

Use of HbA_{1c} in Predicting Progression to Diabetes in French Men and Women

Data from an Epidemiological Study on the Insulin Resistance Syndrome (DESIR)

CELINE DROUMAGUET, MD, MSc¹
 BEVERLEY BALKAU, PHD^{1,2}
 DOMINIQUE SIMON, MD, PHD^{1,2,3}
 EMILE CACES, MSc⁴

JEAN TICHET, MD⁴
 MARIE ALINE CHARLES, MD^{1,2}
 EVELINE ESCHWEGE, MD^{1,2}
 THE DESIR STUDY GROUP*

OBJECTIVE — Early identification of subjects at high risk for diabetes is essential, and random HbA_{1c} (A1C) may be more practical than fasting plasma glucose (FPG). The predictive value of A1C, in comparison to FPG, is evaluated for 6-year incident diabetes.

RESEARCH DESIGN AND METHODS — From the French cohort study Data from an Epidemiological Study on the Insulin Resistance Syndrome (DESIR), 1,383 men and 1,437 women, aged 30–65 years, were volunteers for a routine health check-up. Incident diabetes was defined by FPG ≥ 7.0 mmol/l or treatment by antidiabetic drugs. Multivariate logistic regression models were used to predict diabetes at 6 years. Receiver operating characteristic curves compared the predictive values of A1C and FPG.

RESULTS — At 6 years, 30 women (2.1%) and 60 men (4.3%) had developed diabetes. Diabetes risk increased exponentially with A1C in both sexes ($P < 0.001$). After stratifying on FPG, A1C predicted diabetes only in subjects with impaired fasting glucose (IFG) (FPG ≥ 6.10 mmol/l): the odds ratio (95% CI) for a 1% increase in A1C was 7.20 (3.00–17.00). In these subjects, an A1C of 5.9% gave an optimal sensitivity of 64% and specificity of 77% to predict diabetes.

CONCLUSIONS — A1C predicted diabetes, even though the diagnosis of diabetes was based on FPG, but it was less sensitive and specific than FPG. It could be used as a test if fasting blood sampling was not available or in association with FPG. In subjects with IFG, A1C is better than glucose to evaluate diabetes risk, and it could be used to select subjects for intensive early intervention.

Diabetes Care 29:1619–1625, 2006

The prevalence of type 2 diabetes is increasing worldwide, and it is projected that the number of adults with diabetes will double between 2000 and 2030 (1). This means a large burden for the health care system. Recent clinical trials have demonstrated that lifestyle

(2–4) or pharmaceutical (4–6) interventions in individuals with impaired glucose tolerance (IGT) can delay or prevent diabetes; thus high-risk subjects should be identified for early intensive lifestyle counseling or even pharmaceutical treatment (7).

Fasting and 2-h plasma glucose after an oral glucose tolerance test (OGTT) are currently used to identify subjects at high risk of diabetes (8): those with impaired fasting glucose (IFG) and IGT. However, the OGTT is not common in clinical practice, because it is time consuming, costly, and less reproducible (9) than measurement of fasting plasma glucose (FPG).

HbA_{1c} (A1C), an indirect measure of mean blood glucose over the previous 2–3 months, is correlated with FPG and 2-h plasma glucose (10–12). A1C is more reproducible than FPG (13) and within-subject coefficients of variation are 1.7 and 5.7%, respectively (14). Moreover, measurement of A1C does not require that the subject is fasting. The use of A1C could better integrate chronic hyperglycemia than FPG.

Few studies have investigated predicting diabetes using A1C and none in a Caucasian population. Moreover, previous investigations were in populations at high risk of diabetes. A study in Pima Indians (15) reported that A1C was an independent predictor for diabetes only in individuals with IGT, not in subjects with normal 2-h plasma glucose. The same relation was found in a Chinese study (16).

To determine whether A1C predicted incident diabetes after a 6-year follow-up in a Caucasian population, we analyzed data from the prospective French cohort study: Data from an Epidemiological Study on the Insulin Resistance Syndrome (DESIR).

RESEARCH DESIGN AND METHODS

The 3,854 subjects studied, aged 30–65 years, were included in 1994–1996 in the DESIR Study, a follow-up study on the development of the insulin resistance syndrome. They were volunteers for a routine health check-up in seven Health Examination Centers financed by the French Social Security in the center-west of France. All participants gave informed consent, and the study was approved by an ethics committee (Comité Consultatif de Protection

From the ¹Institut National de la Santé et de la Recherche Médicale U258, Villejuif, France; the ² Faculté de Médecine, University of Paris-Sud, Villejuif, France; the ³Service de Diabétologie et Métabolisme, Groupe Hospitalier Pitié-Salpêtrière, Paris, France; and the ⁴Institut inter-Régional pour la Santé, La Riche, France.

Address correspondence and reprint requests to Beverley Balkau, INSERM U258-IFR69, 16 avenue Paul Vaillant-Couturier, 94807 Villejuif cedex, France. E-mail: balkau@vjf.inserm.fr.

Received for publication 23 December 2005 and accepted in revised form 13 April 2006

*A complete list of DESIR Study Group members can be found in the APPENDIX.

