a population with type 1 diabetes [ $>0.280$ ]). Ab, antibody.

characteristics were similar across all categories of T1D GRS.

When comparing the characteristics of GADA-positive and -negative participants (Table 2), GADA-positive participants had a higher T1D GRS (median 0.251 vs. 0.231;  $P < 0.001$ ) and a lower BMI (median 28.73 vs. 30.48;  $P < 0.001$ ) but similar age of diagnosis (median 59 vs. 60 years;  $P = 0.052$ ).

### T1D GRS and GADA Are Predictors of Rapid Insulin Requirement and Are Independent of Age and BMI

Table 3 shows the Cox proportional hazards regression model for time to insulin (censored at 5 years) controlled for effects of cohort differences. As expected, the presence of GADA was a significant predictor of time to insulin (hazard ratio [HR] 3.43 [2.50, 4.71];  $P < 0.001$ ). T1D GRS was independently associated with time to insulin, but only in the presence of GADA (HR per 1-SD increase in T1D GRS 1.48 [1.15, 1.90];  $P = 0.002$ ). These associations were independent of age at diagnosis and BMI.

### A 10 SNP T1D GRS and HLA Type Alone Are Predictive of Future Insulin Requirement

The association between the 10 SNP T1D GRS and rapid insulin requirement was consistent with our findings using the full 30 SNP T1D GRS. The 10 SNP T1D GRS was

associated with rapid insulin requirement in the GADA-positive risk groups ( $P < 0.001$ ) but was not associated in the GADA-negative groups ( $P = 0.4$ ) (Supplementary Fig. 2). In the Cox proportional hazards regression model (Supplementary Table 5), the 10 SNP T1D GRS was independently associated with future insulin treatment in GADA-positive participants (HR per 1-SD increase in T1D GRS 1.34 [1.05, 1.71];  $P = 0.02$ ). Kaplan-Meier plots for HLA-DR3/DR4 genotype risk groups, individually and grouped by number of at-risk alleles, are shown in Supplementary Figs. 3 and 4.

### CONCLUSIONS

In this large study of participants with a clinical diagnosis of type 2 diabetes, we have found that type 1 genetic susceptibility alters the clinical implications of a positive GADA when predicting rapid time to insulin. GADA-positive participants with high T1D GRS were more likely to require insulin within 5 years of diagnosis, with 48% progressing to insulin in this time in contrast to only 18% in participants with low T1D GRS. The T1D GRS was independent of and additive to participant's age of diagnosis and BMI. However, T1D GRS was not associated with rapid insulin requirement in participants who were GADA negative.

To our knowledge, this is the first study to assess the association between an integrated assessment of type 1 genetic risk and GADA in patients with type 2 diabetes or LADA. A key strength of this study is the use of large, predominantly population-based, cohorts of participants diagnosed with type 2 diabetes and, to date, is the largest cohort with measured GADA in a Western population. This means our results are likely to reflect true associations in patients seen in clinical practice. An additional key strength is the use of a single laboratory and assay for measuring GADA across cohorts, with a very robustly defined threshold for positive GADA based on a large, predominantly adult control population. We have demonstrated that our results are independent of and additive to participants' clinical features.

A limitation of our study is that time to insulin has been self-reported in the Exeter cohorts at a single visit, in contrast to other cohorts in which electronic health care records were available. Insulin commencement was also based on clinical decision making rather than a trial protocol. Both of these aspects may introduce imprecision but because both clinicians and participants were unaware of the results, systematic bias would be unlikely. An additional limitation of cross-sectional study design is that GADA was measured at a median 6.1 years diabetes duration, which could result in a lower prevalence than if measurement was undertaken at diagnosis. However, in adult populations, the difference is likely to be small, with GADA positivity being stable over the first 6 years in the UK Prospective Diabetes Study (UKPDS) participants (adult-onset type 2 diabetes) (24) and a modest reduction in prevalence (from 72 to 63%) observed after 8 years in adult-onset type 1 diabetes (25). The results of this study can only be applied to white European populations, and we do not have measurement of other islet autoantibodies in this cohort—the interaction between genetic risk and other islet autoantibodies would be an area of interest for future research (26).

