Birthweight and mortality in adulthood: a systematic review and meta-analysis

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Background** Small birth size may be associated with increased risk of cardiovascular diseases (CVD), whereas large birth size may predict increased risk of obesity and some cancers. The net effect of birth size on long-term mortality has only been assessed in individual studies, with conflicting results.

**Methods** The Meta-analyses of Observational Studies in Epidemiology (MOOSE) guidelines for conducting and reporting meta-analysis of observational studies were followed. We retrieved 22 studies that assessed the association between birthweight and adult mortality from all causes, CVD or cancer. The studies were systematically reviewed and those reporting hazard ratios (HRs) and 95% confidence intervals (95% CIs) per kilogram (kg) increase in birthweight were included in generic inverse variance meta-analyses.

**Results** For all-cause mortality, 36,834 deaths were included and the results showed a 6% lower risk (adjusted HR = 0.94, 95% CI: 0.92–0.97) per kg higher birthweight for men and women combined. For cardiovascular mortality, the corresponding inverse association was stronger (HR = 0.88, 95% CI: 0.85–0.91). For cancer mortality, HR per kg higher birthweight was 1.13 (95% CI: 1.07–1.19) for men and 1.04 (95% CI: 0.98–1.10) for women (Pinteraction = 0.03). Residual confounding could not be eliminated, but is unlikely to account for the main findings.

**Conclusion** These results show an inverse but moderate association of birthweight with adult mortality from all-causes and a stronger inverse association with cardiovascular mortality. For men, higher
that occur later in life. Evidence from animal studies suggests that fetal undernutrition may increase susceptibility to diseases throughout the life course to determine disease susceptibility. 

High birthweight has also been associated with increased risk of adverse adult health outcomes, such as overweight and type 1 diabetes, which are important determinants of adult mortality. There is also evidence that high birthweight is associated with higher risk of some adult cancers. In particular, birth size has been positively associated with breast cancer risk in many studies. In Western countries, birthweights have increased over recent decades and it is important to assess whether a high birthweight may be associated with adverse effects on long-term survival.

Birthweight appears to be inversely associated with adult mortality, but there is conflicting evidence related to the importance of potentially confounding factors such as socio-economic status and gestational age. Also, there is conflicting evidence of possible sex differences and in some studies, the results suggest increased adult all-cause mortality for the highest birthweights instead of a linear inverse association across the spectrum of birthweights.

The aim of this systematic review and meta-analysis was to determine whether birthweight, or other measures of birth size, is associated with adult mortality from all-causes, CVD and all cancers.

Methods

The proposal for conducting and reporting meta-analyses of observational studies in epidemiology (MOOSE) was followed.

Study inclusion criteria

Included studies have assessed the association of birthweight with one or more of the following causes of adult mortality: all-causes, CVD or cancer. Historical and prospective cohort studies were eligible, retrospective studies were not. Inclusion was restricted to studies where the majority (>80%) of deaths occurred after 15 years of age. Studies that predominantly included twins were excluded. For studies where information had been published in more than one report, we used data that included the longest follow-up. All-causes of death were defined according to the International Classification of Diseases (ICD); deaths from CVD were defined according to ICD8 and ICD9: 390–458 and ICD10: 100–199. Mortality from cancer was defined according to ICD8 and ICD9: 140–290 and ICD10: C00–C99.

Search strategy

Searches were conducted through October 2010 using MEDLINE from 1950 (via ISI web of science), PubMed (from 1966), Ovid EMBASE (from 1980) and Google Scholar (www.google.com). The following terms were used as keywords to locate studies: ‘birthweight’ and ‘birth weight’, ‘birth length’, ‘head size’, ‘ponderal index’, ‘birth size’, all separately combined with: ‘all-cause mortality’, ‘adult mortality’, ‘mortality’, ‘cohort study’, ‘etiology’, ‘cardiovascular mortality’ and ‘cancer mortality’. We also used an alternative search strategy conducted by a medical librarian with special training in searching the medical literature (details available upon request).

References of studies eligible from full text, including cross references, were searched without restriction by language of publication. Unpublished abstracts were excluded. Searches were performed independently by the lead author and a librarian specialized in medical database searches. Titles identified from searches were first checked for relevance by the lead author. Abstracts of studies regarded as relevant on the basis of the title were assessed for eligibility. The full text of studies regarded as potentially eligible by abstract were assessed to decide whether the studies were eligible.
should be included. Two reviewers (K.R.R. and G.W.J.) worked independently to search eligible studies for inclusion, data extraction and methodological quality. Data extraction and assessment of quality were performed using a structured form, and differences resolved by consensus and discussion with a third reviewer (M.B.B.).

Methodological quality assessment, including assessment of potential biases, were based on 'The Newcastle–Ottawa quality assessment scale for cohort studies' (http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp), which is evaluated to be one of few useful tools in assessing quality of observational studies. For each study, a maximum of 8 points could be achieved and a priori, high-quality studies had ≥6 points and medium- or low-quality studies <6 points.

