POLICE PRESENCE, RAPID RESPONSE RATES, AND CRIME PREVENTION

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Abstract—This paper estimates the impact of police presence on crime using a unique database that tracks the exact location of Dallas Police Department patrol cars throughout 2009. To address the concern that officer location is often driven by crime, my instrument exploits police responses to calls outside their allocated coverage beat. This variable provides a plausible shift in police presence within the abandoned beat that is driven by the police goal of minimizing response times. I find that a 10% decrease in police presence at that location results in a 7% increase in crime. This result sheds light on the black box of policing and crime and suggests that routine changes in police patrol can have a significant impact on criminal behavior.

I. Introduction

DOES police presence deter crime? While it was once generally accepted that the role of police officers was apprehending criminals after they committed a crime, today a growing body of research shows that increased investment in policing results in lower crime rates. Specifi-
cally, previous papers have found that larger police forces and high doses of police presence in small areas result in lower crime rates (see Levitt, 1997; Evans & Owens, 2007; Di Tella & Schargrodsky, 2004; Draca, Machin, & Witt, 2011). However, the literature has largely ignored the fact that the rapid response philosophy where police officers are spread thinly throughout the city and much of an officer’s time is dedicated to responding to emergency calls remains the domi-
Policing has been a constant feature of American life, but the way it is performed has changed significantly over time. The rapid response philosophy, which originated in the 1930s, has been widely adopted in the United States and other countries as a way to increase police visibility and deter crime. However, the effectiveness of this approach has been a subject of debate, with some critics arguing that it is not effective at reducing crime and that it diverts police resources away from other important tasks.

This paper contributes to the literature by estimating the impact of police presence on crime. Using a unique database that tracks the exact location of Dallas Police Department patrol cars throughout 2009, the author finds that a 10% decrease in police presence at a location results in a 7% increase in crime. This result sheds light on the black box of policing and crime and suggests that routine changes in police patrol can have a significant impact on criminal behavior.

The paper's methodology is innovative, using police responses to calls outside their allocated coverage beats as an instrument for police presence. This approach allows the author to control for unobserved factors that may affect both police presence and crime rates. The results are robust to a variety of specifications and robustness checks.

In conclusion, this paper provides a valuable contribution to the literature on policing and crime. It highlights the importance of routine changes in police patrol and suggests that police presence has a significant impact on crime. The findings have important implications for policymakers and those responsible for allocating police resources.
The Endogenous Relationship Between Policing & Crime

Police Presence by police allocation

Crime by police allocation

and stored.5 I focus on the beat (a geographic patrol area averaging 1.7 square miles in size) each car was allocated to patrol, as well as where these officers were actually present throughout the day. Information on incidents of crime was acquired from a separate database that tracks calls for service (911 calls) placed by local citizens to the police department.6 Thus, the current project is not motivated by a specific policing experiment; rather, it takes advantage of a large amount of data (roughly 100 million pings of information) to provide an estimate of the social returns of an additional hour of police patrol in the current policing system.

A deterrence mechanism that is based on police interactions would imply that areas or times of day with higher levels of police presence will report less crime. However, this ignores both the allocation of officers to riskier locations during riskier periods and the fact that the occurrence of a crime is likely to increase police presence as officers are called to respond to the incident. Indeed, each year, beat-level police allocation in Dallas is based on calls for service and crime from previous years (using Staff Wizard software). Division commanders then adjust these allocation decisions based on weekly Compstat meetings where current crime concerns are discussed and areas in need of additional police attention are identified. This type of well-managed police allocation would result in a positive correlation between policing and crime, which is illustrated in figure 1. Generally, areas and times with higher levels of allocated patrol tend to have higher levels of both police presence and crime.7 Thus, while this data set provides a unique picture of police presence across a city, the location of officers may be determined by factors unobserved by the econometrician and correlated with crime.

5The AVL data do not include the location of officers on motorcycle and horseback (mounted division). The motorcycle patrol unit consists of 42 officers, and the mounted division consists of 17 police officers.

6I separate calls that relate to crime into the following categories: violent crimes, burglaries, thefts, and public disturbances. I focus on 911 calls as they are less likely to suffer from reporting bias than reported crimes and are more likely to provide the exact time at which the incident occurred.

7While there are 873 Dallas patrol vehicles tracked in this study, on average 132 cars are on active patrol per hour. These cars are allocated among 232 beats at the beginning of their shift. Thus, the most common allocation points are either 0 or 1 car allocated per hour. The reason that police presence (defined by the location of police vehicle throughout the shift) does not have a one-to-one relationship with police allocation is that officers often spend time outside their allocated coverage beat.
My identification strategy stems from the two distinct responsibilities facing police patrol cars: proactive and reactive policing. While police may be allocated to a certain area (beat) in order to create a deterrence effect and lower the expected benefit of committing a crime, they are also responsible for answering emergency calls within their larger division quickly—generally, in under 8 minutes. I use incidents that can result in patrol officers being assigned to calls outside their area of patrol to capture an element of randomness regarding whether police are present at a given location and time. Thus, I introduce the outside calls ratio (OCR) instrument, which is equal to the number of calls an officer patrolling this beat is likely to address outside the beat. It is calculated as the ratio of outside calls to allocated outside patrol. The instrument is motivated by the idea that assignment to an outside beat is a function of both the density of outside calls and outside police presence. This measure is different from realized outside assignment, which could be driven by crime risks at the beat level.

In order to ensure that the OCR is not correlated with crime at the beat, I focus on outside 911 calls reporting incidents related to mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses. Reports of crime are not included in these outside calls in order to avoid the concern that crimes occurring across beats may be correlated. I also focus on out-of-beat patrol allocation as opposed to out-of-beat police presence. Out-of-beat police allocation is determined at the start of the shift as opposed to out-of-beat presence that is immediately affected by the distribution of calls and crime. My results are robust to controlling for beat, month, day of week, and weekend × hour of day fixed effects.

