

Variability of annual peak flows in the Beijiang River Basin, South China, and possible underlying causes

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ABSTRACT

Peak flows are the most important flood parameter which relatively reflects the highest level and potential destructive power of a flood. Understanding peak flow changes can effectively capture a flood characteristic and is essential for developing flood control strategies. This study aims to reveal how regional peak flows evolved in recent decades, mainly from a non-linear perspective. The Beijiang River Basin (BRB) was chosen for the analysis, and hydrological data from four hydrologic stations were used. Methods including ensemble empirical mode decomposition and rescaled range analysis were applied to advance the research. Results indicate a non-significant uptrend and a multiple periodicity of peak flows in BRB. However, short periods were more distinct than long ones. In the future, peak flows may continue to increase over time. Such changes in peak flows are possibly due to local reservoir operations and the changing South Asian Summer Monsoon (SASM). The research suggests an increasing flood risk and recommends more regional flood adaptations to avoid flood losses for BRB. Synchronously, it provides a reference for studies regarding periodicity and the future trend of peak flows in other regions.

Key words | Beijiang River Basin, ensemble empirical mode decomposition, periodicity, regional peak flows, South Asian Summer Monsoon

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INTRODUCTION

Flood hazards are well known to have a strong impact on society and the environment and, thus, are currently drawing the attention of the hydrological research community. Previous investigations of flood events usually included studies on their mechanisms, impacts and characteristics (Leviandier *et al.* 2000; Marchi *et al.* 2010; Haddad *et al.* 2012; Munoz *et al.* 2012). Among these studies, the analysis of flood parameters, such as duration, frequency and flow, has enabled a more direct understanding of flood mechanisms. As such, it is crucial to analyze variations of these flood parameters in terms of flood assessment and flood forecasts.

Peak flow, the maximum flow during a flood process, is one of the most important flood parameters. On the one hand, peak flows often reveal the aspects of the intensity

and even the total flow of a flood. By extension, it corresponds to the highest flood level and relatively reflects the potential destructive power of a flood. On the other hand, estimation of the design floods of hydraulic structures, such as dikes, spillways and storm water evacuation canals, requires the determination of peak flows based on the stream flow fluctuations (Fill & Steiner 2003; Munoz *et al.* 2012). Consequently, flood control and disaster mitigation planning usually depend a great deal on peak flow examination.

It is an accepted fact that abnormal climate and human activities have triggered extreme precipitation, causing flood events (IPCC 2012). Flood peak flows are expected to become more variable and unpredictable, due to the impacts of global change on climate, storm-weather systems and river

discharge conditions. Recent studies have attempted to discuss how flood peak flows will respond to changes in precipitation and land use. They have proven the close relationship between flood peak flows and precipitation (Gottschalk & Weingartner 1998; Calenda *et al.* 2005; Daniels 2007). Furthermore, these studies support the view that human activities, such as urbanization, deforestation and afforestation, affect flood peak flows (Beschta *et al.* 2000; Guillemette *et al.* 2005; El Alfy 2016; Mei *et al.* 2016). In addition to these issues, several researchers have focused on the variability of flood peak flows. For instance, previous studies have revealed that peak flow variability would decrease when average flows are expected (Pearce *et al.* 1980; Hewlett 1982). The variation coefficient of annual flood peak flows tends to decrease as the catchment area increases (Kuzuha *et al.* 2009). An example of applying a distribution model to flood peak flows is seen in the study of the River Tiber in Rome; the Gumbel distribution best fits the high return period quantiles of peak flows in the River Tiber Basin (Calenda *et al.* 2009).

Many of the studies referenced above focus on peak flow trends and influential factors, rather than periodicities and fractal properties of annual peak flows. A thorough understanding of the periodicity of peak flows can provide a reference for regional flood prediction and help to further explore the influential factors, whilst studying the fractal properties of peak flows can provide a reference for relevant researches of peak flows from a non-linear perspective. In addition, it is particularly necessary for small regions with prosperous economies and large populations, where increased flood peak flows are expected to threaten local residents (Kuzuha *et al.* 2009). However, up to now few studies have focused on changes in peak flows in the Beijiang River Basin (BRB), a prosperous area in South China where major floods often occur in the rainy season (Zhang *et al.* 2007; Wang *et al.* 2012, 2015).

Taking BRB as a study case, the main objectives of this study are to: (1) detect whether basin-scale peak flows showed significant trends in recent decades; (2) reveal how basin-scale peak flows evolved from a non-linear perspective; (3) discuss whether climate change and human activity have impacts on regional peak flow changes and, if so, how. The study would be useful for local regulators to manage and can provide a reference for relevant researches on periodicity and future trends of peak flows across other regions.

The paper is organized as follows. The study area and data sources are described first; methodologies such as Ensemble Empirical Mode Decomposition (EEMD) and R/S are presented next; and results are listed afterwards. Finally, the discussion and conclusion sections are presented.

STUDY AREA AND DATA SOURCES

Study area

The BRB, which includes economically developed cities such as Guangzhou, Zhuhai and Foshan located in the downstream, is one of the three major drainages of the Pearl River Basin, South China (Figure 1). The basin is one of the most prosperous areas in China (Zhou *et al.* 2011). It has 13 tributaries and its watershed area is more than 1,000 km². The annual mean temperature in the BRB is approximately 21 °C. Precipitation in the BRB is usually concentrated in the warm season from April to September, which accounts for nearly 70% of the annual total (Luo *et al.* 2008). During spring and summer, warm air masses from the sea are blocked by the Nanling Mountains, causing frequent rainstorms and leading to subsequent floods in the BRB. The extraordinary historic flood that occurred in June 1994 had a peak flow of up to 19,700 m³/s at Shijiao station and caused serious damage to the basin, including the loss of 371 lives (see www.gdbjdd.com.cn/). The recent historic flood that occurred in July 2006 also brought enormous damage to the local society.

Data sources

The Hydrology Bureau of Guangdong Province (HBGP) collected the flow data used in the study from four hydrological stations located in the main stream of Beijiang River (Figure 1 and Table 1). The quality of the flow data is strictly controlled by HBGP, which is officially responsible for the regional hydrological data. For each station, the annual peak flow series was obtained using the annual maximum value method (Mitosek *et al.* 2006). Therefore, there is only one maximum instant peak flow for each year for each station.