### Measures of exposure

For individual studies, hazard ratios (HRs) with 95% confidence intervals (95% CIs) per unit increase in each birth size measure were cited from the published report. The estimate most completely adjusted for confounders in the original study was used in the meta-analysis. Authors of the studies who did not report estimates that could be used for the meta-analyses were contacted to provide this information. For studies where HRs per kilogram (kg) birthweight could not be provided by one of the authors, the reported results were sought to be converted into appropriate estimates.

In one study, results were reported per units of standard deviation (SD) in grams (g) and HRs and CIs could be calculated per kg. Since many published studies have assessed these associations by sex and some have suggested sex differences, the main analyses were stratified by sex. For the meta-analyses, sex-specific HRs per 1 kg increase in birthweight were calculated for each study. Sex interactions were tested using a chi-squared test between the sex-specific pooled results for each outcome of interest. An additional categorical analysis was performed for all-cause mortality according to the birthweight categories (<3000 g and >4000 g, compared with the referent category 3000–4000 g). For this analysis, each HR from individual studies was adjusted for sex before the combined result for men and women was entered into the meta-analysis. To assess the associations of alternative birth size measurements with mortality, sex-specific HRs per centimetre increase in birth length and per 1 U (kg/m³) increase in ponderal index were sought. In two studies, the association of birth length with adult all-cause mortality was assessed and suggested a possible inverse association of birth length. There were no available data to justify the pooled analysis for birth size measures other than birthweight and any of the outcomes of interest.

### Data synthesis

In the analyses, studies were weighted according to the inverse variance of the regression coefficient, and the weighted regression coefficients were combined assuming a fixed effect model. An alternative approach, assuming a random effects model, was evaluated and this revealed the same results. For reports where the variance was not reported, the variance of each study was calculated by deriving the standard error from the reported 95% CI. Separate analyses were performed including the crudest and the most adjusted estimate from each report.

Funnel plots were assessed for symmetry to evaluate the impact of possible publication bias. Given sufficient number of included studies, heterogeneity was assessed using the I² statistic. Sensitivity analyses were performed when the I² statistic showed evidence of moderate or high heterogeneity, corresponding to values of >40% or 75%, respectively. We also performed sensitivity analyses to evaluate potential bias due to exclusion of studies where exact results per kg birthweight were not available. In these analyses, approximate HRs calculated from results in the published reports were added to the meta-analyses. All analyses were performed in Review Manager 5 from the Cochrane Collaboration (www.cc-ims.net/revman).

### Results

#### Availability of data

Initial searches retrieved 1522 articles (Figure 1). After screening titles and abstracts, we identified 22 reports that assessed the association of birthweight with adult mortality from one or more of the following causes: all-causes, CVD and total cancer. All these studies were systematically reviewed. Authors of 10 of the studies contributed additional information by clarifying methodology and providing results that could be included in the meta-analyses. The included studies were all prospective or longitudinal cohort studies and included populations from the UK, Denmark, Sweden, Finland, Norway, The Netherlands, Australia and Israel. The oldest cohort included populations born in the 1850s and the youngest dated from the 1960s and 1970s, with roughly half of all participants born after 1940.

Table 1 shows study characteristics of the 15 reports that assessed the association of birthweight with adult all-cause mortality. Three reports, including 781 deaths, were secondary reports from studies that were already included and these were excluded from the analysis. Two eligible studies, including 222 deaths (3.6% of total deaths), were excluded from the meta-analysis because risk estimates per kg birthweight were unavailable. In total, the meta-analysis for all-cause mortality (Supplementary Figure S1) included 10 studies; 9 studies with
23,839 deaths among men and 9 studies29,32,35–38,40,41,47 including 12,995 deaths among women. The association of birthweight with cardiovascular mortality was reported in 14 eligible reports (Supplementary Table S1). Four secondary reports were excluded4,44,48,49 and 2 studies including 372 male deaths 2,45 were excluded from the meta-analysis of deaths from CVD due to unavailability of data. The meta-analysis for cardiovascular deaths (Supplementary Figure S2) included 9 studies31,32,34,36,38–41,50 with 8570 cardiovascular deaths in men and 2796 cardiovascular deaths in women.

Five studies32,36,38,40,41 that assessed the association of birthweight with adult cancer mortality are presented in Supplementary Table S2. The meta-analysis included 4208 cancer deaths in women and 4176 cancer deaths in men (Supplementary Figure S3).