Our findings are consistent with previous research in a population of participants diagnosed with diabetes between 20 and 40 years of age, in whom the same T1D GRS was predictive of insulin-dependent diabetes (9), and other work

**Table 2—Participant characteristics by risk group**

	<5th T1D GRS centile for type 1 diabetes† (low)	5th–50th T1D GRS centile for type 1 diabetes† (medium)	>50th T1D GRS centile for type 1 diabetes† (high)	<i>P</i> value
<b>GADA negative</b>				
<i>n</i> (% of GADA negative)	4,484 (54)	3,369 (40)	475 (6)	
Sex (% male)	56.2	57.0	56.2	>0.1
Age at diagnosis (years)	60 (52, 68)	60 (52, 68)	60 (51, 68)	>0.1
BMI (kg/m <sup>2</sup> )*	30.5 (27.3, 34.9)	30.4 (27.1, 34.6)	30.7 (27.4, 34.3)	>0.1
Duration of diabetes at latest follow-up (years)	10.6 (6.1, 14.4)	10.5 (5.8, 14.1)	10.6 (6.0, 14.4)	>0.1
Duration of diabetes at GADA (years)	6.3 (3.3, 10.0)	6.0 (3.3, 10.0)	6.3 (3.6, 10.2)	>0.1
HbA <sub>1c</sub> (%)†	7.0 (6.4, 8.0)	7.0 (6.4, 8.0)	6.9 (6.4, 8.0)	>0.1
HbA <sub>1c</sub> (mmol/mol)†	53 (46, 64)	53 (46, 64)	52 (46, 64)	>0.1
Insulin treated within 5 years (when observed ≥5 years) (%)	6.7	7.4	8.1	>0.1
GADA (units/mL)	4.9 (4.9, 5.0)	4.9 (4.9, 5.0)	4.9 (4.9, 5.0)	>0.1
<b>GADA positive</b>				
<i>n</i> (% of GADA positive)	96 (34)	135 (48)	49 (18)	
Sex (% male)	51.0	57.0	49.0	>0.1
Age at diagnosis (years)	61 (50, 69)	59 (51, 67)	54 (49, 63)	0.06
BMI (kg/m <sup>2</sup> )*	29.6 (26.7, 34.1)	28.7 (25.6, 32.5)	27.7 (25.4, 30.4)	0.01
Duration of diabetes at latest follow-up (years)	11.1 (9.0, 13.8)	10.4 (6.7, 14.9)	11.8 (9.1, 15.0)	>0.1
Duration of diabetes at GADA (years)	5.2 (3.1, 9.5)	5.6 (3.0, 10.1)	8.9 (4.9, 11.1)	0.01
HbA <sub>1c</sub> (%)†	7.3 (6.6, 9.1)	7.8 (6.7, 9.0)	8.1 (7.1, 9.1)	0.04
HbA <sub>1c</sub> (mmol/mol)†	56 (49, 76)	62 (50, 75)	66 (55, 77)	0.04
Insulin treated within 5 years (when observed ≥5 years) (%)	18.4	27.8	40.5	0.03
GADA (units/mL)	77.6 (24.3, 1,191.9)	111.4 (28.8, 1,354.9)	175.9 (38.6, 1,218.2)	>0.1

Data are median (interquartile range) or percentage. *P* values given for continuous variables are Wilcoxon-type test for trend and Pearson  $\chi^2$  for categorical variables. \*Closest to diagnosis (median 3 years diabetes duration). †At latest follow-up. ‡Centile of participants with type 1 diabetes from the Wellcome Trust Case Control Consortium.

that has shown this risk score to be additive to islet autoantibodies in predicting future type 1 diabetes (27,28). It is also consistent with previous research showing patients defined as LADA who have HLA type associated with type 1 diabetes susceptibility have more rapid progression to insulin (8), and with research showing a combination of positive islet cell autoantibodies and high-risk HLA is associated with low C-peptide in a cohort diagnosed as having type 2 diabetes in contrast to either of these features alone (29). Although the relationship between integrated genetic risk of type 1 diabetes and progression of type 2 diabetes or LADA has not been previously assessed, it has previously

been shown that a type 2 diabetes genetic risk score covering 61 established type 2 diabetes risk variants is not associated with time to insulin (17) and that a 69 SNP type 2 diabetes genetic risk score has very limited utility in discriminating patients with type 1 from type 2 diabetes (9).