For investigating the possible climatic causes of changes in the annual peak flows in BRB, this study considers five climatic indices: the Multivariate El Niño-Southern

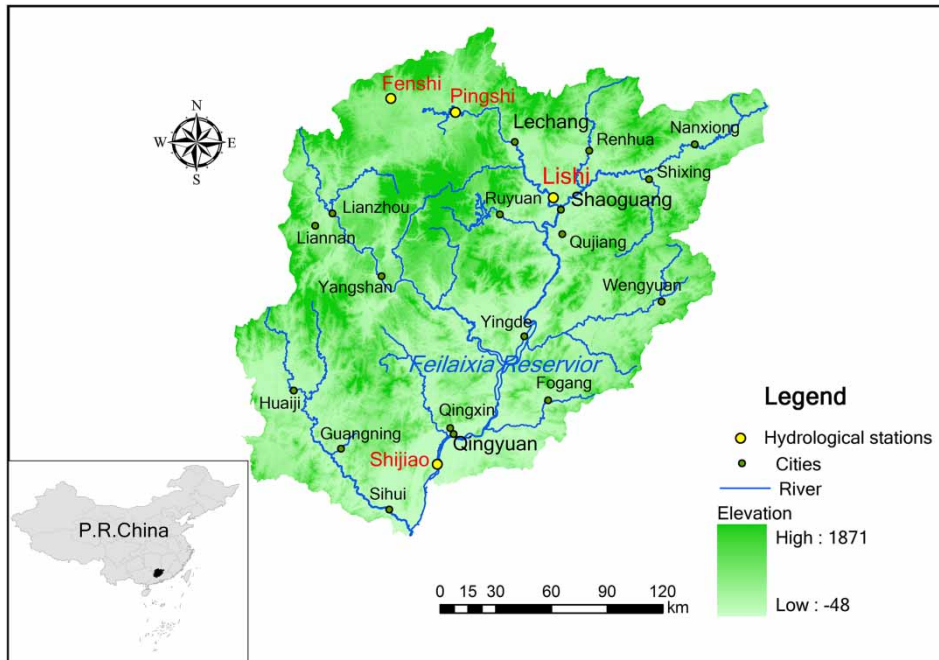


Figure 1 | Study region and the four hydrological stations chosen for the study.

Table 1 | The four chosen hydrological stations in BRB

Station	Fenshi	Pingshi	Lishi	Shijiao
Measurement period	1970–2007	1964–2007	1955–2007	1954–2007
Control area (km ²)	880	3,567	6,976	38,363

Oscillation index (MEI), the East Asian Summer Monsoon Index (EASMI), the South Asian Summer Monsoon Index (SASMI), the South China Sea Summer Monsoon Index (SCSSMI), and the Dipole Mode Index of India Ocean (DMI). EASMI/SASMI/SCSSMI is defined as an area-averaged seasonally dynamical normalised seasonality at 850/850/925 hPa within the East Asian monsoon domain (10–40°N, 110–140°E)/South Asian domain (5–22.5°N, 35–97.5°E)/South China Sea monsoon domain (0–25°N, 100–125°E) (Li & Zeng 2005). These monsoon indices were developed by Li & Zeng (2005). The MEI data set was obtained from the US National Oceanic and Atmospheric Administration Earth System Research Laboratory (www.cdc.noaa.gov/people/klaus.wolter/MEI/). The EASMI, SASMI and SCSSMI data sets were obtained from the United States NCEP/NCAR monthly mean data sets

(<http://jlp.lasg.ac.cn/dct/page/65544>), whilst the DMI data set was obtained from the Japan Marine Science and Technology Center (www.jamstec.go.jp/frcgc/research/d1/iod/HTML/Dipole%20Mode%20Index.html).

METHODOLOGY

Linear regression

Linear regression analysis is one of the most concise methods for detecting trends in a time series. Its main objective is to detect shifts in time series and to describe potential generating processes underlying a given sequence of observations (Nyeko-Ogiramoi *et al.* 2013). The linear regression

equation is described as:

$$Y = aX + b \quad (1)$$

where X is the independent variable, Y is the dependent variable, a is the regression coefficient that reflects the rate of changes of the time series, and b is the regression constant. In this study, the parametric t -test was employed and the P values for the t -test were computed to identify whether the trends in the peak flow series are statistically significant (Yue & Pilon 2004).

Multiple linear regression

Relationships between the peak flows in BRB and some climatic predictors are assessed using the multiple linear regression (MLR). The MLR is an approach which fits a multivariate linear function between a response variable and a set of predictors (Jencso & McGlynn 2011). The multivariate linear function is given as:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_n X_n \quad (2)$$

where Y is the response variable and X_1, X_2, \dots, X_n are the predictors, and $\beta_1, \beta_2, \dots, \beta_n$ are the coefficients corresponding to each of the predictors. β_0 is a constant. For a predictor, X_i , the contribution of X_i to Y is determined by the following equation:

$$\eta_i = \frac{|\beta_i|}{|\beta_1| + |\beta_2| + \dots + |\beta_n|}, i = 1, 2, \dots, n \quad (3)$$

Mann-Kendall analysis

The Mann-Kendall (M-K) method is widely used for identifying abrupt changes in a time series. The advantage of M-K is that the series does not need to comply with a certain distribution of samples, avoiding interference from abnormal values. A detailed algorithm has been described elsewhere (Mosmann et al. 2004; Tabari & Hosseinzadeh Talaei 2013; Wu et al. 2015).

Define a rank statistic S_k , such that:

$$s_k = \sum_{i=1}^k r_i, k = 2, 3, \dots, n \quad (4)$$

where n is the data set record length, r_i is given by:

$$r_i = \begin{cases} 1, & x_i > x_j \\ 0, & x_i \leq x_j \end{cases} \quad (5)$$

where x_i and x_j are data values at times i and j , then define a statistic UF_k which is computed by:

$$UF_k = \frac{[S_k - E(S_k)]}{\sqrt{Var(S_k)}}, k = 1, 2, \dots, n \quad (6)$$

where $E(S_k)$ and $Var(S_k)$ are the mean value and variance of S_k , respectively.

Then define another statistic UB_k . Similarly, the values of UB_k are computed backward, starting from the end of the time series.

By drawing the UF_k and UB_k curves on the same graph, the point of abrupt change in the series can be identified. If the UF_k and UB_k curves intersect, and then diverge, acquiring a specific threshold value of 1.96 at a 0.05 significance level, there is a statistically significant trend. The point where they intersect shows the approximate year at which the abrupt change occurred.

Ensemble empirical mode decomposition

This specific method is derived from the method, Empirical Mode Decomposition (EMD). Before looking at the EEMD, a brief description of the original EMD is proposed. Presented in 1998, EMD is a method for the time series analysis of non-stationary and non-linear signals (Huang et al. 1998). It decomposes non-stationary data into a collection of intrinsic mode function (IMF) components with a residue component (Res.) using different time scales (Bi et al. 2010). The original series is decomposed into different time scale fluctuations, from which the periodicities of the original series can be detected (Ding et al. 2010). The procedure for extracting IMFs from a signal is as follows (Kaluzynski 2014; Kim et al. 2014):

1. Identify all local extrema over the entire time period of the signal, $x(t)$.
2. Create an envelope of the local maxima, $e_{up}(t)$, and minima, $e_{low}(t)$.
3. Calculate the mean of the upper and lower envelopes, i.e.

$$m_1(t) = \frac{(e_{up}(t) + e_{low}(t))}{2} \quad (7)$$

4. Subtract $m_1(t)$ from the signal, i.e.

$$d_1(t) = x(t) - m_1(t) \quad (8)$$

5. Check whether $d_1(t)$ satisfies the following criteria of an IMF:
 - the number of zeros and extrema are equal or differ by no more than 1;
 - the sum of the envelopes of maxima and of minima of an IMF is zero.