Abbreviations: DESIR, Data from an Epidemiological Study on the Insulin Resistance Syndrome; FPG, fasting plasma glucose; IFG, impaired fasting glucose; IGT, impaired glucose tolerance; OGTT, oral glucose tolerance test; ROC, receiver operating characteristic.

A table elsewhere in this issue shows conventional and Système International (SI) units and conversion factors for many substances.

DOI: 10.2337/dc05-2525

© 2006 by the American Diabetes Association.

The costs of publication of this article were defrayed in part by the payment of page charges. This article must therefore be hereby marked "advertisement" in accordance with 18 U.S.C. Section 1734 solely to indicate this fact.

Table 1—Comparison of baseline characteristics between men and women who developed or did not develop diabetes at 6 years*

	Men			Women		
	Nondiabetic	Diabetic	Age-adjusted P value	Nondiabetic	Diabetic	Age-adjusted P value
n	1,323	60		1,407	30	
Age (years)	46.9 ± 9.9	50.5 ± 9.6		47.4 ± 9.9	53.2 ± 8.2	
A1C (%)	5.4 ± 0.4	5.8 ± 0.5	<0.0001	5.3 ± 0.4	5.9 ± 0.6	<0.0001
Fasting plasma glucose (mmol/l)	5.54 ± 0.48	6.17 ± 0.50	<0.0001	5.17 ± 0.46	6.00 ± 0.61	<0.0001
Insulin (pmol/l)	41.1 ×/÷ 1.7	57.9 ×/÷ 1.8	<0.0001	39.2 ×/÷ 1.6	64.5 ×/÷ 1.8	<0.0001
BMI (kg/m ²)	25.1 ± 2.9	27.6 ± 4.3	<0.0001	23.6 ± 3.7	29.4 ± 6.0	<0.0001
Waist circumference (cm)	89 ± 9	98 ± 11	<0.0001	77 ± 10	91 ± 13	<0.0001
HDL cholesterol (mmol/l)	1.51 ± 0.38	1.47 ± 0.44	0.3	1.82 ± 0.42	1.48 ± 0.38	<0.0001
Triglycerides (mmol/l)	1.12 ×/÷ 1.7	1.67 ×/÷ 1.8	<0.0001	0.84 ×/÷ 1.64	1.50 ×/÷ 1.57	<0.0001
Systolic blood pressure (mmHg)	132 ± 14	139 ± 16	0.001	126 ± 14	138 ± 16	0.0005
Diastolic blood pressure (mmHg)	82 ± 10	84 ± 9	0.1	77 ± 9	83 ± 6	0.003
Hypertension (%)†	42	63	0.008	29	63	0.006
Family history of diabetes (%)	17	25	0.08	20	47	0.0006
Sporting activity (%)						
None	46	53		50	77	
Moderate (<1 session per week)	20	22	0.5	15	13	0.03
Frequent (≥1 session per week)	34	25		35	10	
Smoking (%)	25	28	0.2	12	13	0.2
Alcohol intake (%)‡						
None	13	7		36	27	
Moderate	73	65	0.02	41	50	0.5
Elevated	14	28		23	23	
Child with birth weight >4 kg	—	—	—	15	31	0.03

Data are means ± SD, geometric mean ×/÷ SD, or column %. *Comparisons are adjusted on age. †Arterial hypertension was defined by a systolic blood pressure ≥140 mmHg and/or a diastolic blood pressure ≥90 mmHg and/or an antihypertensive treatment (30). ‡Alcohol intake coding: no consumption, moderate (1–20 g/day for women and 1–40 g/day for men), and elevated (>20 g/day for women and >40 g/day for men).

des Personnes pour la Recherche Biomédicale of Kremlin-Bicêtre).

Of the 3,854 subjects, we excluded 114 with known diabetes or FPG ≥7.0 mmol/l at baseline, 92 who had an unknown glucose status, and 21 without A1C, waist circumference, or BMI information. Among the 3,627 remaining, 2,924 (81%) were reexamined 6 years after inclusion and the 104 subjects with incomplete information to evaluate glucose status were excluded: 1,383 men and 1,437 women were studied.

Participants were examined at inclusion and at 6 years. A medical interview provided information about lifestyle, use of medication, and personal and familial history of diabetes. Weight and height were measured with subjects lightly clothed, and a tape measure was used around the waist at the smallest circumference between the lower ribs and the iliac crests. Blood pressure was measured by a doctor with a sphygmomanometer on the right arm, with subjects lying at rest for at least 5 minutes.

Blood was drawn after a 12-h fast. A1C was measured in a central laboratory

by high-performance liquid chromatography, using a L9100 automated ion-exchange analyzer (Hitachi/Merck-VWR). This assay was standardized to the National Glycohemoglobin Standardization Program. Intra- and interassay coefficients of variation (CVs) for A1C were 1.5 and 1.8%. FPG was assayed by the glucose oxidase method on fluoro-oxalated plasma, using a Technicon RA 1000 analyzer (Bayer). Fasting insulin was measured by a microenzyme immunoassay and total cholesterol, HDL cholesterol, and triglycerides by enzymatic methods.

