A Family Longevity Selection Score: Ranking Sibships by Their Longevity, Size, and Availability for Study

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Initially submitted February 10, 2009; accepted for publication September 1, 2009.

Family studies of exceptional longevity can potentially identify genetic and other factors contributing to long life and healthy aging. Although such studies seek families that are exceptionally long lived, they also need living members who can provide DNA and phenotype information. On the basis of these considerations, the authors developed a metric to rank families for selection into a family study of longevity. Their measure, the family longevity selection score (FLoSS), is the sum of 2 components: 1) an estimated family longevity score built from birth-, gender-, and nation-specific cohort survival probabilities and 2) a bonus for older living siblings. The authors examined properties of FLoSS-based family rankings by using data from 3 ongoing studies: the New England Centenarian Study, the Framingham Heart Study, and screenees for the Long Life Family Study. FLoSS-based selection yields families with exceptional longevity, satisfactory sibship sizes and numbers of living siblings, and high ages. Parameters in the FLoSS formula can be tailored for studies of specific populations or age ranges or with different conditions. The first component of the FLoSS also provides a conceptually sound survival measure to characterize exceptional longevity in individuals or families in various types of studies and correlates well with later-observed longevity.

Aged, 80 and over; family data; longevity; Shannon information

Abbreviations: FLoSS, family longevity selection score; FRS, family risk score.

Exceptional longevity strongly aggregates in families (1–4). Thus, studies of exceptionally long-lived families have the potential to identify genetic variants and other factors contributing to longevity, particularly factors that are too rare to detect in a population-based study (5, 6). Both a family’s exceptionality of survival and its number of living siblings contribute to its value for a genetic study of longevity. A challenge in selecting families for such studies is that the most long-lived individuals may have few or no living siblings to provide DNA and other biologic material. To select families for a study of environmental and genetic factors contributing to longevity, we developed the family longevity selection score (FLoSS). The FLoSS combines a measure of the exceptionality of a family’s survival with a bonus for the presence of living old members; it is computed from information that can be collected relatively easily from an immediate family member.

We compare the FLoSS with an alternative scoring method, the family risk score (FRS). The FRS generalizes the family history score (7) to quantitative phenotypes, creating an average score among family members adjusted for the distribution of the phenotype in the general population (8, 9). We compare the FLoSS and FRS with regard to characteristics of sibships that would be selected by choosing the highest scoring sibships according to the following: sibship size, numbers of living siblings, mean age, and exceptionality of sibling survival. We compare these characteristics in 2 US populations selected for longevity (the Long Life Family Study and the New England Centenarian Study) and in 1
not selected for longevity (the Framingham Heart Study), a geographically based US cohort.

**MATERIALS AND METHODS**

**Metric of estimated survival exceptionality for a family**

We define both a family measure of survival exceptionality that requires knowing the age at death of all members and a method for estimating this measure when all family members have not yet died.

We characterize an individual’s “exceptionality of longevity” at age A via \( p(A) \), the probability that a random person in the same birth cohort survives to at least age A. We use \(-\ln(p(A))\), the information content of the rarity of survival measured by “Shannon information” (10), as the basis for an individual’s exceptionality score. The increment in this function for each additional year of survival rises with \( \ln\) (as \( e\) – 1). This suggests that each person’s score should be of the form, \( \pi \ln(1/\pi) \). Thus, each additional family member, even one who dies younger, substracts. The desired value of \( k \) is 1 because, for each cohort, the probabilities \( p \) in a life table are uniformly distributed on the interval from 0 to 1, and the integral (i.e., mean value) over \(-\ln(p)\) over this interval equals 1. Thus, we choose \(-\ln(p(A)\mid C)\) to measure each member’s survival exceptionality, while one who lives younger subtracts. The desired value of \( k \) is 1 because, for each cohort, the probabilities \( p \) in a life table are uniformly distributed on the interval from 0 to 1, and the integral (i.e., mean value) over \(-\ln(p)\) over this interval is 1. Thus, we choose \(-\ln(p(A)\mid C)\) to measure each member’s survival exceptionality for including in a family score, defined as follows:

\[
SE_f = \sum_{all\ family\ members} (-\ln(p(A\mid C)) - 1),
\]

in which all A’s are ages at death.

For a living person of age A, we can calculate \( A^* \), an expected age at death calculated from life tables. Specifically, \( A^* \) is a cohort-specific expected age at death, conditional on survival to age A; then, \(-\ln(p(A^*\mid C))\) is an estimated measure of the person’s exceptionality of survival. Because exceptional longevity is strongly familial (1, 2, 4–6), those with long-lived siblings are likely to live longer than life-table calculations suggest. (In this sense, \(-\ln(p(A^*\mid C))\) is a conservative estimate of expected exceptionality in families of interest in longevity studies.)