If $d_1(t)$ does not satisfy the criteria of an IMF, set $d_1(t)$ to $x(t)$, then repeat steps (1)–(5) until $d_1(t)$ satisfies the criteria. The first signal to satisfy the criteria is termed as the first IMF, $h_1(t)$.

1. Compute the residual, i.e.

$$r(t) = x(t) - \sum_{i=1}^N h_i(t) \quad (9)$$

2. Reiterate steps (1)–(6) until all IMFs are identified. The process is completed when the residual is constant, monotonic or has a single extremum.

Suggested by Wu & Huang (2009), mode mixing can be the significant drawback of EMD (Wu & Huang 2009; Wang et al. 2012). To overcome the shortcoming of EMD, the EEMD method is proposed. It is a new noise assisted analysis method to obtain the actual time–frequency distribution of the seismic signal. With the procedure of extracting IMFs by EMD in mind, the EEMD is developed as follows (Wu & Huang 2009):

1. Add a white noise series to the targeted data.
2. Decompose the data with added white noise series into IMFs.

3. Repeat steps (1) and (2), but with different white noise series each time.
4. Calculate the means of corresponding IMFs of the decompositions as the final results.

The added white noise series would cancel each other in the final mean of the corresponding IMFs. The mean IMFs stay within the natural dyadic filter windows, and therefore reduce significantly the chance of mode mixing and preserve the dyadic property. In this study, the ratio between the standard deviation of the white noise series and that of the targeted data is 0.2, whilst the repeat times of steps (1) and (2) is 20.

Rescaled range analysis

The R/S is used to estimate the fractal properties of a time series. The main idea of the R/S is that one looks at the scaling behavior of the rescaled cumulative derivations either from the mean or from the distance the system travels as a function of time (Karakasidis & Liakopoulos 2004).

For a time series, $\xi(t)$, over a period of time, τ , the range of the accumulated deviation, $R(\tau)$, is given as:

$$R(\tau) = \max_{1 \leq t \leq \tau} X(t, \tau) - \min_{1 \leq t \leq \tau} X(t, \tau) \quad (10)$$

where $X(t, \tau)$ is the accumulated deviation from the mean value of the time series.

To normalize the range relative to the input fluctuations in the series, a standard deviation denoted as $S(\tau)$ is used, calculated by the following equation:

$$S(\tau) = \sqrt{\frac{1}{\tau} \sum_{t=1}^{\tau} [\xi(t) - \langle \xi \rangle_{\tau}]^2} \quad (11)$$

where $\langle \xi \rangle_{\tau}$ is the mean value of $\xi(t)$.

In the case of a fractional Brownian motion in the limit of large τ , the equation:

$$\frac{R(\tau)}{S(\tau)} \propto \tau^H \quad (12)$$

with $0 \leq H \leq 1$, is used.

Equation (7) presents a fitted straight line in a log-log plot of R/S as a function of τ , and H is the slope of the

straight line. In the case of only short-range correlations (uncorrelated), the walk profile displays the properties of a standard random walk with $H = 0.5$. If fluctuations in subsequent values are positively correlated, then $H > 0.5$; if fluctuations in subsequent values are negatively correlated, $H < 0.5$.

RESULTS

Trends

Using linear regressions, trends of the annual peak flows of the four hydrological stations in BRB were analyzed (Figure 2), and the statistical eigenvalues of the four series were calculated (Table 2). As seen in Figure 2, the annual peak flows exhibited large interannual fluctuations and upward trends. The growth rates of the series of Fenshi, Pingshi, Lishi and Shijiao stations are 4.77, 15.46, 19.09 and 27.42 $\text{m}^3/\text{s}\cdot\text{a}$, respectively. The maximum, minimum and mean values of peak flows in Table 2 indicate that the annual peak flows tended to increase from the Fenshi to

the Shijiao station. The P values of the series are respectively 0.375, 0.098, 0.105 and 0.324 (Table 2), all of which are greater than 0.05. In other words, none of the upward trends are statistically significant. The maximum eigenvalues of the time series at the Fenshi, Pingshi, Lishi and Shijiao stations are 1,990, 4,830, 8,800 and 17,400 m^3/s , respectively. With the exception of the maximum at the Fenshi station that occurred in 1994, all the stations' maxima occurred in 2006. Furthermore, the maxima of the series increased from the Fenshi to the Shijiao station. The coefficients of deviation (C_v) of the four series, on the other hand, decrease in value from the Fenshi to the Shijiao station. This signifies that the annual peak flow series of the Fenshi station was more discrete, though with a lower magnitude, compared with the other three stations.

Abrupt changes

The M-K method was applied to identify abrupt changes in annual peak flows in BRB. As described in Figure 3, the UF_k and UF_B curves for the four series intersect at more than one time point, though none of the UF_k curves

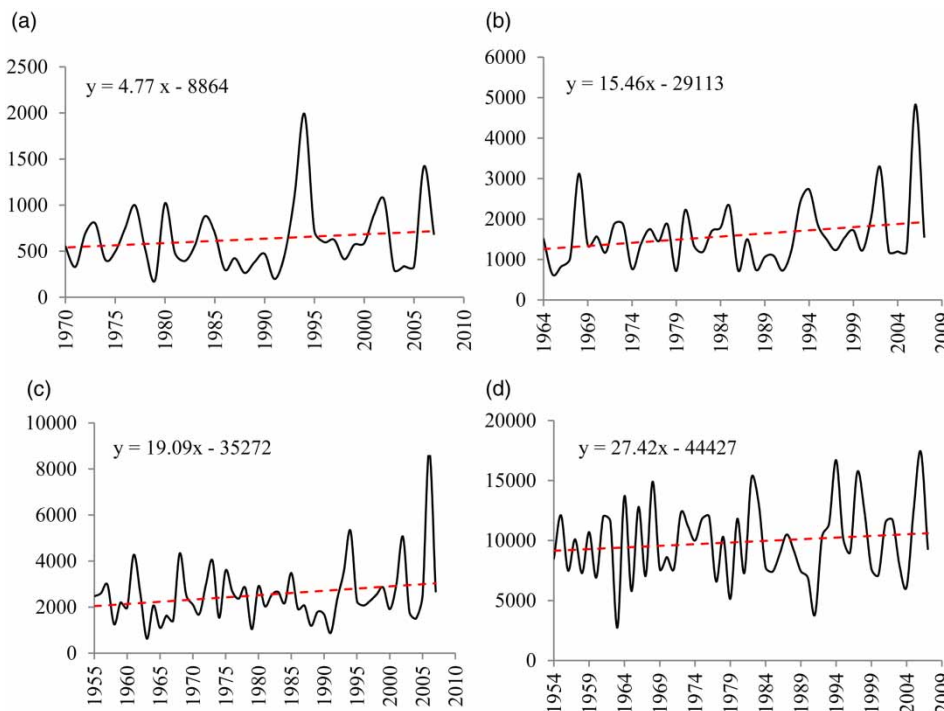


Figure 2 | Inter-annual changes in annual peak flows at the (a) Fenshi, (b) Pingshi, (c) Lishi and (d) Shijiao stations. The dashed straight lines with linear equations denote linear trends of the annual peak flows.