Definitions

Diabetes was defined according to the 1997 American Diabetes Association criteria (17): FPG ≥7.00 mmol/l or treatment by oral antidiabetic drugs or insulin. At inclusion, subjects were classified into three FPG groups: <5.60, 5.60–6.09, and ≥6.10 mmol/l (IFG). A family history of diabetes was coded if there was at least one diabetic first-degree relative. Sporting activity was assessed according to the number of sessions per week, alcohol in-

take was assessed by grams of alcohol per day, and subjects were classified as smokers or nonsmokers at inclusion.

Statistical analyses

The symmetry of the distributions of quantitative variables was assessed graphically: logarithms of insulin and triglycerides were used. Results are given as means ± SD or as geometric means ×/÷ SD, in the case of logarithmic transformation. Means were compared by Student's *t* tests or by ANCOVA, after age adjustment. For qualitative variables, results are expressed as percentages and compared by χ^2 or Fisher's tests or by logistic regression after age adjustment. The relation between A1C and FPG was evaluated by a Pearson correlation coefficient.

The role of A1C in the risk of developing diabetes at 6 years was evaluated by logistic regression models adjusted for age. Analyses were by sex, and in a second phase, sexes were pooled. For each continuous risk factor, linearity was studied by plotting the middle of each quartile group against the β coefficient of each quartile, obtained from logistic regression

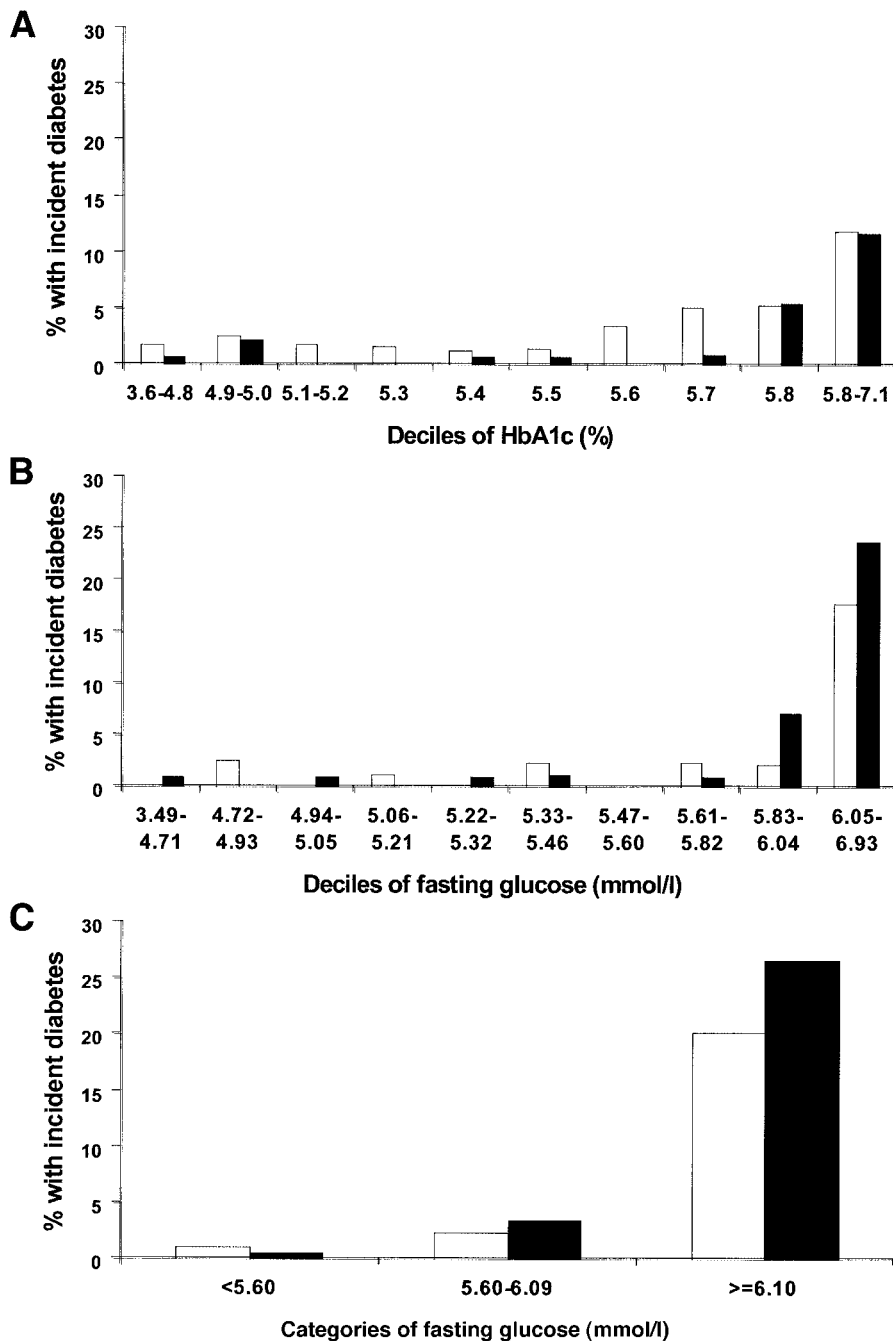


Figure 1—Percentage of men and women with diabetes after 6 years of follow-up according to A1C deciles (A), FPG deciles (B), and fasting plasma glucose (C) categories. □, men; ■, women.