The main threat to the exclusion restriction for this proposed instrument is that even keeping outside calls constant, a decrease in out-of-beat patrol allocation will increase the OCR. While this is likely to increase the probability that a local patrol officer will abandon his or her beat, one could be concerned that the decrease in out-of-beat patrol allocation was a result of a general decrease in crime concerns at the division level (both inside and outside the beat). In this case, applying the OCR could bias the estimated deterrence effect toward zero. To address this concern, I introduce an additional instrument of outside calls (OC). This instrument is equal to the numerator of the OCR and is a weighted sum of calls, unrelated to crimes, occurring within the division outside the beat being currently patrolled. While this instrument has a weaker first stage in terms of predicting police presence in the beat, the exclusion restriction is more compelling. My results suggest that the number of officers patrolling a beat has a significant impact on the probability of crime. I first demonstrate that as reported in previous studies, there is a positive correlation in the data between police presence and crime. This positive correlation remains significant even when controlling for beat and time fixed effects. This suggests that police departments may be able to quickly adjust police presence to changing crime risks within locations over time. It is only when instrumenting for actual police presence with either the OCR or OC that I can identify a deterrence effect. I estimate that a 10% decrease in police presence results in a 7% increase in crime. I provide evidence that this result cannot be explained by displacement of crime to neighboring beats.

While police departments often consider rapid-response times (minimizing the elapsed time between receiving an emergency call and responding to that call) to be one of the most important tools for solving crimes, criminologists argue that no evidence exists to support that claim (Sherman, 2013). Not only have few studies examined the impact of rapid-response times on solving crimes, but also no attempt has been made to measure the impact of rapid response tools on the deterrence capacity of the police. My results provide an estimate of both the deterrence created by routine police activities and the possible community safety costs of police officers dividing their time between preventing future crimes and responding to past crimes.

These results join an empirical literature on measuring deterrence that focuses on applying techniques to mitigate simultaneity bias. My estimated elasticities of −0.9 regarding the impact of a change in police presence on violent crime and −0.6 on property crime fall at the higher end of the range of elasticities of between −0.4 and −1 (violent) −0.3 and −0.5 (property) reported in previous work (see Levitt, 1997, 2002, and Evans & Owens, 2007).

Both Levitt (1997, 2002) and Evans and Owens (2007) applied instrumental variable strategies to estimate the elasticity of crime to police force size. Chalfin and McCrary (2017) raise concerns regarding weak instruments in these papers and point out that these studies show a wide range of estimates that are often not statistically significant at conventional levels.

In the analysis, I interact the instrument with an indicator for division. This strengthens the first stage of the instrument so that it does not fall under the weak instruments category.

The general embracement of rapid-response policing is evident in the summary of “best practices in police performance measurement” provided by the Rand Corporation (Davis, 2012). Using data from the Kansas City Preventative Patrol Experiment, Kelling et al. (1974) found no impact of response times (minimizing the elapsed time between receiving an emergency call and responding to that call) to be one of the most important tools for solving crimes, criminologists argue that no evidence exists to support that claim (Sherman, 2013). Not only have few studies examined the impact of rapid-response times on solving crimes, but also no attempt has been made to measure the impact of rapid response tools on the deterrence capacity of the police. My results provide an estimate of both the deterrence created by routine police activities and the possible community safety costs of police officers dividing their time between preventing future crimes and responding to past crimes.

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The general embracement of rapid-response policing is evident in the summary of “best practices in police performance measurement” provided by the Rand Corporation (Davis, 2012). Using data from the Kansas City Preventative Patrol Experiment, Kelling et al. (1974) found no impact of response times on solving crimes. However, Blanes i Vidal and Kirchmaier (2018) find that faster response times increase the likelihood of detecting crimes when using an instrumenting strategy based on the distance of the incident from police headquarters. Mastrobuoni (2019) reaches a similar conclusion when analyzing the outcomes of quasi-experimental variations in police presence in the city of Milan.

confidence intervals. The instruments used in this research avoid this critique with first-stage $F$-statistics of above 40 for both the OCR and OC instruments. The deterrence estimates reported in this paper also fall within much smaller confidence intervals.

An additional branch of the literature focuses on exogenous changes in police presence that are driven by threats or actual acts of terrorism (Di Tella & Schargrodsky, 2004; Klick & Tabarrok, 2005; Draca et al., 2011; Gould & Stecklov, 2009). Di Tella and Schargrodsky (2004) and Draca et al. (2011) report smaller elasticities of crime with respect to police presence of between $-0.3$ and $-0.4$. Similarly, MacDonald, Fagan, and Geller (2015) report an elasticity of crime with respect to police presence of $-0.33$ when examining an increase of 200% in police presence in the area surrounding the University of Pennsylvania campus. This estimate shrinks further when focusing on randomized experiments. Sherman and Weisburd (1995) found that doubling police patrol at hot-spot locations in Minneapolis resulted in a 6% to 13% decrease in crime. Blattman et al. (2018) report that an intervention that included doubling police presence at high street segments in Bogota led to a decrease in city-wide crime of below 2% when accounting for spatial spillovers.

This paper offers a bridge between the detailed location-specific data that are analyzed in randomized experiments and the aggregate data that are usually available at the city level. To the best of my knowledge, Blanes i Vidal and Mastrobuoni (2018) is the only other paper that has attempted to look at the geographic distribution of police officers throughout an entire city using precise GPS-level data on police location. They take advantage of a natural experiment where police spent an extra 10 minutes per week in a 200 meter radius of an area where a burglary was reported for the week following a burglary. Interestingly, they find no effect of this change in weekly level police patrol on crime. I focus on police presence at the hourly level within Dallas beats (averaging 2.7 km$^2$) and examine whether routine changes in police behavior can have significant impacts on crime. One important difference between these two studies is that this project focuses on moving police away from an area they usually be significantly faster.

Dallas is an ideal location for research using AVL data since it is mostly flat and thus is able to provide fairly precise latitude and longitude points with minimal missing data. Dallas police patrol is divided into seven patrol divisions (Central, North Central, Northeast, Northwest, South Central, Southeast, Southwest), each commanded by a deputy chief of police. Figure 2 provides a map of the city divided into divisions and beats. There is some variation in the characteristics of beats across different divisions in the city, as illustrated by table 1. Beats in the Central division are smaller, averaging 0.6 square miles, while beats in other divisions average 1.8 square miles. Beats in the South Central division have a higher percentage of black residents, while beats in the Southwest have the highest percentage of Hispanic residents. Residents of the North Central division report higher incomes. These characteristics highlight the importance of focusing on crime outcomes at the beat level as different parts of the city may require different levels of police presence and face different crime risks.

The analysis is conducted on geographic beats at hour-long time intervals. I use the call database to count the number of crimes called into 911 for each beat $b$ and hour $h$. Focusing on 911 calls as opposed to crime reports is expected to lower concerns regarding selective reporting of incidents. While I cannot rule out the possibility that in certain beats, crimes may not be called in to the police, this should not affect my results when controlling for beat fixed effects. A larger concern is whether the presence of a police car in an area may reduce calls to 911 as people can speak directly with the patrolling officer. Importantly, 911 is the generally accepted protocol for reporting crime. Beats average 1.7 square miles and roughly 40 minutes of patrol per hour, thus reporting via 911 will usually be significantly faster.