Table 2 | Statistical eigenvalues of the annual peak flows at the four stations in BRB

Station	Fenshi	Pingshi	Lishi	Shijiao
Maximum (m ³ /s)	1,990	4,830	8,800	17,400
Minimum (m ³ /s)	184	636	630	2,740
Mean (m ³ /s)	629	1,591	2,538	9,876
P value	0.375	0.098	0.105	0.324
Coefficient of deviation (C_v)	0.569	0.495	0.521	0.319
Coefficient of skew (C_s)	1.810	1.992	2.402	0.234

$P \leq 0.05$ denotes the trend in the series is beyond the 0.05 significance level.

exceed the critical value of ± 1.96 . Therefore, none of the abrupt change points meet the 0.05 significance level, i.e. the abrupt changes of the annual peak flows in BRB were non-significant for their own measurement periods.

Stationary characteristics and periodicities

Although the trends and abrupt changes in peak flows at the four stations were non-significant, the stationarity characteristics should be further examined. We used the Augmented Dickey–Fuller (ADF) unit root test to analyze the stationarities of peak flow series for the four stations. The ADF method is a classic method which can examine the stationarity of a time series (Zhao & Chen 2015). Table 3 lists

the P value statistics from ADF. Note that the P value statistics in the Fenshi, Pingshi and Lishi series are greater than the critical value 0.05, implying the nonstationary nature of peak flows at these stations. On the contrary, the P value statistic in the Shijiao series is smaller than 0.05, indicating the stationarity of the series. As the Feilaixia reservoir operation may alter, changing the properties of peak flows in the Shijiao station, we omitted the records after 1998, when the Feilaixia reservoir was constructed, and reanalyzed the stationarity of the new series. Results show, as expected, a nonstationary characteristic of the peak flow series covering 1954–1997 at the Shijiao station (Table 3). These results on the one hand can explain the nonstationarity of peak flows in BRB, and on the other hand infer a possible influence of the Feilaixia reservoir operation on the changing property of peak flows at the Shijiao station.

Subsequently, we analyzed the periodicities of annual peak flows at the four stations in BRB using EEMD. For the Shijiao station, to eliminate the effect of the Feilaixia reservoir, only the records covering 1954 to 1997 were used for the analysis. Figure 4 shows the results of IMF and Res., and Figure 5 depicts the power spectra of the corresponding IMFs. The power spectra are used for identifying whether a certain periodicity of an IMF is statistically significant. Note that there are five IMFs and one Res. in each of the series (Figure 4). In terms of EEMD, the first IMF

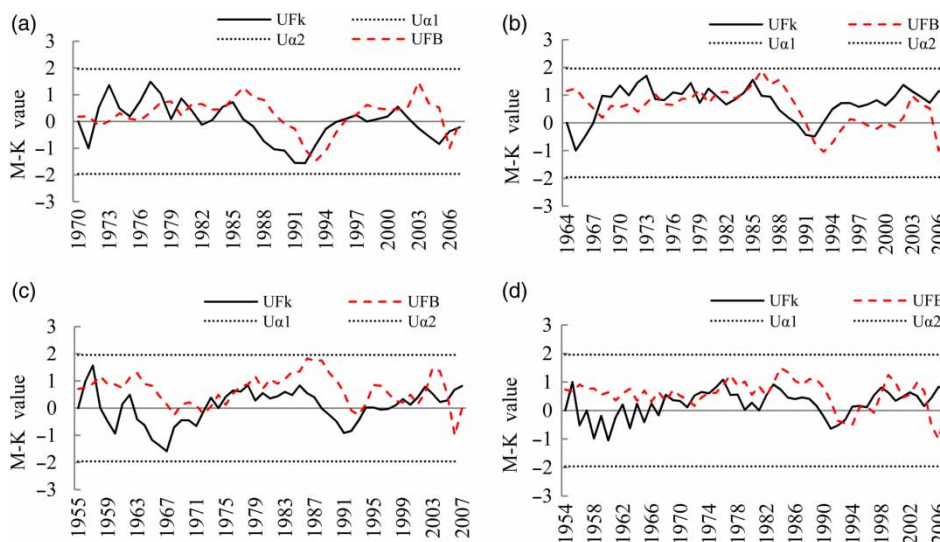


Figure 3 | M-K test results for the annual peak flows of the (a) Fenshi, (b) Pingshi, (c) Lishi and (d) Shijiao stations in BRB. $U_{\alpha 1}$ and $U_{\alpha 2}$ denote the 0.05 significance level lines, whose values are respectively $+1.96$ and -1.96 .

Table 3 | ADF test results of the peak flow series for the four stations in BRB

Station	Dickey-Fuller	Lag order	P value
Fenshi	-3.03	3	0.17
Pingshi	-3.08	3	0.15
Lishi	-3.15	3	0.11
Shijiao (during 1954–2007)	-3.74	3	0.03
Shijiao (during 1954–1997)	-3.46	3	0.06