The prevalence of positive GADA in our cohorts was lower than in much of the previous literature, with previous multicenter studies reporting widely varying prevalence of positive GADA in populations with type 2 diabetes ranging from 4 to 14% (30,31). In addition to diabetes duration, differences in the prevalence of GADA positivity between our and other studies may be explained by our use of

an assay with higher specificity than used in many other studies (30–34), our lack of an upper age limit (with lower GADA prevalence seen at older ages [4,34,35]), and our use of predominantly population cohorts not selected from secondary care in which treatment with insulin is more frequent. We have used a robustly defined high specificity (97.5%) threshold to define positive GADA in line with current clinical laboratory practice using a large control population. Detectable GADA are commonly found in healthy adult populations without diabetes, and therefore, a threshold based on a control population is recommended to robustly define GADA positivity (32–34). An additional potential reason for low autoantibody prevalence is that we have excluded a small number of cohort participants who had GADA tested in clinical practice, which may have influenced treatment choice. However, only 47 participants were excluded, of whom only 13 were GADA positive, so the effect on overall prevalence is small.

Our findings have clear implications for clinical practice. The T1D GRS represents a novel clinical test that can be used to enhance the prognostic value of GADA testing. For predicting

**Table 3—HR from Cox proportional regression model (adjusted for cohort) for time to insulin censored at 5 years (30 SNP T1D GRS)**

Variable	HR (95% CI)	<i>P</i> value
GADA negative	1	
GADA positive	3.43 (2.50, 4.71)	<0.001
GADA negative: T1D GRS (per 1-SD increase in T1D GRS)	1.02 (0.94, 1.12)	>0.1
GADA positive: T1D GRS (per 1-SD increase in T1D GRS)	1.48 (1.15, 1.90)	0.002
Age at diagnosis (per 1 year)	0.97 (0.96, 0.97)	<0.001
BMI (per kg/m <sup>2</sup> unit)*	1.00 (0.98, 1.01)	>0.1

\*Closest to diagnosis.

future insulin requirement in patients with apparent type 2 diabetes who are GADA positive, T1D GRS may be clinically useful and can be used as an additional test in the screening process. However, in patients with type 2 diabetes who are GADA negative, there is no benefit gained from genetic testing. This is unsurprising, as the prevalence of underlying autoimmunity in patients with a clinical phenotype of type 2 diabetes who are GADA negative is likely to be extremely low; therefore, most GADA-negative participants with high T1D GRS will have nonautoimmune diabetes. The use of this two-step testing approach may facilitate a precision medicine approach to patients with apparent type 2 diabetes; patients who are likely to progress rapidly are identified for targeted management, which may include increased monitoring, early therapy intensification, and/or interventions aimed at slowing progression (36,37).

The costs of analyzing the T1D GRS are relatively modest and may fall further, as genetic testing is rapidly becoming less expensive (38). Although the test cost could potentially be reduced further by using 10 SNPs or imputing HLA type alone, the majority of test costs are attributable to DNA extraction, sample handling, and test interpretation, with cost for genotyping additional SNPs as low as 10 cents per SNP. Savings would therefore be modest, and, although this study does not have sufficient statistical power to directly compare different risk scores in antibody-positive participants, this may come at a cost of reduced test accuracy. The use of a risk score approach has an additional advantage over using HLA alone, as it provides genetic information expressed as a simple-to-use continuous variable.

Although using the T1D GRS alone in those who are GADA positive may have clinical utility, approaches that go beyond single tests and thresholds to integrate the T1D GRS with other predictors of insulin requirement may be the optimal way to use this investigation. The T1D GRS provides a continuous measure of risk, and therefore, dichotomizing the result to use thresholds will lose predictive value. It is additive to other predictive features (for example, age of diagnosis, which is freely available). Although the negative predictive value of a low T1D GRS in

participants with GADA is high (<5th centile 82%), positive predictive values are modest, with the majority of high T1D GRS participants not requiring insulin by 5 years. Therefore, approaches that combine different predictive features on a continuous basis, using prediction models (clinical calculators), may have the greatest utility in accurately predicting future insulin requirement in this group and are an important area for future research (39). Additional areas for future research include the association between T1D GRS and progression in which multiple islet autoantibodies have been tested and assessment in a prospective setting in which islet autoantibodies have been measured at diabetes diagnosis.

In conclusion, a T1D GRS alters the clinical implications of a positive GADA test in patients with clinical type 2 diabetes and is independent of and additive to clinical features. This therefore represents a novel test for identifying patients with rapid progression in this population.