Birthweight and all-cause mortality

Of the 12 primary reports reviewed, 431,32,38,41 suggested inverse associations of birthweight with adult mortality for men and women. One study suggested an inverse association for men and a positive association for women,29 whereas inverse associations for women and no associations for men were reported in three studies.32,37,46 In several studies, the number of deaths was low and could not yield strong evidence for any association. In the cohort from Australia including individuals born before 1900,47 in a Swedish cohort of women born after 1938,35 in the Dutch famine study40 and in a follow-up study of British men,45 there were no associations with adult mortality. The largest study reported strong evidence for a U-shaped association using five categories of birthweight, whereas there was an over-all inverse association per kg increase in birthweight provided for the present meta-analysis.36

The summary results of the meta-analyses are presented in Figure 2. The results show a reduction in the risk of death from all-causes per kg increase in birthweight; the sex-specific HRs were 0.95 (95% CI: 0.92–0.97) for men and 0.93 (95% CI: 0.90–0.96) for women. Funnel plots did not suggest evidence of publication bias for these associations and the I² statistics suggest evidence of low heterogeneity: I² was 13% for men and 26% for women.

We assessed all-cause mortality related to categories of birthweight using birthweight between 3000 and 4000 g as the reference category. These analyses showed higher mortality for individuals born relatively small (birthweight <3000 g); the sex-adjusted HR was 1.11 (95% CI: 1.08–1.15) for this group (Figure 3A). There was no evidence that relatively high birthweight (>4000 g) was associated with a further reduction or an increase in all-cause mortality.
Table 1 Characteristics of studies reporting on the association between birthweight (BW): and adult all-cause mortality in men (M) and women (W)

<table>
<thead>
<tr>
<th>Study</th>
<th>Country and birth year</th>
<th>No of participants</th>
<th>Deaths</th>
<th>Age*</th>
<th>Published association for birthweight with 95% CI</th>
<th>HRb with 95% CI per kg increase in BW</th>
<th>Adjusted HRc with 95% CI per kg increase in BW</th>
</tr>
</thead>
<tbody>
<tr>
<td>Leon et al.11</td>
<td>Sweden 1915–29</td>
<td>6162 M 5675 W</td>
<td>2039 M 1109 W</td>
<td>64 (32–80)</td>
<td>HR per kg increase: M: 0.92 (0.84–1.00) F: 0.95 (0.84–1.07)</td>
<td>M: 0.92 (0.84–1.00) W: 0.96 (0.84–1.07)</td>
<td>M: 0.94 (0.87–1.03) W: 0.95 (0.84–1.07)</td>
</tr>
<tr>
<td>Friedlander et al.29</td>
<td>Israel 1964–74</td>
<td>40 774 M 39 776 W</td>
<td>383 M 179 W</td>
<td>&gt;15</td>
<td>HR per SD score increase M: 0.92 (0.83–1.03) F: 1.08 (0.92–1.27)</td>
<td>M: 0.85 (0.69–1.05) W: 1.15 (0.83–1.58)</td>
<td>M: 0.79 (0.64–0.98) W: 1.21 (0.85–1.72)</td>
</tr>
<tr>
<td>Andersen and Osler51</td>
<td>Denmark 1953</td>
<td>10 753 M</td>
<td>810</td>
<td>36.4 (15–49)</td>
<td>HR per 500 g increase: 0.93 (0.87–0.99)</td>
<td>0.83 (0.73–0.97) W: 0.86 (0.76–0.98)</td>
<td></td>
</tr>
<tr>
<td>Kajantie et al.32</td>
<td>Finland 1924–44</td>
<td>7203 M 6627 W</td>
<td>1668 M 671 W</td>
<td>56 (26–74)</td>
<td>HR per 1 kg decrease: M: 1.08 (0.96–1.19) F: 1.25 (1.04–1.49)</td>
<td>M: 0.96 (0.87–1.06) W: 0.81 (0.68–0.96)</td>
<td>M: 0.93 (0.84–1.04) W: 0.80 (0.67–0.95)</td>
</tr>
<tr>
<td>Painter et al.40</td>
<td>The Netherlands 1943–47</td>
<td>2254 M and W</td>
<td>82 M 53 F</td>
<td>47 (18–57)</td>
<td>HR per kilo decrease in birth weight: 0.98 (P=0.95)</td>
<td>M: 1.03 (0.65–1.62) F: 1.01 (0.56–1.80)</td>
<td>0.97 (0.67–1.40)</td>
</tr>
<tr>
<td>Syddal et al.38</td>
<td>England 1911–39</td>
<td>37 615 M and W</td>
<td>5698 M 2218 W</td>
<td>60 (13–88)</td>
<td>HR per SD score increase. Adjusted for year of birth M: 0.97 (0.94-0.99) F: 0.96 (0.92–1.00)</td>
<td>M: 0.94 (0.90–0.99) W: 0.92 (0.85–1.00)</td>
<td>M: 0.94 (0.90–0.99) W: 0.92 (0.85–1.00)</td>
</tr>
<tr>
<td>Baker et al.36</td>
<td>Denmark 1936–79</td>
<td>110 860 M 105 604 W</td>
<td>11 149 M 6609 W</td>
<td>48 (25–68)</td>
<td>Adjusted for period of birth. U shaped association over five categories. Elevated risk for the lowest and the highest category. M: 0.95 (0.92–0.99) W: 0.92 (0.88–0.96)</td>
<td>No</td>
<td></td>
</tr>
<tr>
<td>Lapidus et al.35</td>
<td>Sweden 1938–64</td>
<td>786 W</td>
<td>439</td>
<td>38–100</td>
<td>No association. Estimate not published.</td>
<td>1.11 (0.94–1.31)</td>
<td>1.11 (0.94–1.31)</td>
</tr>
<tr>
<td>McCalman et al.47</td>
<td>Australia 1857–1900</td>
<td>3347 M and W</td>
<td>1767 M 1579 W</td>
<td>65 (18–death)</td>
<td>Mean age at death, by four categories of BW: M = W: F = 0.5</td>
<td>M: 1.02 (0.94–1.17) F: 0.95 (0.85–1.02)</td>
<td>M: 1.04 (0.94–1.14) F: 0.96 (0.87–1.06)</td>
</tr>
<tr>
<td>Jokela et al.37</td>
<td>Great Britain 1958</td>
<td>8231 M 7804 W</td>
<td>243 M 138 W</td>
<td>35 (23–46)</td>
<td>ORm per 100 g increase M = W: OR: 0.99 (0.97–1.02)</td>
<td>M: ORm: 0.96 (0.76–1.22) W: ORm: 0.76 (0.55–1.05)</td>
<td>M: ORm: 1.03 (0.80–1.32) W: ORm: 0.82 (0.59–1.15)</td>
</tr>
</tbody>
</table>