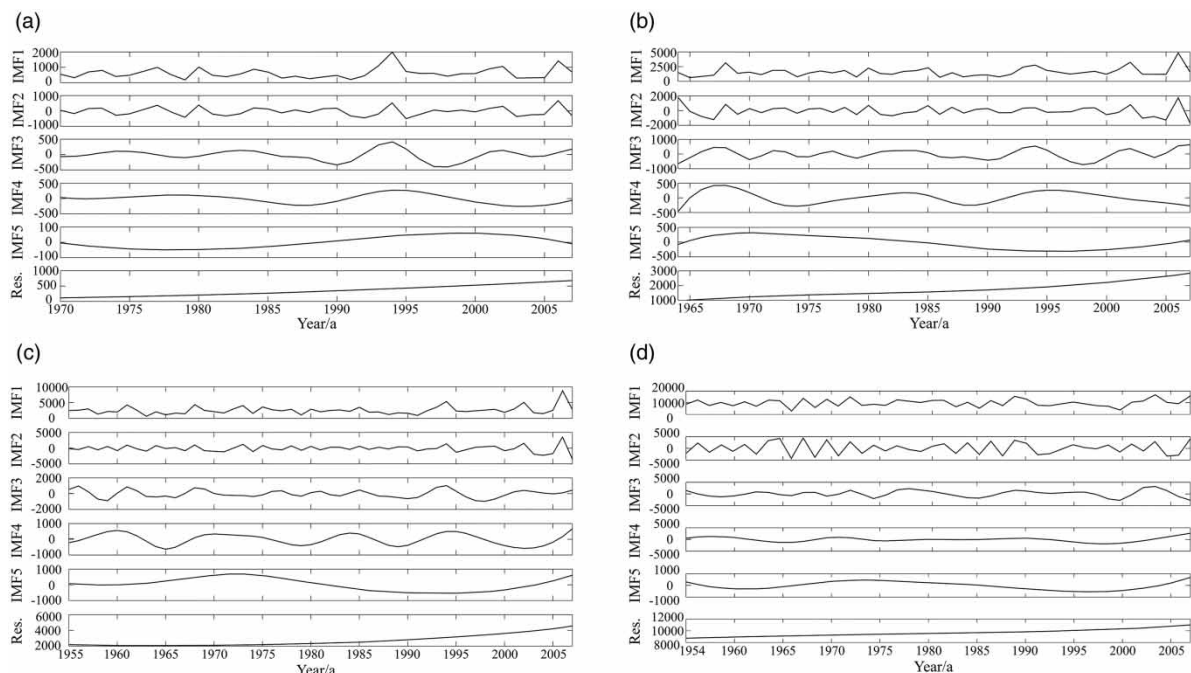
$P \leq 0.05$ and $P > 0.05$ indicate the series are stationary and nonstationary, respectively.

(IMF1) has the largest amplitude and the highest frequency, and the frequency decreases from the first IMF component to the last (Bi et al. 2010). Therefore, the IMFs reflect different periods of the annual peak flow series, i.e. its short, medium and long periods. Taking the Shijiao station as an example, we analyzed the periodicity of peak flow series based on the IMFs, Res. and power spectra.

As shown in Figure 4(d), the Shijiao series showed oscillations with complex frequencies over the entire time interval in the IMF1; however, we saw a 2.1-year period and a 4-year period at the 0.05 significance level in the IMF1 based on the corresponding power spectrum (Figure 5). In the IMF2, the series fluctuates chaotically at first glance,

however, a significant 5.9-year period is detected for IMF2 from the power spectrum. For the IMF3, its power spectrum reveals multiple medium and long periodicities of the series, in which >8-year periods are significant with high spectrum power beyond the 0.05 significance level. The IMF4 remains almost unchanged before 1990, after which it increases gradually. For the IMF5, it decreases before 1960, and during 1970–1990, whereas it increases from 1960 to 1970 and after 1990. By computing the power spectra of IMF4 and 5, however, we could not find any significant periods. In conclusion, the overall IMFs show short, medium and long periodicities of the annual peak flow series at the Shijiao station.

The periodicities of the other three series of the remaining stations are listed in Table 4 based on their own IMFs and Res., and their power spectra. Concisely, the series from the Fenshi station had 3.2, 4.2 and 7-year periodicities, while that of the Pingshi station had 2.2 and 4.2-year periodicities. In addition, the Lishi station series had 2.2, 4 and >6-year periodicities. Specifically, we found that, by comparison, the peak flow series at the Pingshi station was mainly dominated by short periods, while that at the Fenshi station was dominated by short and medium periods. For the Fenshi and Pingshi stations, not only short and medium

**Figure 4** | IMF1–5 and Res. components of the annual peak flows series for the (a) Fenshi, (b) Pingshi, (c) Lishi and (d) Shijiao stations using EEMD.