The original database included 684,584 calls recorded throughout 2009 in Dallas. My final call database consists of 556,978 calls after removing duplicate calls and excluding calls that were classified as hang-ups. The main analysis focuses on 289,030 calls reporting incidents of crime. These crimes are classified into the following categories: public

II. The Data

Dallas, Texas, is the ninth largest city in the United States, with roughly 1.2 million residents and 3,266 sworn police officers spread over 385 square miles. I use two separate Dallas Police Department (DPD) databases that provide information on the precise location of both crime and police in 2009. The DPD call database records the time and location of each report of crime to the department. The AVL database tracks the location of police cars throughout the day. Together they provide an opportunity to understand how police presence has an impact on crime.$^{13}$

$^{13}$ Using geographic mapping software, I collect additional information on population size as well as miles of roads across different areas in Dallas. These data are combined with information on daily temperature, visibility, precipitation, sunrise, and sunset times in order to control for variability in the probability of crime over time.

$^{14}$ Using internal DPD data, a source from the Dallas Police Department estimated that roughly 90% of reported crimes are initiated using 911 calls. The remaining 10% are likely to be a combination of officer-initiated calls (often related to traffic stops), sex assault victims (reporting from trauma centers, hospitals, colleges, and victims’ advocates), as well as residents reporting both to patrol officers and arriving directly at the police station.

$^{15}$ Details of the data cleaning process are in online appendix A.
disturbances, burglaries, violent crimes, and theft. Generally, crime in Dallas tends to peak in May and plummet in December, but there are significant fluctuations in the crime rate throughout the year.

A crime is classified as a burglary if it involves entering a structure with the intent to commit a crime inside. Stabbings, shootings, robberies, assaults, kidnappings, and armed encounters are classified as violent crimes. Public intoxication, illegal parking, suspicious behavior, prostitution, loud music, gunfire, speeding, road rage, and panhandlers are classified as public disturbances.

Beginning in the year 2000, Dallas police cars were equipped with AVL (873 tracked vehicles), which create pings roughly every 30 seconds with the latitudinal and longitudinal coordinates of these vehicles. Each ping includes the radio name of the vehicle, which provides information on the allocation of the police vehicle. Thus, a ping with radio name A142 refers to a car that was allocated to patrol beat 142 during patrol A (during the first watch that takes place between 12 a.m. and 8 a.m.).
The AVL data also include a report indicator for vehicles that are responding to a call for service. This indicator provides information on whether the vehicle is on general patrol or responding to a call. In contrast to an aggregate count of police officers per city, these data present an opportunity to map the activity of each individual squad car throughout the day.

The call database and AVL data set are the actual data that the Dallas Police Department uses to assign officers to 911 incidents. Thus, when a 911 call is placed in Dallas, the call taker records basic information on the incident (location, caller name, classification of incident, and time) and defines the severity of the incident. This information is loaded into the computer-aided dispatch (CAD) system, which is the source for the call data. This incident then appears on the computer of the police dispatcher for the relevant division where the incident was reported. The police dispatcher assigns the incident to a patrol car based on priority (as recorded in call data) and distance of car (as tracked in AVL data). When possible, the incident is assigned to the police officer allocated to the beat of the incident, but since this officer can be otherwise occupied and most calls require two officers, officers are often assigned to out-of-beat calls. While it is possible during severe emergencies for police dispatchers to coordinate and allocate officers across divisions, police are generally dispatched to calls within their division.

To create a database of police location, I divide the city of Dallas into 232 geographic beats of analysis and map each to have data on the exact allocation, this can still provide insight into the general area of allocation.

Table 1 summarizes the mean hourly values for crime, police allocation, and police presence by beat at the division level. The majority of crimes occur in beats that are located generally higher in the North. My identification strategy builds around the idea that actual police presence over time is not fully determined by the allocation of officers.

20The study focuses on 232 out of 234 beats in Dallas. Two beats were excluded from the analysis because they are composed primarily of water.
21I set a lower bound of presence at 5 minutes in order to focus the analysis on cars that were likely to be patrolling the given beat and not simply driving through the area.
22See figures 2 and 3 of online appendix C, which separately graph police allocation and actual presence in Dallas over 2009.
beats for 60% to 80% of each hour. The highest level of police allocation is in the North Central division where, on average, each beat has an assigned officer for over 80% of each hour, while in the Northwest division, a patrol officer is allocated to a beat for only about 60% of every hour. However, police allocation only refers to whether there was an active patrol officer at this hour of the day whose radio name referred to the given beat. Actual police coverage varies significantly from allocated coverage, with the largest average difference observed in the Southeast division. While allocated coverage is determined at the start of an officer’s shift, police presence is a function of the events and crime concerns that develop throughout the day.

The simultaneous relationship between police presence and crime is already made apparent in table 1. Beats in the Northeast division average 30% less police presence than beats in the Southwest division, yet beats in the Southwest division report a higher crime rate. In order to identify a causal effect of policing on crime, I focus on an instrument that affects the level of police presence in a given beat but should not have a direct impact on crime.

Outside Calls (\(OC_{bh}\)) are calculated for each beat (\(b\)) and hour (\(h\)) as a weighted average of the number of calls occurring in division \(D_b\) outside of beat \(b\). Hour \(h\) is a time variable beginning at 0 at 12:00 a.m. on January 1, 2009, and culminating at \(h = 8,736\) at 11:00 p.m. on December 30, 2009. Thus, I sum the number of 911 calls received in division \(D_b\) outside of beat \(b\) during hour \(h\) reporting incidents related to mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses. Importantly, instead of calls occurring in all \(b\) beats of the division being counted equally, \(OC_{bh}\) is a weighted sum:

\[
OC_{bh} = \sum_{k \neq b \in D_b,h} n_{kh}w_{hk},
\]

where \(n_{kh}\) is the number of calls that occurred in beat \(k\) during hour \(h\) and \(w_{hk}\) is the estimated probability that an officer in beat \(b\) will be assigned to a call in beat \(k\) (See online appendix B for a description of this process).

The Outside Calls Ratio \(OCR_{bh}\) divides the OC instrument from equation (1) by the number of officers allocated to patrol in the sectors surrounding the outside call beats. In order to create this instrument, I calculate aggregate measures of police presence at the beat and sector levels.