(continued)
Table 1 Continued

<table>
<thead>
<tr>
<th>Study</th>
<th>Country and birth year</th>
<th>No of participants</th>
<th>Deaths</th>
<th>Age</th>
<th>Published association for birthweight with 95% CI</th>
<th>HRb with 95% CI per kg increase in BW</th>
<th>Adjusted HRc with 95% CI per kg increase in BW</th>
</tr>
</thead>
<tbody>
<tr>
<td>Not included in meta-analysis</td>
<td></td>
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</tr>
<tr>
<td>Frankel et al.45</td>
<td>England 1920–38</td>
<td>1258 M</td>
<td>130 M</td>
<td>45–69</td>
<td>Mortality percentage over quartiles of birthweight: $P_{\text{trend}} = 0.5$, age adjusted</td>
<td>NA</td>
<td></td>
</tr>
<tr>
<td>Power and Li42</td>
<td>England 1958</td>
<td>17000 M and W</td>
<td>143 M</td>
<td>16–38</td>
<td>Per gestational age and sex-specific SD score: $P_{\text{trend}} = M: 0.94, F: 0.91$</td>
<td>SR</td>
<td></td>
</tr>
<tr>
<td>Roseboom et al.44</td>
<td>The Netherlands 1943–47</td>
<td>2155 M and W</td>
<td>97</td>
<td>18–50</td>
<td>No association. Mortality percentage over four categories of birthweight $P &gt; 0.2$</td>
<td>SR</td>
<td></td>
</tr>
<tr>
<td>Osler et al.43</td>
<td>Denmark 1953</td>
<td>7493 M</td>
<td>541</td>
<td>15–49</td>
<td>Three-level category of birthweight. &lt;2400 g compared with referent &gt;3500 g: HR: 1.69 (1.14–2.59)</td>
<td>SR</td>
<td></td>
</tr>
<tr>
<td>Nilsson et al.46</td>
<td>Sweden 1964–67</td>
<td>2010 M 1982 W</td>
<td>54 M</td>
<td>1–38</td>
<td>SMRo over six categories of birthweight. Men: $F_{\text{trend}} = 0.65$, Women: Inverse: $P_{\text{trend}} = 0.0006$</td>
<td>NA</td>
<td></td>
</tr>
</tbody>
</table>

**Footnotes:**

- Age is the age (years) when participants were followed. Mean age at death (range).
- The less adjusted estimate, only adjusted for age and period of birth (if applicable). Estimates are provided for meta-analysis.
- The most fully adjusted estimate provided from study, provided for meta-analysis.
- Calculated from data in original report; HR per birthweight SD score increase (SD score: M: 552g, W: 520g).
- Adjusted for age, origin and education of mother, SES, pregnancy complication, previous fetal loss.
- Adjusted for paternal occupational status, paternal and maternal lifespan, maternal marital status.
- Adjusted for gestational age (weeks).
- Estimate for men and women combined, adjusted for sex and exposure to famine.
- Adjusted for weight at 1 year.
- Adjusted for education, social class, adult BMI (quadratic).
- Adjusted for birth cohort (1857–70, 1871–80, 1881–90, 1891–00), first child (yes, no), maternal marital status (supported, unsupported), socio-economic status (1–3, based upon paternal occupation), preterm birth (yes, no), residential area at death (four groups).
- Analysis using discrete-time survival analysis.
- Adjusted for fathers' social class, mother smoking before pregnancy and mothers' age.
- Adjusted for maternal smoking, age, parity, placental weight, disparity index.
- NA, not available results (HR per kg) for meta-analysis; SMR, standard mortality ratio; SR, secondary report of study included in meta-analysis; report not eligible for meta-analysis.
compared with the reference group (sex-adjusted HR = 1.02, 95% CI: 0.99–1.05) (Figure 3B). There was no evidence of heterogeneity among studies for these analyses.