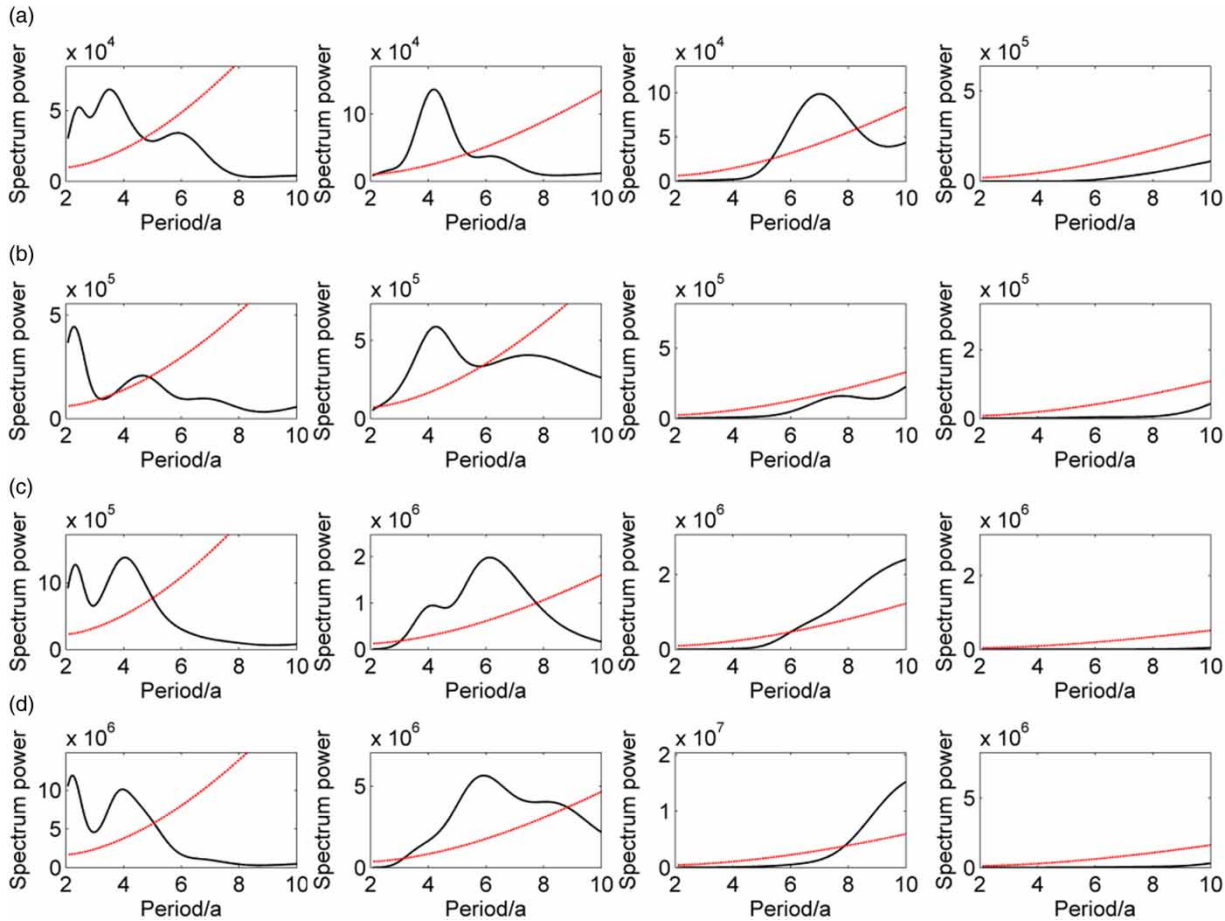


Figure 5 | Power spectra of the IMF1–4 (shown from left to right) of the peak flow series at the (a) Fenshi, (b) Pingshi, (c) Lishi and (d) Shijiao stations in BRB. The dark curve line is the spectrum power line and the red curve line is the 0.05 significance level line beyond which the periods are statistically significant. The power spectra of IMF5 for each station are not presented, as the spectrum power lines of IMF5 are quite similar to those of IMF4. The full colour version of this figure is available in the online version of this paper, at <http://dx.doi.org/10.2166/nh.2016.228>.

periods were remarkable in the peak flow series, but also long periods.

Fractal property analysis

The R/S method has the advantage of determining whether there is a long-term correlation in a time series (Hurst 1951).

Table 4 | Results from Fourier transforms of IMF1–5 for the series at the four stations (Unit: year)

Station	IMF1	IMF2	IMF3	IMF4	IMF5
Fenshi	3.2	4.2	7	/	/
Pingshi	2.2	4.2	/	/	/
Lishi	2.2 and 4	6	>6	/	/
Shijiao	2.1 and 4	5.9	>8	/	/

// denotes no significant periodicity is detected.

An existing long-term correlation provides information on future trends in the series. If a time series is detected as persistent behavior, it indicates the future trend of the series will tend to follow that of the past, whereas anti-persistent behavior indicates the opposite trend. If a series is detected as pure random behavior, it is difficult to identify its future trends since there is little correlation between past and future trends (Karakasidis & Liakopoulos 2004). Results of the R/S are presented in Figure 6 and Table 5. The calculated Hurst exponents indicate that the Fenshi, Pingshi and Lishi series exhibited preferably correlated behaviors. Thus, future trends of the annual peak flows in the Fenshi, Pingshi and Lishi stations may follow those of the past. Note that the H value of the Shijiao station data is 0.503, which is very close to 0.5, demonstrating a nearly pure random walk characteristic of the series. Thus, it is unsure whether the future trend

of the series will follow that of the past or not. On the other hand, by comparing the H values of the four stations, it can be concluded that correlations between past and future trends gradually weaken from the Fenshi to the Shijiao station.

DISCUSSION

Reviews and implications

The characterization of the annual peak flows across BRB reveals the temporal patterns of its variability, which can not only enable comparison between such changes and those reported in other regions, but can also be useful for local farmers and regulators (Pisaniello 2016). Previous research has revealed an increase in annual peak stream flow in BRB (Zhang *et al.* 2015), and our study supports this point of view. Yang *et al.* (2012) implied some shifts together with an approximate six-month periodicity in peak flows during 1957–1967 in the Hailiutu River, Northwest China, whereas our study suggests no abrupt changes but multiple periodicities of peak flows in BRB. This may be ascribed to different precipitation regimes between these two regions (Wu *et al.* 2015). There is an increased

necessity for flood control strategies in BRB during the past decades, as well as in the future, in light of the detected upward trends and long-term correlation characteristic in the recorded data. The unique multi-scale periodic oscillation characteristics of the peak flows, on the other hand, may provide a reference for flood prediction.

Possible climatic and anthropogenic causes

In the current study, distinct differences between the variabilities of annual peak flows at the four stations are observed. Specifically, from the Fenshi to the Shijiao station, both the coefficients of deviation (Table 2) and long-term correlation (Table 5) of peak flows decrease in value. Such geographical contrasts may, in part, be due to the geomorphology (Whitaker *et al.* 2002). Geomorphologic factors are known to affect flood occurrences through two main mechanisms: orographic effects that augment precipitation, and topographic reliefs that promote the rapid concentration of flow (Costa 1987). BRB is considered to be a small watershed but with topographic contrast (Figure 1). Topographic reliefs in this watershed diminish from the upstream to the downstream, causing peak flows to vary spatially. In addition to these physiographic factors, human activities are assumed to be responsible for such

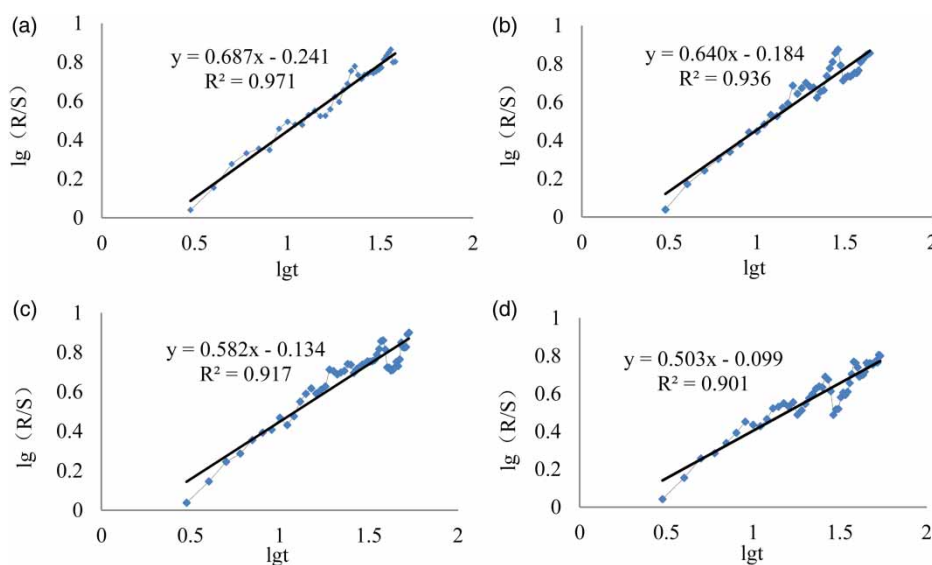


Figure 6 | Results from the R/S analysis of the annual peak flows series of the (a) Fenshi, (b) Pingshi, (c) Lishi and (d) Shijiao stations in BRB. The slope of the straight line obtained by the least square method denotes the H exponent value. R^2 is the coefficient of autocorrelation of the time series.