I define \(PA_{bh}\) as the amount of time (in hours) patrol cars were allocated to spend in beat \(b\) during hour \(h\). The allocation of police vehicles (\(PA_{bh}\)) is determined by assignment at the start of their shift and is different from police presence (\(P_{bh}\)), a measure of where they actually patrol. For example, \(PA_{bh}\) would be equal to 1 at a beat and hour where one car was allocated to patrol, even when actual police presence (\(P_{bh}\)) was only 0.5 as the car was only physically patrolling the beat for 30 minutes of that hour. Sector \(PA_{kh} = \sum_{j \in S_k} PA_{jh}\) is the amount of time (in hours) patrol cars were allocated to spend in the sector (\(S_k\)) surrounding beat \(k\) during hour \(h\). \(OPatrol_{bh}\) is calculated for each beat (\(b\)) and hour (\(h\)) as a weighted average of allocated patrol cars in the sectors within division \(D_b\) where calls took place:

\[
OPatrol_{bh} = \sum_{k \neq b \in D_b,h} Sector_{PA_{kh}} \times n_{kh} \times w_{hk}.
\]

It is important to note that \(Sector_{PA_{kh}}\) in equation (2) is a count of police allocation outside the given beat \(b\). Thus, when the sector of the outside call (\(k\)) is the same as that of the given beat \(b\), \(Sector_{PA_{kh}} = \sum_{j \in S_k} PA_{jh} - PA_{bh}\). The variable \(OPatrol_{bh}\) provides a weighted measure of the number of officers who are in the area surrounding beats where relevant outside calls occurred. This may affect the likelihood that an officer from beat \(b\) will abandon his beat.

I define the \(OCR_{bh}\) as

\[
OCR_{bh} = \frac{OC_{bh}}{OPatrol_{bh}}.
\]

In the next section, I lay out my empirical strategy for estimating the deterrence effect of police presence on crime. I discuss unobserved factors that can create bias in estimating this effect and explain how the instruments address these concerns. My results illustrate that even with very detailed microdata, absent an exogenous shift in police presence, policing and crime remain positively correlated.

### III. Empirical Strategy and Results

In equation (4), I model the occurrence of a crime (\(C_{bh}\)) as a function of police presence (\(P_{bh}\)):

\[
C_{bh} = \gamma_0 + \beta_1 P_{bh} + \gamma_1 T + \eta_b + \varepsilon_{bh}.
\]

\(C_{bh}\) is a count of the number of 911 calls reporting incidents of crime (violent crimes, burglaries, thefts, and public disturbances) at beat \(b\) during hour \(h\). The variables included in \(x_{bh}\) capture time-varying environment characteristics that could have an impact on the costs and benefits of crime (e.g., weather, visibility). The focus of my analysis is \(P_{bh}\), a count of the amount of time police officers spent patrolling inside beat \(b\) at hour \(h\). If one police vehicle was present for a full hour \(h\) at beat \(b\), then \(P_{bh} = 1\). A single patrol car in the beat that was present for only 30 minutes will result in a \(P_{bh}\) value of 0.5; alternatively, and two cars that were present over the entire hour will result in \(P_{bh} = 2\). The time and beat fixed effects \(\gamma_1\) and \(\eta_b\) account for the differential probabilities in crime across different times and beats. If policing is uncorrelated with the remaining unobserved factors affecting crime (\(\varepsilon_{bh}\)), then \(\beta_1\) estimates the amount of deterrence created when police coverage is increased by one car.

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23The Dallas police department divides the city of Dallas into seven divisions, and each division is split into sectors. Each sector consists of roughly seven beats.
My main concern is the endogeneity of policing $P_{bh}$. It has been well documented that police allocation is far from exogenous. In a well-functioning police department, officer allocation will be highly correlated with crime. Using detailed geographic data can complicate the relationship, as one would expect that when a crime occurs in a given hour, more police will immediately enter the beat in response to that crime. Even after removing cars that are specifically assigned to respond to the call, I cannot rule out a situation where additional officers may be drawn to the location of the crime incident for backup purposes. An additional concern is that there may be seasonal differences in crime risks that are addressed by the police force by means of changing police allocation across beats and time.

The Dallas Police Department has a stated goal of answering all serious 911 calls (priority 1) within 8 minutes and priority 2 calls (e.g., potential for violence or past robbery) within 12 minutes (Eiserer, 2013). Thus, the preplanned allocation of an officer to a beat can be disrupted by an influx of emergency calls. It is exactly this differentiation between the endogenous choice of sending officers to higher-risk crime locations and the plausibly random timing of emergency calls in surrounding areas that provide a first stage for police presence $P_{bh}$:

$$P_{bh} = x_{bh} \alpha_0 + \alpha_1 OC_{bh} + \theta_t + \rho_b + \delta_{bh}. \quad (5)$$

While the allocated level of presence can be determined by the perceived crime risk in that area ($\delta_{bh}$), actual presence is affected by an exogenous factor, $OC$ one as defined in equation (1). The estimated coefficient on the instrument ($\hat{\alpha}_1$) is expected to be negative, since an increase in outside calls should decrease police presence in the beat ($P_{bh}$). Figure 3 shows that beats and intervals of time with a higher measure of $OC$ have lower levels of police presence and higher levels of crime.²⁴

Table 2 presents regression estimates for the general impact of the OCR and OC on police presence as defined in equation (5). I find that increasing the OCR by 1 decreases police coverage by 0.055 (SE 0.003), which is significant at the 1% level. Not surprisingly, without information on outside patrol,

²⁴Outside calls is a weighted sum of outside incidents related to mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses.
TABLE 2.—OUTSIDE CALLS AS PREDICTORS OF POLICE PRESENCE

<table>
<thead>
<tr>
<th>Instrument = Outside Calls Ratioa</th>
<th>Instrument = Outside Callsb</th>
<th>OCR</th>
<th>Interactions</th>
<th>OC</th>
<th>Interactions</th>
</tr>
</thead>
<tbody>
<tr>
<td>Instrument</td>
<td>−0.055*** (0.003)</td>
<td>−0.078*** (0.026)</td>
<td></td>
<td></td>
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<tr>
<td>Instrument × South Central</td>
<td>−0.080*** (0.008)</td>
<td>−0.687*** (0.051)</td>
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<td></td>
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<tr>
<td>Instrument × Southeast</td>
<td>−0.044*** (0.009)</td>
<td>−0.224*** (0.050)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Instrument × Southwest</td>
<td>−0.044*** (0.006)</td>
<td>0.016 (0.055)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Instrument × North Central</td>
<td>−0.070*** (0.008)</td>
<td>0.107 (0.071)</td>
<td></td>
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<tr>
<td>Instrument × Northeast</td>
<td>−0.052*** (0.006)</td>
<td>−0.273*** (0.053)</td>
<td></td>
<td></td>
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<tr>
<td>Instrument × Northwest</td>
<td>−0.055*** (0.006)</td>
<td>−0.032 (0.069)</td>
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<td>Instrument × Central</td>
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<td>0.834*** (0.110)</td>
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<td>Month FE</td>
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<td>Day of Week FE</td>
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<td>Weekend × Hour FE</td>
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<td>Observations</td>
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<td>2,017,676</td>
<td></td>
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</tr>
</tbody>
</table>