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## References

1. U.K. Prospective Diabetes Study Group. U.K. Prospective Diabetes Study 16. Overview of 6 years' therapy of type II diabetes: a progressive disease. *Diabetes* 1995;44:1249–1258
2. Fonseca VA. Defining and characterizing the progression of type 2 diabetes. *Diabetes Care* 2009;32(Suppl. 2):S151–S156
3. Groop LC, Bottazzo GF, Doniach D. Islet cell antibodies identify latent type I diabetes in patients aged 35–75 years at diagnosis. *Diabetes* 1986;35:237–241
4. Turner R, Stratton I, Horton V, et al.; UK Prospective Diabetes Study Group. UKPDS 25: autoantibodies to islet-cell cytoplasm and glutamic acid decarboxylase for prediction of insulin requirement in type 2 diabetes. *Lancet* 1997;350:1288–1293
5. Tuomi T, Groop LC, Zimmet PZ, Rowley MJ, Knowles W, Mackay IR. Antibodies to glutamic acid decarboxylase reveal latent autoimmune diabetes mellitus in adults with a non-insulin-dependent onset of disease. *Diabetes* 1993;42:359–362
6. Pozzilli P, Di Mario U. Autoimmune diabetes not requiring insulin at diagnosis (latent autoimmune diabetes of the adult): definition, characterization, and potential prevention. *Diabetes Care* 2001;24:1460–1467