Table 5 | Hurst exponents of the annual peak flows series at the four stations

Station	Fenshi	Pingshi	Lishi	Shijiao
Hurst exponent	0.687	0.640	0.582	0.503

phenomena. The peak flow series at the Shijiao station, for instance, is stationary for the period 1954–2007, but is non-stationary for the period 1954–1998, i.e. before the Feilaixia reservoir was constructed. Zhang *et al.* (2015) have confirmed the impacts of reservoir regulations on peak flows changes in the Pearl River Basin (with BRB involved), and this finding supports our point of view of Feilaixia reservoir affecting peak flows at the Shijiao station.

Climate change may influence flood occurrences through its effects on the variability of storm characteristics and the seasonality of rainfall and evapotranspiration, which affect the antecedent watershed conditions for individual storm events (Marchi *et al.* 2010). Assessments of the influences of climatic factors on hydrologic variables, such as precipitation and flow, are usually undertaken to interpret changes in these variables. To explore possible climatic causes of the annual peak flow changes in BRB, five climatic indices were selected after considering global and regional climate characteristics, i.e. EASMI, SASMI, SCSSMI, MEI and DMI. We used MLR to investigate the relationships between annual peak flows in the Fenshi, Lishi, Pingshi stations and these indices, and the contributions of these indices to peak flows. We did not consider the Shijiao station series because this station should be influenced by the Feilaixia reservoir, where the effect of the subjective operation may be greater than the climate change impact. The results indicate that the five indices, except the SCSSMI, may positively influence peak

flows (Table 6). Of all the indices, SASMI contributed the most to the peak flow changes; a strengthened/weakened South Asian Summer Monsoon (SASM) may increase/decrease peak flows at the Fenshi, Lishi and Pingshi Stations in BRB.

The overall analyses reveal possible impacts of climatic and anthropogenic factors on annual peak flows in BRB. Nevertheless, the underlying mechanisms of the causes are far more complex, and a further study on this issue should be conducted to substantiate the MLR results.

CONCLUSIONS

This study applied several methods including EEMD and R/S to examining the trends, abrupt changes, periodicities and fractal properties of annual peak flows in BRB, based on the hydrological data from four hydrological stations. In addition, the climatic and anthropogenic causes of peak flow changes were also explored. Below summarizes what can be inferred from this research:

- For the past decades annual peak flows have increased but the uptrend was non-significant in the study region, which suggests a potential increase in regional flood risk. Meanwhile, no abrupt change occurred in peak flow series.
- From a non-linear perspective, peak flows displayed a multi-periodic nature, in which both short and long periods were seen but the short ones were more distinct. Moreover, the fractal property analysis reveals that peak flows may continue to increase in the future, causing a higher flood risk than the past. Therefore, flood

Table 6 | Contributions of the five climatic indices to the peak flows at the Fenshi, Pingshi and Lishi stations

Station	Indices	EASMI	SASMI	SCSSMI	MEI	DMI
Fenshi	Coefficient	0.024	0.334	0.152	0.147	0.089
	Contribution	0.033	0.447	0.204	0.197	0.119
Pingshi	Coefficient	0.203	0.347	-0.101	0.198	0.021
	Contribution	0.234	0.399	0.116	0.227	0.024
Lishi	Coefficient	0.129	0.283	-0.070	0.131	0.075
	Contribution	0.188	0.411	0.102	0.190	0.109

Coefficient means MLR coefficient. Magnitude of contribution is calculated according to Equation (3).

protection might need to be upgraded to avoid unacceptable losses.

- Peak flow variability in BRB may be attributed to both local reservoir operations and the changing SASM. A strengthened/weakened SASM possibly causes an increase/decrease in peak flows, while a local reservoir might alter the stationary characteristics of peak flows.

These findings would be useful for flood prediction and defenses in the study region. More importantly, the research can provide a reference for other regions regarding periodicity and future trend of peak flows, as well as the impacts of human activities and changes in climate.

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REFERENCES

- Beschta, R. L., Pyles, M. R., Skaugset, A. E. & Surfleet, C. G. 2000 Peakflow responses to forest practices in the western cascades of Oregon, USA. *J. Hydrol.* **233**, 102–120.
- Bi, S. B., Xu, Y., Chen, X. & Wang, B. Q. 2010 Research on short-range climatic forecast method based on EMD and SVM. In: *3rd International Conference on Information and Computing*, IEEE Computer Society Washington, DC, USA, 4, 117–120.
- Calenda, G., Mancini, C. P. & Volpi, E. 2005 Distribution of the extreme peak floods of the Tiber River from the XV century. *Adv. Water Resour.* **28**, 615–625.
- Calenda, G., Mancini, C. P. & Volpi, E. 2009 Selection of the probabilistic model of extreme floods: the case of the River Tiber in Rome. *J. Hydrol.* **371**, 1–11.
- Costa, J. E. 1987 Hydraulics and basin morphometry of the largest flash floods in the conterminous United States. *J. Hydrol.* **93**, 313–338.
- Daniels, J. M. 2007 Flood hydrology of the North Platte River headwaters in relation to precipitation variability. *J. Hydrol.* **344**, 70–81.
- Ding, Z., Zhang, J. & Xie, G. 2010 LS-SVM forecast model of precipitation and runoff based on EMD. In: *6th International Conference on Natural Computation*, IEEE, Yantai, China, pp. 1721–1725.
- El Alfy, M. 2016 Assessing the impact of arid area urbanization on flash floods using GIS, remote sensing, and HEC-HMS rainfall–runoff modeling. *Hydrol. Res.* **47** (6), 1142–1160.
- Fill, H. D. & Steiner, A. A. 2003 Estimating instantaneous peak flow from mean daily flow data. *J. Hydrol. Eng.* **8**, 365–369.
- Gottschalk, L. & Weingartner, R. 1998 Distribution of peak flow derived from a distribution of rainfall volume and runoff coefficient, and a unit hydrograph. *J. Hydrol.* **208**, 148–162.
- Guillemette, F., Plamondon, A. P., Prevost, M. & Levesque, D. 2005 Rainfall generated storm flow response to clear cutting a boreal forest: peak flow comparison with 50 world-wide basin studies. *J. Hydrol.* **302**, 67–76.
- Haddad, K., Rahman, A. & Stedinger, J. R. 2012 Regional flood frequency analysis using Bayesian generalized least squares: a comparison between quantile and parameter regression techniques. *Hydrol. Process.* **26**, 1008–1021.
- Hewlett, J. D. 1982 Forests and floods in the light of recent investigation. In: *Proceedings of the Canadian Hydrology Symposium '82*. National Research Council of Canada, Ottawa, pp. 543–559.
- Huang, N. E., Shen, Z., Long, S. R., Wu, M. C., Shih, H. H. & Zheng, Q. 1998 The empirical mode decomposition and the Hilbert spectrum for nonlinear and nonstationary time series analysis. *Proc. R. Soc. Lond. Ser. A Math. Phys. Eng. Sci.* **454**, 903–995.
- Hurst, H. E. 1951 The long-term storage capacity of reservoirs. *Trans. Am. Soc. Civil Eng.* **116**, 770–799.
- IPCC 2012 Summary for policymakers. In: *Managing the Risks of Extreme Events and Disasters to Advance Climate Change Adaptation. A Special Report of Working Groups I and II of the Intergovernmental Panel on Climate Change* (C. B. Barros, V. Stocker, T. F. Qin, D. Dokken, D. J. Ebi, K. L. Mastrandrea, M. D. Mach, K. J. Plattner, G.-K. Allen, S. K. Tignor & M. Midgley, eds). Cambridge University Press, Cambridge, NY, USA, pp. 3–21.
- Jencso, K. G. & McGlynn, B. L. 2011 Hierarchical controls on runoff generation: topographically driven hydrologic connectivity, geology, and vegetation. *Water Resour. Res.* **47**, W11527.
- Kaluzynski, K. 2014 Empirical mode decomposition of simulated and real ultrasonic Doppler signals of periodic fetal activity. *Med. Eng. Phys.* **36**, 859–868.
- Karakasidis, T. E. & Liakopoulos, A. B. 2004 Two-regime dynamical behaviour in Lennard–Jones systems: Spectral and rescaled range analysis. *Physica. A* **333**, 225–240.
- Kim, J., Lee, S. & Lee, B. 2014 EEG Classification in a single-trial basis for vowel speech perception using multivariate empirical mode decomposition. *J. Neural Eng.* **3**, 1–12.
- Kuzuha, Y., Tomosugi, K., Kishii, T. & Komatsu, Y. 2009 Coefficient of variation of annual flood peaks: variability of flood peak and rainfall intensity. *Hydrol. Process.* **23**, 546–588.