Three of the individual studies that were included had assessed departures from linearity in the associations. The largest study showed strong evidence for non-linearity across five categories of birthweight, with the highest risk in the lowest and highest birthweight groups, using a cut-off of 4250 g for the latter group. Omitting that study from the analysis of birthweight categories did not substantially change the overall results. In the large English cohort study, there was no evidence for non-linearity verified by testing a quadratic term of the continuous birthweight measure, and in two studies there were evidence for non-linear associations only for women.

Based on the pre-specified threshold criteria, the quality assessment (Table 2) showed that all included studies except the 19th-century Australian cohort had a score of 6, indicating high quality.

The two largest studies had no information on potential confounders such as socio-economic factors and gestational age, but included adjustment for period of birth. The Danish Metropolitan Study assessed the influence of childhood social indicators (maternal marital status and paternal occupation) and parental lifespan. Although the offspring of unmarried mothers, manual workers and parents with shorter lifespan had relatively higher adult mortality, there was only a marginal attenuation of the association of birthweight with mortality after adjustment for social indicators and parental lifespan. In other studies that included adjustment for indicators of socio-economic status in childhood or adult- hood, the estimates were not substantially changed after adjustment. In the Israeli study, adjustment for maternal education, social class and pregnancy complications strengthened the observed associations. In the Helsinki Birth Cohort Study, adjustment for gestational age in individuals born at term slightly strengthened the associations. Only two studies included information on maternal smoking. One study emphasized this issue and included birthweight and maternal smoking in the same model, suggesting that for men, maternal smoking, but not lower birthweight, may be associated with increased mortality risk in the offspring. In women, however, the results showed that lower birthweight, but not maternal smoking, was
associated with higher mortality. Adjustment for maternal smoking did not substantially alter the results in the British cohort of men born in 1958.

In the present study, subgroup analyses were performed to evaluate confounding. Using adjusted or unadjusted estimates from studies that could adjust for potentially confounding factors resulted in nearly identical pooled estimates (data not shown). In a separate analysis, excluding the two large studies that could not adjust for gestational age or socio-economic factors, the effect estimate [HR per kg increase in birthweight: 0.95 (95% CI: 0.91–0.99)] remained unchanged.

Birthweight and cardiovascular mortality
Consistent inverse associations between birthweight and cardiovascular disease mortality were found across most studies. Three studies, including data from the Dutch famine, a cohort of British men and a cohort of people born in Australia before 1900, did not find any associations of birthweight with cardiovascular mortality. The meta-analysis of cardiovascular mortality (Supplementary Figure S2) showed a reduction in risk per kg increase in birthweight in both men (HR = 0.88, 95% CI: 0.84–0.91) and women (HR = 0.88, 95% CI: 0.82–0.95). For cardiovascular deaths, the $I^2$ statistics were 20% for studies in men and 0% for studies in women. Studies that could adjust for potentially confounding factors, including paternal social class, maternal marital status and gestational age, showed marginal effects of adjustment on the estimated associations.

Birthweight and cancer mortality
The five included studies of birthweight and cancer mortality reported stronger positive associations for men than for women. The meta-analysis
## Table 2 Quality assessment in studies reporting on associations between birthweight and adult mortality from all-causes, CVD and cancer