This leads to the following estimated survival exceptionality for a family \( f \), in which each member’s contribution increases the family score only if \(-\ln(p(A\mid C))\) for dead members or \(-\ln(p(A^*\mid C))\) for living members exceeds 1:

\[
est(SE)_f = \sum_{\text{dead family members}} (-\ln(p(A\mid C)) - 1) + \sum_{\text{living family members}} (-\ln(p(A^*\mid C)) - 1).
\]

**Metric incorporating the added value of old living family members**

The score \( \text{est}(SE)_f \) can be computed with any mix of living and dead family members. For families with living members, it is a plausible current estimate of SE\(_f\), the family’s (ultimate) survival exceptionality. When all have died, it equals SE\(_f\). Although \( \text{est}(SE)_f \) should be useful for many purposes, it may not be ideal for selecting families into a study of familial longevity, because it does not particularly value additional older living siblings, who provide more biologic and phenotypic data. Thus, we add a living siblings’ bonus score that satisfies the following principles:

1. Living family members whose age A is already somewhat exceptional, in the sense that \(-\ln(p(A\mid C))\) is greater than 0, should add to the family’s value for genetic studies on exceptional survival beyond the value already captured in their contribution to \( \text{est}(SE)_f \).
2. The amount of added value conferred by these living family members should be a function of \(-\ln(p(A\mid C))\), where A is their current age, rather than their expected age at death (A\(_*\)), since only survival through age A is certain.
3. Living family members who are still so young that \(-\ln(p(A\mid C))\) is negative should neither add to nor subtract from the bonus.

This leads us to define the following living sibling bonus:

\[
\text{LB}_f = \sum_{\text{living family members}} \max\{0, [-\ln(p(A\mid C)) - 1]\}.
\]

Adding this to \( \text{est}(SE)_f \), produces our family longevity selection score, or FLoSS:

\[
\text{FLoSS}_f = \sum_{\text{dead family members}} (-\ln(p(A\mid C)) - 1) + \sum_{\text{living family members}} \left( -\ln(p(A^*\mid C)) - 1 \right) + \max\{0, -\ln(p(A\mid C)) - 1\}.
\]

To illustrate the FLoSS calculation, consider a sibship with an oldest brother born in 1920 and died (at age 88) in 2008,
while 2 other siblings remain living at the end of 2008: a brother born in 1921 (aged 87 years) and a sister in 1930 (aged 78 years). The US Social Security Administration cohort life tables indicate that only about 12.2% of the brothers’ cohort (males born between 1916 and 1925) lives past 88 and 14.3% lives past 87, while 57.4% of the females of the sister’s cohort (born between 1926 and 1935) lives past 78. Furthermore, the living brother is expected to achieve age $A^* = 91$, while the sister’s $A^*$ is 88. The FLoSS is the sum of their 3 scores. Note that $-\ln(0.122) - 1 = 2.104 - 1 = 1.04$ is the score for the dead sibling, while for each of the 2 living siblings the score is computed as $-\ln(p(A^* | C)) - 1 + \max\{0, -\ln(p(A | C)) - 1\}$. From the US cohort tables, $-\ln(p(A^* = 91 | 1920, m)) - 1 = 1.685$, while $-\ln(p(A = 87 | 1920, m)) - 1 = 0.942$, so that $6 = \max\{0, -\ln(p(A = 87 | 1920, m)) - 1\} = 0.942$. Similarly, $-\ln(p(A^* = 88 | 1930, f)) - 1 = 1.253 - 1 = 0.253$, while $6 = -\ln(p(A = 78 | 1930, f)) - 1 = -0.445$ and $\max\{0, -\ln(p(A = 78 | 1930, f)) - 1\} = 0$. These values lead to an est(SE) for the sibship of 3.042 = 1.104 + 1.685 + 0.253, a living siblings’ bonus of 0.942 + 0, and a FLoSS of 3.984. Additional sibships with various survival patterns and high FLoSS scores are described and illustrated in supplementary material (Web Figure 1). (This information is posted on the Journal’s website (http://aje.oxfordjournals.org/).)