7. Liu L, Li X, Xiang Y, et al.; LADA China Study Group. Latent autoimmune diabetes in adults with low-titer GAD antibodies: similar disease progression with type 2 diabetes: a nationwide, multicenter prospective study (LADA China Study 3). *Diabetes Care* 2015;38:16–21
8. Maioli M, Pes GM, Delitala G, et al. Number of autoantibodies and HLA genotype, more than high titers of glutamic acid decarboxylase autoantibodies, predict insulin dependence in latent autoimmune diabetes of adults. *Eur J Endocrinol* 2010;163:541–549
9. Oram RA, Patel K, Hill A, et al. A type 1 diabetes genetic risk score can aid discrimination between type 1 and type 2 diabetes in young adults. *Diabetes Care* 2016;39:337–344
10. Hebert HL, Shepherd B, Milburn K, et al. Cohort profile: Genetics of Diabetes Audit and Research in Tayside Scotland (GoDARTS). *Int J Epidemiol* 2018;47:380–381j
11. van der Heijden AA, Rauh SP, Dekker JM, et al. The Hoorn Diabetes Care System (DCS) cohort. A prospective cohort of persons with type 2 diabetes treated in primary care in The Netherlands. *BMJ Open* 2017;7:e015599
12. DiabetesGenes. Diabetes Alliance for Research in England (DARE) [Internet]. Available from <http://www.diabetesgenes.org/content/diabetes-alliance-research-england-dare-previously-known-exeter-research-alliance-extra-stud>. Accessed 23 November 2017
13. Jones AG, McDonald TJ, Shields BM, et al.; PRIBA Study Group. Markers of  $\beta$ -cell failure predict poor glycemic response to GLP-1 receptor agonist therapy in type 2 diabetes. *Diabetes Care* 2016;39:250–257
14. Royal Devon and Exeter NHS Foundation Trust. RetroMASTER - Retrospective cohort MRC ABPI Stratification and Extreme Response Mechanism in Diabetes. In: *ClinicalTrials.gov* [Internet]. Bethesda, MD, National Library of Medicine, 2018. Available from <https://www.clinicaltrials.gov/ct2/show/NCT02109978>. NLM Identifier: NCT02109978. Accessed 15 February 2018
15. McDonald TJ, Colclough K, Brown R, et al. Islet autoantibodies can discriminate maturity-onset diabetes of the young (MODY) from type 1 diabetes. *Diabet Med* 2011;28:1028–1033
16. Wellcome Trust Case Control Consortium. Genome-wide association study of 14,000 cases of seven common diseases and 3,000 shared controls. *Nature* 2007;447:661–678
17. Zhou K, Donnelly LA, Morris AD, et al. Clinical and genetic determinants of progression of type 2 diabetes: a DIRECT study. *Diabetes Care* 2014;37:718–724
18. Howie BN, Donnelly P, Marchini J. A flexible and accurate genotype imputation method for the next generation of genome-wide association studies. *PLoS Genet* 2009;5:e1000529
19. Patel KA, Oram RA, Flanagan SE, et al. Type 1 diabetes genetic risk score: a novel tool to discriminate monogenic and type 1 diabetes. *Diabetes* 2016;65:2094–2099
20. Kaplan EL, Meier P. Nonparametric estimation from incomplete observations. *J Am Stat Assoc* 1958;53:457–481
21. Cuzick J. A Wilcoxon-type test for trend. *Stat Med* 1985;4:87–90
22. Cox D. Regression models and life tables. *J R Stat Soc B* 1972;34:187–220
23. Schemper M, Smith TL. A note on quantifying follow-up in studies of failure time. *Control Clin Trials* 1996;17:343–346
24. Desai M, Cull CA, Horton VA, et al. GAD autoantibodies and epitope reactivities persist after diagnosis in latent autoimmune diabetes in adults but do not predict disease progression: UKPDS 77. *Diabetologia* 2007;50:2052–2060
25. Schölin A, Björklund L, Borg H, et al.; Diabetes Incidence Study in Sweden. Islet antibodies and remaining  $\beta$ -cell function 8 years after diagnosis of diabetes in young adults: a prospective follow-up of the nationwide Diabetes Incidence Study in Sweden. *J Intern Med* 2004;255:384–391
26. Bottazzo GF, Bosi E, Cull CA, et al. IA-2 antibody prevalence and risk assessment of early insulin requirement in subjects presenting with type 2 diabetes (UKPDS 71). *Diabetologia* 2005;48:703–708
27. Redondo MJ, Geyer S, Steck AK, et al.; Type 1 Diabetes TrialNet Study Group. A type 1 diabetes genetic risk score predicts progression of islet autoimmunity and development of type 1 diabetes in individuals at risk. *Diabetes Care* 2018;41:1887–1894
28. 53rd EASD Annual Meeting of the European Association for the Study of Diabetes: Lisbon, Portugal, 11 - 15 September 2017 [published correction appears in *Diabetologia* 2017;60:609]. *Diabetologia* 2017;60:1–608
29. Groop L, Miettinen A, Groop P-H, Meri S, Koskimies S, Bottazzo GF. Organ-specific autoimmunity and HLA-DR antigens as markers for  $\beta$ -cell destruction in patients with type II diabetes. *Diabetes* 1988;37:99–103
30. Buzzetti R, Zampetti S, Maddaloni E. Adult-onset autoimmune diabetes: current knowledge and implications for management. *Nat Rev Endocrinol* 2017;13:674–686
31. Laugesen E, Østergaard JA, Leslie RDG; Danish Diabetes Academy Workshop and Workshop Speakers. Latent autoimmune diabetes of the adult: current knowledge and uncertainty [published correction appears in *Diabet Med* 2015;32:1670]. *Diabet Med* 2015;32:843–852
32. Bonifacio E, Yu L, Williams AK, et al. Harmonization of glutamic acid decarboxylase and islet antigen-2 autoantibody assays for National Institute of Diabetes and Digestive and Kidney Diseases consortia. *J Clin Endocrinol Metab* 2010;95:3360–3367
33. Bingley PJ. Clinical applications of diabetes antibody testing. *J Clin Endocrinol Metab* 2010;95:25–33
34. Tuomi T, Carlsson A, Li H, et al. Clinical and genetic characteristics of type 2 diabetes with and without GAD antibodies. *Diabetes* 1999;48:150–157
35. Tuomi T, Santoro N, Caprio S, Cai M, Weng J, Groop L. The many faces of diabetes: a disease with increasing heterogeneity. *Lancet* 2014;383:1084–1094
36. Florez JC. Precision medicine in diabetes: is it time? *Diabetes Care* 2016;39:1085–1088
37. Leslie RD, Palmer J, Schloot NC, Lernmark A. Diabetes at the crossroads: relevance of disease classification to pathophysiology and treatment. *Diabetologia* 2016;59:13–20
38. Christensen KD, Dukhovny D, Siebert U, Green RC. Assessing the costs and cost-effectiveness of genomic sequencing. *J Pers Med* 2015;5:470–486
39. Hattersley AT, Patel KA. Precision diabetes: learning from monogenic diabetes. *Diabetologia* 2017;60:769–777