<table>
<thead>
<tr>
<th>Study</th>
<th>Representativity of cohort, &gt;70% followed (0–1 points)</th>
<th>Birth size from birth record (0–1 points)</th>
<th>Outcome not present at time of study (0–1 points)</th>
<th>Adjusted for listed potential confounders (0–2 points)</th>
<th>Mortality from secure record (0–1 points)</th>
<th>Follow-up &gt;80% (0–1 points)</th>
<th>Follow-up &gt;15 years (0–1 points)</th>
<th>Summary quality assessment (0–8 points)</th>
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<tr>
<td><strong>All-causes, CVD and cancer</strong></td>
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<tr>
<td>Leon et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Birth cohort (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>7</td>
</tr>
<tr>
<td>Painter et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Gestational age, restricted to term births (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>7</td>
</tr>
<tr>
<td>Kajantie et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Gestational age, restricted to term births (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>7</td>
</tr>
<tr>
<td>Syddal et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Birth cohort, weight at 1 year (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>7</td>
</tr>
<tr>
<td>Baker et al.</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>Yes. Birth cohort (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>6</td>
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<tr>
<td><strong>All-causes and CVD</strong></td>
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<tr>
<td>Andersen and Olser</td>
<td>Yes‌&lt;sup&gt;a&lt;/sup&gt;</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Maternal marital status, paternal occupation, maternal and paternal lifespan (2 points)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>8</td>
</tr>
<tr>
<td><strong>All-causes</strong></td>
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<td>Friedlander et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Maternal age, parity, previous fetal losses, pregnancy complications (2 points)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>8</td>
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<tr>
<td>Lapidus et al.</td>
<td>Yes&lt;sup&gt;b&lt;/sup&gt;</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Education, BMI, SES (1 point)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>6</td>
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<td>McCalman et al.</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Marital status, maternal age, parity, SES, year of birth (2 points)</td>
<td>No</td>
<td>No</td>
<td>No</td>
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<td>Jokela et al.</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Paternal SES (1 point)</td>
<td>Yes</td>
<td>No&lt;sup&gt;c&lt;/sup&gt;</td>
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<td><strong>CVD</strong></td>
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<td>Eriksson et al.</td>
<td>Yes&lt;sup&gt;a&lt;/sup&gt;</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Adult diabetes and height restricted to term in subpopulation (2 points)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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<tr>
<td>Risnes et al.</td>
<td>Yes&lt;sup&gt;a&lt;/sup&gt;</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Maternal age, parity, SES. Restricted to term (2 points)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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<tr>
<td><strong>Studies not included in meta-analysis because hazard ratios per kg birthweight were not available</strong></td>
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<td>Barker et al.</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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<td>Frankel et al.</td>
<td>Yes&lt;sup&gt;a&lt;/sup&gt;</td>
<td>No</td>
<td>Yes</td>
<td>Social class, father social class, smoking, CVD risk factors (2 points)</td>
<td>Yes</td>
<td>Yes</td>
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<td>Nilsson et al.</td>
<td>Yes&lt;sup&gt;a&lt;/sup&gt;</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes. Smoke in pregnancy, maternal age and parity (2 points)</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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</table>

<sup>a</sup>Only men included.<br>
<sup>b</sup>Only women included.<br>
<sup>c</sup>Follow-up not continuous.<br>
BMI, body mass index; SES, socio-economic status.
showed that in men, birthweight was positively associated with adult mortality; the HR per kg increase in birthweight was 1.13 (95% CI: 1.07–1.19), whereas in women, there was no strong evidence that birthweight was associated with cancer mortality (HR per kg increase in birthweight = 1.04, 95% CI: 0.98–1.10) (Supplementary Figure S3). These results suggest that the associations may differ by sex ($P_{\text{interaction}} = 0.03$). There was no evidence of high heterogeneity among the studies: $I^2 = 20\%$ for men and 0% for women. In the study from Uppsala, Sweden, adjustment for indicators of adult socio-economic circumstances did not alter the estimated associations.\(^4\) In one study,\(^5\) adjustment for gestational age strengthened the positive association among men. In two other studies, year of birth was included as a covariate.\(^6\) Non-linearity was assessed in the same two studies, but no evidence of non-linear effects was reported. None of the studies of cancer mortality had adjusted for socio-economic factors in childhood or maternal factors.

**Discussion**

This systematic review and meta-analysis included data on nearly 40,000 adult deaths among some 400,000 individuals. The results suggest that birthweight is inversely associated with adult mortality from all-causes; more specifically, we found 6% lower mortality per kg higher birthweight in men and women combined. The association appeared to be stronger at relatively low birthweights and levelled off at higher birthweights. We also found that cardiovascular mortality was 12% lower per 1 kg higher birthweight, whereas for cancer mortality, findings differed by sex. For men, the risk of cancer death was 13% higher per kg higher birthweight, but for women, there was little evidence of any association between birthweight and total cancer mortality.

Although the present study demonstrates consistent results across study populations, there are potential concerns about interpretations of the main findings. The major weakness for all study outcomes was heterogeneity in data on potentially confounding factors such as those related to socio-economic circumstances, gestational age and smoking. In the present analysis, adjustments for confounding made in the original studies were incorporated, but the factors that could be taken into account were limited in some studies.

**All-cause mortality**

Increased risk of CVD associated with lower birthweight is likely to be a main contributor to the inverse association of birthweight with mortality from all-causes. Individual studies with a limited number of cases have reported an association of low birthweight with higher risk of mortality from cirrhosis of the liver\(^3\) and diabetes,\(^3\) whereas for diseases of the nervous system, lung diseases and mental disorders,\(^3\) the direction of the association has differed between studies. Restricted fetal growth may adversely affect growth and development of specific organs, and may result in reduced function of lungs, kidneys, blood vessels, muscles and brain.\(^3\) Associations with other causes of death are also plausible and need to be examined.