An additional consideration is whether to exclude from scoring those family members who died before some age $A_0$ and, if so, what age. Factors to consider include the quality of data on deaths in infancy and the survival phenotype of principal interest, for example, survival from birth, survival conditional on reaching maturity, or survival conditional on reaching a specified older age. In the Long Life Family Study, given our focus on exceptional familial survival over the life span, we calculated survival conditional on reaching age 40 (rather than a younger age), on the basis of several considerations. First, in the early 20th century, deaths below age 2 were often not recorded. Furthermore, in our study cohorts, infectious disease epidemics and wars were prominent contributors to mortality in infancy, adolescence, and early maturity. Moreover, the role of familial factors in survival below age 40 is unclear (2). Hence, we calculated $p(A | C)$ in our analyses as the “probability of survival to age A, conditional upon survival to age $A_0 = 40$.”

We computed this from population birth cohort life tables, using only the subjects alive at age 40 as the referent population. We provide gender- and US birth cohort-specific scores for those who survived to at least age 40 in the online supplementary material (Web Table 1; refer to the Journal’s website) and on the Long Life Family Study website (https://longlifefamilystudy.wustl.edu/FLoSS).

Regarding the actual or expected ages at death required for individuals to make a positive contribution to this score, we note that $-1 \approx \ln(0.37)$, so that $-\ln(p(A | C)) - 1$ is only greater than 0 when $p$ is smaller than 0.37, so that a positive score is achieved only when longevity exceeds the 63rd percentile for members of the same gender/birth year cohort who survived until at least age 40. In the 1900 birth cohort, for example, this corresponds to 77 years for males and 87 for females.

Figure 1. Characteristics of Long Life Family Study sibships in 2007 as ranked by the family longevity selection score (FLoSS) (A) and the family risk score (FRS) (B). These plots show the relation between the percentile ranking of the sibships screened for the Long Life Family Study based on the estimated survival exceptionality, est(SE), on the $y$ axis and the percentile ranking based on the family longevity selection score on the $x$ axis of part A and the family risk score on the $x$ axis of part B. Diagonal lines represent least-square regression lines and 95% confidence bounds. Note that the 3 lines essentially overlap in the plot of part A.

Comparing the FLoSS and the FRS

A FRS can be generated for quantitative phenotypes, adjusting for family size and the distribution of the phenotype in the general population (8, 9). When applied to longevity in sibships, the FRS is the average over all siblings in a family of their individual scores, $S$, where:

Am J Epidemiol 2009;170:1555–1562
S = \frac{(A - \bar{A})}{\sqrt{\bar{A}}}, \text{ if the sibling is dead, and}

= \max(0, \frac{(A - \bar{A})}{\sqrt{\bar{A}}}), \text{ if the sibling is alive,}

where \( A \) denotes the current age (or age at death) of the sibling, and \( \bar{A} \) is the sex/birth year cohort-adjusted life expectancy of that sibling. Note that, in long-lived sibships, each still-living but younger-than-average sibling pulls the FRS toward 0. In addition, the FRS scores a family with 2 exceptional siblings identically to a family with 4 or 8 siblings, each pair of which is just as exceptional as the 2.

Characteristics of populations used to compare FLoSS and FRS

We calculated these scores in 3 sibship samples. Only the first 2 are enriched for longevity.

Sample 1: 660 sibships from the New England Centenarian Study—a nationwide-based sample of centenarians, their siblings, and offspring. In all selected sibships, the proband was aged \( \geq 100 \) years. In 2005, 49% of all siblings were still living, 33% of the sibships contained at least 1 living member, and the average attained age was 91 years.

Sample 2: 1,671 US sibships from the screening pilot phase of the Long Life Family Study—a US–Danish study of long-lived sibships and their offspring (https://longlifefamilystudy.wustl.edu/). Sibships of living probands who were at least 80 years of age were screened in the US component of the Long Life Family Study. In 2007, 58% of all siblings were still living, all sibships contained at least 1 living member (the proband who provided data on the other siblings), and the average age was 83 years.

Sample 3: 766 sibships from the Framingham Heart Study (11)—a sample of longitudinally followed subjects and their families. The Framington Heart Study has enrolled subjects born from 1888 onward; all our sibships had at least 1 sibling born before 1925. In 2004, 13% of all siblings were still living, 21% of the sibships contained at least 1 living member, and the average attained age was 74 years.