(18). A linear trend was observed for BMI, waist circumference, triglycerides (log), HDL cholesterol, insulin (log), and systolic and diastolic pressure, and the continuous variable was chosen. A1C had a nonlinear relation, and its squared term was added to model this relation. Alcohol was coded in classes (Table 1). Diabetes risk according to A1C was modeled by logistic regression, adjusting on age and using those variables with a trend toward being significant ($P = 0.25$) in the age-

adjusted comparison between diabetic and nondiabetic groups. For adjusting variables, in the case of highly correlated variables (e.g., BMI and waist circumference), only the variable most significantly related to diabetes was used. Variable selection was based on likelihood ratio tests. FPG was added to test whether A1C remained significantly predictive. Likelihood ratio tests were used to evaluate interaction terms between A1C and FPG. As this interaction was marginally signif-

icant, we stratified on FPG categories. All models fitted well according to the Hosmer-Lemeshow test (18).

We compared the predictive performances of A1C and FPG as continuous variables, using receiver operating characteristic (ROC) curves, plotting sensitivity against $(1 - \text{specificity})$ at all possible thresholds. STATA software was used to calculate the ROC curves and to estimate areas (95% CIs) under these curves and the statistical significance of differences between these areas (19). The optimal threshold corresponds to the value where sensibility plus specificity is maximized.

All other analyses used the SAS software (SAS Institute, Cary, NC). A P value < 0.05 was regarded as statistically significant for a two-sided test.

RESULTS

Incidence of diabetes at 6 years and baseline characteristic of subjects according to development or not of diabetes

Ninety subjects developed diabetes over the 6 years: 30 women and 60 men, 2.1 and 4.3%, respectively. The incidence of diabetes at 6 years increased in a nonlinear manner with the deciles of A1C and FPG and with the three categories of FPG at inclusion (Fig. 1). Subjects excluded and included were comparable at baseline for all variables except age (44.1 ± 9.9 vs. 47.3 ± 9.9 years) and smoking (32 vs. 18%).

Mean A1C at inclusion was $5.3 \pm 0.4\%$ for the 1,913 subjects with FPG < 5.60 mmol/l, $5.5 \pm 0.4\%$ for the 635 subjects with FPG between 5.60 and 6.09 mmol/l, and $5.7 \pm 0.5\%$ for the 272 individuals with FPG ≥ 6.10 mmol/l, ($P < 0.0001$). The Pearson correlation coefficient between A1C and FPG at inclusion was $0.38 (P < 0.0001)$.

In both sexes, subjects who developed diabetes were older than those who did not ($P = 0.0005$) (Table 1). After age adjustment, mean A1C, FPG, insulin, triglycerides, BMI, waist circumference, and systolic blood pressure were higher in those who developed diabetes, compared with those who did not, whereas HDL cholesterol was significantly lower only in women. A family history of diabetes was more frequent in the diabetic group, significantly in women. In men, alcohol intake was higher in the subjects with incident cases of diabetes and the absence of sporting activity and having a child

Table 2—Predictive value of A1C for diabetes at 6 years, according to successive multiple logistic regression models*

Variables included in model	β Coefficient of		P value of		–2 log likelihood
	A1C	A1C ²	A1C	A1C ²	
Age + sex					768.81
Age + sex + A1C					674.70†
Age + sex + A1C + A1C ²	–13.10	1.39	0.0001	<0.0001	660.69†
Age + sex + A1C + A1C ² + waist circumference	–13.09	1.35	<0.0001	<0.0001	604.72†
Age + sex + A1C + A1C ² + waist circumference + triglycerides (log)	–13.10	1.35	<0.0001	<0.0001	589.30†
Age + sex + A1C + A1C ² + waist circumference + triglycerides (log) + familial history of diabetes	–12.84	1.33	<0.0001	<0.0001	582.99†‡

Data are β coefficients (= ln OR) of A1C and A1C²; $n = 2,820$. *Variables were added successively. †Likelihood ratio test between this model and the preceding nested model was significant at $P < 0.05$. ‡Likelihood ratio test between this model and a model with an interaction term between A1C and sex was not significant. The interaction between A1C and sex was tested only in this final model.

with a birth weight >4 kg were more common in the future diabetic women.

A1C and diabetes prediction

After adjustment for age, A1C predicted incident diabetes in both sexes (data not shown). Addition of insulin, systolic and diastolic pressure, smoking, and sporting activity did not improve the prediction of diabetes significantly in either sex; neither did alcohol in men nor having a child with a birth weight >4 kg in women. As variables included in the models and β coefficients were similar in men and women and as there was no significant interaction between sex and A1C, sexes were pooled (Table 2). A1C predicted diabetes at 6 years, independently of other variables. Diabetes risk increased exponentially with A1C at inclusion: odds ratio (OR) for an A1C increase from 4.5 (mean – 2 SD) to 5.0% was 0.90 (95% CI 0.50–1.50), from 4.5 to 5.5% was 1.50 (0.70–3.40), from 4.5 to 6.0% was 5.0 (2.00–12.80), and from 4.5 to 6.5% was 32.70 (11.50–92.60).