Each observation is a beat and hour in 2009. Standard errors in parentheses account for geographic clustering within a 10 km radius and serial correlation of 5 hours. Standard deviations are presented in brackets. All specifications control for temperature, precipitation, twilight, darkness, and whether it is a holiday.

The OCR is a proxy for police activity outside the beat. It is calculated by dividing the weighted sum of outside unrelated calls by the weighted sum of officers assigned to patrol the sector where the call took place. Outside unrelated calls are defined as those reporting incidents related to mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses.

Outside calls is a weighted sum of 911 calls that are unrelated to crime occurring within the division outside the beat.

Significant at *10%, **5%, and ***1%.

the precision of the instrument in predicting police presence in the beat decreases. Specifically, I find that an additional OC decreases police presence by 0.078 (SE 0.026). One explanation for this nosier estimate is heterogeneity across locations in their response to outside calls. While the OCR takes into account that an officer is more likely to be assigned to an out-of-beat call when that call takes place in an area where police are spread thinly, this does not hold for the OC instrument. A source of heterogeneity could be that in areas and hours when police presence is more saturated, an increase in calls in a neighboring beat could actually bring more officers to the area as officers drive through on their way to respond to the call.

To maximize the flexibility of the OC instrument, I interact it with an indicator for police division. This allows different divisions to follow different protocols or face different constraints regarding beat allocation. After including these interactions, the first-stage $F$-statistic on the OC instrument is 40.81. While an increase in the OCR results in a significant decrease in police presence across all seven divisions, the effect of OC is more varied and exhibits a positive effect on police presence in beats in the Central division that are typically small in size and very close together.

Instrumental variable analysis conducted in this paper always includes this interaction for the OC instrument.

I estimate the impact of police presence on all crimes using equation (4) for fixed effects and 2SLS specifications. I compute heteroskedasticity and autocorrelation-consistent standard errors for all specifications (see Conley, 1999). The focus of this paper is estimating $\beta_1$, the impact of an additional police vehicle in a given beat ($b$) and hour ($h$) on crime outcomes ($C_{bh}$). In the fixed-effect model (column 1 of table 3) I find that an increase in police presence seems to imply an increase in crime even when controlling for weather as well as time fixed effects (month, day of week, and weekend × hour) and beat fixed effects. These results suggest that the presence of an additional police car at a given beat results in a significant 0.012 increase in crime (at an average crime rate of 0.15).

Two-stage, least-squares estimates appear in columns 2 and 3 of table 3. The estimate in column 2 measures the deterrence effect when actual police presence ($P_{bh}$) is instrumented with the OCR; column 3 provides an estimate of the effect when applying the alternative OC instrument. These two-stage, least-squares estimates provide an opportunity to average causal effect of police presence on crime is the same for the compliers (those for whom outside calls decrease police presence) and the defiers (those for whom outside calls increase police presence) (following Angrist, Imbens, & Rubin, 1996).

While this does not affect the exclusion restriction, it could raise questions regarding the monotonicity assumption of the instrument. Note, however, that this will not create additional bias under the assumption that the

25While this does not affect the exclusion restriction, it could raise questions regarding the monotonicity assumption of the instrument. Note, however, that this will not create additional bias under the assumption that the
measure deterrence without the simultaneity bias concerns in the OLS estimates (if more police are present at locations and times with increased crime risks, this will result in a positive bias on the estimated deterrence effect ($\beta_1$)). The instrument allows me to focus on changes in police presence that were not a direct outcome of changes in perceived crime risks at the given beat and hour.

In column 2, when instrumenting for police presence with the OCR instrument, I find a significant negative effect of police presence on crime equal to $-0.185 (0.032)$. While $\beta_1$ in equation (4) represents the effect of an additional police vehicle ($P_{bh}$) on crime, what is driving the estimate is the reality that cars are often withdrawn from their patrol beat when assigned to an outside call. Accordingly, a real-world interpretation of this effect is that removing 60 minutes of presence from a given beat at a given hour results in a 123% increase in crime ($100 \times 0.185 = 185$). If I focus on average police presence per hour (36 minutes), a 10% decrease in police presence implies a 7.4% increase in crime (elasticity of $-0.74$).$^{27}$ I estimate a similar deterrence effect when applying the OC instrument (column 3).

Table 3 also provides information on how different weather and time characteristics have an impact on crime outcomes. I find that crime is more likely to occur during twilight. Higher temperatures increase the occurrence of crime, and bad weather lowers the probability of crime.

In table 4, I separately examine the impact of police on different types of crimes (violent crimes, public disturbances, burglaries, and theft).$^{28}$ I first report the measured effect of police presence in an OLS model that controls for month, day of week, beat × hour, and beat fixed effects, as well as temperature, precipitation, twilight, holiday, and darkness. In specifications 2 and 3, I report results when instrumenting for police presence with the OCR and OC instruments used in table 3.$^{29}$ All crime types exhibit a significant positive correlation between police presence and crime (see column 1) that disappears when instrumenting for police presence with the OCR and OC (see columns 2 and 3).

The estimated deterrence effect of police presence on violent crime after instrumenting for police presence with the OCR is similar to that of OC. These deterrence estimates of $-0.094$ (SE 0.017) and $-0.098$ (SE 0.028) translate to an elasticity of roughly 0.9.$^{30}$ While both instruments provide similar estimates of the effect of police presence on public disturbances and burglaries, it is more noisily measured when applying the OC instrument. I find that a 10% increase in police presence decreases public disturbances by 6% to 7% and burglaries by 5% to 6%. The effect of police presence on theft in Row D is much smaller and not statistically significant from 0 when applying either of the instruments.