- Leviandier, T., Lavabre, J. & Arnaud, P. 2000 Rainfall contrast enhancing clustering processes and flood analysis. *J. Hydrol.* **240**, 62–79.
- Li, J. P. & Zeng, Q. C. 2005 A new monsoon index, its interannual variability and relation with monsoon precipitation. *Clim. Environ. Res.* **10**, 351–365.
- Luo, Y., Liu, S., Fu, S., Liu, J., Wang, G. & Zhou, G. 2008 Trends of precipitation in Beijiang River Basin, Guangdong Province, China. *Hydrol. Process.* **22**, 2377–2386.
- Marchi, L., Borga, M., Preciso, E. & Gaume, E. 2010 Characterisation of selected extreme flash floods in Europe and implications for flood risk management. *J. Hydrol.* **394**, 118–135.
- Mei, X. F., Dai, Z. J., Wei, W. & Gao, J. J. 2016 Dams induced stage–discharge relationship variations in the upper Yangtze River basin. *Hydrol. Res.* **47**, 157–170.
- Mitosek, H. T., Strupczewski, W. G. & Singh, V. P. 2006 Three procedures for selection of annual flood peak distribution. *J. Hydrol.* **323**, 57–73.
- Mosmann, V., Castro, A., Fraile, R., Dessens, J. & Sanchez, J. L. 2004 Detection of statistically significant trends in the summer precipitation of mainland Spain. *Atmos. Res.* **70**, 43–53.
- Munoz, E., Arumi, J. L. & Vargas, J. 2012 A design peak flow estimation method for medium-large and data-scarce watersheds with frontal rainfall. *J. Am. Water Resour. Assoc.* **48**, 439–448.
- Nyeko-Ogiramo, P., Willems, P. & Ndirane-Katashaya, G. 2013 Trend and variability in observed hydrometeorological extremes in the Lake Victoria basin. *J. Hydrol.* **489**, 56–73.
- Pearce, A. J., Rowe, L. K. & O'Loughlin, C. L. 1980 Effects of clearfelling and slash-burning on water yield and storm hydrographs in evergreen mixed forests, western New Zealand. *Int. Assoc. Sci. Hydrol. Publ.* **130**, 119–127.
- Pisaniello, J. D. 2016 Helping farmers and regulators manage and assure the cumulative flood safety of agricultural dams: a cost-effective regionalised review/design tool from Australia. *Hydrol. Res.* **36**, 225–230.
- Tabari, H. & Hosseinzadeh Talaei, P. 2013 Moisture index for Iran: spatial and temporal analyses. *Glob. Planet. Change* **100**, 11–19.
- Wang, T., Zhang, M., Yu, Q. & Zhang, H. 2012 Comparing the applications of EMD and EEMD on time–frequency analysis of seismic signal. *J. Appl. Geophys.* **83**, 29–34.
- Wang, Z. L., Lai, C. G., Chen, X. H., Bing, Y., Zhao, S. W. & Bai, X. Y. 2015 Flood hazard risk assessment model based on random forest. *J. Hydrol.* **527**, 1130–1141.
- Whitaker, A., Alila, Y. & Beckers, J. 2002 Evaluating peak flow sensitivity to clear-cutting in different elevation bands of a snowmelt-dominated mountainous catchment. *Water Resour. Res.* **38** (9), 1172.
- Wu, Z. & Huang, N. E. 2009 Ensemble empirical mode decomposition: a noise-assisted data analysis method. *Adv. Adapt. Data Anal.* **1**, 1–41.
- Wu, X., Wang, Z., Zhou, X., Lai, C., Lin, W. & Chen, X. 2015 Observed changes in precipitation extremes across 11 basins in China during 1961–2013. *Int. J. Climatol.* **36** (8), 2866–2885.
- Yang, Z., Zhou, Y., Wenninger, J. & Uhlenbrook, S. 2012 The causes of flow regime shifts in the semi-arid Hailu River, Northwest China. *Hydrol. Earth Syst. Sci.* **16**, 87–103.
- Yue, S. & Pilon, P. 2004 A comparison of the power of the *t*-test, Mann–Kendall and bootstrap tests for trend detection. *Hydrol. Sci. J.* **49**, 21–37.
- Zhang, Q., Xu, C., Jiang, T. & Wu, Y. 2007 Possible influence of ENSO on annual maximum streamflow of the Yangtze River, China. *J. Hydrol.* **333**, 265–274.
- Zhang, Q., Gu, X., Singh, V. P., Xiao, M. & Xu, C. 2015 Flood frequency under the influence of trends in the Pearl River basin, China: changing patterns, causes and implications. *Hydrol. Process.* **29**, 1406–1417.
- Zhao, X. & Chen, X. 2015 Auto regressive and ensemble empirical mode decomposition hybrid model for annual runoff forecasting. *Water Resour. Manage.* **29**, 2913–2926.
- Zhou, N., Westrich, B., Jiang, S. & Wang, Y. 2011 A coupling simulation based on a hydrodynamics and water quality model of the Pearl River Delta, China. *J. Hydrol.* **396**, 267–276.

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