FPG was then included in this model, and as there was a marginally significant interaction between A1C and FPG ($P = 0.07$), models were stratified on FPG categories, adjusted on the same variables as the last model in Table 2. The addition of a squared A1C term did not improve the models, and there was no interaction between A1C and sex. A1C was only a significant risk factor for diabetes at 6 years in subjects with FPG ≥ 6.10 mmol/l: OR for a 1% A1C increase was 0.78 (95% CI 0.20–3.07) ($P = 0.7$) for FPG < 5.60 mmol/l, 1.47 (0.36–5.80) ($P = 0.6$) for FPG 5.60–6.09 mmol/l, and 7.20 (3.00–17.00) ($P < 0.0001$) for FPG ≥ 6.10 mmol/l (IFG).

For FPG ≥ 6.10 mmol/l, FPG as a

continuous variable was not predictive of diabetes, whether A1C was included in the model (OR for 1 mmol/l increase 1.0 [95% CI 0.2–4.8], $P = 0.9$) or not (2.4 [0.6–9.4], $P = 0.2$). In contrast, A1C was predictive of diabetes in these subjects (7.22 [2.95–17.80], $P < 0.001$, for a 1% A1C increase), independently of whether or not FPG was in the model.

ROC curves

In the entire study population of 2,820 subjects (Fig. 2), the area under the curve for FPG to predict incident diabetes at 6 years of follow-up was significantly greater than that for A1C: 0.85 (95% CI 0.80–0.90) versus 0.78 (0.72–0.83) ($P = 0.005$). The optimal value for FPG was 5.9 mmol/l, with sensitivity and specificity of 76 and 86%, respectively. For A1C, the optimal value was 5.7% with corresponding sensitivity and specificity of 66 and 88%.

In contrast, in subjects with IFG (FPG ≥ 6.10 mmol/l) (Fig. 2B), the area under the curve for A1C was significantly greater than that for FPG: 0.75 (95% CI 0.69–0.82) versus 0.60 (0.52–0.68) ($P = 0.001$). The optimal cutoff for A1C was 5.9%, with sensitivity and specificity of 64 and 77% and a positive predictive value of 44%.

If the test is based on both glucose and A1C, the optimal cutoffs were 5.9 mmol/l and 5.0%, respectively, with sensitivity of 73% and specificity of 87%. Alternatively, for either glucose or A1C, the corresponding optimal cutoffs were 5.9 mmol/l and 6.3%, with corresponding sensitivity and specificity of 77 and 86%, respectively.

CONCLUSIONS— A1C predicted diabetes in this Caucasian population and

the risk increased exponentially in both sexes. As for FPG, Fig. 1 provides evidence that there is an A1C threshold below which the incidence of diabetes is very low and above which the incidence is considerably higher. However, A1C is less sensitive and specific than FPG for diabetes in the entire population (Fig. 2A).

For screening in a nonfasting state, A1C could be used in a general population. At the present time, this use is limited because FPG is still needed for the diagnosis of diabetes. However the diagnosis of diabetes by A1C has been debated for many years. In a study using an external gold standard, diabetic microvascular complications, McCance et al. (20) showed that A1C was as predictive as FPG and 2-h glucose. Moreover, several authors have recommended the use of A1C, in combination with FPG or random plasma glucose, to diagnose diabetes (13,21–23): a FPG ≥ 7.00 mmol/l or a random plasma glucose ≥ 11.10 mmol/l associated with an A1C exceeding the mean + 2 SD would better determine the diagnosis of treatment-requiring diabetes.

After stratifying on FPG classes, A1C was predictive only in subjects with IFG (FPG ≥ 6.1 mmol/l), with an OR of 7.2 (95% CI 3.00–17.00); thus subjects with IFG and high A1C could be identified for preventive care. Moreover within this IFG category, FPG did not predict diabetes risk, in contrast with A1C: subjects with A1C $> 5.9\%$ had a 50% risk of progressing to diabetes within 6 years. This optimal threshold of 5.9% remains to be confirmed in other populations.

It is important to note that the diagnostic criteria for diabetes in this study were based on FPG, giving it an advantage over A1C to predict diabetes. The best

way to compare the two parameters would be with an external diagnostic criterion, such as diabetic retinopathy.

This study was in volunteers presenting for a health check-up, which may constitute a healthier population than the general population, at low risk for type 2 diabetes. Diabetes prevalence in this cohort was 2.7% (24), slightly lower than national data from the French Health Care study of 3.4% (25) in the same age-group. Even if the population studied is not representative of the general French population, it probably represents the population, and, in particular, subjects with IFG, who would accept screening for diabetes and then preventive measures.

One limitation of our study is the small number of incident cases of diabetes and the resulting lack of power, especially in stratified analyses. A1C may also be significantly predictive of diabetes in subjects with the newer definition of IFG: FPG 5.60–6.09 mmol/l with a larger sample, but it would carry a lower OR than IFG (FPG ≥ 6.10 mmol/l).