### IV. A Closer Look at the Mechanisms of Deterrence

My estimates suggest that police presence at the beat level can have a significant impact on crime. The next step is to understand the mechanism by which police presence changes behavior. What are patrol officers doing to prevent crime? Does police presence also affect noncrime-related incidents? Are police officers more or less effective when allocated to certain areas? Does an increase in police presence at this beat displace crime to a neighboring beat?

Police officers engage in both active patrol (e.g., stops, questioning, frisks) and passive patrol (e.g., car patrol, paperwork) when working a beat. In order to correctly interpret my deterrence results, it is relevant to understand the extent to which outside calls affect active police patrol. This differentiation is important for gaining insight into whether an empty patrol car (or an officer who is simply filling out paperwork in his or her car) can have the same deterrence effect as an officer actively patrolling the streets. I therefore use arrests

| Table 4.—The Effect of Police Presence on Different Types of Crimes |
|---------------------|---------------------|---------------------|
|                     | OLS                 | IV = Outside                  | IV = Outside                  |
|                     | (1)                 | Calls Ratio$^b$       | Calls$^a$                     |
| A. Dependent Variable = Violent Crimes (mean of dependent variable 0.065, SD 0.258) | | | |
| Police Vehicles | 0.005$^*$           | $-0.094^***$        | $-0.098^***$                   |
|                   | (0.0003)            | (0.017)              | (0.028)                       |
| B. Dependent Variable = Public Disturbances (mean of dependent variable 0.053, SD 0.234) | | | |
| Police Vehicles | 0.004$^**$          | $-0.061^***$        | $-0.057^**$                   |
|                   | (0.0002)            | (0.017)              | (0.025)                       |
| C. Dependent Variable = Burglaries (mean of dependent variable 0.032, SD 0.181) | | | |
| Police Vehicles | 0.003$^***$         | $-0.030^***$        | $-0.026^*$                    |
|                   | (0.0002)            | (0.011)              | (0.014)                       |
| D. Dependent Variable = Theft (mean of dependent variable 0.012, SD 0.109) | | | |
| Police Vehicles | 0.001$^*$           | $-0.006$             | 0.003                         |
|                   | (0.0001)            | (0.007)              | (0.008)                       |
| Beat FE           | Yes                 | Yes                  | Yes                           |
| Month FE          | Yes                 | Yes                  | Yes                           |
| Day of Week FE    | Yes                 | Yes                  | Yes                           |
| Weekend × Hour FE | Yes                 | Yes                  | Yes                           |
| First-Stage F-Statistic  | 321.73              | 40.81                |
| Observations      | 2,017,676           | 2,017,676            | 2,017,676                     |

*Each observation is a beat and hour in 2009. Average police presence is 0.065 (SD 0.108). Standard errors in parentheses account for geographic clustering within a 10 km radius, and serial correlation of 5 hours. I include controls for temperature, precipitation, twilight, darkness, and whether it is a holiday.

$^{a}$The number of police vehicles patrolling the beat at a given hour (60 minutes of presence = 1 vehicle).

$^{b}$The OCR is a proxy for police activity outside the beat. It is calculated by dividing the weighted sum of outside unrelated calls by the weighted sum of officers assigned to patrol the sector where the call took place. Unrelated calls are defined as those reporting incidents related to mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses.

$^{c}$OC is a weighted sum of 911 calls unrelated to crime occurring within the division outside the beat.

This instrument is interacted with division in order to allow out-of-beat calls to have different effects in different divisions. Significant at "1%", "5%", and "1%".

27This is calculated as $0.065 \times 0.185 \times 100 / 0.077$.

28I classify violent crimes as stabbings, shootings, robberies, assaults, kidnappings, and armed encounters. I classify public intoxication, illegal parking, suspicious behavior, prostitution, loud music, gunfire, speeding, road rage, and panhandlers as public disturbances.

29Outside calls are defined in this paper as those reporting mental health, child abandonment, fire, animal attacks, dead people, suicides, abandoned property, fireworks, and drug houses.

30The deterrence impact on violent crimes was calculated by taking the estimated impact of an additional police vehicle on violent crime ($-0.094$ OCR and $-0.098$ OC) relative to the average violent crime rate of 0.065. Thus, the OCR (OC) instrument specification estimate implies that an additional police car results in a 144% (151%) decrease in violent crime. Since the average amount of police presence is 0.6, a 10% increase in police presence requires dividing the full hour impact by 16.7. This calculation is also applied to all other crime types.
The estimated coefficient on \( P_{\text{inh}} \) is \(-0.163 (SE 0.028)\) when applying the OCR instrument and \(-0.141 (SE 0.034)\) when applying the OC instrument. When police are present, I cannot rule out the possibility that some of this effect is driven by an arrest taking place before the individual had time to commit a crime.

In 2009, DPD received over 10,000 911 calls related to fires, suicides, abandoned children, and drug houses. While these incidents require police involvement, we would not expect them to respond to changes in police presence. Thus, these call categories provide an opportunity for a placebo test to ensure that the deterrence estimates reported in tables 3 and 4 are driven by changes in police presence. Table 6 illustrates that unrelated calls that should not be sensitive to police officer patrol (suicides, abandoned children, fires, and drug houses) are not significantly affected by police presence when applying both instruments.

If police presence affects crime by providing a visual reminder of the costs of crime, I would expect larger beats, where officers are less likely to be seen, to be less affected by losing a police vehicle than smaller beats. In table 7, I split the data into three groups of roughly equal sizes: small beats (less than 4.6 miles of roads), midsize beats (4.6 to 8 miles of roads), and large beats (more than 8 miles of roads). I find that police vehicles have a smaller impact on crime in large areas versus midsized areas when using the OCR and OC instruments.

When instrumenting for police presence with OCR, I find that each additional car reduces crime by 0.116 (0.035) in the larger beats versus 0.246 (0.060) in midsize beats and 0.253 (0.075) in the smaller beats. Interestingly, this implies that a 10% increase in police presence in either a large or a midsized beat results in a 7.6% decrease in crime \( (100 \times 0.094 \times 0.146 \approx 100 \times 0.053 \times 0.246) \), with a smaller marginal effect of 5.7% at smaller beats \( (100 \times 0.034 \times 0.152) \). This similarity at the margin is driven by the significant difference in average police presence between beats of different size, where small beats average 20 minutes of presence, midsize beats average 30 minutes of presence, and large beats average 57 minutes of presence per hour. Differences in the baseline rate of police presence per beat may also contribute to the size of the deterrence effect. In other words, taking an officer away from a beat that averages little to no police presence may be more detrimental to crime control than taking an officer from a beat with relatively high levels of police presence.