Cohort, survivorship, and sibship data

For the New England Centenarian Study, Framington Heart Study, and Long Life Family Study sibships, we estimated gender/birth-year survival functions and calculated the
expected age at death ($A^*$) using the Social Security Administration cohort life tables for successive birth decades, with the 1900 tables for births from 1895 through 1904, the 1910 tables for births from 1905 through 1914, and so on. These tables are available at http://www.ssa.gov/OACT/NOTES/as116/as116LOT.html. In the Long Life Family Study data, siblings with unknown vital status were considered to be dead, with the age at death given (conservatively) as the age at last contact. For an unknown birth year, we used the youngest sibling’s birth year. Both siblings and half-siblings were included in FLoSS score calculations.

Statistical analysis

We estimated Spearman correlations between the percentile ranking based on $\text{est}(SE)_r$ and the FLoSS and FRS. We used summary statistics to illustrate differences between sibships in the lowest 40 percentile groups and in the top 5%, as defined by each score, and Wilcoxon and $t$ statistics to test for their significance. Reported $P$ values are 2 sided.

RESULTS

As shown in Figure 1, the percentile ranking of families screened for the Long Life Family Study based on $\text{est}(SE)_r$, the estimated “survival exceptionality,” is more strongly correlated with the rankings based on the FLoSS ($r = 0.98$) (Figure 1A) than with those based on the FRS ($r = 0.71$) (Figure 1B). As seen in Table 1, the sibships ranked in the top 5% by the FLoSS and FRS have quite different characteristics. Sibships ranked in the top 5% by the FRS average between 4 and 7 years older but are far smaller than those ranked in the top 5% by the FLoSS. Note that 79% of the sibships ranked in the top 5% by the FLoSS have more than 1 living sibling compared with only 3% of the sibships ranked in the top 5% by the FRS.

Figure 2 shows the distribution of the FLoSS and FRS in the sibships of the Framingham Heart Study (left), the Long Life Family Study (middle), and the New England Centenarian Study (right). For comparability, the FLoSS and FRS scores in the Long Life Family Study and the New England Centenarian Study are each standardized to their distributions in the Framingham Heart Study, a sample that is not enriched for longevity. In contrast, the New England Centenarian Study recruited families with centenarians, resulting in a strongly shifted distribution of the standardized FLoSS scores. The FLoSS distribution for the Long Life Family Study screening population is intermediate, because it includes families with some evidence of longevity. The standardized FRS score distributions also shift to the right in the Long Life Family Study and New England Centenarian

![Figure 2](https://academic.oup.com/aje/article-abstract/170/12/1555/157703)
vival exceptionality (SE estimated survival exceptionality score predict actual survival of families with a FLoSS of 7 or more.

In the Long Life Family Study, we set the minimum FLoSS for a family to be eligible at 7. This threshold was chosen by observing that only 0.2% of the FLoSS sibships of the Framingham Heart Study meet this threshold, in contrast to over 30% of the Long Life Family Study screening families and over 40% of families enrolled in the New England Centenarian Study. Thus, families with a FLoSS as large as 7 are extremely rare but findable. Calculation of FLoSS scores in families from additional population-based samples can provide further guidance about appropriate selection thresholds.

The FLoSS as used in the Long Life Family Study can be viewed as a member of a class of family scores that combine selecting long-lived families for studies of genetic and non-genetic factors contributing to longevity. In particular, the FLoSS can identify desirable sibships among families being screened for genetic epidemiologic studies of exceptional longevity. Sibships with a high FLoSS have high ages of the total sibship and of living siblings, as well as high numbers of total and living members.

A high FLoSS can be achieved by the presence of a single individual with extremely long survival (if the other siblings are not short lived) or by the presence of many long (but not extremely)-lived siblings. The modes of transmission of longevity (if any) associated with these types of patterns may differ. For example, the former might reflect a recessive trait; the latter, a dominant. The fact that a high FLoSS value captures both of these patterns is one of its strengths, because it does not exclude longevous families who might not be ascertained by other methods (e.g., setting a single extremely exceptional age for inclusion of individuals or siblings). However, the additional potential heterogeneity of genetic factors captured in high-FLoSS families implies a need for attention to potential differences among subgroups in genetic transmission.

We used the FLoSS to score families defined as members of a single sibship. Although it could be used to score multi-generational groups of relatives, gathering complete information on the parents of old siblings or on widely dispersed groups of near relatives could be challenging. Furthermore, young people contribute little information regarding ultimate survival.