Our results apply to diabetes diagnosed by FPG and could be different if the 2-h glucose measurement following an OGTT was also available, as a large percentage of subjects diagnosed as diabetic by a 2-h glucose value have normal FPG (26). We cannot exclude the possibility that some individuals, particularly subjects with IFG, were already diabetic at inclusion, which might overestimate the relation between A1C and diabetes. However, the OGTT is seldom used in clinical practice in France (27), and the American Diabetes Association encourages the use of FPG for screening (17). Therefore, this study reflects daily practice.

In participants with IFG, a small difference in FPG did not predict diabetes, in contrast to A1C. This suggests that A1C integrates blood glucose variations during the day, reflecting postload glucose abnormalities not shown by FPG (28). A prospective study, comparing 2-h glucose versus FPG and their association with A1C in predicting diabetes diagnosed by specific diabetic microvascular complications, is needed.

Our results are similar to those of Little et al. (15), except that we studied fasting and not 2-h postload plasma glucose. In 381 Pima Indians, A1C predicted diabetes diagnosed by 2-h glucose concentrations at 3.3 years follow-up only in subjects with IGT, with an OR for 1% increase in A1C of 6.76 (95% CI 1.77–25.8).

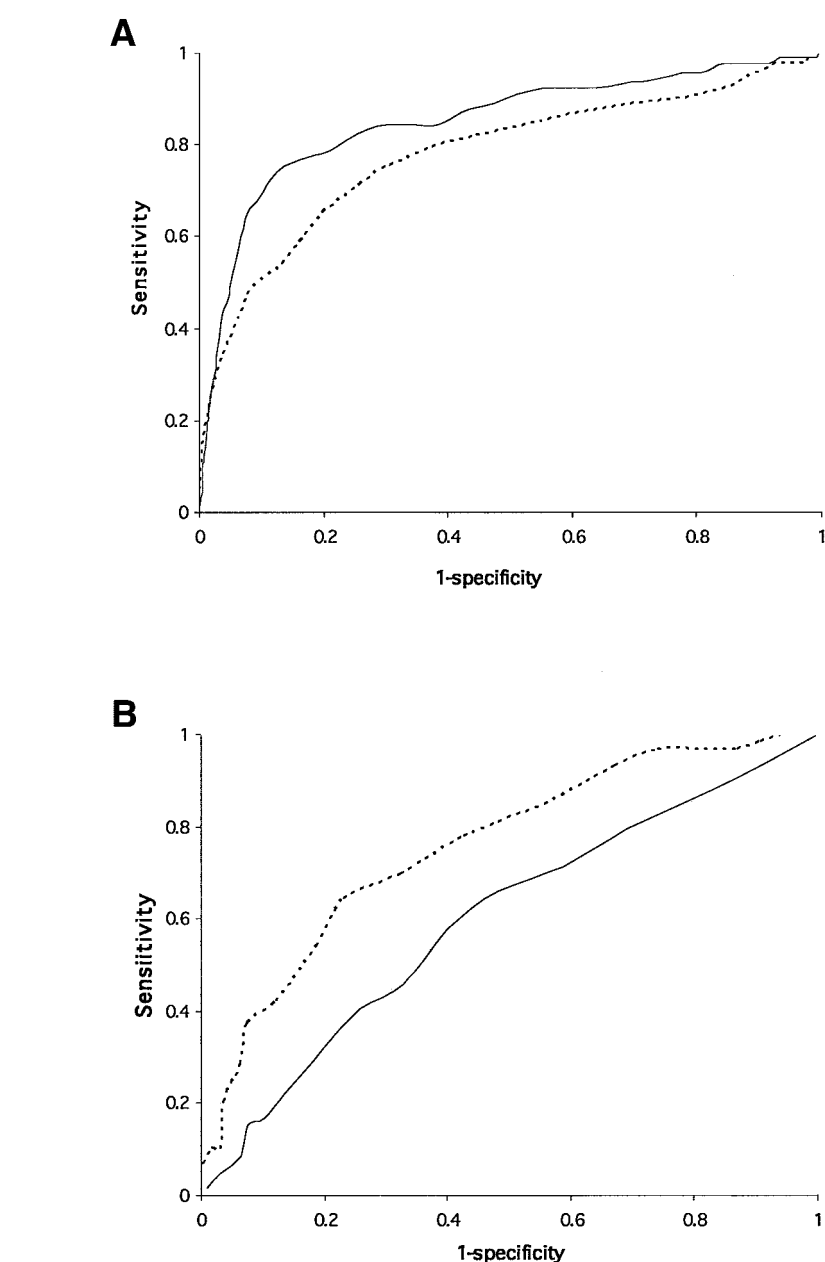


Figure 2—ROC curves for incident diabetes after 6 years of follow up: A1C (dashed line) and fasting plasma glucose (solid line) in the entire study population (A) and in subjects with IFG (B).