The question of deterrence versus geographic displacement is an important issue. My findings suggest that increasing the size of the patrol force would decrease crime (as this could hypothetically allow an increase in police presence in all locations). However, if increasing police presence in one location simply shifts crime to the next location, it could raise significant concerns about increasing police presence.

### Table 5. The Impact of Police Presence on Arrests

<table>
<thead>
<tr>
<th></th>
<th>IV = Outside Calls Ratio***</th>
<th>IV = Outside Calls**</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Police Vehicles</strong></td>
<td>0.025 (SE 0.034)</td>
<td>0.079 (SE 0.0003)</td>
</tr>
<tr>
<td><strong>Temperature</strong></td>
<td>0.0003 (SE 0.00001)</td>
<td>0.00039 (SE 0.00001)</td>
</tr>
<tr>
<td><strong>Precipitation</strong></td>
<td>-0.001 (SE 0.0003)</td>
<td>-0.001 (SE 0.0003)</td>
</tr>
<tr>
<td><strong>Twilight</strong></td>
<td>-0.002 (SE 0.0002)</td>
<td>-0.002 (SE 0.001)</td>
</tr>
<tr>
<td><strong>Holiday</strong></td>
<td>0.00001 (SE 0.0004)</td>
<td>0.005 (SE 0.004)</td>
</tr>
<tr>
<td><strong>Dark</strong></td>
<td>-0.004 (SE 0.0002)</td>
<td>-0.005 (SE 0.0002)</td>
</tr>
<tr>
<td><strong>Beat FE</strong></td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td><strong>Month FE</strong></td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td><strong>Day of Week FE</strong></td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td><strong>Weekend × Hour FE</strong></td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td><strong>First-Stage F</strong>-Statistic</td>
<td>321.73 (SE 0.075)</td>
<td>40.81 (SE 0.035)</td>
</tr>
<tr>
<td><strong>Observations</strong></td>
<td>2,017,676</td>
<td>2,017,676</td>
</tr>
</tbody>
</table>

Each observation is a beat and hour in 2009. The average arrest rate is 0.03 (SD 0.29), average police presence is 0.605 (SD 1.078). Standard errors in parentheses account for geographic clustering within a 10 km radius and serial correlation of 5 hours.

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TABLE 6.—THE EFFECT OF POLICE PRESENCE ON OTHER TYPES OF CALLS

<table>
<thead>
<tr>
<th></th>
<th>IV = Outside Calls Ratio(^a)</th>
<th>IV = Outside Calls(^a)</th>
</tr>
</thead>
<tbody>
<tr>
<td>A. Dependent Variable = Suicide Reports (mean of dependent variable 0.001, SD 0.037)</td>
<td>−0.001 (0.002)</td>
<td>−0.003 (0.003)</td>
</tr>
<tr>
<td>B. Dependent Variable = Abandoned Child Reports (mean of dependent variable 0.001, SD 0.027)</td>
<td>−0.003 (0.002)</td>
<td>−0.001 (0.003)</td>
</tr>
<tr>
<td>C. Dependent Variable = Fire Reports (mean of dependent variable 0.002, SD 0.039)</td>
<td>−0.001 (0.002)</td>
<td>0.005(^c) (0.003)</td>
</tr>
<tr>
<td>D. Dependent Variable = Drug House Reports (mean of dependent variable 0.001, SD 0.037)</td>
<td>−0.001 (0.002)</td>
<td>−0.002 (0.003)</td>
</tr>
<tr>
<td>E. Dependent Variable = Placebo Categories A-D (mean of dependent variable 0.005, SD 0.071)</td>
<td>−0.005 (0.004)</td>
<td>−0.006 (0.006)</td>
</tr>
</tbody>
</table>

\(^a\) The number of police vehicles patrolling the beat at given hour (60 minutes of presence).

\(^b\) The OCR is a proxy for police activity outside the beat. It is calculated by dividing the weighted sum of unrelated calls by the weighted sum of officers assigned to patrol the sector where the call took place.

\(^c\) OC is a weighted sum of unrelated 911 calls occurring within the division outside the beat. This instrument is interacted with division in order to allow out-of-beat calls to have different effects in different divisions. Significant at \(\ast\) 10%; \(\ast\ast\) 5%; and \(\ast\ast\ast\) 1%.

TABLE 7.—THE DETERRENCE EFFECT OF POLICE ON CRIME BY BEAT SIZE

<table>
<thead>
<tr>
<th></th>
<th>Small (1)</th>
<th>Medium (2)</th>
<th>Large (3)</th>
<th>Small (4)</th>
<th>Medium (5)</th>
<th>Large (6)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Police Vehicles(^a)</td>
<td>−0.253*** (0.075)</td>
<td>−0.246*** (0.060)</td>
<td>−0.116*** (0.035)</td>
<td>−0.047 (0.078)</td>
<td>−0.278*** (0.099)</td>
<td>−0.152*** (0.046)</td>
</tr>
<tr>
<td>First-Stage F-Statistic</td>
<td>113.87</td>
<td>182.86</td>
<td>172.37</td>
<td>13.73</td>
<td>12.53</td>
<td>24.15</td>
</tr>
<tr>
<td>Mean Level of Police Presence</td>
<td>[0.615] [0.856] [1.474]</td>
<td>[0.615] [0.856] [1.474]</td>
<td>[0.615] [0.856] [1.474]</td>
<td>[0.615] [0.856] [1.474]</td>
<td>[0.615] [0.856] [1.474]</td>
<td>[0.615] [0.856] [1.474]</td>
</tr>
<tr>
<td>Mean Level of Crime</td>
<td>[0.404] [0.427] [0.395]</td>
<td>[0.404] [0.427] [0.395]</td>
<td>[0.404] [0.427] [0.395]</td>
<td>[0.404] [0.427] [0.395]</td>
<td>[0.404] [0.427] [0.395]</td>
<td>[0.404] [0.427] [0.395]</td>
</tr>
<tr>
<td>Observations</td>
<td>669,570</td>
<td>669,651</td>
<td>678,455</td>
<td>669,570</td>
<td>669,651</td>
<td>678,455</td>
</tr>
<tr>
<td>Beat FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Month FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Day of Week FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Weekend x Hour FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>

Each observation is a beat and hour in 2009. Standard errors in parentheses account for geographic clustering within a 10 km radius and serial correlation of 5 hours. Standard deviations are presented in brackets.

\(^a\) The number of police vehicles patrolling the beat at given hour (60 minutes of presence = 1 vehicle).