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The FLoSS is the sum of est(SE)_f and LB_f (the living bonus). Its first component is a current estimate of a family’s ultimate, “fully observed,” survival exceptionality, SE_f. Although the living bonus enables the FLoSS to select families that are more desirable for genetic epidemiologic studies of exceptional longevity, est(SE)_f may be a better intrinsic measure of family longevity. As such, est(SE)_f may be particularly useful for examining relations of genotypes and other risk factors to phenotypes of family members and their family’s exceptionality of survival, or for finding subsets of populations that are similar with respect to the exceptionality of their families. Although in our comparison the FLoSS and est(SE)_f correlated equally well with the observed SE_f 10 years later, this may not apply over longer intervals. Note that, when all family members have died, all 3 measures coincide.

The FLoSS as used in the Long Life Family Study can be viewed as a member of a class of family scores that combine an estimated exceptionality of survival with a bonus for living siblings, as in the following:

**DISCUSSION**

A sibship’s FLoSS is a good predictor of its true (not yet fully known) exceptionality of survival. Thus, it is useful for
Longevity selection score = \( \text{est}(SE)_j + w \times LB_j \),  

where \( \text{est}(SE)_j \) is of the form,  
\[
\sum_{\text{dead family members}} (-\ln(p(A|C(a))) - k) 
+ \sum_{\text{living family members}} (-\ln(p(A'|C(a))) - k).
\]

Here, \( w \) and \( k \) are nonnegative constants, \( C(a) \) is a person-specific reference cohort of those born around the same time who survived to some minimum age, \( a \), and \( LB_j \) is a nonnegative bonus for older living siblings. Each parameter can be tailored to the particular needs of other studies.

For the FLoSS, we chose \( w = 1, C(a) = \) all people in the same birth-year, gender, and national cohort who survived to at least age 40, \( k = 1 \), and

\[
LB_j = \sum_{\text{living family members}} \max\{0, [ -\ln(p(A|C(a))) - 1]\}.
\]

An age threshold other than 40 years could be chosen depending on a study’s focus (e.g., age 70 to study factors influencing survival only in advanced age). Choosing \( w = 1 \) weights the estimated survival exceptionality and living bonus equally; \( w = 0 \) yields \( \text{est}(SE)_j \), a measure of family survival exceptionality alone. Note that the expected value of \( \text{est}(SE)_j \) for a “pseudofamily” constructed by grouping \( N \) randomly selected individuals from the US population is \( N \times (1 - k) \). Choosing \( k = 1 \) makes \( \text{est}(SE)_j \) neutral to sibship size, because its expected value should be 0 for randomly selected (not particularly long-lived) people. The “neutrality” of the FLoSS was validated by the fact that its average in the geography-based Framingham Heart Study cohort was quite close to 0. Because \( -\ln(0.37) \) is approximately equal to 1, only ages in the top 37th percentile add to \( \text{est}(SE)_j \). More generally, values of \( k \) smaller than 1 favor larger families (because each “typical” person’s expected score is positive), while choosing \( k \) larger than 1 favors smaller ones. If greater exceptionality is sought, a larger cutoff for the FLoSS could be used.

We chose \( w = 1 \), giving equal weight to the estimated exceptionality score and the “bonus” for living older family members. This reflected our interest in both survival exceptionality and the availability of old living subjects. With this choice, the FLoSS still correlates strongly with \( \text{est}(SE)_j \), as we wanted. Larger \( w \)’s will give more weight to the living bonus and reduce the correlation with \( \text{est}(SE)_j \).

In summary, we have introduced and examined the consequences of a conceptually attractive framework for family longevity studies. These include 1) a measure of an individual’s exceptionality of survival, 2) a feasible way to estimate that exceptionality for those still alive, 3) a size-neutral way to combine individual scores into a family score, 4) a plausible bonus measure for the additional value of already exceptional living family members, and 5) a way to balance interest in older living relatives and family survival exceptionality in a single score such as the FLoSS. This framework should be useful in many settings. We also note that formulas for the FLoSS and \( \text{est}(SE)_j \) could be adapted to measure exceptionality of survival until other events besides death. Thus, we could quantify family risk for the onset of conditions (such as stroke or onset of diabetes or disability) whose incidence rises with age. We are investigating this idea.

ACKNOWLEDGMENTS

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Funded by grants U01AG023755 (P.S., T. P., A. S. A.), U01AG023746 (M. P.), and U01AG023712 (K. C.). The Framingham Heart Study is conducted and supported by the National Heart, Lung, and Blood Institute in collaboration with Boston University.

The authors thank the staff and investigators of the Framingham Heart Study for providing access to their data.

This article was not prepared in collaboration with investigators of the Framingham Heart Study and does not necessarily reflect the opinions or views of the Framingham Heart Study, Boston University, or the National Heart, Lung, and Blood Institute.

Conflict of interest: none declared.

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