In another study, 208 Chinese subjects at high risk of diabetes, with FPG < 7.0 mmol/l and 2-h glucose < 11.1 mmol/l, were followed for 1.6 years (16). In subjects with IFG subjects, the progression to diabetes, diagnosed by 2-h glucose, was 44.1% for A1C $> 6.1\%$ versus 17.4% for A1C $< 6.1\%$. In subjects with normal FPG, the conversion rate to diabetes was 13.7 versus 8.1%.

Until 10 years ago, using A1C as a screening tool was limited by poor comparability between measures and between laboratories. Most routine A1C assays are now standardized against one of the local

standardization schemes, such as the National Glycohemoglobin Standardization Program. A new more specific reference measure was developed in 2003 (29). This prompted the reevaluation of A1C as a predictive test for diabetes.

In summary, in the absence of FPG, for example, in the middle of the day, A1C could be used to screen for diabetes in a general population. It could also be used in association with FPG, as in subjects with IFG, A1C better screened those at risk of diabetes, subjects who could be targeted for intensive prevention intervention. Moreover, it could help them to

become aware of their metabolic abnormalities and motivate asymptomatic subjects to modify their lifestyle.

Acknowledgments—This work was supported by cooperative contracts between INSERM, Caisse Nationale d'Assurance Maladie des Travailleurs Salariés, and Novartis Pharma, by INSERM Réseaux en Santé Publique and INSERM Interactions entre les Déterminants de la Santé; by the Association Diabète Risque Vasculaire, the Fédération Française de Cardiologie, La Fondation de France, de l'Association de Langue Française pour l'Etude du Diabète et des Maladies Métaboliques, and Office National Interprofessionnel des Vins; and by Ardis Medical, Bayer Diagnostics, Becton Dickinson, Cardionics, Lilly, Merck Santé, Novo Nordisk, Pierre Fabre, Roche, and Topcon.

APPENDIX

Members of the DESIR Study Group

INSERM U258: B. Balkau, P. Ducimetière, and E. Eschwège; INSERM U367: F. Alhenc-Gelas; Chu d'Angers: Y. Gallois and A. Girault; Hôpital Bichat: F. Fumeron and M. Marre; Centres d'Examens de Santé: Alençon, Angers, Caen, Chateauroux, Cholet, Le Mans, and Tours; Institut de Recherche en Médecine Générale: J. Cogneau; Médecins Généralistes des Départements; Institut Inter Régional pour la Santé: C. Born, E. Cacès, M. Cailleau, J.G. Moreau, F. Rakotozafy, J. Tichet, and S. Vol.