\(^b\) The OCR is a proxy for police activity outside the beat. It is calculated by dividing the weighted sum of unrelated calls by the weighted sum of officers assigned to patrol the sector where the call took place.

\(^c\) OC is a weighted sum of unrelated 911 calls occurring within the division outside the beat. This instrument is interacted with division in order to allow out-of-beat calls to have different effects in different divisions. Significant at \(\ast\) 10%; \(\ast\ast\) 5%; and \(\ast\ast\ast\) 1%.

In a specific beat. I therefore consider the impact of police presence at larger geographic levels in table 8, where I would expect to find a smaller impact of police presence on crime if criminals are shifting their activities to neighboring beats.

In Dallas, beats are grouped into sectors, with each sector comprising roughly seven beats. The measured effect of police on violent crime and burglaries at the sector level when applying the OCR instrument is very similar to the estimates reported in table 4 from analyzing deterrence at the beat level. However, when applying this same instrument, I do not find a statistically significant effect of police presence on public disturbances at the sector level. When the analysis was run at the beat level, I measure an elasticity of −.7 which is significant at the 1% level. These estimates suggest that most crimes do not easily displace to neighboring areas.

V. Conclusion

While there exists an abundance of research and views regarding the deterrent effects of policing on crime, there has yet to be a detailed analysis using information on how the exact location of police officers affects behavior. In a survey conducted in May 2010, 71% of city officials reported...
The results presented in this paper raise concerns that frequently assigning officers to out-of-beat 911 calls may have significant costs in terms of deterring future crimes. I estimate that a 10% decrease in police presence at a given beat and hour increases crime at that location by 7%. This estimate is especially relevant to 911 calls as my instruments focus on shifts in police presence that are created because officers are assigned to incidents outside their beat. This paper asks. What happens when a police car leaves its allocated area to fulfill other departmental duties? I find that shortening response times may have a direct impact on the deterrence effect of patrol officers. This problem will only increase as fewer police officers are hired.

Despite the concern that deterrence is negatively affected by the assignment of officers to out-of-beat calls, the flip side of this finding is that the thin allocation of officers across large areas (which is driven by the rapid-response philosophy) can have crime prevention benefits. The prevalent assumption that there is a tension between the rapid-response philosophy and deterrence is not borne out by my research. In other words, the fact that the movement of these allocated officers affects crime implies that allocating officers in an effort to provide fast response times can be wedded to a deterrence policy. While the allocation of officers to beats may be driven by the demands of providing fast response times, in reality, the presence of these cars reduces the probability of crime. This may imply that police executives can “have your cake and eat it too” but does not rule out the possibility that stronger deterrence may be achieved by a hot-spots versus rapid-response allocation strategy. Either way, my results highlight the caution that must be taken in order to maximize the deterrence benefits of a rapid-response system. While arriving quickly at the scene of an incident may help to lower the expected benefit of committing a crime (see Blanes i Vidal & Kirchmaier, 2018, and Mastrobuoni, 2019), it can also disrupt preemptive police activity.

My results suggest that optimizing the impact of policing on crime requires weighing the costs and benefits of assigning officers to out-of-beat calls. But what is the aggregate impact of outside calls on crime? If beats where 911 incidents occur gain the minutes lost at the assigned officer’s patrol beat, we might expect a decrease in crime that matches the increase at the abandoned beat and no aggregate effect of outside calls on crime. However, this does not take into consideration the commuting time lost, as well as the fact that an officer responding to an incident may not be as visible as an officer on general patrol. Additionally, losing an officer assigned to patrol a beat that he or she is familiar with may have a very different effect from the added presence of a new officer in a possibly unknown environment. Indeed, recent research on increases in patrol-surrounding areas where crime occurred has shown limited deterrence effects (see Blanes i Vidal & Mastrobuoni, 2018, and Blattman et al., 2018).

In addition to providing a measure of the crime costs of decreasing police force size throughout the United States, this paper provides insight into the mechanism through which decreases in the number of police personnel in order to deal with budget cuts resulting from the economic downturn. With lower budgets, police departments are being forced to make tough decisions regarding police activities and deployment. Understanding how these deployment techniques affect crime is key for optimizing outcomes given the current budgets.

Police department performance measures are often a function of crime rates, arrests, response times, and clearance rates (the proportion of crimes reported that are cleared by arrests). Some deterrence programs may take time to develop and see results. Thus, a police department that is very involved in neighborhood-based crime reduction activities may get little reward for its effort in terms of decreased crime rates. Additionally, as crime rates and clearance rates are influenced by outside factors and their outcomes are an imprecise reflection of investment, departments may prefer to focus on shortening response times, an easily measured police activity. Indeed, the Dallas Morning News reported in 2013 that after criticism of rising response times to 911 calls, the Dallas Police Department “temporarily reassigned dozens of officers who normally spend much of their time targeting drug activity to duties where they respond to 911 calls” (Eiserer, 2013).

Information released in U.S. Department of Justice (2011).

See Davis (2012) for a more in-depth discussion regarding police outcomes and outputs (police investment).
POLICE PRESENCE, RAPID RESPONSE RATES, AND CRIME PREVENTION

POLICE PRESENCE, RAPID RESPONSE RATES, AND CRIME PREVENTION

police reduce crime. My outcomes are particularly interesting given recent studies that imply that policing is effective only when focused at specific high-crime locations. One interpretation of my results is that police do not need to be micromanaged, and simply assigning them to a fairly large geographic area (beats average 1.7 square miles) will reduce crime. However, the Dallas Police Department is known to follow a directed patrol data–driven strategy that attempts to direct patrol specifically to hot-spot areas (street blocks with very high crime rates). Thus, within the beat, allocated police may be focused on specific hot-spot areas that they are forced to abandon when answering a call.

This paper attempts to shed light on what police are doing in order to lower crime. My results show that their geographic presence alters crime outcomes. The next natural step is to understand how the activities of patrol officers create these crime impacts. I find that assigning officers to out-of-beat calls not only reduces police presence but also lowers arrest rates. Thus, it is plausible that part of the deterrence effect discussed in this paper is driven by an incapacitation effect, where crime decreases because a criminal is arrested before he or she can commit multiple offenses. However, as this effect occurs immediately (within the same hour), it also suggests a separate deterrence channel where increased police visibility has a direct impact on crime outcomes.

REFERENCES


Blattman, Christopher, Donald Green, Daniel Ortega, and Santiago Tobon, “Place Based Interventions at Scale: The Direct and Spillover Effects of Policing and City Services on Crime,” NBER working paper 23941 (2018).