Socio-economic status in childhood is inversely associated with adult mortality, especially from cardiovascular disease, smoking-related cancers, cancer of the stomach and possibly chronic lung disease as well as psychiatric, alcoholic and drug-related deaths.\(^6\) Adjustment for parental social class at the time of birth was performed in five studies with no or minimal attenuation of the observed associations. Information on social class was abstracted from records as opposed to self-report, but measures of social class may have been crude, often limited to manual vs non-manual work and the possibility of residual confounding by socio-economic factors cannot be ruled out. Nevertheless, given the minimal change in associations when socio-economic status could be controlled, it seems unlikely that confounding by socio-economic status can account for the main findings of this analysis.

Given the strong relation between length of gestation and birthweight, it is possible that the inverse association of birthweight with all-cause mortality is, at least in part, explained by shorter gestation. Pre-term birth is associated with cognitive\(^5\) and medical disabilities\(^7\) that are associated with increased mortality in adulthood.\(^8\) Adjustment for gestational age in individuals born at term strengthened the association in one study\(^9\) and this may support the interpretation that the higher long-term mortality associated with relatively low birthweight may be attributed to fetal growth restriction. There was, however, no information in the included studies to examine whether the association of birthweight with all-cause mortality could be confounded by premature birth.

The results related to large birth size should be cautiously considered. Women who develop gestational diabetes are more likely to give birth to large babies who are at increased risk of developing diabetes later in life.\(^10\) High birthweight is also associated with increased risk of obesity,\(^11\) which is an important determinant of mortality.\(^12\) Although there was no evidence from the present analyses of an increased mortality risk for individuals born large, the results may be compatible with the results from individual studies suggesting that the association is positive for the highest birthweights when using a higher cut-off for high birthweight than the 4000 g used in our study. The present study did not include appropriate data to test for non-linearity in a meta-regression model, and the use of a standardized meta-regression model would be necessary to confirm the observed non-linearity.
rather than birthweight itself, to be associated with that found indicators of fetal growth restriction, animal studies and may also be supported by studies approach or an individual patient data meta-analysis would be more appropriate to assess this issue.

Cardiovascular mortality

Many studies report an inverse association of birthweight with adult CVD. An earlier meta-analysis of ischaemic heart disease showed 15–20% risk reduction per kg higher birthweight, with no clear differences between estimates for incidence and mortality. Our findings related to cardiovascular mortality are consistent with that study. The strong inverse association of birthweight with the risk of cardiovascular death is also compatible with studies that have suggested a strong inverse association of birthweight with stroke. It has been suggested that the inverse association of birth size is stronger for haemorrhagic stroke than for ischaemic stroke. However, differences between subtypes of stroke were not found in the follow-up of US nurses, nor in a Scottish cohort study. Few cases of haemorrhagic stroke yield imprecise results and further studies are needed.

There is a growing body of evidence that relatively short length of gestation may increase cardiovascular risk later in life. Unfortunately, we could not address this issue, since the included studies had not assessed effects of pre-term birth. Studies including incident cases have found that the effect of gestational age may differ between coronary heart disease and stroke. For stroke, lower gestational age may be associated with increased risk, but it was recently suggested that the association of birthweight with coronary heart disease may be explained by restricted fetal growth, rather than by low gestational age.

Our main findings for CVD mortality were unlikely to be attributed to confounding by socio-economic status early in life. A similar interpretation was made from a previous meta-analysis on birthweight and ischaemic heart disease. This may support an interpretation that the inverse association between birthweight and adult cardiovascular disease is not attributed to confounding by socio-economic status early in life.

Several mechanisms could explain the increased risk of cardiovascular disease associated with being born small. The dominant hypothesis is that intrauterine under-nutrition causes fetal adaptations that are related to adverse cardiovascular risk later in life. This interpretation has been supported by numerous animal studies and may also be supported by studies that found indicators of fetal growth restriction, rather than birthweight itself, to be associated with CVD. However, common genetic factors could be associated with both small birth size and cardiovascular risk. Also, genetic and non-genetic factors are likely to interact over the life course, possibly through epigenetic mechanisms and cause CVD.

Cancer mortality

Higher birthweight was associated with higher risk of cancer in men, but not in women. That the positive association of birthweight was restricted to men may be surprising, since there is strong evidence that birthweight is positively associated with breast cancer risk. The result are compatible with main findings of a follow-up study of individuals born in Uppsala, Sweden, between 1915 and 1929. That study included almost 3000 primary cancers and reported a positive association between birthweight and all cancers combined. For men, the adjusted HR for all cancers associated with 1 SD gestational age-adjusted birthweight was 1.08 (95% CI: 1.02–1.14). In women, the corresponding HR for women <50 years was 1.23 (95% CI: 1.06–1.44), whereas for women aged ≥50 years no association was observed (HR = 0.98 (95% CI: 0.91–1.05)). The lack of association in the older age group of women was explained by the lack of any association of birthweight with the risk of post-menopausal breast cancer in that study. Another follow-up study of more than 200 000 Danes born between 1936 and 1975 included more than 12 000 cases of invasive cancer and reported a 7% increased total cancer risk per kg increase in birthweight. The main result in this study did not differ by sex and/or by menopausal status.