References

- Wild S, Roglic G, Green A, Sicree R, King H: Global prevalence of diabetes: estimates for the year 2000 and projections for 2030. *Diabetes Care* 27:1047–1053, 2004
- Pan XR, Li GW, Hu YH, Wang JX, Yang WY, An ZX, Hu ZX, Lin J, Xiao JZ, Cao HB, Liu PA, Jiang XG, Jiang YY, Wang JP, Zheng H, Zhang H, Bennett PH, Howard BV: Effects of diet and exercise in preventing NIDDM in people with impaired glucose tolerance: the Da Qing IGT and Diabetes Study. *Diabetes Care* 20:537–544, 1997
- Tuomilehto J, Lindstrom J, Eriksson JG, Valle TT, Hamalainen H, Ilanne-Parikka P, Keinänen-Kiukkaanniemi S, Laakso M, Louheranta A, Rastas M, Salminen V, Uusitupa M: Prevention of type 2 diabetes mellitus by changes in lifestyle among subjects with impaired glucose tolerance. *N Engl J Med* 344:1343–1350, 2001
- Knowler WC, Barrett-Connor E, Fowler SE, Hamman RF, Lachin JM, Walker EA, Nathan DM: Reduction in the incidence of type 2 diabetes with lifestyle intervention or metformin. *N Engl J Med* 346:393–403, 2002
- Chiasson JL, Josse RG, Gomis R, Hanefeld M, Karasik A, Laakso M: Acarbose for prevention of type 2 diabetes mellitus: the STOP-NIDDM randomised trial. *Lancet* 359:2072–2077, 2002
- Buchanan TA, Xiang AH, Peters RK, Kjos SL, Marroquin A, Goico J, Ochoa C, Tan S, Berkowitz K, Hodis HN, Azen SP: Preservation of pancreatic β -cell function and prevention of type 2 diabetes by pharmacological treatment of insulin resistance in high-risk Hispanic women. *Diabetes* 51:2796–2803, 2002
- Sherwin RS, Anderson RM, Buse JB, Chin MH, Eddy D, Fradkin J, Ganiats TG, Ginsberg HN, Kahn R, Nwankwo R, Rewers M, Schlessinger L, Stern M, Vinicor F, Zimman B: Prevention or delay of type 2 diabetes. *Diabetes Care* 27 (Suppl. 1):S47–S54, 2004
- de Vegt F, Dekker JM, Jager A, Hienkens E, Kostense PJ, Stehouwer CD, Nijpels G, Bouter LM, Heine RJ: Relation of impaired fasting and postload glucose with incident type 2 diabetes in a Dutch population: the Hoorn Study. *JAMA* 285:2109–2113, 2001
- Ko GT, Chan JC, Woo J, Lau E, Yeung VT, Chow CC, Cockram CS: The reproducibility and usefulness of the oral glucose tolerance test in screening for diabetes and other cardiovascular risk factors. *Ann Clin Biochem* 35 (Pt. 1):62–67, 1998
- Rohlfing CL, Wiedmeyer HM, Little RR, England JD, Tennill A, Goldstein DE: Defining the relationship between plasma glucose and HbA_{1c}: analysis of glucose profiles and HbA_{1c} in the Diabetes Control and Complications Trial. *Diabetes Care* 25:275–278, 2002
- Simon D, Senan C, Garnier P, Saint-Paul M, Papoz L: Epidemiological features of glycated haemoglobin A1c-distribution in a healthy population: the Telecom Study. *Diabetologia* 32:864–869, 1989
- Woerle HJ, Pimenta WP, Meyer C, Gosmanov NR, Szoke E, Szombathy T, Mitrakou A, Gerich JE: Diagnostic and therapeutic implications of relationships between fasting, 2-hour postchallenge plasma glucose and hemoglobin A1c values. *Arch Intern Med* 164:1627–1632, 2004
- Barr RG, Nathan DM, Meigs JB, Singer DE: Tests of glycemia for the diagnosis of type 2 diabetes mellitus. *Ann Intern Med* 137:263–272, 2002
- Rohlfing C, Wiedmeyer HM, Little R, Grotz VL, Tennill A, England J, Madsen R, Goldstein D: Biological variation of glycohemoglobin. *Clin Chem* 48:1116–1118, 2002
- Little RR, England JD, Wiedmeyer HM, Madsen RW, Pettitt DJ, Knowler WC, Goldstein DE: Glycated haemoglobin predicts progression to diabetes mellitus in Pima Indians with impaired glucose tolerance. *Diabetologia* 37:252–256, 1994
- Ko GT, Chan JC, Tsang LW, Cockram CS: Combined use of fasting plasma glucose and HbA_{1c} predicts the progression to diabetes in Chinese subjects. *Diabetes Care* 23:1770–1773, 2000
- Expert Committee on the Diagnosis and Classification of Diabetes Mellitus: Report of the Expert Committee on the Diagnosis and Classification of Diabetes Mellitus. *Diabetes Care* 20:1183–1197, 1997
- Hosmer D, Lemeshow S: *Applied Logistic Regression*. New York, Wiley, 1989
- DeLong ER, DeLong DM, Clarke-Pearson DL: Comparing the areas under two or more correlated receiver operating characteristic curves: a non parametric approach. *Biometrics* 44:837–845, 1988
- McCance DR, Hanson RL, Charles MA, Jacobsson LTH, Pettitt DJ, Bennett PH, Knowler WC: Comparison of tests for glycated haemoglobin and fasting and two hour plasma glucose concentrations as diagnostic methods for diabetes. *BMJ* 308:1323–1328, 1994
- Peters AL, Davidson MB, Schriger DL, Hasselblad V: A clinical approach for the diagnosis of diabetes mellitus: an analysis using glycosylated hemoglobin levels: Meta-analysis Research Group on the Diagnosis of Diabetes Using Glycated Hemoglobin Levels. *JAMA* 276:1246–1252, 1996
- Davidson MB, Schriger DL, Peters AL, Lorber B: Relationship between fasting plasma glucose and glycosylated hemoglobin: potential for false-positive diagnoses of type 2 diabetes using new diagnostic criteria. *JAMA* 281:1203–1210, 1999
- Davidson MB: Tests of glycemia. *Ann Intern Med* 138:517, 2003 [author reply 138:517, 2003]
- Balkau B, Eschwège E, Tichet J, Marre M: Proposed criteria for the diagnosis of diabetes: evidence from a French epidemiological study (D.E.S.I.R.). *Diabetes Metab* 23:428–434, 1997
- Ricordeau P, Weill A, Vallier N, Bourrel R, Fender P, Allemand H: Epidemiology of diabetes in metropolitan France. *Diabetes Metab* 26 (Suppl. 6):11–24, 2000
- Harris MI, Eastman RC, Cowie CC, Flegal KM, Eberhardt MS: Comparison of diabetes diagnostic categories in the U.S. population according to the 1997 American Diabetes Association and 1980–1985 World Health Organization diagnostic criteria. *Diabetes Care* 20:1859–1862, 1997
- Biolam: Les actes de biologie remboursés en 2001 et 2002 par le Régime Général d'assurance maladie. Available from <http://www.ameli.fr/244/DOC/1531/article.html>. Accessed 16 May 2005
- Monnier L, Lapinski H, Colette C: Contri-

butions of fasting and postprandial plasma glucose increments to the overall diurnal hyperglycemia of type 2 diabetic patients. *Diabetes Care* 26:881–885, 2003

29. Miedema K: Towards worldwide standardisation of HbA1c determination. *Diabetologia* 47:1143–1148, 2004
30. Guidelines Subcommittee of the WHO/ISH Mild Hypertension Liaison Commit-

tee: 1993 guidelines for the management of mild hypertension: memorandum from a World Health Organization/International Society of hypertension meeting. *Hypertension* 22:392–403, 1993