The individual studies included in the present analysis did not report associations for specific cancer types. Mean age at death from cancer was between 50 and 60 years, indicating that the findings are based on relatively early cancer deaths. The results from studies of cancer incidence show conflicting evidence based to specific cancer types. In the Swedish study, birthweight was positively associated with risks for lymphatic and haematopoietic cancers and for colorectal cancer. In the Danish study, there were positive associations of birthweight with the risk of kidney and lung cancer, multiple myeloma and malignant melanoma. Other studies have suggested positive associations of birthweight with prostate and testicular cancer, but results have not been consistent.

The present meta-analysis included all cohort studies reporting associations of birthweight with adult cancer mortality. Consistent findings across studies and a large number of cases limit the possibility that the main findings could be due to chance. The study quality was generally high, there was no evidence of heterogeneity and the estimated associations were homogenous across studies. The individual studies included limited information on potential confounding factors. However, different mechanisms of confounding may apply to the association of birthweight with cancer than with CVD. There is no evidence that adverse socio-economic circumstances in childhood are associated with increased risk of all cancers. In the Uppsala cohort, adjustment
for socio-economic and maternal factors did not substantially change the estimated associations of birth-weight with cancer mortality. However, it is possible that maternal conditions associated with large birth size, such as gestational diabetes or obesity, may influence long-term cancer risk.

The biological mechanisms for the association of pre-natal factors with adult cancer are poorly understood. Birthweight is a marker for different factors related to intrauterine growth, such as maternal factors, growth factors, hormones, epigenetic changes and genetic variations. The positive association of birth size with breast cancer risk observed in many studies has been explained by high levels of hormones that both increase fetal growth and long-term susceptibility to cancer. Factors that promote fetal growth may influence long-term cancer risk by alterations in the number of stem cells or expression of genes that regulate stem cell behaviour. Finally, gene–environment interactions may be an alternative explanation. Environmental influences that enhance fetal growth may cause altered gene expression that increase susceptibility to cancers. Such epigenetic mechanisms have been demonstrated in pregnant mice fed with a high-fat diet resulting in increased incidence of tumours in the offspring, whereas poor maternal nutrition has been suggested to increase breast cancer risk.

Studies are needed to identify which cancers are related to large birth size and more research is required to understand prenatal processes that could be important for the pathogenesis of cancer in adulthood.

Study strengths and limitation

Search strategies for observational studies pose a challenge and incomplete identification of studies is a potential source of bias. We sought to reduce this potential bias by completing two different search strategies that selected the studies separately. Despite very high agreement in the searches, we cannot be certain that some studies were not missed.

The large number of deaths is a strong feature of these analyses, and <1% of all reported deaths were excluded from the analyses due to unavailable data. Very modest heterogeneity among studies suggests robust associations of birthweight with mortality across different study populations. However, all studies were from high-income countries, mainly in northern Europe and we cannot generalize results to societies where health-care resources are scarce and nutrition may be less satisfactory. Participants in the included studies were born across a considerable time span, mainly from around the turn of the 20th century up to 1979, but findings were largely consistent across different birth cohorts. We could not evaluate whether associations differed by age at death, but mean age at death varied between 35 and 60 years in the included studies, indicating that birthweight is a predictor of mortality at fairly young to middle adult age. One study reported constant HRs across the life course.

The included studies were of high quality and individual studies depended on assessment of outcomes from official registries. Misclassification of causes of death may still introduce bias, although it is unlikely that this misclassification is related to birth size and would therefore mainly underestimate associations.

Conclusion

This systematic review reveals evidence that lower birthweight is associated with increased all-cause mortality in men and women. The results also show strong evidence of an inverse association of birthweight with cardiovascular mortality that do not differ by sex. For cancer mortality, there was a strong positive association of birthweight with cancer mortality in men, but not in women. The findings suggest that birthweight is an indicator of developmental processes that influence long-term health. However, the available data cannot determine whether social factors, genetic factors, the intrauterine environment or life course exposures are more influential in explaining the observed associations.

Supplementary Data

Supplementary data are available at IJE online.

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Conflict of interest: The first and second authors were first and last authors of one of the included studies. Assessment of bias for that study was done by the two independent authors (G.W.J. and M.B.B.). There is no conflict of interest related to this work. The lead author has checked references for accuracy and completeness and confirms that this material has not been published previously in a substantially similar form.

References


